Development of a Wilks Feature Importance Method with Improved Variable Rankings for Supporting Hydrological Inference and Modelling

Kailong Li[1], Guohe Huang[1], Brian Baetz[2]

[1]Faculty of Engineering, University of Regina, Regina, Saskatchewan, Canada S4S 0A2
[2]Department of Civil Engineering, McMaster University, Hamilton, Ontario, Canada L8S 4L8.
Correspondence to: Guohe Huang (huangg@uregina.ca)

Abstract. Feature importance has been a popular approach for machine learning models to investigate the relative significance of model predictors. In this study, we developed a Wilks feature importance (WFI) method for hydrological inference. Compared with conventional feature importance methods such as permutation feature importance (PFI) and mean decrease impurity (MDI), the proposed WFI aims to provide more reliable variable rankings for hydrological inference. To achieve this, WFI measures the importance scores based on Wilks Λ (a test statistic that can be used to distinguish the differences between two or more groups of variables) throughout an inference tree. Compared with PFI and MDI methods, WFI does not rely on any performance measures to evaluate variable rankings, which can thus result in less biased criteria selection during the tree deduction process. The proposed WFI was tested by simulating monthly streamflows for 673 basins in the United States and applied to three interconnected irrigated watersheds located in the Yellow River Basin, China, through concrete simulations for their daily streamflows. Our results indicated that the WFI could generate stable variable rankings in response to the reduction of irrelevant predictors. In addition, the WFI selected predictors helped RF achieve its optimum predictive accuracy, which indicates the proposed WFI could identify more informative predictors than other feature importance measures.

1 Introduction

Machine learning (ML) has been used for hydrological forecasting and examining modeling processes underpinned by statistical and physical relationships. Due to the rapid progress in data science, increased computational power, and the recent advances in ML, the predictive accuracy of hydrological processes has been greatly improved (Reichstein et al., 2019; Shortridge et al., 2016). Yet, the explanatory power of ML models for hydrological inference has not increased apace with their predictive power for forecasting (Konapala and Mishra, 2020). Previous studies have indicated that purely pursuing predictive accuracy may not be a sufficient reason for applying

a certain hydrological model to a given problem (Beven, 2011). The ever-increasing data sources allow ML models to incorporate potential driving forces that cannot be easily considered in physically-based hydrological models (Kisi et al., 2019). The increasing volume of input information has left one challenge as "how to extract interpretable information and knowledge from the model." Even though obtaining exact mappings from data input to prediction is technically infeasible for ML models, previous research has shown opportunities to understand the model decisions through either post-hoc explanations or statistical summaries of model parameters (Murdoch et al., 2019). Nevertheless, the reliability of the interpretable information is not well understood. Therefore, quality interpretable information from ML models is much desired for evolving our understanding of nature's laws (Reichstein et al., 2019).

The main idea of model interpretation is to understand the model decisions, including the main aspects of (i) identifying the most relevant predictor variables (i.e., predictors) leading to model predictions and (ii) reasoning why certain predictors are responsible for a particular model response. Interpretability can be defined as the degree to which a human can understand the cause of a decision (Miller, 2019). The model interpretation for ML is mainly achieved through feature importance, which relies on techniques that quantify and rank the variable importance (i.e., a measure of the influence of each predictor to predict the output) (Scornet, 2020). The obtained importance scores can be used to explain certain predictions through relevant knowledge. Moreover, Gregorutti et al. (2017) pointed out that some irrelevant predictors may have a negative effect on the model accuracy. Therefore, eliminating irrelevant predictors might improve the predictive accuracy. Feature importance methods can be categorized as model-agnostic and model-specific (Molnar, 2020). The model-agnostic methods refer to extracting post-hoc explanations by treating the trained model as a black box (Ribeiro et al., 2016a). Such methods usually follow a process of interpretable model learning based on the outputs of the black-box model (Craven and Shavlik, 1996) and perturbing inputs and seeing the response of the black-box model (Ribeiro et al., 2016b). Such methods mainly include permutation feature importance (PFI) (Breiman, 2001a), partial dependence (PD) plots (Friedman, 2001), individual conditional expectation (ICE) plots (Goldstein et al., 2015), accumulated local effects (ALE) plots (Apley and Zhu, 2016), local interpretable model-agnostic explanations (LIME) (Ribeiro et al., 2016b), Morries method (Morris, 1991) and Shapley values (Lundberg and Lee, 2017; Shapley, 1953). In hydrology, Yang and Chui

40

45

50

55

60

(2020) used Shapley values to explain individual predictions of hydrological response in sustainable drainage systems at fine temporal scales. Kratzert et al. (2019a) used the Morries method to estimate the rankings of predictors for a long short-term memory (LSTM) model. Worland et al. (2019) used the LIME to infer the relation between basin characteristics and the

70   predicted flow duration curves. Konapala and Mishra (2020) used partial dependence plots to understand the role of climate and terrestrial components in the development of hydrological drought. Compared with the above model-agnostic methods, PFI is more widely used in hydrological inference due to its high efficiency and ability to take global insights into model behaviors (Molnar, 2020). Recent applications of PFI include inferring the relationship between

75   basin characteristics and predicted low flow quantiles (Ahn, 2020) and comparing the interpretability among multiple machine learning models in the context of flood events (Schmidt et al., 2020). The above model-agnostic methods are useful for comparative studies of ML models with exceedingly complex (such as deep neuron networks) algorithmic structures to extract the interpretable information.

80

On the other hand, the model-specific methods (also known as interpretable models), such as decision trees and sparse regression models, can inspect model components directly (Ribeiro et al., 2016a; Yang, 2020). For instance, the weights (or coefficients) of a linear regression model can directly reflect how the predictions are produced, thus providing critical information for ranking

85   the model predictors. Due to the oversimplified input-output relationships, linear regression models may be inadequate to approximate the complex reality. As a consequence, these models may hardly achieve satisfactory predictive accuracy and obtain quality interpretable information. As one of the essential branches of interpretable models, tree-structured models such as classification and regression trees (CART) (Breiman et al., 1984) have been an excellent

90   alternative to linear regression models for solving complex non-linear problems. The principle of CART is to successively split the training data space (i.e., predictors and response) into many irrelevant subspaces. These subspaces and the splitting rules will form a decision/regression tree, which asks each of the new observations a series of "Yes/No" questions and guides it to the corresponding subspaces. The model prediction for a new observation shares the same value as the

95   average value for the training responses in that particular subspace. Mean decrease impurity (MDI) is the feature importance method for CART, and it summarizes how much a predictor can improve

100 the model performance through the paths of a tree. Compared with linear regression models, trees are more understandable for inferring a particular model behavior because the transparent decision-making process functions similarly to how the human brain makes decisions for a series of questions (Murdoch et al., 2019). Based on CART, Breiman (2001a) proposed an ensemble of trees named random forest (RF), which significantly improved the predictive accuracy compared
105 with CART. Previous studies reported that RF could outperform many other ML models in predictive accuracy (Fernández-Delgado et al., 2014; Galelli and Castelletti, 2013; Schmidt et al., 2020). The high predictive accuracy allowed RF to become very useful in interpretation, especially in hydrology (Lawson et al., 2017; Worland, 2018). As Murdoch et al. (2019) argued, higher predictive accuracy can lead to a more reliable inference.

110

Owing to its widespread success in prediction and interpretation, Breiman's RF has been under active development during the last two decades. For instance, Athey et al. (2019) presented generalized random forests for solving heterogeneous estimating equations. Friedberg et al. (2020) proposed a local linear forest model to improve the conventional RF in terms of smooth signals.
115 Ishwaran et al. (2008) introduced random survival forests, which can be used for the analysis of right-censored survival data. Wager and Athey (2018) developed a nonparametric causal forest for estimating heterogeneous treatment effects (HTE). Du et al. (2021) proposed another variant of random forests to help HTE inference through estimating some key conditional distributions. Katuwal et al. (2020) proposed several variants of heterogeneous oblique random forest employing
120 several linear classifiers to optimize the splitting point at the internal nodes of the tree. These new variants of RF are primarily focused on handling various regression and classification tasks or improving the predictive accuracy, yet their usefulness for interpretation has been minimally studied.

| | Deleted: the |
| Deleted: is still less |

125 In fact, many studies have reported that the feature importance methods used in Breiman's RF (including PFI and MDI) are unstable (i.e., a small perturbation of training data may significantly change the relative importance of predictors) (Bénard et al., 2021; Breiman, 2001b; Gregorutti et al., 2017; Strobl et al., 2007). Such instability has become one of the critical challenges for the practical use of current feature importance measures. Yu (2013) defined that statistical stability
130 holds if statistical conclusions are robust or stable to appropriate perturbations. In hydrology,

stability is critical in terms of interpretation and prediction. For interpretation, if a distinctive set of variable rankings was observed after a small perturbation of training data, one is unable to conclude realistic reasonings of hydrological processes. For prediction, there is no guarantee that the predictors with low rankings do not bear more valuable information than the higher ones. This problem challenges the selection of a subset of predictors for the optimum predictive accuracy (Gregorutti et al., 2017). Strobl et al. (2008) and Scornet (2020) disclosed that positively correlated predictors would lead to biased criteria selection during the tree deduction process, which further amplifies such instability. To address the issues mentioned above, Hothorn et al. (2006) proposed an unbiased node splitting rule for criteria selection. The proposed method showed that the predictive performance of the resulting trees is as good as the performance of established exhaustive search procedures used in CART. Strobl et al. (2007) examined Hothorn's method under the RF framework, which was called Cforest. They found that the bias of criteria selection can be further reduced if their method is applied using subsampling without replacement. Nevertheless, Xia (2009) found that Cforest only outperformed Breiman's RF in some extreme cases and concluded that RF was able to provide more accurate predictions and more reliable PFI compared to Cforest. A similar finding was also achieved by Fernández-Delgado et al. (2014), who reported RF was likely to be the best among 179 ML algorithms (including Cforest) in terms of predictive accuracy based on 121 data sets. More recently, Epifanio (2017) proposed a feature importance method called intervention in prediction measure (IPM), which was reported as a competitive alternative to other PFI and MDI. Since the proposed IPM was specifically designed for high-dimensional problems (i.e., the number of predictors is much larger than the number of observed samples), it is not suitable for most hydrological problems. Bénard et al. (2021) proposed a stable rule learning algorithm (SIRUS) based on RF. The algorithm (which aimed to remove the redundant paths of a decision tree) has indicated stable behavior when data is perturbed, while the predictive accuracy was not as good as the Breiman's RF. To sum up, the existing approaches do not guarantee stable and reliable variable ranking for robust interpretability and optimum predictive accuracy.

Therefore, as an extension of the previous efforts, the objective of this study is to develop a Wilks feature importance (WFI) method with improved variable rankings for supporting hydrological inference and modelling. WFI is based on an advanced splitting procedure, stepwise cluster

analysis (SCA) (Huang, 1992), which employed statistical significance of *F*-test, instead of least square fitting (used in CART), to determine the optimum splitting points. These points, in combination with the subsequent sub-cluster mergence, can eventually lead to the desired inference tree for variable rankings. The importance scores of predictors can then be obtained according to the values of Wilks Λ for reflecting the significance of differences between two or more groups of response variables. Compared with MDI and PFI, WFI does not rely on any performance measures (e.g., least-square errors in MDI or mean square errors in PFI), and can thus result in less biased criteria selection during the tree deduction process. Comparative assessment of WFI, PFI and MDI performances under the RF framework will then be undertaken through efforts in simulating monthly streamflows for 673 basins in the United States. With a finer temporal resolution, the proposed approach has also been applied to three irrigated watersheds in the Yellow River Basin, China, through concrete simulations of their daily streamflows.

2 Related Works

2.1. Random Forest

RF is an ensemble of decision trees, each of which is grown in accordance with a random subset of predictors and a bootstrapped version of the training set. As the ensemble members (trees) increase, the non-linear relationships between predictors and responses become increasingly stable. The prediction can thus be more robust and accurate (Breiman, 2001a; Zhang et al., 2018). The training set for building each tree is drawn randomly from the original training dataset with replacement. Such bootstrap sampling process will leave about 1/3 of the training dataset as out-of-bag (OOB) data, which thus can be used as a validation dataset for the corresponding tree.

There are many variants of RF according to the types of trees (e.g., CART). Based on splitting rules equipped in different types of trees, the resulting RF may use various feature importance measures. In this study, Breiman's RF is selected as the benchmark algorithm to investigate the feature importance measures. The algorithm is implemented using the R package "randomForest" (Liaw and Wiener, 2002). There are three hyperparameters in RF as the number of trees (*Ntree*), the minimum number of samples in a node (*Nmin*) for a splitting action, and the number/ratio of predictors in a subspace (*Mtry*). In addition, Breiman's RF has two feature importance measures: permutation feature importance (PFI) and mean decrease impurity (MDI).

2.2. Permutation Feature Importance

PFI was initially proposed by Breiman (2001a) and can be described as follows: Assume a trained decision tree t (where $t \in \{1,..., \textbf{\textit{ntree}}\}$; $\textbf{\textit{ntree}}$ is the total number of decision trees in the forest) with a subset of predictor $\textbf{\textit{u}}$ (where $\textbf{\textit{u}} \in \textbf{\textit{p}}$; and $\textbf{\textit{p}}$ is complete set of predictors), predictor matrix $\textbf{\textit{X}}$ (with full predictors), response vector $\textbf{\textit{Y}}$, predicted vector $\textbf{\textit{Y'}}$, and an error measure $\textbf{\textit{L(Y, Y')}}$; (1) calculate the original model error based on the OOB dataset of the t_{th} decision tree: $t(e_{orginal}) = L(Y, t(X^u))$ (where X^u is a subset of predictor matrix $\textbf{\textit{X}}$); (2) for each predictor $\textbf{\textit{j}}$ (where $\textbf{\textit{j}} \in \{1,..., \textbf{\textit{p}}\}$), (i) generate permuted predictor matrix $\textbf{\textit{X}}_{\textbf{\textit{perm,j}}}$ by duplicating $\textbf{\textit{X}}$ and shuffling the values of predictor $\textbf{\textit{X}}_{\textbf{\textit{j}}}$, (ii) estimate error for the permuted dataset $t(e_{perm, j}) = L(Y, t(X^u_{perm, j}))$; and (iii) calculate variable importance of predictor $\textbf{\textit{j}}$ for the t_{th} decision tree as $PFI(t)_j = t(e_{perm, j}) - t(e_{orginal})$; (note that $PFI(t)_j = 0$ if predictor $\textbf{\textit{j}}$ is not in $\textbf{\textit{u}}$); (3) calculate the variable importance for the forest by averaging the variable importance over all trees: $\textbf{\textit{PFI}}_{\textbf{\textit{j}}} = \dfrac{1}{ntree} \sum_{t=1}^{ntree} PFI(t)_j$. The error measure $\textbf{\textit{L(Y, Y')}}$ used in this study is mean squared error (MSE), given by:

$$MSE = \frac{1}{n} \sum_{n=1}^{N} \left(y_n - y_n^* \right)^2 \tag{1}$$

where y_n and y_n^* are the $\textbf{\textit{n}}^{\textbf{\textit{th}}}$ observed and predicted quantities, respectively; $\textbf{\textit{N}}$ is the total number of quantities.

2.3. MDI feature importance

The MDI importance measure is based on the CART decision tree, which is illustrated using a hydrological dataset (Figure 1) including 20 instances and 3 predictors as $\textbf{\textit{X}}_{\textbf{\textit{1}}}$ (i.e., precipitation), $\textbf{\textit{X}}_{\textbf{\textit{2}}}$ (i.e., 3-day cumulative precipitation) and $\textbf{\textit{X}}_{\textbf{\textit{3}}}$ (i.e., temperature), and a response $\textbf{\textit{Y}}$ (i.e., streamflow). It starts by sorting the value of $\textbf{\textit{X}}_{\textbf{\textit{j}}}$ in ascending order ($\textbf{\textit{j}}$ indicates the column index of the predictors so that $\textbf{\textit{j}} \in \{1, 2, 3\}$), and the $\textbf{\textit{Y}}$ will be reordered accordingly. Then we go through each instance of $\textbf{\textit{X}}_{\textbf{\textit{j}}}$ from the top to examine each candidate split point. For a sample set with $\textbf{\textit{k}}$ instances, the total number of split points for $\textbf{\textit{X}}_{\textbf{\textit{j}}}$ will be $\textbf{\textit{k-1}}$. Any instance z (where $z \in \{1, ..., \textbf{\textit{k}}\}$) in $\textbf{\textit{X}}_{\textbf{\textit{j}}}$ can split the predictor space into two subspaces as $\textbf{\textit{X}}_{\textbf{\textit{1}}}(\textbf{\textit{i, j}}) = \{X_{1,j}, X_{2,j}, ..., X_{z,j}\}$ (where $\textbf{\textit{i}} \in$

$\{1, …, z\}$); and $X_2(i, j) = \{X_{z+1,j}, X_{z+2,j}, …, X_{k,j}\}$ (where $i \in \{z+1, …, k\}$).The response space Y will be correspondingly divided into two subspaces as $Y_1(i) = \{Y_1, Y_2, …, Y_z\}$ (where $i \in \{1, …, z\}$); and $Y_2(i) = \{ Y_{z+1}, Y_{z+2}, …, Y_k\}$ (where $i \in \{z+1, …, k\}$). To maximize the predictive accuracy, the objective of the splitting process is to find the split point (based on the row and column coordinate z and j, respectively) with the minimum squared errors (SE) of Y_1 and Y_2:

$$SE(z, j) = \sum_{i=1}^{z}\left(Y_1(i) - \overline{Y_1}\right)^2 + \sum_{i=z}^{k}\left(Y_2(i) - \overline{Y_2}\right)^2 ; \ \forall z \text{ in } 1, …, k-1 ; \ \forall j \text{ in } 1, …, 3 \qquad (2)$$

where $\overline{Y_1}$ and $\overline{Y_2}$ indicate the mean value of Y_1 and Y_2, respectively.

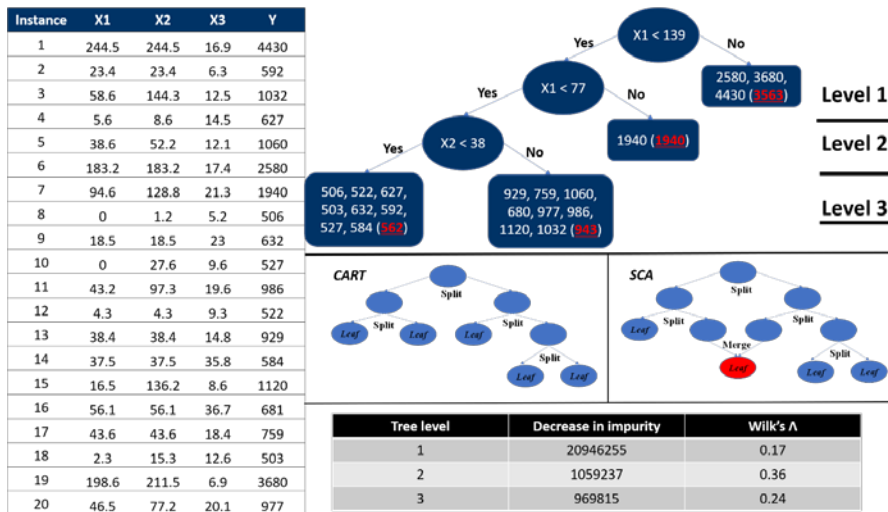



Figure 1: The table on the left is a numeric hydrological dataset; the figure on the top right is the tree deduction process for both CART and SCA with the dataset (note: the highlighted numbers in brackets of the leaf-nodes are the mean response values of those nodes; in this particular case, the two algorithms share the same node splitting rules, however, for most real-world cases, they lead to different decision trees); the figure on the middle right illustrates the distinct difference of deduction process between CART and SCA (not related to the case); the bottom-right table is the statistic summaries for CART and SCA of this synthetic case.

After each split, each of the newly generated subspaces can be further split using the same process as long as the number of instances in a subspace is greater than a threshold. This process will be repeated until reaching a stopping criterion, such as a threshold value by which the square errors must be reduced after each split.

245

The importance score of a particular predictor is measured based on how effective this predictor can reduce the square error in Eq. (1) for the entire tree deduction process (i.e., MDI). In the case of regression, "impurity" reflects the square error of the sample in a subspace (e.g., the larger the square error, the more "impure" the subspace is). The decrease in node impurity (DI) for splitting

250   a particular space s is calculated as:

$$DI(z, j, s) = \sum_{i \in 1,2,\ldots,k} \left(Y(i) - \overline{Y}\right)^2 - \frac{z}{k} \cdot \sum_{i \in 1,2,\ldots,z} \left(Y_1(i) - \overline{Y_1}\right)^2 - \frac{k-z}{k} \cdot \sum_{i \in z+1, z+2,\ldots,k} \left(Y_2(i) - \overline{Y_2}\right)^2 \qquad (3)$$

where z and j are the coordinates for the optimum splitting point of space s, k is the number of instances in space s and $\overline{Y}$ is the mean value of $Y(i)$ in space s. Therefore, the Mean Decrease Impurity (MDI) for the variable X_j computed via a decision tree is defined as:

255    $$MDI(X_j) = \sum_{s \in S; j=j} P_s \cdot DI(z, j, s) \qquad (4)$$

where S is the total spaces in a tree, P_s is the fraction of instances falling into s. In other words, the MDI of X_j computes the weighted DI related to the splits using the j_{th} predictor. MDI computed via RF is simply the average of the MDI computed via each tree of the forest. The ensemble (i.e., average) of important scores from the forest is assumed to be more robust than the individual tree.

260   3. Wilks Feature Importance

WFI is based on the stepwise cluster analysis (SCA) algorithm (Huang, 1992). The fundamental difference between WFI and MDI comes from the split criterion and the tree deduction process. Let us recall the split criterion of CART, in which the optimum split point for X_j is located based on the minimum squared errors of Y_1 and Y_2 as shown in Eq (1). In WFI, this function is achieved

265   by comparing the two subspaces' (i.e., Y_1 and Y_2) likelihood, which is measured through the Wilks Λ statistics (Nath and Pavur, 1985; Wilks, 1967). It is defined as $\Lambda = Det(W)/Det(B+W)$, where $Det(W)$ is the determinant of a matrix, W and B are the within- and between-group sums of squares

and cross-product matrices in a standard one-way analysis of variance, respectively. The *W* and *B* can be given by:

$$W = \frac{z \cdot (k-z)}{k} (\overline{Y}_1 - \overline{Y}_2)' \cdot (\overline{Y}_1 - \overline{Y}_2)$$ (5)

$$B = \sum_{i=1}^{z} \left[Y_1(i) - \overline{Y}_1 \right]' \cdot \left[Y_1(i) - \overline{Y}_1 \right] + \sum_{i=1}^{k-z} \left[Y_2(i) - \overline{Y}_2 \right]' \cdot \left[Y_2(i) - \overline{Y}_2 \right]$$ (6)

The value of Λ is a measure of how effective X_j can differentiate between Y_1 and Y_2. The smaller Λ value representing a larger difference between Y_1 and Y_2. The distribution of Λ is approximated by Rao's *F*-approximation (*R*-statistic), which is defined as:

$$R = \frac{1 - \Lambda^{1/S}}{\Lambda^{1/S}} \cdot \frac{Z \cdot S - d \cdot (m-1)/2 + 1}{d \cdot (m-1)}$$ (7)

$$Z = k - 1 - (d+m)/2$$ (8)

$$S = \frac{d^2 \cdot (m-1)^2 - 4}{d^2 + (m-1)^2 - 5}$$ (9)

where statistic *R* is distributed approximately as an *F*-variate with $n_1 = d \cdot (m-1)$ and $n_2 = d \cdot (m-1)/2 + 1$ degrees of freedom; *m* is the number of groups. Since the number of groups is two in this study, an exact *F*-test is possibly performed based on the following Wilks Λ criterion be:

<!-- Deleted: Wilks' -->

$$F(d, k-d-1) = \frac{1-\Lambda}{\Lambda} \cdot \frac{k-d-1}{d}$$ (10)

Therefore, the two subspaces can be compared for examining significant differences through the *F*-test. The null hypothesis would be $H_0: \mu(Y_1) = \mu(Y_2)$ versus the alternative hypothesis $H_1: \mu(Y_1) \neq \mu(Y_2)$, where $\mu(Y_1)$ and $\mu(Y_2)$ are population means of Y_1 and Y_2, respectively. Let the significance level be α, the split criterion would be: $F_{cal} \geq F_\alpha$ and H_0 are false, which implies that the difference between two subspaces is significant and thus they should be split.

The second difference between the CART and SCA algorithms lies in the tree deduction procedure. In CART, the splitting process will be repeated until any newly generated subspace can no longer be split. In SCA, once all the nodes in the current stage have been examined for splitting, merging will be followed in the next stage, as illustrated in Figure 1. The merging process will compare any pairs of nodes based on the value of Wilks Λ to test if they can be merged (i.e., for $F_{cal} < F_\alpha$ and H_0 are true, which indicates that these two subspaces have no significant difference thus should

<!-- Deleted: Wilks' -->

be merged). Such splitting and merging processes are iteratively performed until no node can be further split or merged. Once an SCA tree is built, the WFI for the variable X_j computed via an SCA tree is defined as:

$$WFI\left(X_j\right)= \sum_{s\in S; j=j} P_s \cdot\left(1\text{-}\Lambda(z, j, s)\right) \tag{11}$$

where S is the total spaces in a tree, P_s is the fraction of instances falling into s, Λ (z, j, s) denotes the value of Λ obtained at the optimum splitting point of space s with row and column coordinates z and j, respectively. Similar to the calculation of MDI in Eq. (3), the WFI for X_j computes the weighted (1-Λ) value related to the splits using the j_{th} predictor.

According to the law of large numbers, WFI is expected to perform better under the RF framework since the randomized predictors ensure enough tree diversity, leading to more balanced importance scores. Therefore, we name the ensemble of SCA as the stepwise clustered ensemble (SCE). In addition to the three hyperparameters (i.e., **Ntree**, **Nmin** and **Mtry**) for Breiman's RF, SCE also requires significance level (α), which is used for the F-test during the node splitting process.

There could be two potential advantages of WFI over MDI. First, the decrease in node impurity (DI) will become smaller and smaller as long as the tree level goes down (as shown in the bottom-right table in Figure 1). Such a mechanism naturally assumes that the predictors considered (for node splitting) in lower levels of the tree are less significant than those in upper levels. This effect is even aggravated by the existence of predictor dependence, which will depress the importance scores of independent predictors and increase the positively dependent ones (Scornet, 2020). As a consequence, some critical predictors may only receive small importance scores. In comparison, Wilks Λ is a measure of the separateness of two subspaces, which could avoid the above-mentioned issue for MDI because values of (1-Λ) do not necessarily decline as long as the tree level goes down (as shown in the bottom-right table in Figure 1). Therefore, the predictors that are primarily considered in later splits may have higher importance scores than those in earlier splits. As a consequence, some critical predictors might be identified by WFI but overlooked by MDI. Second, the node splitting mechanism of WFI is based on the F-test, which, therefore, may significantly reduce the probabilities that the two child-nodes are split due to chance. Such a mechanism could be helpful to build more robust input-output relationships for prediction and

inference by reducing overfitting. The above-mentioned potential advantages of WFI will be tested with a large number of hydrological simulations in the following two sections.

4. Comparative studies over the NCAR CAMELS dataset

4.1. Dataset description

335   The Catchment Attributes and Meteorological (CAMELS) dataset (version 1.2) (Addor et al., 2017; Newman et al., 2015) was used to evaluate the WFI performance. The dataset contains daily forcing and hydrologic response data for 673 basins across the contiguous United States that spans a very wide range of hydroclimatic conditions (Figure 2) (Newman et al., 2015). These basins range in size between 4 and 25,000 km^2 (with a median basin size of 336 km^2) and have relatively

340   low anthropogenic impacts (Kratzert et al., 2019b).

In attempting to demonstrate the relative importance of meteorological data and large-scale climatic indices on streamflow, we used monthly mean values of meteorological data in CAMELS dataset and 4 commonly used large-scale climatic indices (including Nino3.4 (Trenberth, 1997),

345   Pacific decadal oscillation (PDO) (Mantua et al., 1997), interdecadal Pacific oscillation (IPO) (Mantua et al., 1997) and Pacific North American index (PNA) (Leathers et al., 1991)) to simulate the monthly streamflows. To reflect the initial catchment conditions and lagged impact of climatic indices, the 2-month moving average meteorological data and climatic indices of the preceding 2 months were incorporated as model predictors. Therefore, the input-output structure (with 22

350   predictors) for each of these basins can be written as follows:

$$Q_t = f \begin{pmatrix} Pr_t, Rad_t, Tmax_t, Tmin_t, Vp_t, \left(Pr_t + Pr_{t-1} \right)/2, \left(Rad_t + Rad_{t-1} \right)/2, \\ \left(Tmax_t + Tmax_{t-1} \right)/2, \left(Tmin_t + Tmin_{t-1} \right)/2, \left(Vp_t + Vp_{t-1} \right)/2, \\ Nino3.4_t, Nino3.4_{t-1}, Nino3.4_{t-2}, PDO_t, PDO_{t-1}, PDO_{t-2}, \\ IPO_t, IPO_{t-1}, IPO_{t-2}, PNA_t, PNA_{t-1}, PNA_{t-2} \end{pmatrix} \quad (12)$$

where Q_t represents streamflow of month t. Pr, Rad, $Tmax$, $Tmin$ and Vp represent monthly values of precipitation, short-wave radiation, maximum temperature, minimum temperature and vapor pressure, respectively.

355

Hydrological Region	Number of basins
1	27
2	75
3	92
4	31
5	45
6	17
7	33
8	12
9	9
10	72
11	31
12	37
13	7
14	17
15	19
16	18
17	91
18	40

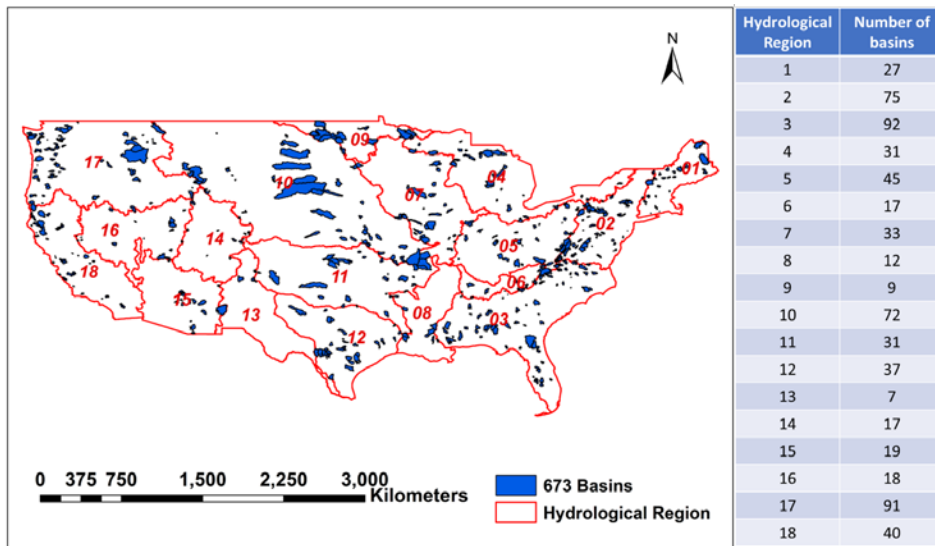



Figure 2. Overview of the basin locations and corresponding hydrological regions. This map was created using ArcGIS software (Esri Inc. 2020).

360

4.2. Evaluation procedures and metrics

The model training was performed based on January 1980 to December 2005, while the testing was done based on the period of January 2006 to December 2014. The hyperparameters for both RF and SCE were set as follows: *Ntree* was set as 100, *Nmin* was set as 5, and *Mtry* was set as 0.5 as suggested by Barandiaran (1998), indicating half of the predictors were selected in each tree. In addition, the significance level (α) was set as 0.05 for the *F*-test in SCE.

The performance of WFI will be evaluated and compared against PFI (applied to RF and SCE) and MDI (applied to RF). To improve the stability of the PFI results, previous studies have suggested repeating and average the PFI over repetitions (Molnar, 2020). In this study, the PFI process was repeated 10 times and then averaged for stabilizing the results. To facilitate the comparisons among different variable rankings, importance scores from the three feature importance methods were scaled into the [0,1] range. All the feature importance methods will be evaluated through recursive feature elimination (RFE) (Guyon et al., 2002) as follows: (1) train

SCE and RF models with all predictors; (2) calculate the importance scores using the three interpretation methods embedded in their corresponding models; (3) exclude the three least relevant predictors for each set of the importance scores obtained in step 2; (4) retrain the models using the remaining predictors in step 3; (5) repeat step 2 to 4 until the number of predictors is less or equal to a threshold (set to 4 in this case study). To directly compare the quality of variable rankings from different feature importance measures, the selected predictors from WFI (after every RFE iteration) were also used to train RF. This procedure allows the effects of different variable rankings to be solely from feature importance methods (i.e., removed the effects from different node splitting algorithms). The same procedure was also performed for SCE-based PFI (i.e., SCE-PFI) to examine whether the differences in variable rankings are from the WFI method or the tree deduction process in SCE.

Two error metrics (i.e., adjusted R^2 and RMSE) were used to evaluate the model performance. Adjusted R^2 has been used instead of R^2 because adjusted R^2 can consider the number of predictors. Adjusted R^2 is defined as:

$$adj\ R^2 = 1 - \frac{\left(1 - R^2\right)\left(N - 1\right)}{N - P - 1} \tag{13}$$

where P is the number of predictors and N is the number of instances.

RMSE is defined as:

$$RMSE = \sqrt{\frac{1}{n}\sum_{n=1}^{N}\left(y_n - y_n^*\right)^2} \tag{14}$$

where y_n and y_n^* are the nth observed and predicted streamflow values, respectively.

To evaluate the stability of a feature importance method, we consider reducing predictors during the RFE iterations as a form of perturbation to the dataset. Suppose the obtained importance scores for a dominant predictor shows an irregular changing pattern during the RFE iterations. In that case, the method is not stable because it leads to many versions of inferences for such a predictor. On the other hand, if such a changing pattern is predictable (e.g., monotonically increasing trend), a stable inference can be achieved among interactions because the predictable pattern can help analyze how a predictor reacts to the change in the dataset. In this study, the monotonicity is

examined by using the Spearman's rank correlation coefficient (i.e., Spearman's ρ), which is commonly used to test the statistical dependence between the rankings of two variables and is defined as:

$$\rho = \frac{\sum_i \left(RX_i - \overline{RX}\right)\left(RY_i - \overline{RY}\right)}{\sqrt{\sum_i \left(RX_i - \overline{RX}\right)^2 \left(RY_i - \overline{RY}\right)^2}} \tag{15}$$

where RX_i is the ranks of variables X for the i_{th} RFE iteration and RY_i is the number of selected predictors for the i_{th} RFE iteration; $\overline{RX}$ and $\overline{RY}$ are the means of RX_i and RY_i, respectively. A larger Spearman's ρ value indicates the importance score for a predictor will increase along with the reduction of irrelevant predictors, leading to stable importance scores.

4.3. Predictive Accuracy and Interpretation Stability

Figure 3 shows the model testing performances (adjusted R^2) for 18 hydrological regions with all 22 predictors. The results show that SCE and RF significantly outperform SCA and CART, respectively. When taking a close look at these two pairs of model performance, SCA and CART are close to each other, while SCE outperforms RF in most hydrological regions (except the 9th region).

The pairwise comparisons of these four algorithms over 673 basins show a high coefficient of determination (0.913) of adjusted R^2 between SCE and RF, and an even higher coefficient of determination (0.965) between SCE and RF (Figure 4). This result indicates that, in general, it is not likely to have a distinct performance gap for a particular simulation task either between SCA and CART, or between SCE and RF. Therefore, SCA/SCE can be a good substitute for CART/RF.

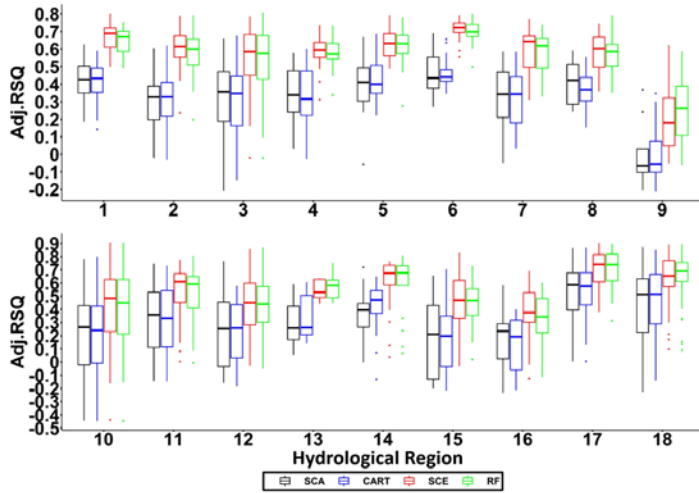



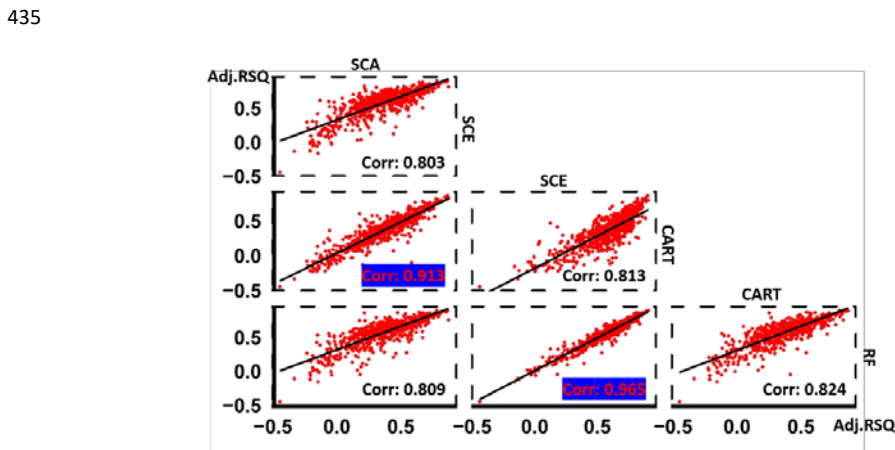
Figure 3. Adjusted R^2 for 18 hydrological regions. Each box indicates statistical summaries (i.e., the bars represent median value; the lower and upper boundaries of a box represent 1st and 3rd quantiles, respectively; dots represent outliers) of adjusted R^2 for all the basins in a particular hydrological region.

435

Figure 4. Pairwise comparison for adjusted R^2 over 673 US basins.

The left column in Figure 5 shows simulation performances based on RFE iterations for three feature importance measures embedded in SCE and RF. In general, both models can improve their

440 simulation performance by eliminating irrelevant predictors. When the number of predictors reduces to 7 (i.e., at the 5th iteration), both models reach their highest predictive accuracy over the OOB and testing dataset. This result indicates that it is plausible to use the OOB dataset to identify the optimum subset of predictors. Comparing the simulation performance for the training period, the simulation performance for SCE is much lower than it for RF, while an opposite result is

445 observed for the testing period. This result highlights the issue of overfitting for RF. One exception where RF outperforms SCE (for the testing period) happens in the last (i.e., 6th) iteration, where RF with MDI selected predictors outperforms SCE with WFI selected ones. We can assume that RF may have a better chance to outperform SCE with insufficient predictors. Nevertheless, SCE owns the overall best performance with PFI-selected predictors.

450 The upper left panel in Figure 6 shows that from 0th to 5th iterations, over 55% to 60% of basins (as indicated in yellow diamonds) simulated by SCE with WFI selected predictors outperform those simulated by RF with MDI selected predictors. In comparison, the number drops to about 40% at the 6th iteration. This result agrees with the results shown in Figure 5. The lower left panel in Figure 6 shows that from the 1st to 5th iterations, there is a higher chance that SCE with PFI

455 selected predictors outperforms RF with MDI selected ones for over 75% of the hydrological regions (as indicated in Figure 6 that the black boxes are above the blue line).

The reliability of variable rankings of WFI is further investigated by using the WFI and SCE-PFI selected predictors from each of the RFE iterations to run RF models. The results are shown in the right column in Figure 5. The RF simulations with WFI selected predictors owned the highest

460 predictive accuracy in most RFE iterations over the training, OOB validation and testing datasets. In particular, the WFI selected predictors have shown significant strength in the last two iterations and facilitated RF to improve its predictive accuracy. It is worth mentioning that even though SCE-PFI selected predictors allowed SCE to achieve its optimum performance, they did not deliver optimum performance for RF. This result shows WFI selected predictors provide a better universal

465 solution than the PFI-selected ones.

The upper left panel in Figure 7 shows that a majority of basins simulated by RF with WFI selected predictors outperform those simulated by RF with MDI selected predictors. In particular, at the 6th iteration, basins in 16 (out of 18) hydrological regions may own better performance with WFI selected predictors than with the MDI-selected predictors. In addition, as the number of predictors

decreases, there are increasing chances that WFI selected predictors could generate better

485 performance than the MDI-selected ones. Based on a two-sided Mann-Kendall (M-K) trend test (Kendall, 1948; Mann, 1945), such an increasing trend is significant (i.e., the Z score equals 2.63 and p-value is smaller than 0.01). Another significant increasing trend (i.e., the Z score equals 1.88 and p-value equals 0.06) also can be observed for the paired studies of WFI and RF-PFI. In contrast, no significant increasing trend can be observed for the pairs of SCE-PFI and MDI, as well as SCE-

490 PFI and RF-PFI. This finding indicates WFI could generate robust variable rankings, based on which informative predictors are more likely to be kept for optimum simulation performance. In contrast, other feature importance measures may lose critical predictors during the RFE process.

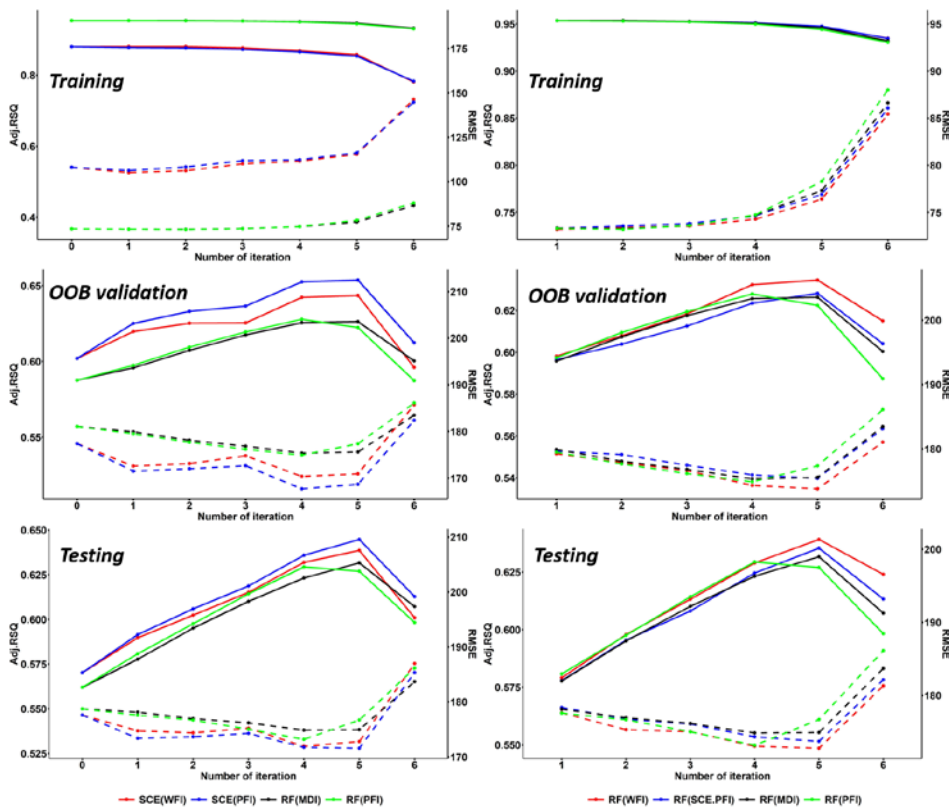



Figure 5. Iterative change in accuracy in mean values of 673 US basins. The solid lines indicate adjusted R^2, while the dashed lines represent RMSE. The Figures on the left column show SCE and RF performances based on variables selected by themselves, while the figures on the right show RF model performances based on variables selected by SCE and RF. The models with an iteration number of 0 represent the model with all 22 predictors.

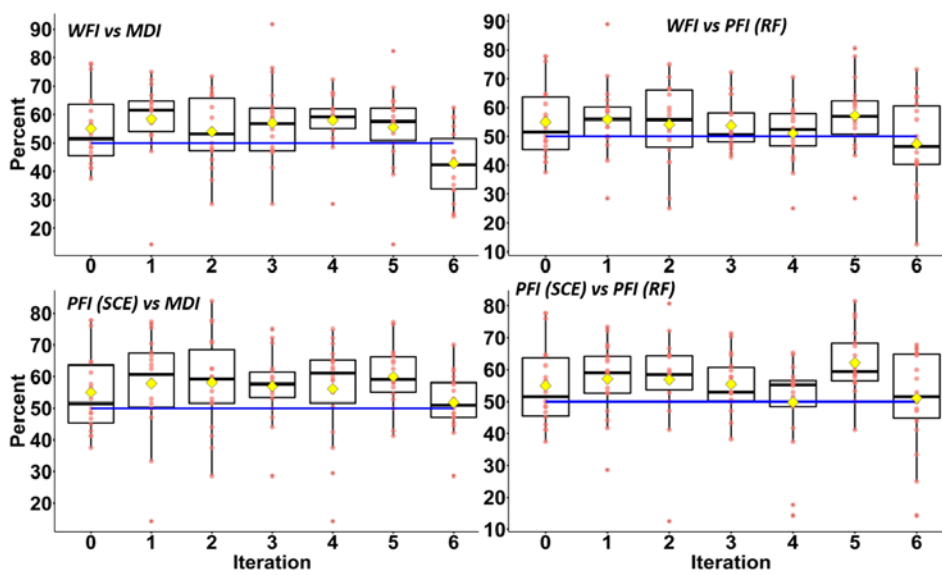



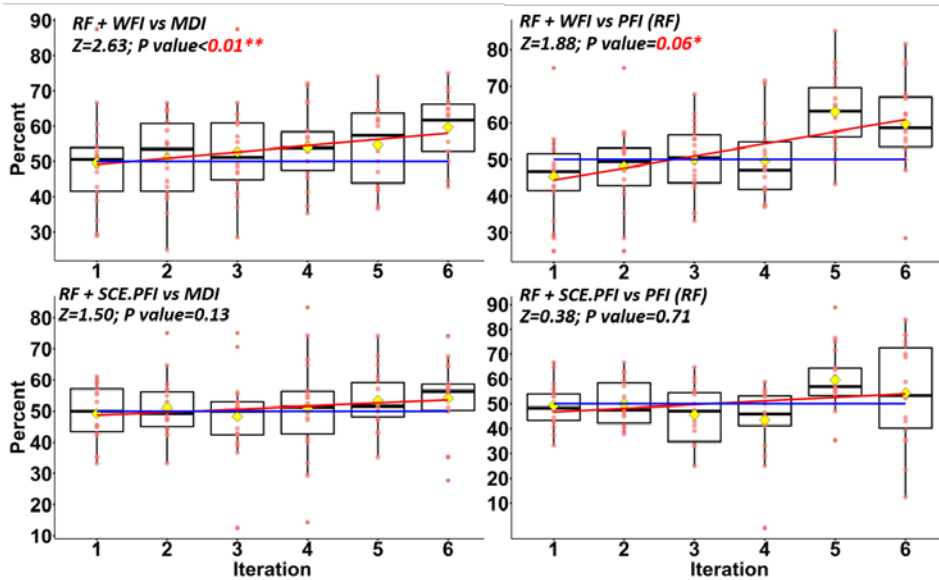
Figure 6. Pairwise comparisons of model performance with different feature importance measures. Each of these red points represents the percentage of basins simulated by model A outperforming model B (based on adjusted R^2), in one particular hydrological region. The blue line represents 50% percent, and the yellow square represents the mean percentage of 18 hydrological regions.

515

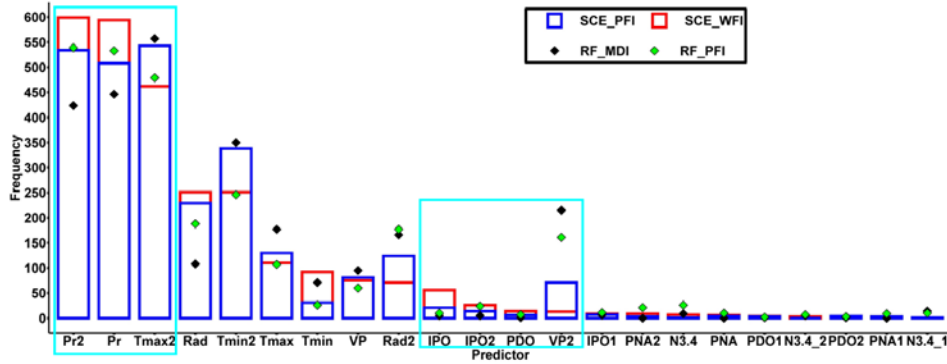
Figure 7. Pairwise comparisons of RF model performance with different feature importance measures. *Z* scores and *P* values are calculated based on the two-sided Mann-Kendall trend test. If the *Z* score is greater than *1.96*, an increasing trend can be assumed with a significance level of *0.05*. If the *Z* score is greater than *1.645*, an increasing trend can be assumed with a significance level of *0.1*. Other notations are the same as those in Figure 6.

520

Figure 8 shows the summaries of selected predictors (in the last iteration) with different feature importance measures. *Pr* (monthly precipitation at time step *t*) and *Pr2* (mean values for monthly precipitation at time step *t* and *t-1*) are considered the two most important predictors for the SCE algorithm with WFI selected predictors. In contrast, MDI considers *Tmax2* (mean values for the monthly maximum temperature at time step *t* and *t-1*) as the most important predictor for monthly

525 streamflow simulation. It is acknowledged that streamflow is more responsive to precipitation than air temperature. Therefore, we can assume that RF may capture more acuate responses of streamflow with WFI selected features than MDI or PFI selected ones. This assumption could be one of the reasons that RF with WFI selected predictors outperforms the others. It should be noted

530 that IPO is considered as an important predictor for 56 out of 673 basins with WFI, while this

predictor has only been employed in 21, 5 and 10 basins with SCE-PFI, MDI and RF-PFI methods, respectively.

 Figure 8. Summaries of the predictors used in last iterations for 673 US basins.

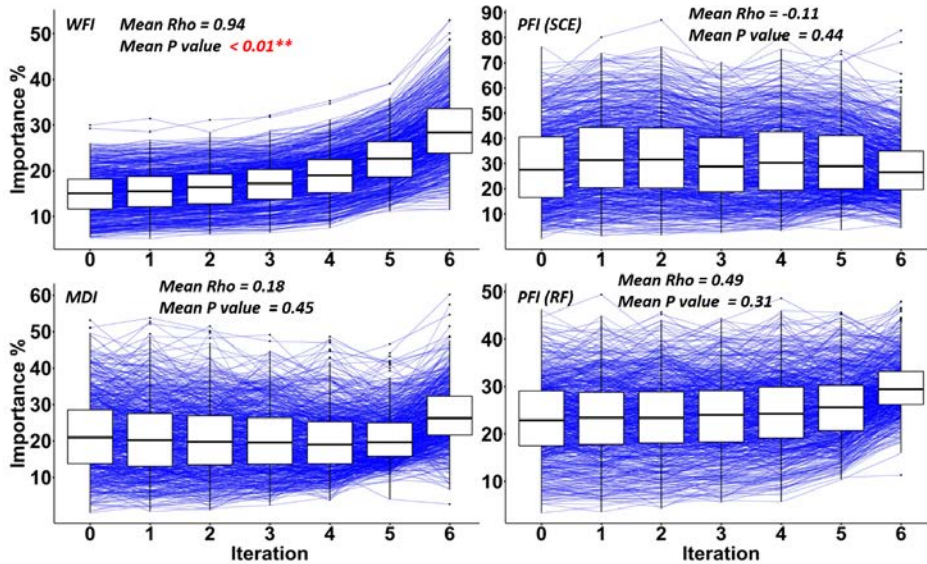
The Spearman's Rho (ρ) values for the predictor with the highest importance score (at the last RFE iteration) illustrate the stability of all three interpretation methods embedded in SCE and RF (Figure 9). The results indicate that the importance score for the predominant predictor increases in response to the reduction of irrelevant predictors. Compared with other feature importance measures, WFI owns the highest ρ values in general with the *p*-value less than 0.01, indicating a significant correlation between the importance score and the reduction of irrelevant features. In comparison, eliminating irrelevant predictors will significantly influence the importance score of predominant predictors obtained by PFI and MDI. This fact challenges the application of the PFI and MDI since the removal of irrelevant predictors cannot guarantee the same or similar level of hydrological inference because the importance score may distinctly vary according to the reduction of irrelevant predictors. In contrast, the WFI method provides more stable importance scores and will lead to more consistent hydrological inferences.

550

Figure 9. Mean Spearman's ρ values for the most important features. The mean p-value means how likely it is that the observed correlation is due to chance. Small *p*-values indicate strong evidence for the observed correlations.

555 **5. Application of WFI over irrigated watersheds in the Yellow River Basin, China**

5.1. Study Area and Data

Daily streamflow simulations for three irrigated watersheds located in the alluvial plain of the Yellow River in China were conducted to test the capability of the proposed WFI method at a finer temporal resolution. These watersheds share a total area of 4,905 km^2, consisting of 52% irrigated
560 land, 17% residential area, 15% desert, 12% forested land, and 4% water surface (Figure 10). The landscape of the study area is characterized by an extremely flat surface with an average slope ranging from 1:4000 to 1:8000, with mostly highly permeable soil (sandy loam). The climatic condition of the study area is characterized by extreme arid environments with annual precipitation ranging from 180 to 200 mm, and annual potential evaporation ranging from 1,100 to 1,600 mm
565 (Yang et al., 2015).

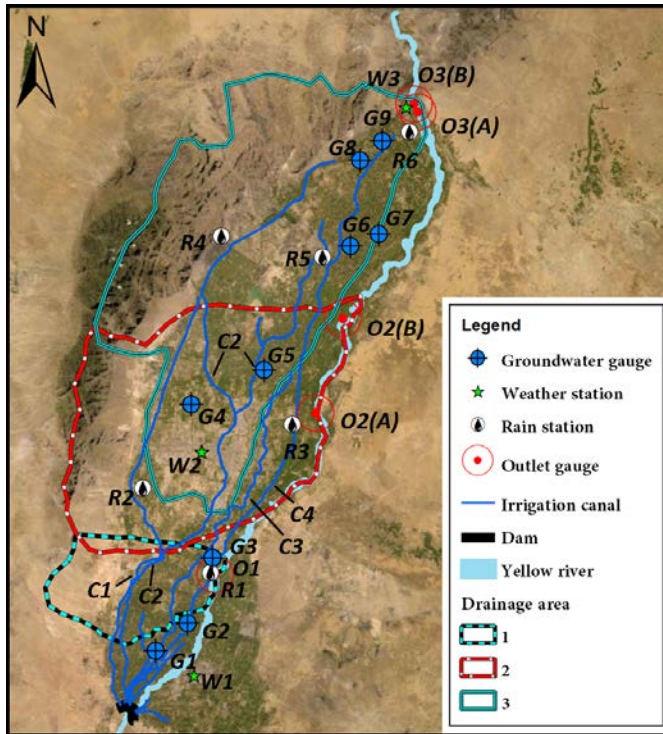



Figure 10: Map of the study area. Note: due to the extremely flat surface, three interconnected irrigated watersheds are approximately delineated. In this map, *G* indicates groundwater gauges, *W* indicates weather stations, *R* indicates rain stations, *C* indicates irrigation canals and *O* indicates drainage outlets. Both 2nd and 3rd irrigated watersheds contain two crisscrossed drainages with strong hydrological connections. The map was created using ArcGIS software (Esri Inc. 2020).

Initial catchment conditions were also considered in this case study to improve the model performance. Specifically, moving sums of daily precipitation, temperature and evaporation time series over multiple time periods $\delta_{P,T,E} = [1, 3, 5]$ prior to the date of predictions were set as predictors to reflect the antecedent watershed conditions. Similarly, the moving window for daily irrigation time series were set as $\delta_I = [1, 3, 5, 7, 15, 30]$. In addition, daily groundwater level data are used as additional predictors to reflect the baseflow conditions of the catchments. The daily time-series data were divided into two subsets: one from 2001/01/01 to 2011/12/31 for model

training and OOB validation and the other from 2012/01/01 to 2015/12/31 for model testing. Table 1 list the weather, rain and groundwater stations used for each basin. The streamflow processes show distinct behaviors in terms of flow magnitude and duration due to the different irrigation schedules in spring and winter. To analyze such temporal variations, daily streamflow for Spring-Summer (April to September) and Autumn-Winter (October to March) were examined separately. In this case study, the hyperparameters for RF and SCE are used as same as those for section 4.

Table 1: Weather, rain and groundwater gauges, and irrigation canals used in each irrigation basin.

Watershed ID	Stations/Canals	Outlets
1	*C1*, *C2*, *C3*, *W1*, *R1*, *G1*, *G2*, *G3*	*O1*
2	*C1*, *C2*, *C3*, *C4*, *W2*, *R2*, *R3*, *R5*, *G4*, *G5*	*O2(A)+ O2(B)*
3	*C1*, *C2*, *C4*, *W2*, *W3*, *R4*, *R5*, *R6*, *G4*, *G5*, *G6*, *G7 G8*, *G9*	*O3(A)+ O3(B)*

Note: Streamflow for each watershed is integrated as the sum of the gauged streamflows within this area.

5.2. Results Analysis

Generally, SCE and RF delivered reasonable predictive accuracy (using all considered predictors) across all watersheds and seasons (Table 2). The SCE approaches the best overall predictive accuracy for the testing dataset. Compared with RF, the SCE has a smaller drop in predictive accuracy from the training to testing period, indicating the SCE algorithm captured a more robust input-output relationship during the training period. This result agrees with those for the large-scale dataset in section 4. The convergence tests for training, OOB validation, and testing datasets were shown in Figures S1, S2 and 11, respectively. The results from the testing period (Figure 11) show that SCE always outperforms RF as the number of trees increases.

Table 2: The adjusted R^2 for SCE and RF with all considered predictors.

Basin	Season	Training		OOB		Testing	
		SCE	RF	SCE	RF	SCE	RF
1st	spring	94.1%	98.2%	86.5%	88.2%	**82.0%**	81.4%
1st	winter	97.9%	99.2%	94.3%	95.1%	**91.3%**	90.0%
2nd	spring	94.0%	98.4%	85.7%	89.0%	**76.7%**	75.7%
2nd	winter	97.6%	99.3%	94.6%	95.7%	**66.0%**	65.1%
3rd	spring	93.8%	98.3%	84.5%	87.7%	**68.5%**	68.1%
3rd	winter	97.8%	99.2%	95.2%	95.3%	**82.7%**	82.1%

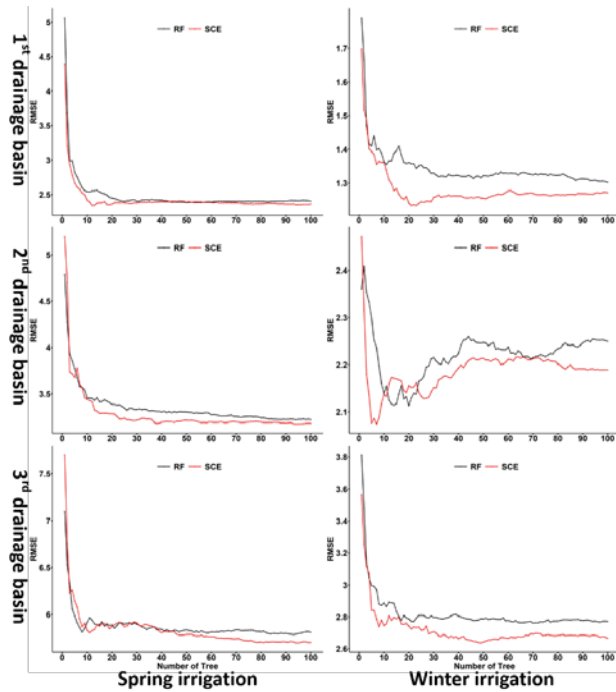



Figure 11: Convergence of the SCE and RF model based on RMSE over the testing period.

695 The iterative reductions in accuracy for training, OOB validation and test datasets are listed in Figure S3, S4 and S5, respectively. The summary (Figure 12) shows that WFI owns the smallest reduction in accuracy (for both adjusted R^2 and RMSE) over the testing period, followed by SCE-PFI, MDI and RF-PFI. A smaller reduction in accuracy means the selected predictors are more informative in describing the complex relationships of hydrological processes. As a consequence,

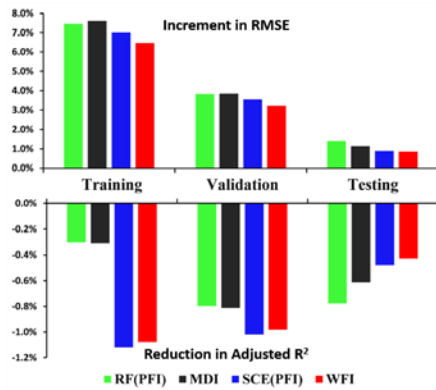
700 WFI can identify the most informative predictors compared with other methods. Figure 12 also shows that over the training period, RF receives a much smaller impact from RFE in terms of adjusted R^2 compared with SCE, because the least-square fittings employed in the CART training process pursue the highest R^2 over the training period.

705

Figure 12: Change in predictive accuracy averaged across three watersheds and two seasons.
Note: the change in predictive accuracy (e.g., RMSE) for a particular case is calculated using
(RMSE (last iteration) - RMSE (full model)) / RMSE (full model).

| | Deleted: as the accuracy for the |
| Deleted: minus it for the |
| Deleted: predictors. |

710 Figure 13 shows the Spearman's ρ values for the most relevant predictor (i.e., with the highest
importance score in the last RFE iteration). The result indicates that WFI has the highest absolute ρ
values for the majority of the cases. This result agrees with those demonstrated in section 4. In fact,
the highest absolute Spearman's ρ values for the rest of the relevant predictors (selected for the last
RFE iteration) mainly belong to the WFI method (as shown in Figure 14), which further illustrates
715 that WFI could provide stable relative importance values for essential predictors for hydrological
inference.

| Deleted: of |
| Deleted: owns |
| Deleted: among |

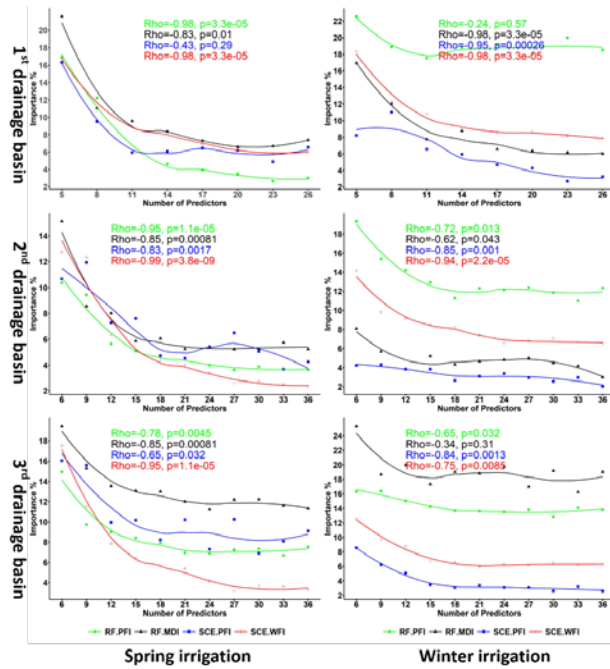



Figure 13: Spearman's ρ values for the most important predictor. Note: the most important predictor is the predictor with the highest importance score in the last RFE iteration. The *p*-value means how likely it is that the observed correlation is due to chance. Small p-values indicate strong evidence for the observed correlations.

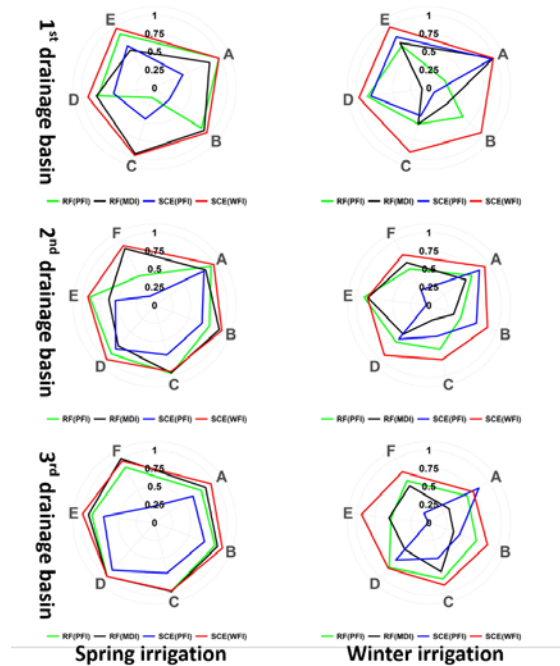



Figure 14: Spearman's ρ values for three watersheds and seasons. Note: the RFE process of this case study keeps at least five and up to seven of the most relevant predictors in the last iteration, according to the remainder of the total considered predictors divided by three. Capital letters from A to F represent the most relevant predictors identified by different feature importance methods.

The importance scores were aggregated and analyzed according to different types (i.e., precipitation, irrigation, evaporation, etc.) to explore the relationships between the hydrological responses and their driving forces. We chose the models with the smallest RMSE (among all the RFE iterations) on the testing dataset for the hydrological inference. The results indicate the importance scores differed significantly according to the algorithms and interpretation methods used (Figure 15). In particular, the aggregated predictor P1 (i.e., daily precipitation for timestep t from all spatial locations) owns positive contributions (in reducing the RMSE) for WFI in the Spring irrigations. At the same time, it has merely no contribution for other feature importance methods. To investigate whether the predictors identified by WFI are also meaningful to other

algorithms, we reinserted the predictors in P1 into the best RF model (in which the set of predictors reaches the smallest RMSE over the testing dataset). Indeed, we found the RF with reinserted predictors showing slightly improved predictive accuracy (i.e., RMSE and adjusted R^2) for Spring irrigations across all watersheds on the testing dataset (Table 3). This result illustrates that even though the predictors in P1 have no contribution in improving the predictive accuracy on the training dataset, it can potentially distinguish different hydrological behavior (i.e., with a small Wilks Λ value) and lead to improved model performance on the testing dataset. In fact, the time of concentration for these basins is usually less than one day if the storm falls near the outlets of the irrigation basins. This fact proves the above hydrological inference is reasonable.

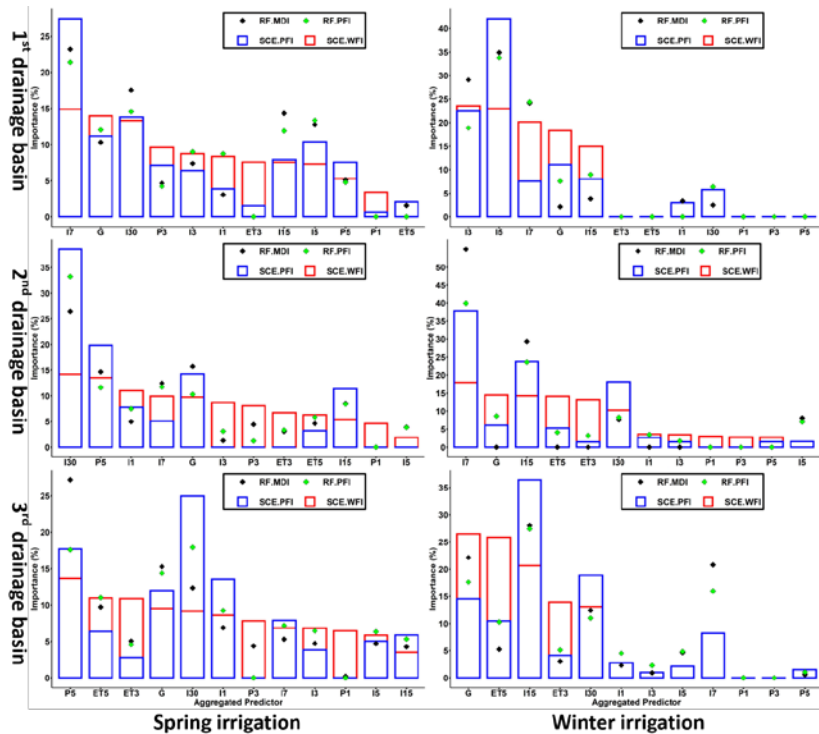



Figure 15: Importance scores aggregated by predictor types. Note: each type of predictor includes predictors from all considered spatial locations. For example, P1 includes predictors for all the

considered climatic stations with 1-day precipitation. Therefore, the importance score of P1 is the average of the importance score from the predictors of P1.

760

Table 3: Predictive accuracy for reinserting the predictors in P1 to the RF model (Spring irrigation).

	Basin	RF with P1	RF without P1
RMSE	1st	**2.42**	2.44
	2nd	**3.16**	3.17
	3rd	**5.81**	5.81
Adjusted R^2	1st	**0.81**	0.81
	2nd	**0.77**	0.76
	3rd	**0.69**	0.69

Note: The RF model was based on the optimum set of predictors in RFE iterations.

6. Discussion

There could be several reasons why WFI can have more robust variable rankings than other feature
765 importance measures. First, WFI does not rely on performance measures to evaluate the variable
importance. Instead, it depends on Wilks Λ, which prevents any node splitting due to chance. In
the node splitting process, a predictor that significantly increases the predictive accuracy may not
necessarily have the ability to differentiate two potential sub-spaces. Therefore, the WFI method
(which evaluates every splitting and merging action based on Wilks test-statistics with the
770 predefined significance level α) is expected to generate more robust variable rankings. Secondly,
WFI considers all the interactions among predictors in the tree deduction process, while PFI can
only consider the effect of one predictor at a time. Thus the interactions between the target
predictor and the rest of the predictors are overlooked. For example, in section 4, the SCE-PFI
selected predictors achieved higher performance (over the testing dataset) than the WFI selected
775 ones. However, these SCE-PFI selected predictors are model-specific, which means when
transferring these predictors to the other model (e.g., RF), they may not deliver the optimum
performance. In contrast, the WFI selected predictors have good transferability: they helped the
RF model achieve optimum predictive accuracy. Similar evidence was also found by Schmidt et
al. (2020), who reported that the variable rankings from PFI might vary significantly according to
780 different algorithms. This fact has been considered a major challenge for hydrological inference
because one cannot reach the same reasoning with different algorithms. Based on the results above,
we can conclude that the WFI could produce more robust variable rankings, which enables a
universal solution rather than a specific one for hydrological inference.

790

RFE was used to identify the most relevant predictors for optimum predictive accuracy. This approach could be quite useful in real-world practice, especially in hydrology, where the simulation problem may involve hundreds of inputs (from climate models, observations or remote sensing, etc.), describing the spatial and temporal variabilities of the system. Each of these inputs may contain useful information, and may also contain noise that will mislead the model (e.g., increase the simulation errors). Therefore, it is critical to eliminate those variables that cannot improve the predictive accuracy. WFI, in combination with the RFE process, can thus be used for facilitating hydrological inference and modelling.

7. Conclusions

WFI was developed to improve the robustness of variable rankings for tree-structured statistical models. Our results indicate that the proposed WFI can provide more robust variable rankings than well-known PFI and MDI methods. In addition, we found that the predictors selected by WFI can replace those selected by RF with its default methods to improve modelling predictive accuracy.

The achievements of the proposed WFI approach are two-fold: firstly, robust variable rankings are provided for a sound hydrological inference. Specifically, some critical predictors that may be overlooked by conventional feature importance methods (PFI and MDI) can be captured through WFI. Secondly, the enhanced variable rankings combined with the RFE process can help identify the most important predictors for optimum model predictive accuracy.

810

The proposed WFI could be a step closer for earth system scientists to get a preliminary understanding of the hydrological process through ML. Future studies may focus on the development of tree-structured hydrological models that may not only be viewed as black-box heuristics but also can be used for rigorous hydrological inference. Even though the focus of this paper is hydrological inference, WFI can also be applied to a variety of other important applications. Moreover, current applications of importance scores are still limited. As interpretable ML continues to mature, its potential benefits for hydrological inference could be promising.

Code and data availability. The climatic data are available on the data repository of China meteorological data service center (http://data.cma.cn/en). The hydrological data and code for the numeric case can be accessed from Zenodo repository (https://doi.org/10.5281/zenodo.4387068). The entire model code for this study can be obtained upon email request to the corresponding author.

Author contribution. Kailong Li designed the research under the supervision of Guohe Huang. Kailong Li carried out the research, developed the model code and performed the simulations. Kailong Li prepared the manuscript with contributions from Guohe Huang and Brian Baetz.

Competing interests: The authors declare that they have no conflict of interest.

Acknolwdgement. We appreciate Ningxia Water Conservancy for offering the streamflow, groundwater and irrigation data, as well as related help. We are also very grateful for the helpful inputs from the Editor and anonymous reviewers.

Financial support. This research was supported by the National Key Research and Development Plan (2016YFA0601502), Canada Research Chair Program, Natural Science and Engineering Research Council of Canada, Western Economic Diversification (15269), and MITACS.

References

Addor, N., Newman, A.J., Mizukami, N., Clark, M.P., 2017. The CAMELS data set: catchment attributes and meteorology for large-sample studies. Hydrology and Earth System Sciences, 21(10): 5293-5313.

Ahn, K.H., 2020. A neural network ensemble approach with jittered basin characteristics for regionalized low flow frequency analysis. Journal of Hydrology, 590: 125501.

Apley, D. W., & Zhu, J. (2020). Visualizing the effects of predictor variables in black box supervised learning models. Journal of the Royal Statistical Society: Series B (Statistical Methodology), 82(4), 1059-1086.

855 Athey, S., Tibshirani, J., Wager, S., 2019. Generalized random forests. Annals of Statistics, 47(2): 1148-1178.

Barandiaran, I., 1998. The random subspace method for constructing decision forests. IEEE Trans. Pattern Anal. Mach. Intell, 20(8): 1-22.

Bénard, C., Biau, G., Veiga, S., Scornet, E., 2021. Interpretable random forests via rule extraction,
860   International Conference on Artificial Intelligence and Statistics. PMLR, pp. 937-945.

Beven, K.J., 2011. Rainfall-runoff modelling: the primer. John Wiley & Sons.

Breiman, L., 2001a. Random forests. Machine learning, 45(1): 5-32.

Breiman, L., 2001b. Statistical modeling: The two cultures (with comments and a rejoinder by the author). Statistical science, 16(3): 199-231.

865 Breiman, L., Friedman, J., Stone, C.J., Olshen, R.A., 1984. Classification and regression trees. CRC press.

Craven, M., Shavlik, J.W., 1996. Extracting tree-structured representations of trained networks, Advances in neural information processing systems, pp. 24-30.

Du, Q., Biau, G., Petit, F., Porcher, R., 2021. Wasserstein Random Forests and Applications in
870   Heterogeneous Treatment Effects, International Conference on Artificial Intelligence and Statistics. PMLR, pp. 1729-1737.

Epifanio, I., 2017. Intervention in prediction measure: a new approach to assessing variable importance for random forests. BMC bioinformatics, 18(1): 1-16.

Fernández-Delgado, M., Cernadas, E., Barro, S., Amorim, D., 2014. Do we need hundreds of classifiers
875   to solve real world classification problems? The journal of machine learning research, 15(1): 3133-3181.

Friedberg, R., Tibshirani, J., Athey, S., Wager, S., 2020. Local linear forests. Journal of Computational and Graphical Statistics: 1-15.

Friedman, J.H., 2001. Greedy function approximation: a gradient boosting machine. Annals of statistics:
880   1189-1232.

Galelli, S., Castelletti, A., 2013. Assessing the predictive capability of randomized tree-based ensembles in streamflow modelling. Hydrology and Earth System Sciences, 17(7): 2669.

885 Goldstein, A., Kapelner, A., Bleich, J., Pitkin, E., 2015. Peeking inside the black box: Visualizing statistical learning with plots of individual conditional expectation. Journal of Computational and Graphical Statistics, 24(1): 44-65.

Gregorutti, B., Michel, B., Saint-Pierre, P., 2017. Correlation and variable importance in random forests. Statistics and Computing, 27(3): 659-678.

890 Guyon, I., Weston, J., Barnhill, S., Vapnik, V., 2002. Gene selection for cancer classification using support vector machines. Machine learning, 46(1-3): 389-422.

Hothorn, T., Hornik, K., Zeileis, A., 2006. Unbiased recursive partitioning: A conditional inference framework. Journal of Computational and Graphical statistics, 15(3): 651-674.

Huang, G., 1992. A stepwise cluster analysis method for predicting air quality in an urban environment.
895 Atmospheric Environment. Part B. Urban Atmosphere, 26(3): 349-357.

Ishwaran, H., Kogalur, U.B., Blackstone, E.H., Lauer, M.S., 2008. Random survival forests. Annals of Applied Statistics, 2(3): 841-860.

Katuwal, R., Suganthan, P.N., Zhang, L., 2020. Heterogeneous oblique random forest. Pattern Recognition, 99: 107078.

900 Kendall, M.G. (1948). Rank correlation methods. Griffin.

Kisi, O., Choubin, B., Deo, R.C., Yaseen, Z.M., 2019. Incorporating synoptic-scale climate signals for streamflow modelling over the Mediterranean region using machine learning models. Hydrological Sciences Journal, 64(10): 1240-1252.

Konapala, G., Mishra, A., 2020. Quantifying climate and catchment control on hydrological drought in
905 the continental United States. Water Resources Research, 56(1): e2018WR024620.

Kratzert, F. et al., 2019a. Toward improved predictions in ungauged basins: Exploiting the power of machine learning. Water Resources Research, 55(12): 11344-11354.

910   Kratzert, F. et al., 2019b. Towards learning universal, regional, and local hydrological behaviors via
            machine learning applied to large-sample datasets. Hydrology & Earth System Sciences, 23(12).

        Lawson, E., Smith, D., Sofge, D., Elmore, P., Petry, F., 2017. Decision forests for machine learning
            classification of large, noisy seafloor feature sets. Computers & Geosciences, 99: 116-124.

        Leathers, D.J., Yarnal, B., Palecki, M.A., 1991. The Pacific/North American teleconnection pattern and
915   United States climate. Part I: Regional temperature and precipitation associations. Journal of
            Climate, 4(5): 517-528.

        Liaw, A., Wiener, M., 2002. Classification and regression by randomForest. R news, 2(3): 18-22.

        Lundberg, S.M., Lee, S.-I., 2017. A unified approach to interpreting model predictions, Advances in
            neural information processing systems, pp. 4765-4774.

920   Mann, H.B., 1945. Nonparametric tests against trend. Econometrica: Journal of the Econometric Society:
            245-259.

        Mantua, N.J., Hare, S.R., Zhang, Y., Wallace, J.M., Francis, R.C., 1997. A Pacific interdecadal climate
            oscillation with impacts on salmon production. Bulletin of the american Meteorological Society,
            78(6): 1069-1080.

925   Miller, T., 2019. Explanation in artificial intelligence: Insights from the social sciences. Artificial
            intelligence, 267: 1-38.

        Molnar, C., 2020. Interpretable Machine Learning. Lulu. com.

        Morris, M.D., 1991. Factorial sampling plans for preliminary computational experiments. Technometrics,
            33(2): 161-174.

930   Murdoch, W.J., Singh, C., Kumbier, K., Abbasi-Asl, R., Yu, B., 2019. Definitions, methods, and
            applications in interpretable machine learning. Proceedings of the National Academy of Sciences,
            116(44): 22071-22080.

        Nath, R., Pavur, R., 1985. A new statistic in the one-way multivariate analysis of variance. Computational
            Statistics & Data Analysis, 2(4): 297-315.

935   Newman, A. et al., 2015. Development of a large-sample watershed-scale hydrometeorological data set

        for the contiguous USA: data set characteristics and assessment of regional variability in

        hydrologic model performance. Hydrology and Earth System Sciences, 19(1): 209.

    Reichstein, M. et al., 2019. Deep learning and process understanding for data-driven Earth system

        science. Nature, 566(7743): 195-204.

940   Ribeiro, M.T., Singh, S., Guestrin, C., 2016a. Model-agnostic interpretability of machine learning. arXiv

        preprint arXiv:1606.05386.

    Ribeiro, M.T., Singh, S., Guestrin, C., 2016b. " Why should I trust you?" Explaining the predictions of

        any classifier, Proceedings of the 22nd ACM SIGKDD international conference on knowledge

        discovery and data mining, pp. 1135-1144.

945   Schmidt, L., Heße, F., Attinger, S., Kumar, R., 2020. Challenges in applying machine learning models for

        hydrological inference: A case study for flooding events across Germany. Water Resources

        Research, 56(5): e2019WR025924.

    Scornet, E., 2020. Trees, forests, and impurity-based variable importance. arXiv preprint

        arXiv:2001.04295.

950   Shapley, L.S., 1953. A value for n-person games. Contributions to the Theory of Games, 2(28): 307-317.

    Shortridge, J.E., Guikema, S.D., Zaitchik, B.F., 2016. Machine learning methods for empirical

        streamflow simulation: a comparison of model accuracy, interpretability, and uncertainty in

        seasonal watersheds. Hydrology & Earth System Sciences, 20(7).

    Strobl, C., Boulesteix, A.-L., Kneib, T., Augustin, T., Zeileis, A., 2008. Conditional variable importance

955   for random forests. BMC bioinformatics, 9(1): 1-11.

    Strobl, C., Boulesteix, A.-L., Zeileis, A., Hothorn, T., 2007. Bias in random forest variable importance

        measures: Illustrations, sources and a solution. BMC bioinformatics, 8(1): 25.

    Trenberth, K.E., 1997. The definition of el nino. Bulletin of the American Meteorological Society,

        78(12): 2771-2778.

960   Wager, S., Athey, S., 2018. Estimation and inference of heterogeneous treatment effects using random forests. Journal of the American Statistical Association, 113(523): 1228-1242.

Wilks, S.S., 1967. Collected papers; contributions to mathematical statistics. Wiley.

Worland, S.C., 2018. Data-driven methods for hydrologic inference and discovery. Vanderbilt University.

Worland, S.C., Steinschneider, S., Asquith, W., Knight, R., Wieczorek, M., 2019. Prediction and

965   Inference of Flow Duration Curves Using Multioutput Neural Networks. Water Resources Research, 55(8): 6850-6868.

Xia, Rong, "Comparison of Random Forests and Cforest: Variable Importance Measures and Prediction Accuracies" (2009). *All Graduate Plan B and other Reports*. 1255. https://digitalcommons.usu.edu/gradreports/1255.

970   Yang, J. et al., 2015. Drought adaptation in the Ningxia Hui Autonomous Region, China: Actions, planning, pathways and barriers. Sustainability, 7(11): 15029-15056.

Yang, Y., Chui, T.F.M., 2020. Modeling and interpreting hydrological responses of sustainable urban drainage systems with explainable machine learning methods. Hydrology and Earth System Sciences Discussions: 1-41.

975   Yang, Z., 2020. DCT-Based Least-Squares Predictive Model for Hourly AQI Fluctuation Forecasting. J. Environ. Inform., 36(1).

Yu, B., 2013. Stability. Bernoulli, 19(4): 1484-1500.

Zhang, Y., Chiew, F.H., Li, M., Post, D., 2018. Predicting Runoff Signatures Using Regression and Hydrological Modeling Approaches. Water Resources Research, 54(10): 7859-7878.

980