Technical Note: Space-Time Statistical Quality Control of Extreme Precipitation Observations

Abbas El Hachem[1], Jochen Seidel[1], Florian Imbery[2], Thomas Junghänel[2], and András Bárdossy[1]

[1]Institute for Modelling Hydraulic and Environmental Systems, University of Stuttgart, D-70569 Stuttgart, Germany
[2]Deutscher Wetterdienst, Offenbach, Germany

Correspondence: Abbas El Hachem (abbas.el-hachem@iws.uni-stuttgart.de)

Abstract. Information about precipitation extremes is of vital importance for many hydrological planning and design purposes. However, due to various sources of error, some of the observed extremes may be inaccurate or false. The purpose of this investigation is to present a quality control of observed extremes using space-time statistical methods. To cope with the highly skewed rainfall distribution a Box-Cox transformation with suitable parameter was used. The value at the location of a potential outlier is estimated using the surrounding stations and the calculated spatial variogram and compared to the suspicious observation. If the difference exceeds the threshold of the test, the value is flagged as possible outlier. The same procedure is repeated for different temporal aggregations in order to avoid singularities caused by convection. Detected outliers are subsequently compared to the corresponding radar and discharge observations and finally, implausible extremes are removed. The procedure is demonstrated using observations of sub-daily and daily temporal resolution in Germany.

1 Introduction

A clear definition of an outlier might be intuitive to many but it has been formulated differently by several researchers. In the work of Barnett and Lewis (1994) an outlier was defined as an observation showing an inconsistent behavior compared to other data values. Hawkins (1980) described an outlier as being an observation that differs substantially from other observations as if it might have been produced by an alternating mechanism. More precisely, for Iglewicz and Hoaglin (1993) an outlier is an observation that arouses suspicion to the analyst and does not belong to the same data distribution. In general, there are two types of outliers, those associated with an error and those associated with a real observation. The reasons for an observation being erroneous could be due to instrumental errors (e.g. use of false instrument, equipment malfunction, false equipment operation) or/and human errors (false reading or recording or even computation of observations). Moreover, errors can occur if the measuring site is falsely chosen, providing a false representativeness of the observed process.

Hydroclimatological data are of unique nature as they occur in a non repetitive manner. If an observation is not registered correctly reconstructing it is very challenging, especially for precipitation values. However, reliable information about precipitation extremes is essential for many design purposes such as flood analysis, extreme value statistics and stationarity analysis to name just a few.

Many quality control (QC) algorithms have been developed and are being used by weather service agencies to minimize and detect false measurements. Durre et al. (2010) established a comprehensive QC algorithm for daily surface meteorological observations (temperature, precipitation, snowfall and snow depth). For precipitation data, the QC method for detecting false observations consisted of several steps. A climatological outlier check is used for flagging values exceeding a certain temperature-dependent threshold and a spatial consistency check based on comparing the target observation to neighboring ones. An observation is eventually flagged if the difference exceeds a certain climatological percent ranks threshold. Qi et al. (2016) implemented a QC algorithm to identify erroneous hourly rain gauge observations by using additional information as radar quantitative precipitation estimates (QPE). A common practice for detecting outliers is to use an interpolation method to estimate the probability of a local observation using the surrounding locations. If this probability is very low then the observation is suspicious. This was mathematically formulated by Ingleby and Lorenc (1993). In the work done by Hubbard et al. (2005) a QC method was developed for daily temperature and precipitation values consisting of four steps. Observations are flagged if they do not fall within $\pm$ 3 standard deviation of the long-term mean and if they differ from the estimated value using a spatial regression technique. Some other QC methods are available but are often limited to time series analysis and tend to disregard the temporal-spatial extent of precipitation.

Precipitation observations have a space-time dimension. Observations are taken at different locations in space and in discrete time intervals. Due to the presence of non-negative and many no-precipitation (zero) values, precipitation data (especially at daily and sub-daily resolution) have a positively non-normal skewed distribution with heavy tails (Klemeš, 2000) and fall under the zero-inflated data. Therefore, an adequate transformation of the data should be performed to reduce the effect of the data skewness. A relatively simple approach to normalizing a variable is the Box-Cox transformation (Box and Cox, 1964).

The following work proposes a statistical space-time methodology based on interpolation in a cross-validation mode to find possible outliers in precipitation observations across several temporal aggregations. An outlier is defined here as an observation that strongly differs for a certain temporal aggregation from its spatial neighboring locations. A difficult task while working with outliers in general and especially in hydrology is distinguishing between correct and false observations. Therefore, to validate detected outliers, the suspected values are additionally compared to independent information such as discharge and radar measurements.

This paper is organized as follows: after the introduction, the data and methodology are presented. Afterward, the results of the outlier detection are presented and four examples of verification via subsequent comparison to radar or discharge data are shown. In the final section of the results, the number of identified outliers for every year (and month) and for each location is mapped and presented.

2 Study Area and Data

This study was done using the German-wide precipitation data set from the German Weather Service DWD which covers an area of approximately 357,000 km^2. The average annual rainfall in Germany is around 800 mm and can reach up to 2100 mm in the higher elevations of the Alps in the South. Currently, the DWD operates a network of rain gauges with different temporal

resolutions ranging from minutely to daily. Hourly and sub-hourly data are available from the 1990s onwards. The number of these stations has been continuously increasing since then. On the other hand, the number of stations with daily observations started to decrease since then as they were replaced by automatic rain gauges. Rain gauges near the border (30-kilometer inland buffer) were not included in this analysis.

In the 1990s, most DWD rain gauges were tipping buckets or drop counters. From 2000 onwards, these were replaced by weighing gauges (OTT Pluvio) and since 2017 these are being replaced by combined tipping bucket and weighing rain gauges (Lambrecht rain[e]).

Precipitation data from the recent DWD observation network go through several quality control steps. The first step is a quality control directly at the automatic stations. Since this is an automatic test, relatively wide thresholds are applied. It includes tests for completeness, thresholds, temporal and internal consistency. Based on these tests, a quality flag is assigned to the data. The data is then submitted to a database. Another test with tighter thresholds is then performed, based on the QualiMet software (Spengler, 2002). This phase of the quality check also tests for completeness as well as climatological, temporal, spatial, and internal consistency. Questionable values are manually checked and corrected and the quality label is adjusted. A final quality check step occurs after all of a month's data are available, focusing on aggregate values. The quality flags are stored in the database and are also made available to users. DWD quality assurance also includes the identification and correction or description of errors in the historical data (Kaspar et al., 2013). Appropriate procedures have been developed for the quality control of historical data. In general, the quality of these values can be considered reasonably good, but there are still doubtful values on the order of about 0.1-1%, especially for the pre-1979 data. The user must keep in mind that the data can be affected by certain non-climatic effects, such as changes in instrumentation or observation time. With few exceptions, the data are reported "as observed", i.e., no homogenization procedure was applied.

As independent data for verification, radar derived rainfall QPE and discharge observations from the state of Bavaria were used. The radar data used is the product RADOLAN-RW that is provided by the DWD in hourly and daily resolutions starting the year 2005 (DWD Climate Data Center (CDC), 2021). These products have been gauge-adjusted with the observed hourly station data. The occurrence (or absence) of precipitation observation in the radar data over the target location is an indication to the quality of the observation. The discharge data were quality checked and provided by the environmental agency of Bavaria with hourly and daily resolutions for approximately 400 gauges within the region of Bavaria (Bayerisches Landesamt für Umwelt, 2022). Different headwater catchments were derived and selected for validating the results. A reaction (within few hours) in the headwater catchment discharge is expected after the event occurred in case of correct rainfall observations.

Figure 1 illustrates the location of the daily and sub-daily rain gauges as well as their spatial density. For the daily data all available locations (historical and present) are displayed. The stations do not have a homogeneous spatial distribution over the country where some locations have a higher network density than others. The spatial density was calculated using a kernel density estimation (KDE) with a Quartic shape and a radius of influence of 30 kilometers. The estimated density value depends on the separating distance between the known and unknown locations and the kernel parameters. These are the bandwidth (h) which is reflected by the radius of influence and the weighting function or kernel function (K). The latter defines the

contribution of each point as a function of the separating distance. Further detail regarding KDE estimation can be found in Yu et al. (2015).
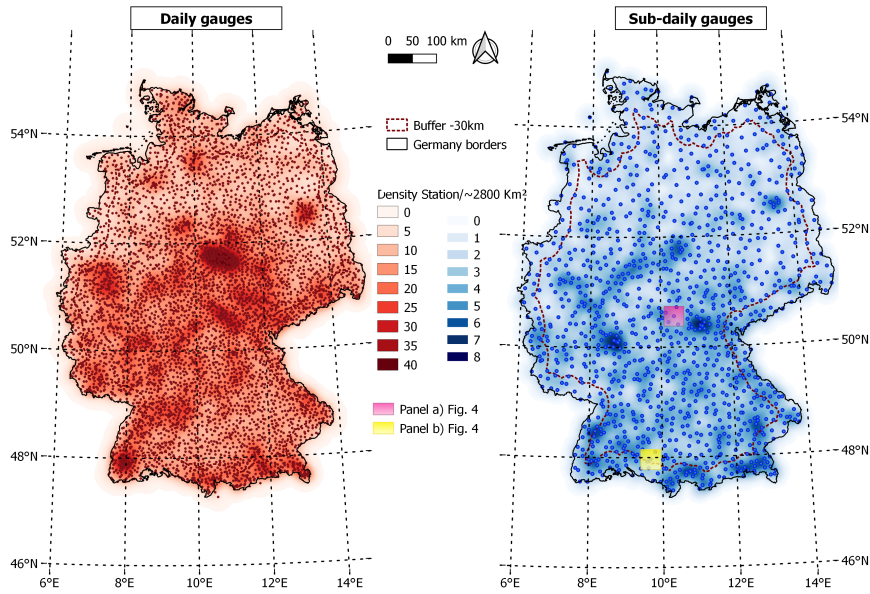
Figure 1. Map of the study area showing the location and density (number of stations per $\sim 2800 \ km^2$) of the DWD gauges with daily (left) and sub-daily (right) resolutions.The yellow and pink boxes in the right map refer to the two locations of the examples presented in Figure 4

3 Methodology

3.1 Data transformation

95 As an initial step, a Box-Cox transformation as described by equation 1 was applied for every variable X and temporal aggregation t to reduce the effect of the skewed precipitation distribution (Box and Cox, 1964).

$$X_t^* = \begin{cases} \frac{(X_t^\lambda - 1)}{\lambda} & \text{if } \lambda \neq 0 \\ \log{(X_t)} & \text{if } \lambda = 0 \end{cases} \tag{1}$$

Where:

$X^* =$ transformed precipitation data at location u and temporal aggregation t

$X =$ original precipitation data at location u and temporal aggregation t

$\lambda =$ transformation factor for temporal aggregation t

To find which transformation factor λ is most suitable, several simulated lower truncated standard normal distribution function (sampling space bounded by $[-\infty < a = p_0,\ b = +\infty]$) were fitted to the original data (Burkardt, 2014). The probability of having a value above or below p_0 is then derived (p_0 probability of having 0 mm precipitation value) .

From this probability (denoted p_{norm}) a new standard normal distribution is generated where ($x < p_{norm} = 0$, $x >= p_{norm} = x$). From this distribution the skewness γ_{norm} is calculated. The goal now is to find which transformation factor minimizes the difference between the original data skewness and γ_{norm}. This was done for each station separately and for all aggregations. Eventually an average transformation factor (denoted hereafter λ) was derived for each temporal aggregation. The results of this procedure can be seen in Table 1.

Table 1. Average transformation factor λ used to transform the original data to the truncated normal space with reduced skewness.

	60 min	120 min	180 min	240 min	360 min	720 min	1440 min
Average transformation factor λ	0.097	0.155	0.219	0.262	0.318	0.427	0.499

Once λ was calculated, the original precipitation data were transformed as in equation 1, and in the newly truncated normalized space the following approach was implemented to find outliers in the precipitation data over several temporal resolutions.

3.2   Outlier detection

The proposed method was initially tested for identifying outliers in groundwater quality data (Bárdossy and Kundzewicz, 1990). In this paper, a similar method was implemented to identify unusual precipitation data and is extended by a validation of the results using external information such as radar or discharge observations. For detecting possible outliers the concept of jackknifing is used, a method initially developed by Quenouille (1949, 1956). The main idea is based on removing one (or each) observation from the data and estimating its value again. In this study, the four largest annual observations for every station are compared to the estimated values at the same location. Each cross-validated value is estimated using the nearest 30 neighboring locations with valid observations. This is needed for a reliable estimation of the variogram and has less influence on the estimated value due to the shading effect in Kriging, ie. the stations further away have smaller weights.

Since many possible faulty observations can only be detected at lower temporal resolution, the procedure was applied over several temporal aggregations. For example, when looking at sub-daily and sub-hourly values a single observation might not be unusual but the accumulation of many values reveals suspicious sums. Furthermore, single events might occur on high temporal scales (e.g. hourly) and are not detected on lower aggregations (e.g. daily).

For estimating the target value Ordinary Kriging (OK) is used as an interpolation technique. It is a regionalization method initially introduced by Krige (1951) and Matheron (1962) to estimate an unknown value at a target location by solving a linear equation system by minimizing the estimation variance and maximizing the accuracy (no systematic error). Each cross-validated value is estimated using the nearest 30 neighboring locations with valid observations. The spatial correlation structure is reflected by the variogram which is derived in the rank space domain and rescaled to the variance of the data. This allows a for variogram calculation in a more robust manner (Lebrenz and Bárdossy, 2019). The target location is calculated by

solving the kriging equation and the estimation variance is noted. For identifying unusual observations the ratio between the absolute value of the difference between the observed and the estimated values and the estimation variance is calculated. This Criteria Ratio, (CR) describes the relative agreement/disagreement between the observed value and the spatial surroundings for the corresponding time step. Larger CR values reflect high spatial-temporal disagreement and low values denote greater agreement. Based on the CR value, different types of events can be identified, namely those occurring on a local scale with high CR values and other on a regional scale with low CR values. Following Bárdossy and Kundzewicz (1990) a CR value of three is initially used to identify suspicious observations. The CR value is derived for every cross-validated event. Eventually, the CR value is related to all of the observed (interpolated) data establishing a possibility to find a suitable CR value for identification of precipitation outliers.

$$CR_i(u) = \frac{|Z_i^*(u) - Z_i(u)|}{\sigma_i(u)} \tag{2}$$

Where:

$$
\begin{aligned}
Z_i^*(u) &= \quad \text{estimated value at location } u \text{ and timestep } i \\
Z_i(u) &= \quad \text{observed value at location } u \text{ and timestep } i \\
\sigma_i &= \quad \text{kriging standard deviation at location } u \text{ and timestep } i
\end{aligned}
$$

Since precipitation events occurring on a local scale might represent an actual small scale event to validate the first or the second case, the suspicious events are compared to the observed radar QPE or discharge values in the corresponding catchment. Despite having their own drawbacks the radar and discharge observations are used here as a qualitative decision support tool.

The flowchart in figure 2 describes the implemented space-time precipitation outlier detection scheme.
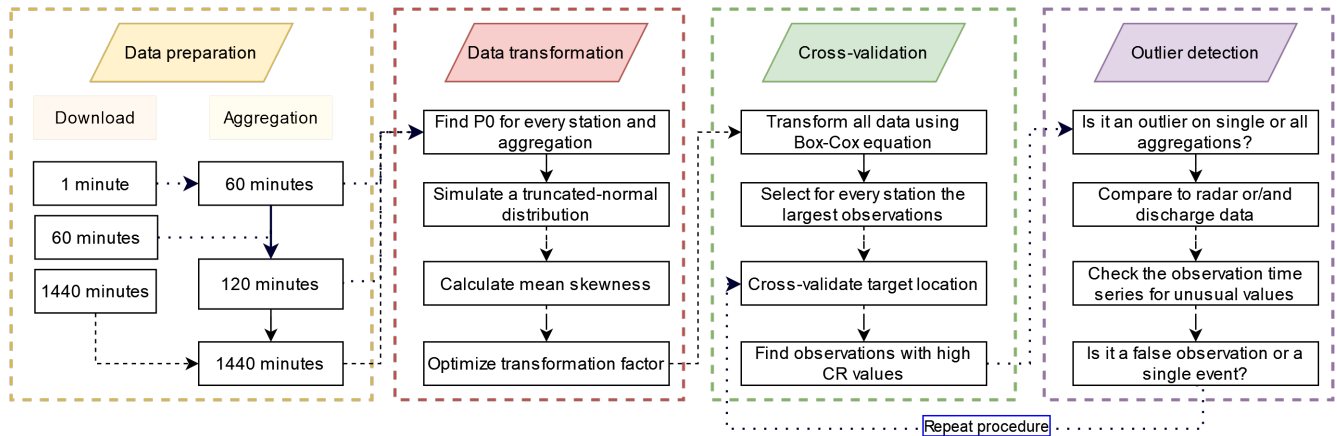
Figure 2. Flowchart summarizing the described method starting with the data download procedure and ending with the identification of suspicious observations.

3.3 Data corruption

To test further the validity of the method, 20 stations without any detected outliers were randomly selected and their data (same events as before) were 'artificially' manipulated such that the transformed observations of each target location were decreased and increased by several percentages (from 25 to 100 %) and the outlier detection method was tested. The results of this procedure can be seen in table 2. By decreasing the observed value until reaching a false zero observation the method was able to identify on the hourly scale around 60 % and on the daily scale 94 % of the cases as being outliers. On the other hand, by increasing the error value to up to 100 %, almost all values were detected on all temporal aggregations. This emphasizes the validity of the method especially regarding identifying false high observations.

Table 2. Number of newly detected events after corrupting by different percentages the cross-validated observations of 20 randomly selected stations with no previous outliers.

	Temporal aggregation	60 min	120 min	180 min	240 min	360 min	720 min	1440 min
	Number of events	**150**	**150**	**150**	**150**	**150**	**150**	**150**
	Minimum of the minima [mm]	5.12	5.16	5.17	5.13	5.26	5.6	5.17
	Average of all averages [mm]	11.05	12.67	14.41	14.41	16.85	19.8	24.03
	Maximum of the maxima [mm]	51.2	50.1	53.47	63.93	71.92	73.6	76.37
Percentage of error	**-100 [%] (false zero)**	88	115	102	125	100	124	141
	-50 [%]	10	29	38	41	33	46	65
	-25 [%]	2	3	13	9	8	8	4
	0 [%]	**0**	**0**	**0**	**0**	**0**	**0**	**0**
	+25 [%]	23	45	48	52	65	46	55
	+50 [%]	74	88	118	121	119	113	116
	+100 [%]	149	150	150	148	150	149	149

4 Results

4.1 Outliers vs single events

The center panel of figure 3 represents the CR value versus the ratio between the interpolated and observed values. All values denoted in red have a CR value above 3. This figure allows identifying the events that are of interest and relating the CR value to the interpolated and observed data. Note that the observed and interpolated values are in the original non-transformed space, only the CR values are calculated from the interpolation of the transformed values.

160   The values in the plot having a ratio of interpolated to observed of 5, are values obtained when a neighboring station (or stations) had simultaneously recorded an outlier (in this case a false high observation). This leads to detecting a false outlier. This can be accounted for by running the method again after all neighbours have been checked.

    In the left and right panels of figure 3 the cumulative distribution function (CDF) from all investigated observations was calculated and the location of the detected outlier ismarked. The events that were detected as being outliers spread over the

165 curve showing that the method can detect not only high values but as well relatively small values that differ highly from their neighboring space.

    The left panel of figure 3 shows the results for the original hourly observations. The right panel shows those for the aggregated minutely observations. By comparing the two, the quality control procedure of the DWD can be investigated. Spatial consistency is checked more intensively by the DWD for higher aggregated precipitation data (≥ 1 h) than for high temporal

170 resolution data (e.g. 1 min). For example, in the hourly data, none of the largest

    sum (> 60 mm h) is detected as an outlier and only one observation is larger than > 80 mm h.

    In the hourly data based on the aggregated minutely data, many values above 80 mm h exist and are mostly all detected as being suspicious. There are even unrealistic values with accumulated sums above 200 mm h which can be caused by several faulty 'small' measurements or a few single large spikes in the data.
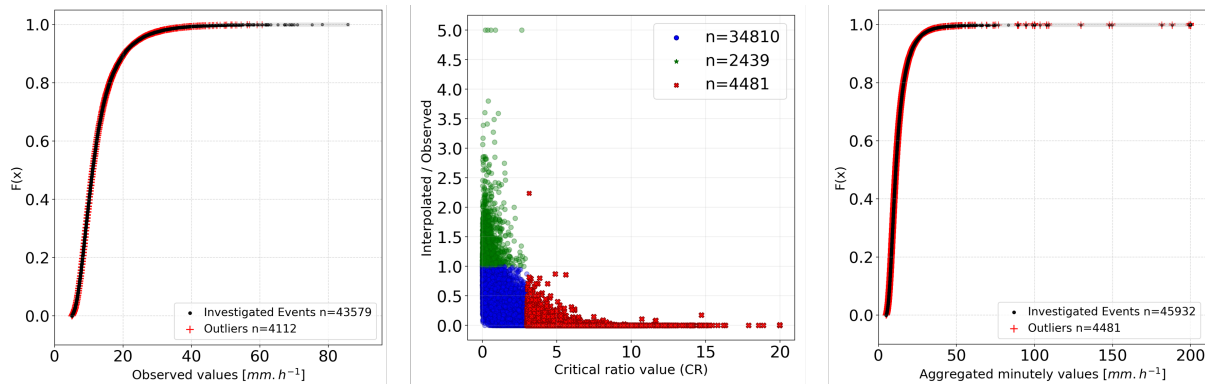
Figure 3. The left panel shows the CDF of all investigated hourly events with the detected outliers marked in red. The center panel shows for the minutely aggregated data the CR values versus the ratio of interpolated and observed hourly values. The right panel shows the CDF of all investigated hourly events with the detected outliers marked in red. The hourly data in the center and right panel were aggregated from the minutely values. Note that an upper limit of 200 mm h^{-1} was set.

175 **4.2   Selected case studies**

The first example in panel a) of figure 4 shows the presence of unusual values in the minutely data of the cross-validated station (> 8 mm min^{-1}). The radar data for that hour do not show such a high-intensity event above the investigated location. The

second example in panel b) of figure 4 shows a similar case in the minutely data but the radar image confirms the occurrence of the event.
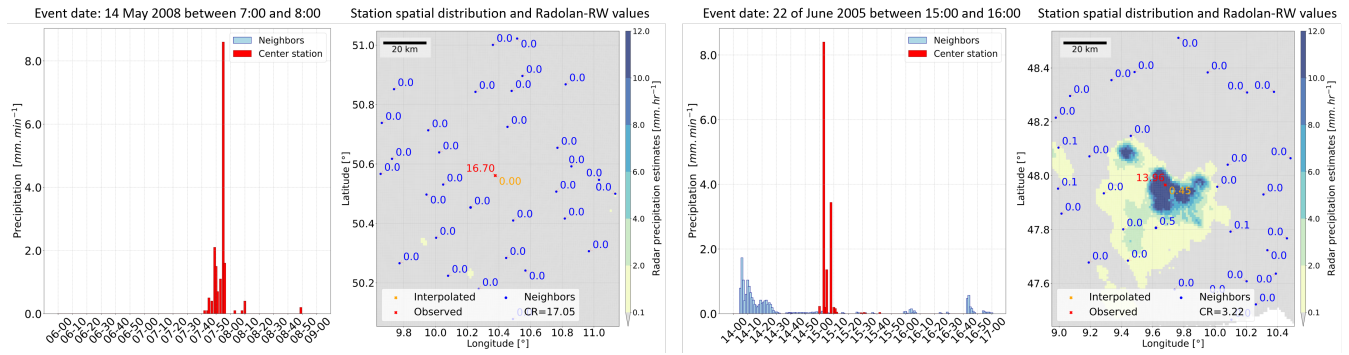
Figure 4. Two examples are shown: Panel (a) of an identified false observation and Panel (b) with a plausible event. The right figures in Panels (a) and (b) show the values at the neighboring stations (in blue), the observed value (in red), the estimated value (in orange), and the Radolan-RW QPE data for that hour.

180   Discharge data from small headwater catchments in the federal state of Bavaria with one (or many) rain gauge stations within the catchment were analysed. If a rain gauge observation was identified as being suspicious the discharge values for the next hours following the event were checked. An example for this is shown in the upper Pegnitz catchment which is located on the northern part of the Bavaria (Fig. 5). Panel b) in figure 5 shows an hourly outlier observation that resulted in a reaction in the corresponding headwater catchment. On the other hand, panel c) in figure 5 shows the opposite case, i.e. an hourly outlier that

185   did not cause any reaction in the Pegnitz catchment.
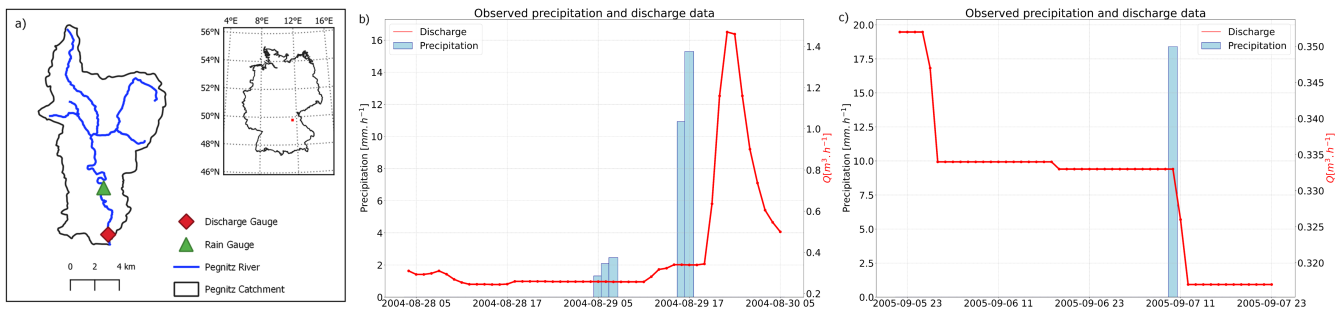
Figure 5. a) Location of discharge and rain gauge station within the Pegnitz headwater catchment (panel a) and observed discharge and precipitation data (+/- 1 day) for detected outliers with b) a discharge increment and c) without a discharge increase

4.3 Results over all stations and aggregations

The method was applied over several temporal aggregations (hourly to daily) and events that are suspicious over single or several aggregations were identified. The result of this can be seen in Table 3. The diagonals show events that are common over the corresponding test and reference temporal aggregation. Some observations are only suspicious until a temporal aggregation is reached or exceeded beyond which they are not detected anymore. The result of this can be seen in the values above and below the diagonals in Table 3.

Table 3. The diagonals show the number of unique days with identified outliers. The values above the diagonals reflect the number of different days between the reference and test aggregation. For example, there are 358 days in the reference 60 min aggregation that are not in the test 120 min aggregation.

		Test aggregation						
		60 min	120 min	180 min	240 min	360 min	720 min	1440 min
Reference aggregation	60 min	**1581**	358	392	414	498	762	898
	120 min	218	**1441**	210	237	354	646	787
	180 min	344	302	**1533**	240	341	657	825
	240 min	437	400	311	**1604**	329	657	837
	360 min	539	535	430	347	**1622**	559	762
	720 min	771	795	714	643	527	**1590**	439
	1440 min	889	918	864	805	712	421	**1572**

The number of active stations (and device quality) affects the number of detected outliers. The red curve in figure A1 represents the ratio of detected outliers to the number of active stations (for every hour) which is shown by the blue curve. As the number of active stations increases the number of detected outliers decreases which is an indication that the quality of the observations is improving with time. In the right panel of figure A1 the effect of seasonality was inspected. The detected outliers were grouped by the month in which they occurred. The results show that in the summer period the number of detected outliers is much larger than in the winter period. This is related to convectional rainfall processes occurring in the summer period leading to more small scale single events. Finally, the percentage of outliers in the investigated events of every station for the hourly aggregated data is presented in the right panel of figure 6. The map does not present any clear structure related to elevation and topography. Moreover, the map shows that outliers can happen everywhere meaning this is not a systematic problem.

5 Conclusion

In this study, a methodology to identify outliers in precipitation data was presented. Due to the high spatio-temporal variability of precipitation, the quality control of precipitation data takes on a special task among the meteorological parameters. Therefore, a transfer of the procedures developed in this paper to other variables is not necessarily reasonable. From an hourly to
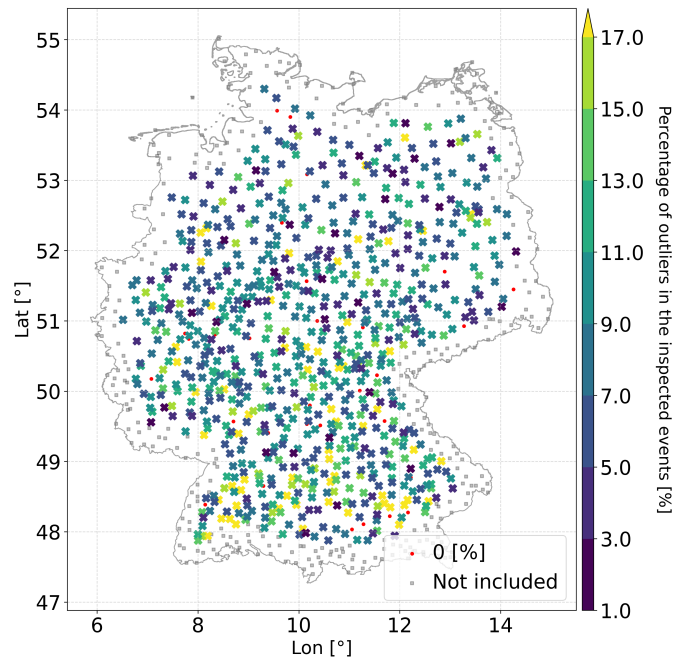
Figure 6. The figure shows the percentage of outliers in the investigated events of every station.

a daily temporal aggregation the largest four yearly values for every station were inspected. To cope with the skewed rainfall distribution a Box-Cox transformation with a suitable parameter was applied. The results revealed different outliers throughout various temporal aggregations To test the robustness of the method, data from stations with no outliers were corrupted with several percentages and checked again. The method was able to identify most events as outliers as the value of the added error

210 increased. Seasonality was seen to play a major role in the number of detected outliers and as the quality of the observations improved, the number of detected outliers (namely the false measurements) decreased. For distinguishing between a false observation and a single event external data was used. Discharge gauge data of corresponding headwater catchments and radar rainfall images were used when available. A final choice regarding flagging an observation is done carefully and individually for every location. Eventually, the flagged observations are kept aside and investigated before being used in further analysis.

215 The current method needs to be extended and modified for temporal aggregations below the hourly scale. Especially the kriging methodology should include time as a third dimension to account for advection and correlation between subsequent steps. Moreover, many events are identified as being an outlier when part of the neighboring stations had zero precipitation values. This can happen in the case of directional-dependent events driven by a frontal system. These cases could be further handled by including anisotropy in the interpolation method.

Data availability. The precipitation data was obtained from the Climate Data Center of the Deutscher Wetterdienst (https://opendata.dwd.de/
climate_environment/CDC). The discharge data are provided by the environmental state agency of Bavaria LfU (https://www.gkd.bayern.de).

Code and data availability. The corresponding code is available on Github in the repository *qcpcp* quality control of precipitation observation (El Hachem).
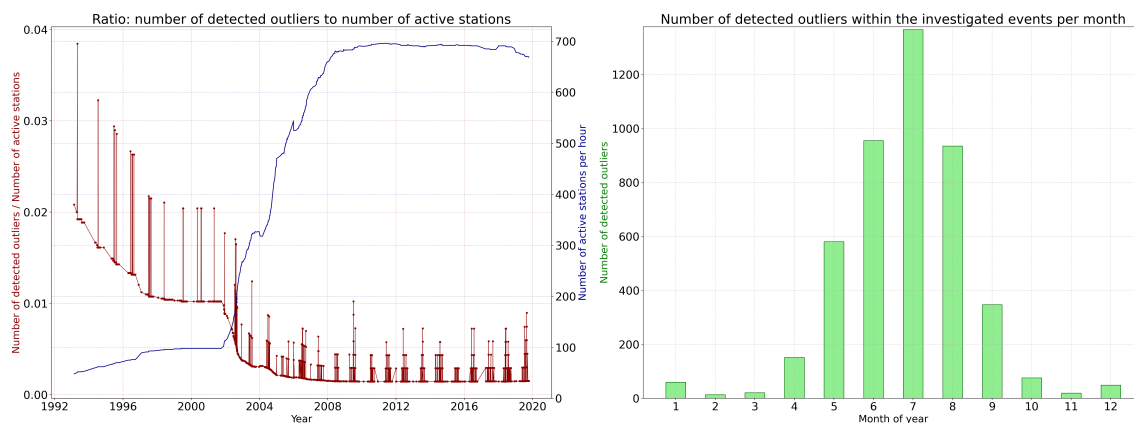
Figure A1. Left figure shows the number of hours with outliers within the investigated hourly events (aggregated from 1 minute) of all stations per year. The right figure shows the number of detected outliers within the investigated events for every month within the hourly data using a CR value of 3.

Author contributions. AEH developed and implemented the algorithm for the study area. JS assisted in the analysis and description of the results. FI and TJ provided valuable information regarding the data processing. AB designed and supervised the study. Moreover, all authors contributed to the writing, reviewing and editing of the manuscript.

Competing interests. The authors declare that they have no conflict of interest.

Acknowledgements. This study is part of the project B 2.7 "STEEP - Space-time statistics of extreme precipitation" (Grant No. 01LP1902P) of the "ClimXtreme" project funded by the German Ministry of Education and Research (Bundesministerium für Bildung und Forschung, BMBF). The authors thank the German Weather Service DWD for providing the precipitation and weather radar data and the Bavarian Environment Agency for providing the discharge data. Moreover we acknowledge all developers of different python core libraries (e.g. numpy, pandas, matplotlib, cython, scipy) for providing open source code. The authors thank the University of Stuttgart for funding this open-access publication and the two anonymous reviewers for their valuable remarks that helped improve the quality of this manuscript.

References

235 Bárdossy, A. and Kundzewicz, Z.: Geostatistical methods for detection of outliers in groundwater quality spatial fields, Journal of Hydrology, 115, 343–359, https://doi.org/10.1016/0022-1694(90)90213-H, 1990.

Barnett, V. and Lewis, T.: Outliers in statistical data, John Wiley and Sons Location, Hoboken, NJ, 1994.

Bayerisches Landesamt für Umwelt: Wasserstand und Abfluss, https://www.lfu.bayern.de/wasser/wasserstand_abfluss/index.htm, 2022.

Box, G. E. and Cox, D. R.: An analysis of transformations, Journal of the Royal Statistical Society: Series B (Methodological), 26, 211–243,
240 1964.

Burkardt, J.: The truncated normal distribution, Department of Scientific Computing Website, Florida State University, 1, 35, 2014.

Durre, I., Menne, M. J., Gleason, B. E., Houston, T. G., and Vose, R. S.: Comprehensive automated quality assurance of daily surface observations, Journal of Applied Meteorology and Climatology, 49, 1615–1633, 2010.

DWD Climate Data Center (CDC): Historical hourly RADOLAN grids of precipitation depth (GIS-readable), version 2.5, 2021.

245 El Hachem, A.: qcpcp, quality control of precipitation observations, https://doi.org/10.5281/zenodo.7310836.

Hawkins, D. M.: Identification of outliers, Springer Dordrecht, https://doi.org/10.1007/978-94-015-3994-4, 1980.

Hubbard, K., Goddard, S., Sorensen, W., Wells, N., and Osugi, T.: Performance of quality assurance procedures for an applied climate information system, Journal of Atmospheric and Oceanic Technology, 22, 105–112, 2005.

Iglewicz, B. and Hoaglin, D.: The ASQC basic references in quality control: statistical techniques, in: How to detect and handle outliers,
250 edited by Mykytka, E., vol. 16, pp. 1–87, ASQC Quality Press Milwaukee, WI, 1993.

Ingleby, N. B. and Lorenc, A. C.: Bayesian quality control using multivariate normal distributions, Quarterly Journal of the Royal Meteorological Society, 119, 1195–1225, 1993.

Kaspar, F., Müller-Westermeier, G., Penda, E., Mächel, H., Zimmermann, K., Kaiser-Weiss, A., and Deutschländer, T.: Monitoring of climate change in Germany – data, products and services of Germany's National Climate Data Centre, Advances in Science and Research, 10,
255 99–106, https://doi.org/10.5194/asr-10-99-2013, 2013.

Klemeš, V.: Tall tales about tails of hydrological distributions. I, Journal of Hydrologic Engineering, 5, 227–231, 2000.

Krige, D. G.: A statistical approach to some basic mine valuation problems on the Witwatersrand, Journal of the Southern African Institute of Mining and Metallurgy, 52, 119–139, 1951.

Lebrenz, H. and Bárdossy, A.: Geostatistical interpolation by quantile kriging, Hydrology and Earth System Sciences, 23, 1633–1648, 2019.

260 Matheron, G.: Traité de géostatistique appliquée, 14, Editions Technip, 1962.

Qi, Y., Martinaitis, S., Zhang, J., and Cocks, S.: A real-time automated quality control of hourly rain gauge data based on multiple sensors in MRMS system, Journal of Hydrometeorology, 17, 1675–1691, 2016.

Quenouille, M. H.: Approximate tests of correlation in time-series, Journal of the Royal Statistical Society: Series B (Methodological), 11, 68–84, 1949.

265 Quenouille, M. H.: Notes on Bias in Estimation, Biometrika, 43, 353–360, http://www.jstor.org/stable/2332914, 1956.

Spengler, R.: The new quality control and monitoring system of the Deutscher Wetterdienst, in: Proceedings of the WMO Technical Conference on Meteorological and Environmental Instruments and Methods of Observation, Bratislava, vol. 330, 2002.

Yu, W., Ai, T., and Shao, S.: The analysis and delimitation of Central Business District using network kernel density estimation, Journal of Transport Geography, 45, 32–47, 2015.