



# <sup>1</sup> Comparing the Impacts of Single- and Multi-Objective

- 2 **Optimization on the Parameter Estimation and Performance of a**
- **3 Land Surface Model**
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#### 10 Abstract

11 In land surface models (LSMs), precise parameter specification is crucial to reduce the inherent uncertainties and 12 enhance the accuracy of the simulation. However, due to the multi-output nature of LSMs, the impact of different 13 optimization strategies (e.g., single- and multi-objective optimization) on the optimization outcome and efficiency 14 remains ambiguous. In this study, we applied a revised particle evolution Metropolis sequential Monte Carlo (PEM-15 SMC) algorithm for both single- and multi-objective optimization of the Common Land Model (CoLM), constrained 16 by latent heat flux (LE) and net ecosystem exchange (NEE) measurements from a typical evergreen needle-leaf 17 forest observation site. The results reveal that the revised PEM-SMC algorithm, demonstrates a robust ability to 18 tackle the multi-dimensional parameter optimization challenge for LSMs. The sensitive parameters for different 19 target outputs can exhibit conflicting optimal values, resulting in single-objective optimization improving the 20 simulation performance for a specific objective at the expense of sacrificing the accuracy for other objectives. For 21 instance, solely optimizing for LE reduced the root-mean-square error (RMSE) of the simulated and observed LE by 22 20% but increased the RMSE of the NEE by 97%. Conversely, multi-objective optimization can not only ensure that 23 the optimized parameter values are physically sound but also balances the simulation performance for both LE and 24 NEE, as evidenced by the decrease in RMSE for LE and NEE of 7.2 W/m2 and 0.19 µmol m-2 s-1, respectively. In 25 conclusion, these findings reveal that comprehensively integrating the various available observational data for multi-

26 objective optimization is preferable for parameter calibration in complex models.

## 27 1 Introduction

28 Global warming has led to increasingly intricate dynamics between terrestrial surfaces and the atmospheric system, 29 necessitating advancements in the precision of the representations utilized within models, particularly those related to 30 terrestrial surface processes. Land surface models (LSMs) play a crucial role in accurately simulating the exchange of 31 water, carbon, and energy between the land surface and atmosphere, understanding biosphere-climate interactions, 32 and evaluating global climate change impacts (Field et al., 2004; Henderson-Sellers et al., 1995; McGuire et al., 2001). 33 However, LSMs are challenged by their complexity and the difficulty in accurately determining the numerous inherent 34 parameters, leading to significant uncertainties in simulating the land-atmosphere flux exchanges at a large scale 35 (Folberth et al., 2019; Li et al., 2020; Luo et al., 2016; Thum et al., 2017). Therefore, parameter calibration/estimation 36 which integrates multi-source observational data to optimize the model parameters has emerged as a fundamental step 37 in diminishing these uncertainties and enhancing the efficacy of the models in simulating the interaction processes 38 between land and atmosphere under the current and future exacerbated climate change conditions (Duan et al., 2017). 39

40 LSMs generally encompass multiple interdependent processes, where the configuration of the parameters has a 41 concurrent impact on the accuracy of the diverse output simulations (Bastrikov et al., 2018). The traditional parameter 42 calibration methods for LSMs predominantly target single-objective optimization, emphasizing parameter adjustment 43 to enhance the simulation performance for specific processes (e.g., surface carbon and water fluxes) (Kato et al., 2013; 44 Li et al., 2018; Ricciuto et al., 2011; Sellers et al., 1989; Xia et al., 2004b). However, given the intricate nature of 45 LSMs and the interactions among the multiple outputs, optimization targeting a single output can inadvertently 46 compromise the simulation accuracy for the other output variables. In recent years, multi-objective optimization 47 algorithms have gained increasing attention for their capacity to provide a balanced solution for multiple conflicting 48 objectives (Bastidas et al., 1999; Saini et al., 2021; Segura et al., 2016). These algorithms adeptly navigate the dynamic 49 input-output competition in complex models, achieving an optimal balance in overall model performance. 50 Consequently, multi-objective optimization is increasingly favored for dealing with the parameter estimation of





51 complex models with numerous outputs and interactive processes, such as hydrological models (Gupta et al., 1999; 52 Vrugt et al., 2003b), LSMs (Gong et al., 2015; Leplastrier et al., 2002; Varejao et al., 2013), and soil-vegetation-53 atmosphere coupled models (Liu et al., 2005; Pollacco et al., 2013). Despite this, most of the existing research 54 primarily focuses on the development and implementation of various multi-objective algorithms for parameter 55 estimation, with limited studies offering comprehensive scientific substantiation and practical validation for the 56 superiority of the multi-objective optimization strategy over single-objective optimization. Critical considerations 57 include the potential detriment of single-objective optimization to non-target outputs and whether multi-objective 58 optimization can concurrently improve the accuracy for multiple outputs. Furthermore, the distinctions between 59 single- and multi-objective optimization in terms of parameter estimation, model performance improvement, and 60 application reliability remain to be clarified.

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62 The introduction of expanded parameter spaces and increased optimization complexity by multi-objective 63 optimization emphasizes the importance of developing more efficient global optimization algorithms. Over the past 64 few decades, numerous optimization algorithms have been employed in the LSMs to obtain appropriate parameter 65 values, including genetic algorithms (D'heygere et al., 2006; Ines et al., 2008), particle swarm optimization (Eberhart 66 et al., 2001; Gill et al., 2006; Zhang et al., 2009), shuffled complex evolution (SCE) (Duan et al., 1993, 1994), the 67 Markov chain Monte Carlo (MCMC) algorithm (Smith et al., 2008; Van et al., 2005; Zhang et al., 2019), and the sequential Monte Carlo (SMC) algorithm (Dong et al., 2023; Jeremiah et al., 2011; Zhu et al., 2018). Among these 68 69 algorithms, the SMC samplers, which are also known as particle filters, are recognized for theoretically providing a 70 direct and effective way of estimating the posterior distribution through a series of gradual approximations and weight 71 redistributions (Doucet et al., 2001; Fan et al., 2008). However, the traditional SMC samplers face the challenge of 72 the particle impoverishment problem, which is a consequence of the resampling step that discards less significant 73 particles in favor of duplicating more promising ones. To combat this, candidate particle generation algorithms using 74 an MCMC transition kernel have been implemented in the SMC moving step to enhance the particle diversity and 75 quality [e.g., the random walk Metropolis (RWM) algorithm (Metropolis et al., 1953) and the adaptive random walk Metropolis (ARM) algorithm (Chopin, 2002; Jeremiah et al., 2011)]. In line with these advancements, our previous 76 77 research introduced the particle evolution Metropolis (PEM) method, which is a novel candidate-generating approach 78 that combines the appealing aspects of genetic and evolutionary algorithms with the robustness of the Metropolis-79 Hasting (M-H) algorithm (Zhu et al., 2018). In the case study on a synthetic multi-dimensional bimodal normal 80 distribution, the PEM-SMC sampler demonstrated a superior efficiency in exploring high-dimensional and complex 81 parameter spaces, compared to other SMC samplers (i.e., RWM-SMC and ARM-SMC). Nevertheless, while the 82 genetic-styled operations (e.g., crossover and mutation) enhance the particle diversity, they simultaneously impose 83 significant computational burdens, necessitating some redundant calculations in the original algorithm for each newly 84 generated particle. This limitation is particularly pronounced in the parameter optimization of complex LSMs with extensive single-run durations. In this paper, based on our previous work, we present a revised PEM-SMC algorithm 85 86 that further refines the particle candidate mechanism in the moving step to improve the computational efficiency while 87 maintaining the optimization efficiency. Moreover, while our previous research focused on the performance of the





- PEM-SMC sampler in simple hydrological model parameter optimization, the current paper explores its applicability
   and potential in the high-dimensional parameter optimization of more complex LSMs.
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91 This paper provides a comprehensive parameter sensitivity analysis and optimization of the Common Land Model 92 (CoLM) based on the latent heat flux (LE) and net ecosystem exchange (NEE) measurements from a FLUXNET 93 observation site and extensively evaluates the impact of different optimization strategies on the model simulation 94 performance. The study encompasses: (a) the application of both qualitative and quantitative sensitivity analysis 95 methods to precisely identify the pivotal parameters for accurately simulating water and carbon processes in the CoLM; 96 (b) a validation of the efficacy of the integrated optimization framework combining sensitivity analysis with the 97 modified PEM-SMC algorithm in optimizing the multi-dimensional parameters of complex LSMs; and (c) the implementation of both single-objective and multi-objective optimization of the model parameters to elucidate the 98 99 distinctions in the optimization outcome and efficiency attributed to different constraints. The novel optimization 100 algorithm proposed in this paper, coupled with the extensive investigation into different optimization strategies, provides methodological insights for the parameter optimization of LSMs. 101

## 102 2 Materials and Methods

## 103 2.1 The CoLM and Adjustable Parameters

The Common Land Model (CoLM), as developed by Dai et al. (2003), has significantly evolved from its initial form into a globally acclaimed third-generation LSM. The CoLM is characterized by its intricate, comprehensive, and precise representation of biophysical, biochemical, ecological, and hydrological processes on the land surface, and has been widely used in the simulation of energy, momentum, water, and carbon transport between the land and atmosphere (Ment et al., 2009; Zeng et al., 2002).

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110 In this study, we strategically selected 40 out of the 46 time-invariant parameters from the CoLM, deliberately 111 excluding certain of the model's internal parameters, to conduct a comprehensive sensitivity analysis and optimization 112 of the model parameters (Table S1). These chosen parameters, which are inherently static, represent the physical 113 properties of vegetation and soil and can be adjusted according to the specific local environmental conditions. For 114 ease of reference, these parameters were indexed from P1 to P40. The predefined initial range of these parameters 115 significantly influences the results of the high-dimensional parameter sensitivity analysis and optimization. Ensuring 116 the objectivity and rationality of the parameter range is crucial for the validity of the final calibration results. 117 Consequently, the range for the 40 parameters was established based on a literature review, the local environmental 118 conditions at the study site, and the biophysical/chemical meaning of each specific parameter (Sellers et al., 1996; Ji 119 et al., 2010; Li et al., 2013). A detailed description of this process is provided in Sect. S2.





#### 120 **2.2 Design of the Parameter Sensitivity Analysis**

121 Prior to the optimization, a global sensitivity analysis of the parameters was conducted to assess the parameter 122 importance via a relatively inexpensive "coarse" sampling of the parameter space. Among the sensitivity analysis 123 techniques, qualitative methods enable the identification of crucial parameters using a relatively small sample size 124 (hundreds to thousands), albeit with significant outcome variability among different methodologies. In contrast, 125 quantitative methods based on variance decomposition provide superior precision but require extensive datasets (from 126 tens to hundreds of thousands). To navigate these challenges, we identified the 10 most sensitive parameters by 127 integrating the results from three qualitative methods and subsequently ranked them by employing a quantitative 128 method for further refinement (Li et al., 2013). Here, we provide a brief description of how these methods are applied 129 in variable selection.

(a) The delta test (DT) method: a noise variance estimator based on the concept of nearest neighbors (NNs) (Eirola et al., 2008; Pi et al., 1994). For a given set of input parameters  $\theta_i$  (i = 1, ..., m) and associated output *Y*, the assumption

132 is that there is a functional dependence between them:

$$Y = f(\theta_i) + \varepsilon_i \tag{1}$$

where  $\varepsilon_i$  is an independent identically distributed random variable with zero mean. Noise variance estimation is the study of how to give an "a priori estimate" for  $\delta(\varepsilon)$ . The NN of a point is defined as the unique point that minimizes

136 a Euclidean distance to that point in the input space:

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$$N(i) := \underset{\substack{j \neq i}}{\arg\min} \left\| \theta_i - \theta_j \right\|^2 \tag{2}$$

138 The DT criterion of a variable subset  $s \subseteq \{\theta_1, ..., \theta_m\}$  is then written as:

139  $\delta(s) = \frac{1}{2N} \sum_{i=1}^{N} \left( Y_{N_s(i)} - Y_i \right)^2$ (3)

where N is the sample size,  $Y_i$  is the function value corresponding to  $\theta_i$ , and  $Y_{N_s(i)}$  is the function value corresponding to the NNs of the input point  $\theta_i$  for subset *s*. Consequently, the variable subset *s* with the smallest DT criterion is the most sensitive parameter.

(b) The multivariate adaptive regression splines (MARS) method: a non-parametric regression technique (Friedman,
144 1991) that employs a specific class of basis functions as predictors, replacing the original input variables. The general

145 form of the MARS model can be expressed as:

$$Y = \beta_0 + \sum_{j=1}^M \beta_j B_j(\theta_i) \tag{4}$$

where  $\theta_i$  (i = 1, ..., m) is the vector of the inputs;  $B_j$  is the *j*-th basis function, which can be a single spline function or a product of two or more basis functions; and the coefficients  $\beta_j s$  are estimated by minimizing the sum of squared residuals (Shahsavani et al., 2010). In fact, the MARS regression model is constructed by fitting these basis functions to various intervals of the independent variables. The final model in MARS is developed through a forward-backward procedure: initially, an over-fitted model is constructed by considering all the variables in the forward step; subsequently, this model is pruned by sequentially eliminating variables in the backward step, thereby creating a new model G. The performance of each model G is evaluated using generalized cross-validation (GCV):





(6)

154	$GCV(G) = \frac{1}{N} \frac{\sum_{i=1}^{N} (O_i - Y_i)^2}{\left[1 - \frac{C(G)}{N}\right]^2}$	(5)
155		(1 1 C(C)) (1

where N is the number of samplers;  $O_i$  and  $Y_i$  are the i-th observed and estimated values, respectively; C(G) is the 155 156 number of effective parameters, and is equal to 1 + c(G)d; d is the effective degrees of freedom; and c(G) is a penalty 157 for adding a basis function. The increase in GCV values between the pruned and over-fitted models is employed as a 158 metric to gauge the importance of the eliminated variables: a larger increase in GCV values signifies greater 159 importance of the removed variable (i.e., a sensitive parameter).

160 (c) The Morris method: a gradient-based sensitivity analysis technique using an individually randomized Morris onefactor-at-a-time (MOAT) design (Campolongo et al., 2007; Morris, 1991). It involves calculating multiple incremental 161 162 ratios, termed elementary effects, for each input variable (parameter) and averaging these effects to assess the overall 163 importance of the input variables. Campolongo et al. (2007) introduced a refined version of the elementary effects

method. In this approach, the model parameters  $\theta_i$  (i = 1, ..., m) are assumed to vary across p specified levels within 164

165 the input factor space, creating an experimental region  $\Omega$  that constitutes an m-dimensional p-level grid. For a given input  $\theta^0 = (\theta_1, \theta_2, ..., \theta_m)$ , the elementary effect of variable  $\theta_i$  is defined as: 166

167 
$$d_j = \frac{f(\theta_1, \dots, \theta_j + \Delta, \dots, \theta_m) - f(\theta_1, \dots, \theta_j, \dots, \theta_m)}{\Lambda}$$

where  $\Delta$  is a value in 1/p - 1, ..., p - 2/p - 1. The sampling strategy entails randomly determining the starting point 168 of each trajectory and perturbing each input variable by either  $+\Delta$  or  $-\Delta$  in random order. At the end of the process, a 169 170 trajectory spanning m+1 points is evaluated to compute the elementary effects for all m input variables. The mean  $(\mu_i)$ 171 and standard deviation ( $\sigma_i$ ) of the elementary effects ( $d_i$ ) serve as indicators of the sensitivity of the input variable  $\theta_i$ : 172

$$\mu_j = \sum_{i=1}^r \left| d_j(i) \right| / r \tag{7}$$

173 
$$\sigma_j = \sqrt{\sum_{i=1}^r \left( d_j(i) - \frac{\sum_{i=1}^r d_j(i)}{r} \right)^2 / r}$$
(8)

174 where  $\mu_i$  assesses the overall influence of  $\theta_i$  on the output, while  $\sigma_i$  estimates the higher-order effects (i.e., effects 175 due to interactions) of  $\theta_i$ .

176 (d) The Sobol' method: a quantitative sensitivity analysis approach based on variance decomposition (Sobol', 1993).

177 It decomposes the total variance of outputs Y into a summation of incrementally dimensional terms:

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 $V = \sum_{i=1}^{m} V_i + \sum_{i=1}^{m-1} \sum_{i=i+1}^{m} (V_{ij} + \dots + V_{1,\dots,m})$ 178 (9) 179 where m is the number of input variables (parameters),  $V_i$  represents the part of the output variance attributable to the individual input parameter  $\theta_i$  (i = 1, ..., m) (first-order sensitivity),  $V_{ii}$  represents the part of the output variance 180 181 resulting from the interaction between input variables ( $\theta_i$  and  $\theta_i$ ,  $i \neq j$ ) (second-order sensitivity), and  $V_1 \dots M_j$ 

182 represents the part of the output variance which can be explained by the interaction of all the variables. In this paper, the total effect of  $\theta_i$  is utilized as the metric for assessing its sensitivity, computed by: 183

$$S_i = 1 - \frac{v_{-i}}{v} \tag{10}$$

185 where  $V_{-i}$  represents the variance computed excluding variable  $\theta_i$ . The Sobol' method rigorously quantifies the

relative contribution of the individual parameters and their interactive effects on the total variance in the model output, 186

187 necessitating an extensive sample dataset (104 to 105 or more).





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189 The specific processes of the sensitivity analysis for the three targets (LE/NEE/LE+NEE) are detailed as follows: (a) 190 Sample generation: an ensemble of 400 samples within the prior range of 40 parameters was generated for the DT and 191 MARS methods using the Latin Hypercube (LH) sampling method (Deutsch et al., 2012). For the MOAT method, a 192 distinct ensemble of 410 (10 multiples of n+1, where n is the number of parameters) was created, utilizing the Monte 193 Carlo (MC) sampling approach (Hastings, 1970). (b) Cost function calculation: Both the 400 and 410 ensembles were 194 used to drive the CoLM, followed by computing the cost function values to measure the discrepancy between the 195 simulations and observations. Given the varying magnitudes of the target variables (LE/NEE), we employed the normalized root-mean-square error (NRMSE) as the cost function: 196

197 
$$NRMSE = \frac{\sqrt{\sum_{t=1}^{T} S(t) - O(t))^2}}{\sum_{t=1}^{T} O(t)}$$
(11)

where *T* is the total number of simulations; and S(t) and O(t) are the simulated and observed values of the target variables, respectively. For multi-objective sensitivity analysis (*LE*+*NEE*), a weighting function-based method was utilized to transform the multiple objectives into a single objective:

 $F = \sum_{i=1}^{m} NRMSE_i$ (12)

where *i* is the index of the target variables, and *m* denotes the number of objectives. (c) Sensitivity analysis: The input 202 203  $(\theta)$ -output (F) sample pairs were analyzed using four qualitative methods to discern the sensitivity of each parameter 204 to the target outputs. (d) Parameter selection: The mean sensitivity for each parameter was calculated across the four 205 qualitative analysis outcomes, leading to the selection of 10 parameters with the highest sensitivity for further 206 quantitative analysis. (e) Sample regeneration and cost function re-calculation: A new set of 100,000 samples for the 207 10 parameters was generated via LH sampling, followed by a repetition of step 2 to calculate the corresponding cost 208 function value. (f) Parameter determination for optimization: The total sensitivity results, derived from the Sobol' 209 method applied to the 100,000 input-output sample pairs, guided the determination of parameters for the subsequent 210 optimization. All the sensitivity analyses were conducted using the Problem Solving Environment for Uncertainty 211 Analysis and Design Exploration (PSUADE) software package (Tong, 2005) at the Supercomputing Center of 212 Lanzhou University in China.

## 213 2.3 Parameter Optimization with the Revised PEM-SMC Algorithm

214 Within the Bayesian single-objective optimization framework, the parameters are conceptualized as probabilistic 215 variables, with the posterior parameter distribution formulated as:

216

$$p(\theta \mid D) \propto p(\theta)p(D \mid \theta)$$
 (13)

where  $D = \{O_{1:T}\}$  is the set of observations of the target variable; *T* is the total number of observed data;  $\theta$  represents the parameters;  $p(\theta)$  and  $p(\theta | D)$  respectively denote the prior and posterior distributions of the parameters; and  $p(D | \theta)$  represents the model likelihood, which can be expressed as follows (Zhu et al., 2014):

220 
$$p(D \mid \theta) = (2\pi\sigma^2)^{-T/2} \prod_{t=1}^T \exp\left\{-\frac{[o(t)-S(X_t;\theta)]^2}{2\sigma^2}\right\}$$
(14)

221 where O(t) and  $S(X_t; \theta)$  denote the observed and simulated sequences of the target variable at each time step

222 t(t=1,2,...,T), respectively. The latter is driven by the forcing data  $X_t$  and parameters  $\theta$ .  $\pi$  is a mathematical constant,





223 and  $\sigma$  denotes the standard deviation of the measurement error, which can be estimated using the analytical method 224 (Braswell et al., 2005):

225

 $\sigma = \sqrt{\frac{1}{T} \sum_{t=1}^{T} [\mathcal{O}(t) - \mathcal{S}(X_t; \theta)]^2}$ (15)

For multi-objective optimization, the posterior parameter distributions are expressed as the product of the prior distribution and multiple likelihood functions, paralleling the approach in the weighting function-based methods:

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 $p(\theta \mid D) \propto p(\theta) \prod_{i=1}^{m} \prod_{t=1}^{T_i} p(O_{ti}^i \mid \theta)$ (16)

where *m* is the number of objective variables, and D=  $\{O_{1:T_1}^1, O_{1:T_2}^2, ..., O_{1:T_m}^m\}$  denotes the observation sets of the ith(i=1,2,..., m) objective.  $p(O_t^i | \theta)$  represents the model likelihood for objective *i* and is calculated by Equations 14 and 15.

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Owing to the unfeasibility of deriving a direct analytical solution for the integral in Equations 14 and 16, the SMC sampler is utilized to generate a sequence of weighted particles, thereby approximating the posterior distribution of parameters  $p(\theta | D)$ . However, since it is difficult to sample directly from  $p(\theta | D)$ , the SMC sampler alternatively samples from a sequence of intermediary distributions  $\pi_s(\theta)$  constructed by the geometric bridge method (Del Moral et al., 2006):

$$\pi_{s}(\theta) \propto p_{0}(\theta)^{1-\beta_{s}} p(\theta \mid D)^{\beta_{s}}$$
(17)

where  $p_0(\theta)$  and  $\pi_s(\theta)$  denote the initial and the s-th distribution in the sequence (s=0, 1, ..., S), respectively.  $\beta_s$  is a sequence of scalar powers, such that  $0 \le \beta_0 \le \beta_1 \le \dots \le \beta_s = 1$ , which allows a gradual transition of  $\pi_s(\theta)$  from the initial distribution  $\pi_0(\theta) \propto p_0(\theta)$  when  $\beta_0 = 0$  to the posterior distribution  $\pi_s(\theta) \propto p(\theta \mid D)$  when  $\beta_s = 1$ . Following Jeremiah et al. (2011, 2012), an exponential ( $\beta_s$ ) sequence is used in the PEM-SMC method.

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244 Upon determining the number of particles (Np) and the number of evolutions (S), the SMC sampler employs a series 245 of steps—reweighting, resampling, and moving—to transition particles from distribution  $\pi_{s-1}(\theta)$  to  $\pi_s(\theta)$ . In the 246 reweighting step, particles more closely aligned with the posterior distribution  $\pi_s(\theta)$  are assigned greater weights  $w_i^s$ , 247 enhancing their influence in the samplers. The subsequent resampling step addresses the issue of less significant 248 particles, termed "bad" particles. These are replaced with exact replicas of more promising particles through the 249 systematic resampling method. This step plays a pivotal role in ensuring the algorithm's convergence, facilitating a 250 gradual transition of particles from the prior to the posterior distribution. However, a notable challenge arises from 251 the resampling step: the potential reduction in particle diversity, leading to insufficient exploration of the parameter 252 space. To mitigate this, in our previous study (Zhu et al., 2018), we introduced a new candidate-generating method 253 named particle evolution Metropolis (PEM). This method integrates genetic algorithm features-crossover and 254 mutation operators-into the M-H algorithm framework. In the crossover operator, each parental chromosome pair 255  $\theta_i^s$  and  $\theta_i^s$  ( $i \neq j, i = 1, 2, ..., N/2, s = 1, 2, ..., S$ ) is selected to create a new offspring pair  $\overline{\theta}_i^s$  and  $\overline{\theta}_i^s$  using the onepoint crossover operator. The new offspring pair is accepted with probability min  $\left\{1, \frac{\pi_s(\bar{\theta}_s^S)\pi_s(\bar{\theta}_s^S)}{\pi_s(\theta_s^S)\pi_s(\theta_s^S)}\right\}$  according to the 256





(18)

- 257 M-H rule; otherwise, the current parental pair remains unchanged. In the mutation operator, each chromosome  $\theta_i^s(i = 0)$
- 1,2,..., N) is used to create a new chromosome  $\bar{\theta}_i^s$  according to the differential evolutionary algorithm:
- 259
- 260 where  $r_1$  and  $r_2$  are integer values without replacement from  $\{1, ..., j 1, j + 1, ..., N\}; \gamma = 2.38/\sqrt{2d}$  denotes the

 $\bar{\theta}_i^s = \theta_i^s + \gamma \left( \theta_{r_1}^s - \theta_{r_2}^s \right) + \zeta_d$ 

- jump rate; and  $\zeta_d \sim N_d(0, b^*)$  is drawn from a normal density with a small standard deviation, say  $b^* = 10^{-6}$ . The
- 262 new chromosome  $\bar{\theta}_i^s$  is accepted with the probability min  $\left\{1, \frac{\pi_s(\hat{\theta}_i^s)}{\pi_s(\hat{\theta}_i^s)}\right\}$ .
- 263

264 While the integration of crossover and mutation operators in the SMC algorithm enriches the particle diversity, it also 265 imposes significant computational burdens. To improve the algorithm's efficiency, we made pivotal modifications to the original PEM-SMC algorithm, which include: (a) Elimination of the crossover operator: In the revised algorithm, 266 267 the moving step exclusively employs the mutation operator to generate new particles, abandoning the crossover 268 operator. This adjustment addresses the inefficiency inherent in the crossover operator, which recombines parameters without value alteration. This often results in the time-intensive generation of particles with low acceptance probability 269 270 under the M-H rule and a risk of encountering the unexplained hyperparameter problem. (b) Modification of the mutation operator execution conditions: Unlike the original PEM-SMC algorithm, where the mutation is conditional 271 upon the effective sample size  $N_{eff} = 1/\sum_{i=1}^{N} (w_i^s)^2$  being less than half the number of particles, the mutation operator 272 273 is now uniformly employed at every evolutionary step for each particle. This strategic alteration promotes greater 274 particle diversity and prevents stagnation at local optima, albeit at the cost of some computational efficiency. (c) 275 Runtime reduction: Through comparative analysis and validation, including a synthetic five-dimensional bimodal 276 normal distribution and a benchmark experiment on the CoLM, it is evident that the revised PEM-SMC algorithm maintains its optimization efficacy while significantly reducing the runtime by over 40% (see Sect. S1). This time 277 278 reduction is crucial, especially for parameter optimization in complex models with lengthy single-run durations. 279 Overall, these modifications substantially enhance the PEM-SMC algorithm, striking a balance between efficiency 280 and a thorough exploration of the parameter space. The pseudo-code of the revised PEM-SMC algorithm is detailed below: 281

282 STEP 1: Initialization

(a) Draw an initial population 
$$\{\theta_j^0\}(j = 1, 2, ..., N)$$
 from the prior distribution  $p_0(\theta)$ , and set weights  $w_i^0 = 1/N$ .

- (b) Determine the exponential sequence  $0 \le \beta_0 \le \beta_1 \le \dots \le \beta_S = 1$ .
- 285 FOR  $s \leftarrow 1, 2, \dots, S$  DO (stage evolution)
- 286 STEP 2: Reweighting

(a) Set  $\theta_j^s = \theta_j^{s-1}$  (j = 1, ..., N), and calculate the weight  $w_j^s = w_j^{s-1} \frac{\pi_s(\theta_j^s)}{\pi_{s-1}(\theta_s^{s-1})}$ 

- (b) Normalize the weight so that  $\sum_{i=1}^{N} w_i^S = 1$ .
- 289 STEP 3: Resampling





290	(a) Calculate the effective sample size $N_{eff}$
291	(b) if $N_{eff} < N/2$ , resample from $\{\theta_i^s, w_i^s\}(j = 1, 2,, N)$ based on the systematic resample procedure and
292	set $w_j^s = 1/N(j = 1, 2,, N)$ ; otherwise, go to the next step.
293	STEP 4: Mutation
294	FOR $j \leftarrow 1,2, \dots, N$ DO (mutation operator)
295	(a) For each chromosome $\theta_j^s$ , create a new chromosome $\bar{\theta}_j^s$ by
296	(b) with probability min $\left\{1, \frac{\pi_s(\theta_i^s)\pi_s(\theta_j^s)}{\pi_s(\theta_i^s)\pi_s(\theta_j^s)}\right\}$ , set $\theta_i^s = \bar{\theta}_i^s$ and $\theta_j^s = \bar{\theta}_j^s$ , else leave $\theta_i^s$ and $\theta_j^s$ unchanged.
297	END FOR (mutation operator)
298	END FOR (stage evolution)
299	
300	After determining the parameters ( $\theta$ ) used for optimization through the sensitivity analysis, we employed the revised
301	PEM-SMC algorithm to automatically calibrate the selected parameters. The two control variables in the PEM-SMC
302	sampler, i.e., the number of particles in the population $Np$ and the number of evolutions $S$ , were set to 200 and 100,
303	respectively. The parameters of the CoLM were optimized by the observed LE and NEE separately and simultaneously.
304	For ease of description, the single-objective and multi-objective simultaneous optimizations constrained by the LE
305	and NEE fluxes are denoted as Opt_LE, Opt_NEE, and Opt_ALL, respectively. The revised PEM-SMC algorithm,
306	written in MATLAB, was deployed for the parameter optimization of the CoLM at the Supercomputing Center of

307 Lanzhou University.

## 308 2.4 Study Site and Model Performance Evaluation

Encompassing approximately one-third of the Earth's forests, the boreal forest ecosystem constitutes the most 309 310 substantial terrestrial biomass reservoir, significantly influencing global climate regulation. The precise modeling of 311 the intricate hydrocarbon dynamics within these ecosystems is pivotal for advancing our understanding of global terrestrial carbon storage and climate dynamics (Pan et al., 2011; Bradshaw et al., 2015). In this study, we focused on 312 313 RU-FY2, which is a typical evergreen needle-leaf forest (ENF) observation station, strategically situated in the Central 314 Forest Reserve of the Tver region of Russia. Positioned at 32°54'E, 56°27'N, this site experiences a warm humid 315 continental climate with an average annual temperature of 4.39 °C and precipitation of 668.53 mm. The predominant vegetation is dry spruce. Data for this site, encompassing conventional meteorological and eddy covariance 316 317 measurements, was sourced from the FLUXNET community. This meteorological dataset, spanning from 2015 to 318 2020 and characterized by high quality with no missing entries, includes half-hourly variables such as downward 319 short-wave and long-wave radiation, precipitation, specific humidity, temperature, atmospheric pressure, and wind 320 speed in eastward and northward directions. The multi-year average energy closure rate of the flux observations from 321 2015 to 2020 was near 100% (Fig. S6), signifying the exceptional quality of the energy flux observations. For NEE, 322 we selected the mean values of two variables- NEE\_CUT\_MEAN and NEE\_VUT\_MEAN-as the observational





metrics. Given the minimal nature of water and carbon fluxes in the non-growing season, coupled with their uncertain influence on parameter optimization effectiveness, our analysis was exclusively concentrated on the modeling of these fluxes during the growing season. The meteorological driver data from 1 June 2015 to 31 August 2019 were repeated 10 times to spin up the CoLM, while model simulations from 1 June 2020 to 31 August 2020 with a half-hour time step were used for the model parameter sensitivity evaluation and optimization.

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332

329 We quantified the difference between the simulated and observed target variables (*LE* and *NEE*) in the different

- 330 optimization scenarios using the root-mean-square error (RMSE), the Nash-Sutcliffe efficiency coefficient (NSE), and
- 331 the Pearson's correlation coefficient (R):

$$RMSE = \sqrt{\frac{1}{T} \sum_{t=1}^{T} (S(t) - O(t))^2}$$
(19)

333 
$$NSE = 1 - \frac{\sum_{t=1}^{T} (S(t) - O(t))^2}{\sum_{t=1}^{T} [O(t) - \bar{O}(t)]^2}$$
(20)

where O(t) and S(t) are respectively the observed and simulated values for each simulation point t(t=1,2,...,T), and  $\overline{O}(t)$  is the mean of the observed data. The NSE serves to quantitatively assess the precision of the model outputs in relation to the observed data, with a range from  $-\infty$  to 1. The closer the NSE is to 1, the more accurate the simulation is.

## 338 3 Results

## 339 3.1 Parameter Sensitivity

340 The results of the qualitative parameter sensitivity analysis among the different methods are shown in Fig. 1. For every 341 target variable, the different sensitivity analysis methods obtained similar results in screening out the sensitive 342 parameters, although there are some discrepancies in the sensitivity scores. For example, in the case of Opt\_LE, all 343 the methods identified the same three most sensitive parameters-P33, P34, and P35-while the MOAT method 344 screened out more moderately sensitive parameters (e.g., P2, P7, P13, P18, etc.). Therefore, it is most reasonable to identify the sensitive parameters according to the sum (or mean) of the sensitivity values obtained by the different 345 346 qualitative analysis methods. In the three optimization scenarios, there are some common most sensitive parameters 347 (e.g., P33, P34, P35, P3, etc.) and some individually sensitive parameters. For example, P8, P7, P36, and P13 are sensitive to Opt\_LE, while they are not the sensitive parameters of Opt\_NEE and Opt\_ALL. Meanwhile, we can see 348 349 that the sensitive parameters of Opt\_NEE and Opt\_ALL are essentially the same, which could indicate that the NEE 350 observations are more restrictive on the parameters than the LE observations at this site. Based on the results of the qualitative analysis, we selected the 10 parameters with the highest sensitivity scores for each target variable to 351 perform the subsequent Sobol' quantitative sensitivity analysis: Opt\_LE (P34, P33, P35, P8, P3, P7, P36, P18, P13, 352 353 P32); Opt\_NEE (P34, P33, P3, P35, P9, P5, P30, P29, P31, P18); and Opt\_ALL (P34, P33, P3, P40, P5, P9, P35, P37, 354 P29, P6).







Paremeter Index



359

360 The screening results obtained by the four qualitative methods indicated that P34 and P33 are the most sensitive parameters for the three target variables, which is consistent with the Sobol' quantitative analysis results, as shown in 361 362 Fig. 2. The Sobol' results showed that P34 and P33 can explain 65%, 93%, and 77% of the total variance between the 363 simulated and observed values of the three target variables, respectively. Based on the principle that the cumulative relative importance of the parameters is greater than 95% (which means that the variance can essentially be explained 364 365 by these parameters), we selected the most sensitive parameters for each target variable to perform the subsequent optimization: Opt\_LE (P34, P33, P35, P8, P36, P3); Opt\_NEE (P34, P33, P35, P30, P3); and Opt\_ALL (P34, P33, P35, 366 P9, P5, P37). Since the selected parameters can explain more than 95% of the total variance of the model output, we 367





- 368 believe that taking them as the optimized parameters instead of all 40 parameters to calibrate the model can improve
- 369 the optimization efficiency without losing effectiveness.



370 371

Figure 2 The relative importance of the top 10 parameters to the target variables obtained by Sobol' sensitivity analysis. The percentage of each parameter is the ratio between the total order Sobol' index of  $\theta_i$  and the sum value of all the parameters  $\theta$ . Each slice of the pie chart indicates the extent to which the changes in parameter  $\theta_i$  can explain the total variance of the model outputs.

## 375 **3.2 Comparison of the Parameter Optimization Results**

376 The particle transitions, evolution, and optimized results of specific parameters in the three optimization scenarios constrained by LE and NEE observations are shown in Table 1 and Fig. S7 to S9. Compared to the default values, the 377 378 optimized values of almost all the parameters have changed significantly (especially the two most sensitive parameters P34 and P33), while very few parameters have remained unchanged (e.g., P36 and P30). At the same time, the posterior 379 380 distributions of the common parameters (i.e., P34, P33, and P35) in the three optimization scenarios are quite different (Fig. 3). For example, the optimal solution of P34 toward two extremes in Opt\_LE (195.79) and Opt\_NEE (58.86) in 381 the prior range ([10,200]). Based on the total error minimization principle, the multi-objective simultaneous 382 383 optimization algorithm makes trade-offs in the simulation performance of the two target variables and calibrates 384 parameter P34 to an intermediate value (82.19). The optimization of parameter P33 is similar (Opt LE: 0.0761, 385 Opt\_NEE: 0.0702, Opt\_ALL: 0.0735). However, this does not mean that the multi-objective optimization values of all the parameters will be between the two single-objective optimization values. For example, the optimized value of P35 386 in Opt\_ALL (8.87) is greater than that in Opt\_LE (6.62) and Opt\_NEE (8.52). This is because the simulation 387 388 performance of the model depends not only on the value of a single parameter but also on the combination effect 389 between parameters. Therefore, the calibrated values of the model parameters must be derived from the simultaneous 390 multi-objective optimization rather than from single-objective optimization or the comparison of the optimized values 391 of multiple single-objective optimizations.

392

Table 1. The default and optimized values of the most sensitive parameters in the three optimization scenarios: *Opt\_LE*,
 *Opt\_NEE*, and *Opt\_ALL*. The optimized results are the median values of the posterior distributions obtained by the PEM SMC algorithm.

Opt_LE			Opt_NEE			Opt_ALL		
Para.	Default	Optimized	Para.	Default	Optimized	Para.	Default	Optimized
P34	100	195.79	P34	100	58.86	P34	100	82.19
P33	0.08	0.0761	P33	0.08	0.0702	P33	0.08	0.0735
P35	9	6.62	P35	9	8.52	P35	9	8.87
P8	131.88	472.01	P30	0.3	0.3	P9	207.34	116.62
P36	0.01	0.01	P3	0.43	0.35	P5	5.77	4.56
P3	0.4348	0.35				P37	0.5	0.67









<sup>399</sup> **3.3 Comparison of the Optimization Effectiveness** 

Three statistical metrics are used here to characterize the global effectiveness of the three optimized parameter combinations in improving the performance of the CoLM in simulating *LE* and *NEE* (Fig. 4a–c). In addition, we present the observations and simulations of *LE* and *NEE* under the different parameter combinations to show the characterization of detailed changes in the two target variables in Fig. 4d–e.

By comparing *Control* and *Opt\_LE*, it can be found that the performance of the CoLM in simulating LE can be significantly improved by applying the optimized parameters of *Opt\_LE*, as evidenced by the RMSE decreasing by 18.05 W/m2 and the NSE and R increasing by 0.18 and 0.07, respectively. The default parameter set for the CoLM significantly underestimates LE, while the *Opt\_LE* optimized parameter set significantly reduces the difference





- between the simulated and observed *LE* (Fig. 4d). Meanwhile, we note that the simulation performance of the CoLM for *NEE* was sacrificed by taking the optimized values of *Opt\_LE*. From the statistical indicators, the R between the simulated and observed *NEE* slightly increases by 0.09, but the RMSE increases by 6.94  $\mu$ mol m-2 s-1 and the NSE decreases to -1.09. From the values, the simulated NEE in the daytime (the negative value in Fig. 4e) under *Opt\_LE* is nearly three times lower than the observed values.
- 414
- 415 By comparing Control and Opt\_NEE, it can be seen that the simulation performance for NEE does not improve
- 416 significantly after the single-objective optimization of *NEE*, as evidenced by the RMSE decreasing from 7.15 μmol m-
- 417 2 s-1 to 6.91 μmol m-2 s-1 and the NSE and R slightly increasing by 0.03 and 0.03, respectively. The simulated *NEE*
- under the *Opt\_NEE* parameter combination is essentially the same as that under *Control*, suggesting that the simulation
- 419 performance of the model for NEE cannot be improved by parameter calibration alone. Unlike Opt\_LE, which
- 420 improves *LE* at the expense of the *NEE* simulation, *Opt\_NEE*'s optimized parameter combination provides a slight
- 421 improvement in *LE* simulation performance (the RMSE is decreased by 1.22 and the NSE and R increase by 0.01).
- 422
- 423 Compared to single-objective optimization (Opt\_LE and Opt\_NEE), multi-objective optimization (Opt\_ALL) can
- 424 simultaneously take into account the enhancement of the simulation performance for multiple variables. Compared to
- 425 Control, the Opt\_ALL optimized parameters decrease the RMSE of LE and NEE by 7.2 W/m2 and 0.19 µmol m-2 s-
- 426 1, respectively, and increase the NSE of LE and NEE by 0.07 and 0.02, respectively (Fig. 4). Although the optimized
- 427 parameters of Opt\_ALL are not as good as those of Opt\_LE in improving the underestimated LE, it does not lose the
- 428 simulation accuracy for NEE. Comparing Opt\_NEE with Opt\_ALL, although they both improve the simulation
- 429 performance for LE and NEE, the latter's simulation accuracy for LE is significantly higher than that of the former.
- 430 In summary, multi-objective simultaneous optimization can improve the simulation performance for specific variables
- 431 without compromising the simulation accuracy of the other objective variables.







432 433

435Figure 4 Comparison of the performance of the CoLM in simulating LE and NEE under the four parameter schemes434(Control,  $Opt\_LE$ ,  $Opt\_NEE$ , and  $Opt\_ALL$ ) using three statistical metrics: (a) RMSE, (b) the Nash efficiency coefficient435(NSE), and (c) the Pearson's correlation coefficient (R). Control denotes the default parameter values, while  $Opt\_LE$ ,436 $Opt\_NEE$ , and  $Opt\_ALL$  represent the optimized values of the parameters under single-objective optimization of LE, single-objective optimization of NEE, and simultaneous multi-objective optimization of LE, respectively. The differences438between the half-hourly observed and simulated (d) LE and (e) NEE in the four parameter schemes over a half-month439period are also displayed.

## 440 4. Discussion

## 441 4.1 The Advantages of the PEM-SMC Algorithm

442 This paper has introduced an enhanced PEM-SMC algorithm anchored in the Bayesian framework, which integrates

443 prior distributions with observational data to closely approximate the posterior distributions of the parameters. Facing

the challenge of the posterior distribution's analytical intractability, the proposed approach adopts an SMC method,





445 representing parameter distributions with a sequence of random particles that are iteratively updated through strategies 446 such as resampling to achieve a closer alignment with the actual posterior distribution. To mitigate the reduction in 447 particle diversity that can result from SMC resampling, a differential evolution mutation operator is introduced, aimed 448 at boosting the search efficiency within the parameter space. Consequently, this revised PEM-SMC algorithm not only 449 maintains the particle diversity but also optimizes the computational efficiency, surpassing its predecessor (Zhu et al., 450 2018). Moreover, its capacity to yield a comprehensive posterior distribution, as opposed to the mere point estimations 451 offered by the deterministic parameter estimation techniques, significantly bolsters the robustness of the parameter 452 estimation (Thiemann et al., 2001; Jeremiah et al., 2011). This feature renders it particularly advantageous for total 453 uncertainty assessment, covering model parameters, inputs, and structural uncertainties.

454

455 Furthermore, the revised PEM-SMC algorithm effectively integrates the analytical power of Bayesian theory with the 456 flexibility of the SMC framework, demonstrating significant potential for practical applicability and structural 457 scalability. By applying Bayesian theory, it incorporates information from multiple sources, including prior knowledge 458 and observational data, into a unified analytical framework and expresses this information in a rigorous mathematical 459 form (Equations 14 and 16). Compared to the traditional metaheuristic multi-objective optimization algorithms (Deb 460 et al., 2002; Mirjalili et al., 2016; Xue et al., 2012), this framework solves multi-objective optimization problems more 461 directly through the joint probability distribution, avoiding the complexities of balancing multiple objectives via cost functions. In addition, the PEM-SMC algorithm merges the strengths of the SMC and MCMC methods into a flexible 462 framework for structural extension (Speich et al., 2021), facilitating the design of adaptive transition kernels, effective 463 464 particle diversity enhancement strategies, and efficient intermediary proposal distributions.

#### 465 **4.2 Differences Between Single- and Multi-Objective Optimization**

The distinction between the single- and multi-objective optimization strategies is first manifested in the parameter 466 467 estimation. This research has revealed a significant phenomenon: a single parameter can necessitate different optimal 468 values depending on the target variable. For instance, in the Opt\_LE optimization scenario, the optimal value of 469 parameter P34 is nearly double its default setting, whereas in the Opt\_NEE scenario, it is halved (Table 1). This discrepancy arises from P34's (i.e., the maximum rate of carboxylation at 25 °C, Vmax25) dual role in the leaf stomatal 470 471 photosynthesis-condunctance model, where it regulates both the rate of leaf photosynthesis and stomatal conductance 472 (see Sect. S3). To better align the simulation outcomes with actual observations, this parameter requires different 473 optimized values for the two target variables (i.e., LE and NEE). Specifically, as the simulated LE under the default 474 parameters falls short of the observed value, the optimization algorithm modifies P34 from 100 to 195.79, thereby 475 enhancing the leaf photosynthesis and reducing the stomatal resistance. Conversely, to minimize the discrepancy 476 between the simulated and observed NEE, the value of P34 is required to be halved, leading to a conflict in P34's 477 optimal values for both the LE and NEE simulations. Therefore, multi-objective optimization emerges as a key strategy 478 to balance the optimization performance of disparate target variables. In practice, the value of P34 is adjusted to 82.19, 479 situated between the Opt\_LE and Opt\_NEE values, rendering this balanced value more congruent with the default 480 empirical value of 100 and physically sound. Gong et al. (2015) also reported similar findings, where the "bsw"





481 parameter in the CoLM exhibited high optimal values under the constraints of observed sensible heat, latent heat, and 482 soil moisture, but showed low optimal values when constrained by upward long-wave radiation, net radiation, and soil 483 temperature observations. In fact, due to the complex interactions of the processes within LSMs, there are many such 484 "contradictory" parameters that can simultaneously affect multiple output variables. Although these sensitive 485 parameters can vary across different LSMs, ecological processes, or ecosystems, a common phenomenon is evident: 486 to maximize the reduction of discrepancies between the simulated values and observations across multiple target 487 variables, these parameters often demonstrate conflicting optimal values.

488

489 Furthermore, the difference between single- and multi-objective optimization is particularly evident in enhancing the simulation performance across various output variables. While single-objective optimization, such as targeting solely 490 491 on LE, can improve the accuracy for that specific variable, it can adversely affect the simulation performance for other 492 variables, such as NEE. This highlights the complex interactions of the numerous processes (such as radiative transfer, 493 energy exchange, water transition, carbon cycling, etc.) in LSMs, where the impact of optimizing a single output on 494 the others is unpredictable. In contrast, multi-objective optimization has been proven to be more effective in improving 495 model performance across multiple outputs, as evidenced by the simultaneous improvement of LE and NEE in the 496 Opt\_ALL scenario. Therefore, comprehensively integrating the various available observational data for multiobjective optimization is preferable for parameter calibration in complex models. It is noteworthy that multi-objective 497 optimization may not achieve as high an accuracy for individual variables as single-objective optimization. For 498 499 instance, the discrepancy in the LE simulation of the Opt\_ALL scenario (Fig. 4a: RMSE = 86 W/m2) compared to the 500 Opt\_LE scenario (Fig. 4a: RMSE = 75 W/m2) indicates a trade-off. However, this trade-off is justified, as it results in 501 a more balanced and overall enhanced model performance at the expense of a slight sacrifice in simulation accuracy 502 improvement. If necessary, this limitation can be mitigated by adjusting the weights in the objective weighting method 503 to prioritize certain variables. In summary, the multi-objective optimization strategy is recommended for calibrating 504 complex models with multiple interrelated outputs, as it not only ensures that the optimized parameter values adhere 505 to objective physical constraints, but also balances the simulation performance among the multiple outputs.

#### 506 **4.3 Defects in the Model Structure**

507 It is imperative to acknowledge that the potential for enhancing a model's simulation accuracy through parameter 508 optimization critically hinges on the robustness of the model structure and the quality of the driving data (Duan et al., 509 2006). Our findings indicate that the imposition of constraints on the model parameters based on the NEE observations does not yield a significant increase in the simulation performance for NEE, particularly concerning the 510 511 underestimated nocturnal respiration (Figure 4). This limitation is not attributable to the deficiency of the optimization algorithm, but rather to the inadequate representation of the soil respiration processes within the model. In the CoLM, 512 soil respiration is quantified based on the exponential empirical equation (i.e.,  $R = R_{10} e^{E_0 \left(\frac{1}{283.15-T_0} - \frac{1}{T_0-T_0}\right)}$ , where R 513 514 and T are the soil respiration and soil temperature, respectively;  $R_{10}$  denotes the basal respiration at 10 °C; and  $E_o$  is 515 the active energy; Lloyd et al., 1994). Theoretically, this approach is potentially more appropriate for estimating soil 516 respiration over annual or more extended temporal scales (Raich et al., 2002; Chen et al., 2013). A substantial body





517 of literature corroborates the existence of a phase lag (hysteresis) between the temporal dynamics of soil temperature 518 and soil respiration at hourly and seasonal timescales (Tang et al., 2005; Liu et al., 2006; Riveros-Iregui et al., 2007; Ma et al., 2020). Such a lag leads to the empirical representation of soil respiration that diverges from the precise 519 modeling requirements of LSMs at hourly to daily intervals for carbon cycle simulation. This discrepancy highlights 520 521 the necessity for a more holistic consideration of the processes, encompassing soil heat and moisture dynamics, 522 microbial decomposition, and canopy photosynthesis (Hanson et al., 2000; Ryan et al., 2005; Davidson et al., 2006). 523 Therefore, the CoLM could be strengthened by integrating more detailed mechanistic process modules, exemplified 524 by the soil autotrophic and heterotrophic respiration modules featured in the Community Land Model (CLM4.5 and 525 5.0) (Lawrence et al., 2019), to substantially improve the accuracy and robustness of terrestrial carbon cycle simulation. Accordingly, the simple calibration of model parameters falls short of addressing the inherent structural inadequacies 526 527 in the CoLM's representation of the soil respiration process. Therefore, parameter calibration (optimization) should be viewed as a multifaceted tool-not only does it enhance the local-scale applicability of the model, but it also plays 528 529 a crucial role in uncovering the model's structural deficiencies and providing guidance for model refinement and 530 development.

## 531 4.4 Limitations and Future Work

532 This research has two principal limitations. Firstly, the PEM-SMC algorithm faces constraints in execution time, as each evolutionary iteration of the particle swarm necessitates running the original dynamic model to evaluate the 533 likelihood function values of the different parameter combinations. Consequently, the cumulative runtime of the PEM-534 SMC algorithm is quantified as  $3 \times N_p \times S \times T$  (where  $N_p$  represents the number of particles, S is the number of 535 536 evolutionary iterations, and T is the duration of a single dynamic model run), resulting in a computational demand 537 exceeding ten thousand operations. In response, our focus will be on refining the execution mode (parallel computing), 538 enhancing the sequential characteristics, and optimizing the resampling mechanisms, which are all aimed at strengthening the efficiency of the PEM-SMC algorithm for complex model parameter optimization. 539

540

541 Secondly, the proposed approach for multi-objective optimization essentially converts the different objective variables 542 into a single-objective framework through a goal-weighting strategy. However, despite equalizing the weights for two 543 variables (LE and NEE), the subjectivity and uncertainty in the weight allocation could potentially restrict the diversity 544 of the optimal solutions. Therefore, in the future, we will work on developing a Bayesian-inspired multi-objective 545 parameter estimation algorithm. This algorithm will synergize the autonomous optimization capabilities of the PEM-546 SMC algorithm with the non-dominant and diverse characteristics inherent in Pareto optimal solution theory, thereby 547 substantially augmenting the effectiveness and applicability of the PEM-SMC algorithm in complex multi-objective 548 optimization scenarios.





#### 549 5 Conclusions

In this study, we employed the revised PEM-SMC algorithm for single- and multi-objective optimization of the sensitive parameters in the water-carbon process of the CoLM and conducted a comparative analysis of different optimization strategies concerning parameter estimation, model performance enhancement, and applicability reliability. The key findings include:

554

555 Firstly, the revised PEM-SMC algorithm demonstrates a robust ability to tackle the multi-dimensional, multi-objective 556 parameter optimization challenge for complex dynamics models. Secondly, significant differences were observed 557 between single- and multi-objective optimization in parameter estimation. The optimization values of the three 558 sensitive parameters present conflicts after single-objective optimization, while the values after multi-objective 559 optimization appear more rational. Moreover, the multi-objective optimization demonstrated superiority over single-560 objective optimization in enhancing the simulation performance for the multiple output variables. Although single-561 objective optimization can improve the simulation performance for specific objectives, it can adversely affect the other target variables. For instance, optimizing for LE reduced the RMSE of the simulated and observed LE by 20%, but 562 563 increased the RMSE of NEE by 97%. Conversely, the multi-objective optimization concurrently improved the 564 simulation performance for both LE and NEE, evidenced by decreases in RMSE for LE and NEE of 7.2 W/m2 and 565 0.19 µmol m-2 s-1, respectively. Finally, through the parameter optimization process, we identified the structural deficiencies in the CoLM's soil respiration calculations. Consequently, we suggest that the CoLM modeling 566 community consider integrating more precise mechanistic process models to enhance the accuracy and robustness of 567 568 terrestrial carbon cycle simulation.

### 569 Code and Data Availability

570 The detailed information for RU-FY2 can be accessed via FLUXNET at https://fluxnet.org/sites/siteinfo/RU-Fy2. The 571 flux data employed in this research were obtained from FLUXNET2015, available at https://fluxnet.org/. The 572 sensitivity analysis software package, PSUADE version 1.7.8a, is accessible at 573 https://computing.llnl.gov/projects/psuade/software. The CoLM codes (2014 version) can be found through the Land-574 Atmosphere Interaction Research Group at Sun Yat-sen University, detailed at 575 http://globalchange.bnu.edu.cn/research/models. Comprehensive access to the data, code (inclusive of the CoLM 576 model program, sensitivity analysis software package, and the revised PEM-SMC program), and associated outcomes 577 from this study are available at https://doi.org/10.5281/zenodo.10900461.

#### 578 Author contributions

579 Under the guidance of GFZ, CX was responsible for conceptualization, algorithm development, data handling, and 580 drafting the initial manuscript. YZ and KZ revised the initial draft of the manuscript and provided improvement 581 suggestions. All authors discussed the results throughout the research period and approved the final version of the 582 manuscript for publication.





## 583 Competing interests

584 The contact author has declared that none of the authors has any competing interests.

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### 593 References

- 594 Bastidas, L. A., Gupta, H. V., Sorooshian, S., Shuttleworth, W. J., & Yang, Z. L.: Sensitivity analysis of a land surface
- scheme using multicriteria methods, J Geophys Res., 104(D16), 19481-19490, <u>https://doi.org/10.1029/1999JD900155</u>,
  1999.
- 597 Bastrikov, V., MacBean, N., Bacour, C., Santaren, D., Kuppel, S., & Peylin, P.: Land surface model parameter
- 598 optimisation using in situ flux data: Comparison of gradient-based versus random search algorithms (a case study

using ORCHIDEE v1. 9.5. 2), Geosci Model Dev., 11(12), 4739-4754, https://doi.org/10.5194/gmd-11-4739-2018,

- 600 2018.
- 601 Braswell, B. H., Sacks, W. J., Linder, E., & Schimel, D. S.: Estimating diurnal to annual ecosystem parameters by
- synthesis of a carbon flux model with eddy covariance net ecosystem exchange observations, Glob Chang Biol., 11(2),
- 603 335-355, <u>https://doi.org/10.1111/j.1365-2486.2005.00897.x</u>, 2005.
- 604 Campolongo, F., Cariboni, J., & Saltelli, A.: An effective screening design for sensitivity analysis of large models,
- 605 Environmental modelling & software, 22(10), 1509-1518. https://doi.org/10.1016/j.envsoft.2006.10.004, 2007.
- Chen, S., Huang, Y., Xie, W., Zou, J., Lu, Y., & Hu, Z.: A new estimate of global soil respiration from 1970 to 2008,
- 607 Sci Bull (Beijing), 58, 4153-4160, <u>https://doi.org/10.1007/s11434-013-5912-1</u>, 2013.
- Chopin, N.: A sequential particle filter method for static models, Biometrika., 89(3), 539-552,
  https://doi.org/10.1093/biomet/89.3.539, 2002.
- 610 Cowles, M. K., & Carlin, B. P.: Markov chain Monte Carlo convergence diagnostics: a comparative review, J Am
- 611 Stat Assoc., 91(434), 883-904, <u>https://doi.org/10.1080/01621459.1996.10476956</u>, 1996.
- 612 D'heygere, T., Goethals, P. L., & De Pauw, N.: Genetic algorithms for optimisation of predictive ecosystems models
- 613 based on decision trees and neural networks, Ecol Model., 195(1-2), 20-29,
- 614 <u>https://doi.org/10.1016/j.ecolmodel.2005.11.005</u>, 2006.
- 615 Dai, Y., Zeng, X., Dickinson, R. E., Baker, I., Bonan, G. B., Bosilovich, M. G., ... & Yang, Z. L.: The common land
- 616 model, Bull Am Meteorol Soc., 84(8), 1013-1024, https://doi.org/10.1175/BAMS-84-8-1013, 2003.





- 617 Davidson, E. A., & Janssens, I. A.: Temperature sensitivity of soil carbon decomposition and feedbacks to climate
- 618 change. Nature., 440(7081), 165-173, <u>https://doi.org/10.1038/nature04514</u>, 2006.
- 619 Deb, K., Pratap, A., Agarwal, S., & Meyarivan, T. A. M. T.: A fast and elitist multiobjective genetic algorithm: NSGA-
- 620 II, IEEE Trans Evol Comput., 6(2), 182-197, <u>https://doi.org/10.1109/4235.996017</u>, 2002.
- 621 Del Moral, P., Doucet, A., & Jasra, A.: Sequential Monte Carlo samplers, J R Stat Soc Series B Stat Methodol., 68(3),
- 622 411-436, <u>https://doi.org/10.1111/j.1467-9868.2006.00553.x</u>, 2006.
- 623 Deutsch, J. L., & Deutsch, C. V.: Latin hypercube sampling with multidimensional uniformity, J Stat Plan Inference.,
- 624 142(3), 763-772, <u>https://doi.org/10.1016/j.jspi.2011.09.016</u>, 2012.
- 625 Dong, J., Akbar, R., Feldman, A. F., Gianotti, D. S., & Entekhabi, D.: Land Surfaces at the Tipping-Point for Water
- and Energy Balance Coupling, Water Resour Res., 59(2), e2022WR032472, <u>https://doi.org/10.1029/2022WR032472</u>,
- 627 2023.
- 628 Doucet, A., De Freitas, N., & Gordon, N.: An introduction to sequential Monte Carlo methods, Sequential Monte
- 629 Carlo methods in practice, 3-14, <u>https://doi.org/10.1007/978-1-4757-3437-9\_1</u>, 2001.
- 630 Duan, Q. Y., Gupta, V. K., & Sorooshian, S.: Shuffled complex evolution approach for effective and efficient global
- 631 minimization, J Optim Theory Appl., 76, 501-521, <u>https://doi.org/10.1007/BF00939380</u>, 1993.
- 532 Duan, Q., Di, Z., Quan, J., Wang, C., Gong, W., Gan, Y., ... & Fan, S.: Automatic model calibration: A new way to
- 633 improve numerical weather forecasting, Bull Am Meteorol Soc., 98(5), 959-970, <u>https://doi.org/10.1175/BAMS-D-</u>
- 634 <u>15-00104.1</u>, 2017.
- 635 Duan, Q., Schaake, J., Andréassian, V., Franks, S., Goteti, G., Gupta, H. V., ... & Wood, E. F.: Model Parameter
- 636 Estimation Experiment (MOPEX): An overview of science strategy and major results from the second and third
- 637 workshops, J Hydrol (Amst), 320(1-2), 3-17, <u>https://doi.org/10.1016/j.jhydrol.2005.07.031</u>, 2006.
- 638 Duan, Q., Sorooshian, S., & Gupta, V. K.: Optimal use of the SCE-UA global optimization method for calibrating
- 639 watershed models, J Hydrol (Amst), 158(3-4), 265-284, <u>https://doi.org/10.1016/0022-1694(94)90057-4</u>, 1994.
- Eirola, E., Liiti änen, E., Lendasse, A., Corona, F., & Verleysen, M.: Using the Delta Test for Variable Selection, In
  ESANN (pp. 25-30), 2008.
- Fan, Y., Leslie, D. S., & Wand, M. P.: Generalised linear mixed model analysis via sequential Monte Carlo sampling,
   <a href="https://doi.org/10.1214/07-EJS158">https://doi.org/10.1214/07-EJS158</a>, 2008.
- 644 Field, C. B., & Raupach, M. R. (Eds.).: The global carbon cycle: integrating humans, climate, and the natural world
- 645 (Vol. 62), Island Press, 2004.
- 646 Folberth, C., Elliott, J., Müller, C., Balkovič, J., Chryssanthacopoulos, J., Izaurralde, R. C., ... & Wang, X.:
- 647 Parameterization-induced uncertainties and impacts of crop management harmonization in a global gridded crop
- 648 model ensemble, PLoS One., 14(9), e0221862, <u>https://doi.org/10.1371/journal.pone.0221862</u>, 2019.
- 649 Friedman, J. H.: Multivariate adaptive regression splines, Ann Stat., 19(1), 1-67,
   650 <u>https://doi.org/10.1214/aos/1176347963</u>, 1991.
- 651 Gill, M. K., Kaheil, Y. H., Khalil, A., McKee, M., & Bastidas, L.: Multiobjective particle swarm optimization for
- 652 parameter estimation in hydrology, Water Resour Res., 42(7), https://doi.org/10.1029/2005WR004528, 2006.





Gong, W., Duan, Q., Li, J., Wang, C., Di, Z., Dai, Y., ... & Miao, C.: Multi-objective parameter optimization of
common land model using adaptive surrogate modeling, Hydrol Earth Syst Sci., 19(5), 2409-2425,
https://doi.org/10.5194/hess-19-2409-2015, 2015.

- 656 Gupta, H. V., Bastidas, L. A., Sorooshian, S., Shuttleworth, W. J., & Yang, Z. L.: Parameter estimation of a land
- 657 surface scheme using multicriteria methods, J Geophys Res., 104(D16), 19491-19503,
   658 <u>https://doi.org/10.1029/1999JD900154</u>, 1999.
- 659 Hanson, P. J., Edwards, N. T., Garten, C. T., & Andrews, J. A.: Separating root and soil microbial contributions to
- 660 soil respiration: a review of methods and observations, Biogeochemistry, 48, 115-146, 661 https://doi.org/10.1023/A:1006244819642, 2000.
- Hastings, W. K.: Monte Carlo sampling methods using Markov chains and their applications,
   https://doi.org/10.1093/biomet/57.1.97, 1970.
- 664 Henderson-Sellers, A., Pitman, A. J., Love, P. K., Irannejad, P., & Chen, T. H.: The project for intercomparison of
- land surface parameterization schemes (PILPS): Phases 2 and 3, Bull Am Meteorol Soc., 76(4), 489-504,
   <a href="https://doi.org/10.1175/1520-0477(1995)076%3C0489:TPFIOL%3E2.0.CO;2">https://doi.org/10.1175/1520-0477(1995)076%3C0489:TPFIOL%3E2.0.CO;2</a>, 1995.
- Ines, A. V., & Mohanty, B. P.: Near-surface soil moisture assimilation for quantifying effective soil hydraulic
  properties using genetic algorithm: 1. Conceptual modeling, Water Resour Res., 44(6),
  https://doi.org/10.1029/2007WR005990, 2008.
- 670 Jeremiah, E., Sisson, S. A., Sharma, A., & Marshall, L.: Efficient hydrological model parameter optimization with
- Sequential Monte Carlo sampling, Environ Modell Softw., 38, 283-295, <u>https://doi.org/10.1016/j.envsoft.2012.07.001</u>,
  2012.
- 673 Jeremiah, E., Sisson, S., Marshall, L., Mehrotra, R., & Sharma, A.: Bayesian calibration and uncertainty analysis of
- 674 hydrological models: A comparison of adaptive Metropolis and sequential Monte Carlo samplers, Water Resour Res.,
- 675 47(7), <u>https://doi.org/10.1029/2010WR010217</u>, 2011.
- Ji, D., & Dai, Y.: The common land model (CoLM) technical guide, GCESS, Beijing Normal University, Beijing,
  China, 2010.
- 678 Kato, T., Knorr, W., Scholze, M., Veenendaal, E., Kaminski, T., Kattge, J., & Gobron, N.: Simultaneous assimilation
- 679 of satellite and eddy covariance data for improving terrestrial water and carbon simulations at a semi-arid woodland
- 680 site in Botswana, Biogeosciences, 10(2), 789-802, <u>https://doi.org/10.5194/bg-10-789-2013</u>, 2013.
- Lawrence, D. M., Fisher, R. A., Koven, C. D., Oleson, K. W., Swenson, S. C., Bonan, G., ... & Zeng, X.: The
- 682 Community Land Model version 5: Description of new features, benchmarking, and impact of forcing uncertainty, J
- 683 Adv Model Earth Syst., 11(12), 4245-4287, <u>https://doi.org/10.1029/2018MS001583</u>, 2019.
- Leplastrier, M., Pitman, A. J., Gupta, H., & Xia, Y.: Exploring the relationship between complexity and performance
- in a land surface model using the multicriteria method, J Geophys Res., 107(D20), ACL-11,
   https://doi.org/10.1029/2001JD000931, 2002.
- 687 Li, J., Chen, F., Lu, X., Gong, W., Zhang, G., & Gan, Y.: Quantifying contributions of uncertainties in physical
- parameterization schemes and model parameters to overall errors in Noah-MP dynamic vegetation modeling, J Adv
- 689 Model Earth Syst., 12(7), e2019MS001914, <u>https://doi.org/10.1029/2019MS001914</u>, 2020.





- 690 Li, J., Duan, Q. Y., Gong, W., Ye, A., Dai, Y., Miao, C., ... & Sun, Y.: Assessing parameter importance of the Common
- Land Model based on qualitative and quantitative sensitivity analysis, Hydrol Earth Syst Sci., 17(8), 3279-3293,
- 692 <u>https://doi.org/10.5194/hess-17-3279-2013</u>, 2013.
- Li, J., Duan, Q., Wang, Y. P., Gong, W., Gan, Y., & Wang, C.: Parameter optimization for carbon and water fluxes in
- two global land surface models based on surrogate modelling, Int J Climatol., 38, e1016-e1031,
   https://doi.org/10.1002/joc.5428, 2018.
- Liu, Q., Edwards, N. T., Post, W. M., Gu, L., Ledford, J., & Lenhart, S.: Temperature-independent diel variation in
- soil respiration observed from a temperate deciduous forest, Glob Chang Biol., 12(11), 2136-2145,
   <u>https://doi.org/10.1111/j.1365-2486.2006.01245.x</u>, 2006.
- Liu, Y., Gupta, H. V., Sorooshian, S., Bastidas, L. A., & Shuttleworth, W. J.: Constraining land surface and
   atmospheric parameters of a locally coupled model using observational data, J Hydrometeorol., 6(2), 156-172,
- 701 https://doi.org/10.1175/JHM407.1, 2005.
- Lloyd, J., & Taylor, J. A.: On the temperature dependence of soil respiration, Funct Ecol., 315-323,
   <u>https://doi.org/10.2307/2389824</u>, 1994.
- Luo, Y., Ahlström, A., Allison, S. D., Batjes, N. H., Brovkin, V., Carvalhais, N., ... & Zhou, T.: Toward more realistic
- projections of soil carbon dynamics by Earth system models, Glob Biogeochem Cycle, 30(1), 40-56,
   https://doi.org/10.1002/2015GB005239, 2016.
- 707 Ma, T., Zhu, G., Ma, J., Zhang, K., Wang, S., Han, T., & Shang, S.: Soil respiration in an irrigated oasis agroecosystem:
- linking environmental controls with plant activities on hourly, daily and monthly timescales, Plant Soil., 447, 347364, https://doi.org/10.1007/s11104-019-04354-w, 2020.
- 710 McGuire, A. D., Sitch, S., Clein, J. S., Dargaville, R., Esser, G., Foley, J., ... & Wittenberg, U.: Carbon balance of the
- terrestrial biosphere in the twentieth century: Analyses of CO2, climate and land use effects with four process-based
- 712 ecosystem models, Glob Biogeochem Cycle., 15(1), 183-206, https://doi.org/10.1029/2000GB001298, 2001.
- 713 Meng, C. L., Li, Z. L., Zhan, X., Shi, J. C., & Liu, C. Y.: Land surface temperature data assimilation and its impact
- on evapotranspiration estimates from the Common Land Model, Water Resour Res., 45(2),
   https://doi.org/10.1029/2008WR006971, 2009.
- 716 Metropolis, N., Rosenbluth, A. W., Rosenbluth, M. N., Teller, A. H., & Teller, E.: Equation of state calculations by
- 717 fast computing machines, J Chem Phys., 21(6), 1087-1092. <u>https://doi.org/10.1063/1.1699114</u>, 1953.
- 718 Mirjalili, S., Saremi, S., Mirjalili, S. M., & Coelho, L. D. S.: Multi-objective grey wolf optimizer: a novel algorithm
- for multi-criterion optimization, Expert Syst Appl., 47, 106-119, https://doi.org/10.1016/j.eswa.2015.10.039, 2016.
- 720 Morris, M. D.: Factorial sampling plans for preliminary computational experiments, Technometrics., 33(2), 161-174,
- 721 https://doi.org/10.1080/00401706.1991.10484804, 1991.
- 722 Pan, Y., Birdsey, R. A., Fang, J., Houghton, R., Kauppi, P. E., Kurz, W. A., ... & Hayes, D.: A large and persistent
- 723 carbon sink in the world's forests, Science., 333(6045), 988-993. <u>https://doi.org/10.1126/science.1201609</u>, 2011.
- 724 Pi, H., & Peterson, C.: Finding the embedding dimension and variable dependencies in time series, Neural Comput.,
- 725 6(3), 509-520, https://doi.org/10.1162/neco.1994.6.3.509, 1994.





- 726 Pollacco, J. A., Mohanty, B. P., & Efstratiadis, A.: Weighted objective function selector algorithm for parameter
- restination of SVAT models with remote sensing data, Water Resour Res., 49(10), 6959-6978,
  https://doi.org/10.1002/wrcr.20554, 2013.
- 729 Raich, J. W., Potter, C. S., & Bhagawati, D.: Interannual variability in global soil respiration, 1980–94, Glob Chang
- 730 Biol., 8(8), 800-812, <u>https://doi.org/10.1046/j.1365-2486.2002.00511.x</u>, 2002.
- 731 Ricciuto, D. M., King, A. W., Dragoni, D., & Post, W. M.: Parameter and prediction uncertainty in an optimized
- terrestrial carbon cycle model: Effects of constraining variables and data record length, J Geophys Res Biogeosci.,
- 733 116(G1), <u>https://doi.org/10.1029/2010JG001400</u>, 2011.
- 734 Riveros-Iregui, D. A., Emanuel, R. E., Muth, D. J., McGlynn, B. L., Epstein, H. E., Welsch, D. L., ... & Wraith, J. M.:
- 735 Diurnal hysteresis between soil CO2 and soil temperature is controlled by soil water content, Geophys Res Lett.,
- 736 34(17), <u>https://doi.org/10.1029/2007GL030938</u>, 2007.
- Ryan, M. G., & Law, B. E.: Interpreting, measuring, and modeling soil respiration, Biogeochemistry, 73, 3-27,
   <u>https://doi.org/10.1007/s10533-004-5167-7</u>, 2005.
- 739 Saini, N., & Saha, S.: Multi-objective optimization techniques: a survey of the state-of-the-art and applications: Multi-
- objective optimization techniques, Eur Phys J Spec Top., 230(10), 2319-2335, <u>https://doi.org/10.1140/epjs/s11734-</u>
- 741 <u>021-00206-w</u>, 2021.
- 742 Segura, C., Coello, C. A. C., Miranda, G., & León, C.: Using multi-objective evolutionary algorithms for single-
- objective constrained and unconstrained optimization, Ann Oper Res., 240, 217-250, <u>https://doi.org/10.1007/s10479-</u>
   015-2017-z, 2016.
- 745 Sellers, P. J., Shuttleworth, W. J., Dorman, J. L., Dalcher, A., & Roberts, J. M.: Calibrating the simple biosphere
- 746 model for Amazonian tropical forest using field and remote sensing data. Part I: Average calibration with field data,
- 747 J Appl Meteorol Climatol., 28(8), 727-759, <u>https://doi.org/10.1175/1520-0450(1989)028<0727:CTSBMF>2.0.CO;2</u>,
- 748 1989.
- 749 Sellers, P. J., Tucker, C. J., Collatz, G. J., Los, S. O., Justice, C. O., Dazlich, D. A., & Randall, D. A.: A revised land
- surface parameterization (SiB2) for atmospheric GCMs. Part II: The generation of global fields of terrestrial
- 751 biophysical parameters from satellite data, J Clim., 9(4), 706-737, https://doi.org/10.1175/1520-
- 752 0442(1996)009%3C0706:ARLSPF%3E2.0.CO;2, 1996.
- 753 Shahsavani, D., Tarantola, S., & Ratto, M.: Evaluation of MARS modeling technique for sensitivity analysis of model
- 754 output, Procedia Soc Behav Sci., 2(6), 7737-7738, https://doi.org/10.1016/j.sbspro.2010.05.204, 2010.
- 755 Shi, Y.: Particle swarm optimization: developments, applications and resources, In Proceedings of the 2001 congress
- on evolutionary computation (IEEE Cat. No. 01TH8546) (Vol. 1, pp. 81-86). IEEE, 2001.
- 757 Smith, T. J., & Marshall, L. A.: Bayesian methods in hydrologic modeling: A study of recent advancements in Markov
- chain Monte Carlo techniques, Water Resour Res., 44(12). <u>https://doi.org/10.1029/2007WR006705</u>, 2008.
- 759 Sobol', I. M.: Sensitivity analysis for nonlinear mathematical models, Mathematical Modeling & Computational
- 760 Experiment, 1, 407-414, 1993.
- 761 Speich, M., Dormann, C. F., & Hartig, F.: Sequential Monte-Carlo algorithms for Bayesian model calibration-A
- 762 review and method comparison, Ecol. Model., 455, 109608, <u>https://doi.org/10.1016/j.ecolmodel.2021.109608</u>, 2021.





- 763 Tang, J., Baldocchi, D. D., & Xu, L.: Tree photosynthesis modulates soil respiration on a diurnal time scale, Glob
- 764 Chang Biol., 11(8), 1298-1304, <u>https://doi.org/10.1111/j.1365-2486.2005.00978.x</u>, 2005.
- 765 Thiemann, M., Trosset, M., Gupta, H., & Sorooshian, S.: Bayesian recursive parameter estimation for hydrologic
- 766 models, Water Resour Res., 37(10), 2521-2535, <u>https://doi.org/10.1029/2000WR900405</u>, 2001.
- 767 Thum, T., MacBean, N., Peylin, P., Bacour, C., Santaren, D., Longdoz, B., ... & Ciais, P.: The potential benefit of
- vising forest biomass data in addition to carbon and water flux measurements to constrain ecosystem model parameters:
- 769 case studies at two temperate forest sites, Agric For Meteorol., 234, 48-65,
  770 https://doi.org/10.1016/j.agrformet.2016.12.004, 2017.
- 771 Tong, C.: PSUADE User's Manual, Lawrence Livermore National Laboratory, LLNL-SM-407882, 2005.
- 772 Van Oijen, M., Rougier, J., & Smith, R.: Bayesian calibration of process-based forest models: bridging the gap
- 773 between models and data, Tree Physiol., 25(7), 915-927, <u>https://doi.org/10.1093/treephys/25.7.915</u>, 2005.
- 774 Varejao, C. G., Costa, M. H., & Camargos, C. C. S.: A multi-objective hierarchical calibration procedure for land
- <sup>775</sup> surface/ecosystem models, Inverse Probl Sci Eng., 21(3), 357-386, <u>https://doi.org/10.1080/17415977.2011.639453</u>,
- 776 2013.
- 777 Vrugt, J. A., Gupta, H. V., Bastidas, L. A., Bouten, W., & Sorooshian, S.: Effective and efficient algorithm for
- multiobjective optimization of hydrologic models, Water Resour Res., 39(8), https://doi.org/10.1029/2002WR001746,
- 779 2003.
- 780 Xia, Y., Yang, Z. L., Jackson, C., Stoffa, P. L., & Sen, M. K.: Impacts of data length on optimal parameter and
- uncertainty estimation of a land surface model, J Geophys Res., 109(D7), <u>https://doi.org/10.1029/2003JD004419</u>,
  2004.
- 783 Xue, B., Zhang, M., & Browne, W. N.: Particle swarm optimization for feature selection in classification: A multi-
- 784 objective approach, IEEE Trans Cybern., 43(6), 1656-1671, <u>https://doi.org/10.1109/TSMCB.2012.2227469</u>, 2012.
- 785 Zeng, X., Shaikh, M., Dai, Y., Dickinson, R. E., & Myneni, R.: Coupling of the common land model to the NCAR
- 786 community climate model, J Clim., 15(14), 1832-1854, <u>https://doi.org/10.1175/1520-</u>
- 787 <u>0442(2002)015%3C1832:COTCLM%3E2.0.CO;2</u>, 2002.
- Zhang, K., Zhu, G., Ma, J., Yang, Y., Shang, S., & Gu, C.: Parameter analysis and estimates for the MODIS
  evapotranspiration algorithm and multiscale verification, Water Resour Res., 55(3), 2211-2231,
  https://doi.org/10.1029/2018WR023485, 2019.
- 791 Zhang, X., Srinivasan, R., Zhao, K., & Liew, M. V.: Evaluation of global optimization algorithms for parameter
- 792 calibration of a computationally intensive hydrologic model, Hydrol Process., 23(3), 430-441,
   793 <u>https://doi.org/10.1002/hyp.7152</u>, 2009.
- 794 Zhu, G., Li, X., Ma, J., Wang, Y., Liu, S., Huang, C., ... & Hu, X.: A new moving strategy for the sequential Monte
- 795 Carlo approach in optimizing the hydrological model parameters, Adv Water Resour., 114, 164-179,
- 796 https://doi.org/10.1016/j.advwatres.2018.02.007, 2018.
- 797 Zhu, G., Li, X., Su, Y., Zhang, K., Bai, Y., Ma, J., ... & He, J.: Simultaneous parameterization of the two-source
- revapotranspiration model by Bayesian approach: application to spring maize in an arid region of northwest China,
- 799 Geosci Model Dev, 7, 741-775, <u>https://doi:10.5194/gmdd-7-741-2014</u>, 2014.





## 800 References From the Supporting Information

- 801 Landsberg, J., & Sands, P.: Physiological ecology of forest production: principles, processes and models (Vol. 4),
- 802 Amsterdam, 2011.
- 803 Wallace, J. M., & Hobbs, P. V.: Atmospheric science: an introductory survey (Vol. 92), Elsevier, 2006.