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Response to Reviewers comments – report #3

Dear Editor,

Following please find our point-by-point response to the two Referees' comments. Their comments are in regular font and our responses are in **bold**. The page and line numbers refer to the revised manuscript that will be submitted with this response (with "all mark up" display for review).

Referee #3, Report #2

This is my second review of this manuscript. While I think it has improved considerably with respect to the previous version, in particular in the description and treatment of the Budyko-parameter, I remain somewhat skeptical about the generality of the results. I generally like scaling analysis, but I think one should always make a clear distinction between an analysis that is supposed to represent the real-world, or a hypothetical ("first-order"/potential/etc) analysis that shows the potential effect under certain restrictive assumptions (such as the use of a constant Budyko parameter in this case). The authors do acknowledge the limitations of their approach, but I feel this could (and should) be better reflected in the text. In my view, the manuscript would read much better if "heterogeneity bias" is replaced everywhere by "potential/theoretical/hypothetical heterogeneity bias" (or similar), so the suggestion of it being an analysis of the complete system including (varying) land surface properties is removed. I leave it up to the authors to make the relatively minor textual changes needed to accommodate this final comment.

The term "heterogeneity bias" occurs 79 times in the text, and we do not think that adding "potential", "theoretical", or "hypothetical" 79 times would be an improvement. Where appropriate, we already refer to "estimated heterogeneity bias", "realistic estimates of the heterogeneity bias", "numerical estimates of the heterogeneity bias", "heterogeneity bias estimates", "heterogeneity bias in ET estimates", "heterogeneity bias in a hypothetical two-column model", and so forth. We also make clear that our goal is to identify spatial patterns in the heterogeneity bias, not its absolute magnitude.

We have reviewed all 79 occurrences of the phrase "heterogeneity bias", looking for any cases where it is not already clear from context that we are referring to a hypothetical calculation (or, alternatively, the general concept of heterogeneity bias, for which "hypothetical" would not apply anyhow). In the 12 cases where we could imagine that this was not already completely clear, we modified the text accordingly.

Referee #1, Report#3

I highlight below few more points where, in my opinion, authors response is still not fully convincing. I think that an additional effort to improve the manuscript along these points could eventually make the work suitable for publication in HESS.

- In the first and second review iteration I made clear about the weak link between this work and ESMs. A pragmatic way of avoiding any misunderstanding on this point is to start the abstract (lines 22-25 of the track-changes version of the manuscript) and the first paragraph of the introduction (lines 41-48) in a different way. In any case, I would remove any link to ESMs that could potentially mislead the reader.

We have removed any mention of ESM's from the abstract and have removed the first paragraph of the introduction almost entirely, line 42-49, (transferring only half a sentence into the second paragraph of the old introduction, which is now the first paragraph of the new introduction). These introductory statements now focus on "estimates of evapotranspiration" rather than ESM's.

- I suggested extending the analysis of the heterogeneity bias by clustering the results over different climate zones over the globe. I think this analysis is still useful even if using one single dataset for P (i.e., WClim) and two datasets for PET (i.e., WClim and MODIS). This effort could lead to additional discussion that eventually elevate the scientific significance of the work.

We appreciate the reviewer's perspective, but our own assessment of the usefulness of a global analysis is different. The advantages of confining this part of our analysis to the US are clear: we can compare the two precipitation data products (Prism and WorldClim) instead of having only one, and the observational constraints on both the P and PET data products are better in the US than over most of the globe. While we could of course ALSO do this analysis at global scale (but without Prism), we would prefer not to make the paper longer and more complicated at this stage.

- I appreciate authors effort in providing additional insights on the implications of using different n values. However, I would expect some more explanations on the physical mechanisms leading to larger biases by higher n values. Finally, I do not think that repeating the analysis with spatially-distributed n values could create "artifacts". This additional analysis will show the interplay between scale-dependency in P, PET, and n with respect to the heterogeneity bias in ET. Again, this could be an effective way of bringing the scientific aspects of the work on the front.

We now explain the effects of "n" as follows: "as expected from Eqs. 3 and 4, higher values of n lead to larger heterogeneity biases, because higher values of n localize the curvature of the Budyko function more strongly at the transition between the energy and water limits (Fig. 1b), increasing the heterogeneity bias for P/PET values near this transition." (Lines 393-400)

The problem with doing an analysis with spatially distributed n is this: how, and on what basis, should we assume that n varies from place to place? "n" is an empirical parameter, typically estimated by comparing mass balances from many catchments. Thus, it is an ensemble estimate over a group of catchments, and we do not have a solid basis for attributing different n values to individual catchments, let alone individual points on the landscape within catchments. Given that we have little or no information on how, and how much, n actually varies across real-world landscapes, we would prefer not to perform analyses based on arbitrary assumptions about spatial variability in n.

1	Global assessment of how averaging over spatial heterogeneity in precipitation and potential evapotranspiration
2	affects modeled evapotranspiration rates
3	
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5	
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13	
14	Short summary
15	Evapotranspiration (ET) rates and the properties that regulate them are spatially heterogeneous. Averaging over
16	spatial heterogeneity in precipitation and potential evapotranspiration as main drivers of ET may lead to biased
17	estimates of energy and water fluxes from the land surface to the atmosphere. Here we show that this bias will be
18	largest in mountainous terrain, in regions with temperate climates and dry summers, and in landscapes where
19	spatial variations in precipitation and potential evapotranspiration are inversely correlated.

21 Abstract

22 Accurately estimating large-scale evapotranspiration rates is essential to understanding and predicting The major 23 goal of large scale Earth System Models (ESMs) is to understand and predict global change. Evapotranspiration models that are applied at continental scale typically However, computational constraints require ESMs to operate 24 25 on relatively large spatial grids - (typically ~1 degree or ~100 km in size), with the result that the heterogeneity in land surface properties and processes at smaller spatial scales cannot be explicitly represented. Averaging over this 26 27 spatial heterogeneity may lead to biased estimates of energy and water fluxes. Here we estimate how averaging 28 over spatial heterogeneity in precipitation (P) and potential evapotranspiration (PET) may affect grid-cell-averaged 29 evapotranspiration (ET) rates, as seen from the atmosphere over heterogeneous landscapes across the globe. Our 30 goal is to identify where, under what conditions, and at what scales this "heterogeneity bias" could be most 31 important, but not to quantify its absolute magnitude. We use Budyko curves as simple functions that relate ET to 32 precipitation (P) and potential evapotranspiration (PET). Because the relationships driving ET are nonlinear, 33 averaging over sub-grid heterogeneity in P and PET will lead to biased estimates of average ET. We examine the 34 global distribution of this bias, its scale dependence, and its sensitivity to variations in P versus PET. Our analysis 35 shows that this "heterogeneity bias" is more pronounced in mountainous terrain, in landscapes where spatial 36 variations in P and PET are inversely correlated, and in regions with temperate climates and dry summers. We also 37 show that this heterogeneity bias increases on average, and expands over larger areas, as the grid cell size 38 increases. 39

41 1. Introduction

42	Earth System Models (ESMs) are designed to understand interactions between the land surface, atmosphere, and
43	oceans and to predict global environmental changes. However, the Earth system and its underlying physical
44	processes are highly heterogeneous across orders of magnitude in scale below the scale of typical ESM grids (e.g.,
45	1° by 1°). Despite increasing recognition of the need to mechanistically represent physical processes in ESMs,
46	currently even the most disaggregated large scale ESMs cannot explicitly represent the spatial heterogeneity of
47	land surface hydrological properties at scales that are important to atmospheric fluxes. Averaging over land surface
48	properties at the scale of ESM model grid cells may have important implications for water and energy flux estimates
49	(Avissar and Pielke, 1989; Giorgi and Avissar, 1997; Ershadi et al., 2013; Lu et al., 2014).
50	

51 Estimates of evapotranspiration (ET) fluxes have significant implications for future temperature predictions. Smaller 52 ET fluxes imply greater sensible heat fluxes and, therefore, drier and warmer conditions in the context of climate 53 change (Seneviratne et al., 2010). Surface evaporative fluxes (and thus energy partitioning over land surfaces) are 54 nonlinear functions of available water and energy, and thus are coupled to spatially heterogeneous surface 55 characteristics (e.g., soil type, vegetation, topography) and meteorological inputs (e.g., radiative flux, wind, and 56 precipitation; Kalma et al., 2008; Shahraeeni and Or, 2010; Holland et al., 2013). These characteristics are spatially 57 variable on length scales of <1 m to many kilometers. Even the highest-resolution continental-scale 58 evapotranspiration models, such as those that are embedded in Earth System Models (ESMs), typically cannot 59 explicitly represent the spatial heterogeneity of land surface hydrological properties at scales that are important to 60 atmospheric fluxes. Instead, these models usually - well below typical ESM grid scales of ~100 km. ESMs-calculate 61 grid-averaged surface and atmosphericevapotranspiration fluxes based on using parameterizations that correspond 62 to grid-averaged properties of the land surface (Sato et al., 1989; Koster et al., 2006; Santanello and Peters-Lidard, 63 2011). Thus, ET estimates that are derived from spatially-averaged land surface properties do not capture ET 64 variations driven by the underlying surface heterogeneity (McCabe and Wood, 2006). These spatially averaged ET 65 estimates may differ from the average of the actual spatially heterogeneous ET flux, because the relationships 66 driving ET are nonlinear, Because the relationships driving ET are nonlinear, the average ET flux from a 67 heterogeneous landscape may be different from an ET estimate calculated from spatially averaged inputs

68 (Rouholahnejad Freund and Kirchner, 2017).

69

70 Several studies have quantified the effects of land surface heterogeneity on potential evapotranspiration (PET) and

71 latent heat (LH) fluxes, and have found that averaging over land surface heterogeneity can potentially bias ET

72 estimates either positively or negatively. For example, Boone and Wetzel (1998) studied the effects of soil texture

73 variability within each pixel in the Land-Atmosphere-Cloud Exchange (PLACE) model, which has a spatial resolution

- 74 of approximately 100 by 100 km. They reported that accounting for sub-grid variability in soil texture reduced
- 75 global ET by 17%, increased total runoff by 48%, and increased soil wetness by 19%, compared to using a

76 homogenous soil texture to describe the entire grid cell. Kollet (2009) found that heterogeneity in soil hydraulic

77 conductivity had a strong influence on evapotranspiration during the dry months of the year, but not during 78 months with sufficient moisture availability. Hong et al. (2009) reported that aggregating radiance data from 30 m 79 to 60, 120, 250, 500, and 1000 m resolution (input upscaling) and then calculating ET from these aggregated inputs 80 at these grid scales using Surface Energy Balance Algorithm for Land (SEBAL, Bastiaanssen et al., 1998a) yields 81 slightly larger ET estimates as compared to ET calculated with finer resolution inputs and then aggregated at the 82 desired grid scales (output upscaling). The discrepancy between ET estimated with the output upscaling method 83 and the input upscaling method grows as the size of the grid cell increases (the difference between ET calculated 84 from the input and output upscaling methods is ~20% more at a grid scale of 1 km by 1 km compared to a grid scale 85 of 120 m by 120 m). Aminzadeh et al. (2017) investigated the effects of averaging surface heterogeneity and soil 86 moisture availability on potential evaporation from a heterogeneous land surface including bare soil and vegetation 87 patches. They found that if the heterogeneity length scale is smaller than the convective atmospheric boundary 88 layer (ABL) thickness, averaging over heterogeneous land surfaces has only a small effect on average potential 89 evaporation rates. Averaging over larger-scale heterogeneities, however, led to overestimates of potential 90 evaporation.

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92 Heterogeneity biases have also been identified in ET calculation algorithms that use remote sensing data as inputs. 93 McCabe and Wood (2006) found that remote sensing retrievals of ET are larger than the corresponding in-situ flux 94 estimates and characterized the roles of land surface heterogeneity and remote sensing resolution in the retrieval 95 of evaporative flux. McCabe and Wood (2006) used Landsat (60 m), Advanced Space borne Thermal Emission and 96 Reflection Radiometer (ASTER) (90 m), and MODIS (1020 m) independently to estimate ET over the Walnut Creek 97 watershed in Iowa. They compared these remote sensing estimates to eddy covariance flux measurements and 98 reported that Landsat and ASTER ET estimates had a higher degree of consistency with one another and correlated 99 better to the ground measurements (r=0.87 and r=0.81, respectively) than MODIS- based ET estimates did. All three 100 remote sensing products overestimated ET as compared to ground measurements (at 12 out of 14 tower sites). 101 Upon aggregation of Landsat and ASTER retrievals to MODIS scale (1 km), the correlation with the ground 102 measurements decreased to r=0.75 and r=0.63 for Landsat and ASTER, respectively. 103 104 Contrary to overestimation bias, many remotely sensed ET estimates that include parameters related to aerodynamic resistance are significantly affected by heterogeneity, and underestimate ET as the scale increases

aerodynamic resistance are significantly affected by heterogeneity, and underestimate ET as the scale increases
(Ershadi et al., 2013). Because aerodynamic resistance is significantly affected by land surface properties (e.g.,
vegetation height, roughness length, and displacement height), decreases in aerodynamic resistance at coarser
resolutions could lead to smaller estimates of evapotranspiration. Ershadi et al. (2013) showed that input
aggregation from 120m to 960 m in Surface Energy Balance System (SEBS, Su, 2002) leads to up to 15 %
underestimation of ET at the larger grid resolution in a study area in the south-east of Australia.

112 Rouholahnejad Freund and Kirchner (2017) quantified the impact of sub-grid heterogeneity on grid-average ET 113 using a simple Budyko curve (Turc, 1954; Mezentsev, 1955) in which long-term average ET is a non-linear function 114 of long-term averages of precipitation (P) and potential evaporation (PET). They showed mathematically that averaging over spatially heterogeneous P and PET results in overestimation of ET within the Budyko framework (Fig. 115 116 1). Their analysis implies that large-scale ESMs that overlook land surface heterogeneity will also yield biased 117 evapotranspiration estimates due to the inherent nonlinearity in ET processes. They did not, however, determine 118 where around the globe, and under what conditions, this heterogeneity bias is likely to be most important. 119 120 The recognition that spatial averaging can potentially lead to biased flux estimates has prompted methods for 121 representing sub-grid-scale heterogeneities and processes within large scale land surface models and ESMs. 122 Accounting for land surface heterogeneity in large-scale ESMs is not merely constrained by limitations in both 123 computational power (Baker et al. 2017) and the availability of high-resolution forcing data, but also by the fact 124 that the atmospheric and land surface components of some ESMs operate at different resolutions. There have been 125 several attempts to integrate sub-grid heterogeneity in ESMs while keeping the computational costs affordable. In 126 "mosaic" approaches, the model is run separately for each surface type in a grid cell, and then the surface-specific 127 fluxes are area-weighted to calculate the grid-cell average fluxes (e.g., Avissar and Pielke, 1989; Koster and Suarez, 128 1992). The "effective parameter" approach (e.g., Wood and Mason, 1991; Mahrt et al., 1992), by contrast, seeks to 129 estimate effective parameter values at the grid cell scale that subsume the effects of sub-grid heterogeneity. 130 Estimating these effective parameters can be challenging because the relevant land-surface processes typically 131 depend nonlinearly on multiple interacting parameters, and land-surface signals at different scales are propagated 132 and diffused differently in the atmosphere. Alternatively, the "correction factor" approach (e.g., Maayar and Chen, 133 2006) uses sub-grid information on spatially heterogeneous land-surface processes and properties to estimate 134 multiplicative correction factors for fluxes that are originally calculated from spatially averaged inputs at the grid-135 cell scale. All three approaches try to reduce the heterogeneous problem to a homogeneous one that has 136 equivalent effects on the atmosphere at the grid-cell scale. 137 138 There is a growing need to understand how sub-grid heterogeneity (and the atmosphere's integration of it) affect grid-scale water and energy fluxes, and to develop effective methods to incorporate these effects in ESMs (Clark et

139 grid-scale water and energy fluxes, and to develop effective methods to incorporate these effects in ESMs (Clark et 140 al., 2015, Fan et al., 2019). In a previous study, we proposed a general framework for quantifying systematic biases 141 in ET estimates due to averaging over heterogeneities (Rouholahnejad Freund and Kirchner, 2017). We used the 142 Budyko framework as a simple estimator of ET, and demonstrated theoretically how averaging over heterogeneous 143 precipitation and potential evapotranspiration can lead to systematic overestimation of long-term average ET 144 fluxes from heterogeneous landscapes. In the present study, we apply this analysis across the globe and highlight 145 the locations where the <u>resulting</u> heterogeneity bias is largest. Our hypotheses, derived from the Budyko 146 framework as summarized in Eq. (4) below, are that (1) strongly heterogeneous landscapes, such as mountainous terrain, will exhibit greater heterogeneity bias, (2) this bias will be larger in climates where P and PET are inversely
correlated in space, and (3) heterogeneity bias will decrease as the spatial scales of averaging decrease.

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$150 \qquad \hbox{2. Effects of sub-grid heterogeneity on ET estimates in the Budyko framework}$

151 Budyko (1974) showed that long-term annual average evapotranspiration is a function of both the supply of water

152 (precipitation, P) and the evaporative demand (potential evapotranspiration, PET) under steady-state conditions

and in catchments with negligible changes in storage (Eq. 1; Turc, 1954; Mezentsev, 1955):

154
$$ET = f(P, PET) = \frac{P}{\left(\left(\frac{P}{PET}\right)^n + 1\right)^{1/n}}$$

155 where ET is actual evapotranspiration, P is precipitation, PET is potential evaporation, and n (dimensionless) is a 156 catchment-specific parameter that modifies the partitioning of P between ET and discharge. 157 158 Evapotranspiration rates are inherently bounded by energy and water limits. Under arid conditions ET is limited by 159 the available supply of water (the water limit line in Fig. 1b), while under humid conditions ET is limited by 160 atmospheric demand (PET) and converges toward PET (the energy limit line in Fig. 1b). Budyko showed that over a 161 long period and under steady-state conditions, hydrological systems function close to their energy or water limits. 162 These intrinsic water and energy constraints make the Budyko curve downward-curving. 163 164 In a heterogeneous landscape, like the simple example of two model columns in Fig. 1a, P and PET vary spatially. 165 The two columns with heterogeneous P and PET are represented by the two solid black circles on the Budyko curve 166 in Fig. 1b. In this hypothetical two-column example, the true average of ET values calculated from individual 167 heterogeneous inputs (the solid black circles) lies below the curve (the grey circle, labeled "true average"). 168 However, if we aggregate the two columns and consider the system as one column with average properties, the 169 function of average inputs (averaged P and PET over the two columns) lies on the Budyko curve (the open circle) 170 which is larger than the true average of the two columns. In short, in any downward curving function, the function 171 of the average inputs (the open circle) will always be larger than the average of the individual function values (the 172 true average; grey circle). The difference between the two can be termed the "heterogeneity bias". 173 174 In a previous study (Rouholahnejad Freund and Kirchner, 2017) we showed that when nonlinear underlying 175 relationships are used to predict average behaviour from averaged properties, the magnitude of the resulting 176 heterogeneity bias can be estimated from the degree of the curvature in the underlying function and the range 177 spanned by the individual data being averaged. Here we summarize theses findings as building blocks of the current 178 study. The second-order, second-moment Taylor expansion of the ET function f(P,PET) (Eq. 1) around its mean

179 directly yields:

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(1)

180
$$\bar{f}(P, PET) = \overline{ET} \approx f(\overline{P}, \overline{PET}) + \frac{1}{2} \frac{\partial^2 f}{\partial P^2} var(P) + \frac{1}{2} \frac{\partial^2 f}{\partial PET^2} var(PET) + \frac{\partial^2 f}{\partial P \partial PET} cov(P, PET) , \qquad (2)$$

where $\overline{f}(P, PET)$ is the true average of the spatially heterogeneous ET function, $f(\overline{P}, \overline{PET})$ is the ET function evaluated at its average inputs \overline{P} and \overline{PET} , and the derivatives are calculated at \overline{P} and \overline{PET} . Evaluating the

derivatives using Eq. (1) and reshuffling the terms, Rouholahnejad Freund and Kirchner (2017) obtained the

following expression for the heterogeneity bias, the difference between the average ET, $\overline{f}(P, PET)$, and the ET function evaluated at the mean of its inputs, $f(\overline{P}, \overline{PET})$:

186
$$f(\overline{P}, \overline{PET}) - \overline{f}(P, PET) \approx (n+1) \frac{\overline{P}^{n+1} \overline{PET}^{n+1}}{(\overline{P}^n + \overline{PET}^n)^{2+1/n}} \left[\frac{1}{2} \frac{var(P)}{\overline{P}^2} + \frac{1}{2} \frac{var(PET)}{\overline{PET}^2} - \frac{cov(P, PET)}{\overline{P} \, \overline{PET}} \right].$$
(3)

187 To more clearly show the effects of variations in P and PET, Eq. (3) can be reformulated as follows:

188
$$f(\overline{P}, \overline{PET}) - \overline{f}(P, PET) \approx (n+1)\frac{\overline{P}^{n+1}\overline{PET}^{n+1}}{(\overline{P}^n + \overline{PET}^n)^{2+1/n}} \left[\frac{1}{2}\left(\frac{SD(P)}{\overline{P}}\right)^2 + \frac{1}{2}\left(\frac{SD(PET)}{\overline{PET}}\right)^2 - r_{P,PET}\left(\frac{SD(P)}{\overline{P}}\right)\left(\frac{SD(PET)}{\overline{PET}}\right)\right] .$$
(4)

189 Equation (4) shows that the heterogeneity bias depends on only four quantities: the fractional variation (i.e., the coefficient of variation) in precipitation $\left(\frac{SD(P)}{\bar{P}}\right)$ and in potential ET $\left(\frac{SD(PET)}{PET}\right)$, the correlation between precipitation 190 and potential ET $(r_{P,PET})$, and the function $(n + 1) \frac{\bar{P}^{n+1}\bar{PET}^{n+1}}{(\bar{P}^n + PET^n)^{2+1/n}}$, which quantifies the curvature in the ET function 191 192 in Budyko space. As shown by Fig. 1b and Eq. (2), the discrepancy between average of the ET function and the ET 193 function of the average inputs (the heterogeneity bias) is proportional to both the degree of nonlinearity in the 194 function, as defined by its second derivatives, and the variability of P and PET. Equation (4) allows one to estimate 195 how much the curvature of the ET function and the fractional variability (standard deviation divided by mean) of P 196 and PET will affect estimates of ET. However, to the best of our knowledge, the consequences of these nonlinearities for global evaporative flux estimates have not previously been quantified. 197 198

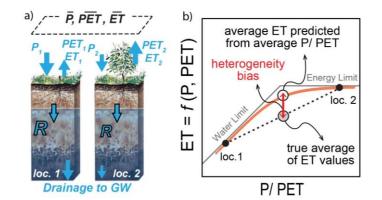


Figure 1. Heterogeneity bias in a hypothetical two-column model in the Budyko framework. The true average ET of the columns (gray circle) lies below the curve and is less than the average ET estimated from the average P/PET of the two columns (open circle). The heterogeneity bias depends on the curvature of the function and the spread of its inputs. Both panels are adapted from Rouholahnejad Freund and Kirchner (2017).

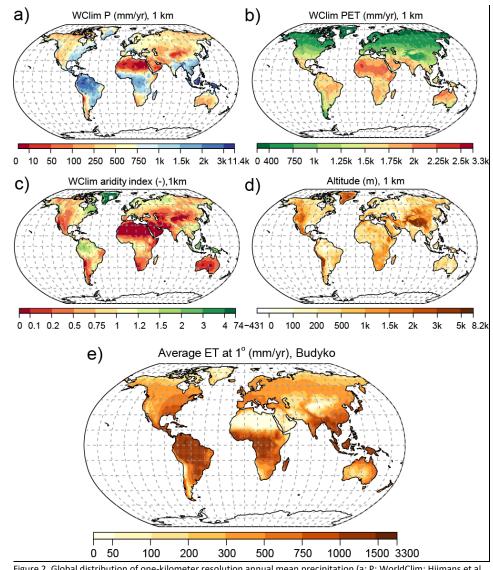
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205 3. Effects of sub-grid heterogeneity on ET estimates at 1° by 1° grid scale across the globe

206 Across a landscape of similar size to a typical ESM grid cell (1° by 1°), soil moisture, atmospheric demand (PET) and 207 precipitation (P) will vary with topographic position; hillslopes will typically be drier, and riparian regions will be 208 wetter. To map the spatial pattern in the heterogeneity bias that results could result from averaging over this land 209 surface heterogeneity, we applied the approach outlined in section 2 to the global land surface area at 1° by 1° grid 210 scale. Within each 1° by 1° grid cell, we used 30 arc-second values of P (WorldClim; Hijmans et al., 2005) and PET 211 (WorldClim; Hijmans et al., 2005) to examine the variations in small-scale climatic drivers of ET. Because 30 arc-212 seconds is nearly 1 km, hereafter we refer to the 30 arc-second data as 1km values for simplicity. The spatial 213 distribution of long-term annual averages (1960-1990) of P and PET values at 1 km resolution, along with 1km 214 values of the aridity index (AI=P/PET), are shown in Fig 2a-c. ET values calculated from these 1km P and PET values 215 using Eq. (1) are then averaged at 1° by 1° scale ("true average", Fig. 2e). We also averaged the 1km values of P and 216 PET within each grid cell and then modeled ET using the Budyko curve (Eq. 1) applied to these averaged input 217 values. The difference between these two ET estimates is the heterogeneity bias. 218 219 We also calculated the heterogeneity bias using Eq. (4), which describes how the nonlinearity in the governing 220 equation and the heterogeneity in P and PET jointly contribute to the heterogeneity bias. The heterogeneity bias 221 estimates obtained by Eq. (4) were functionally equivalent (R²=0.97, root mean square error of 0.17%) to those 222 obtained by direct calculation using Eq. (1) as described above. 223

224 Fig. 3a-d illustrates the variability (quantified by standard deviation) of 1km values of P, PET, aridity index, and 225 altitude at the 1° by 1° grid scale. The heterogeneity bias in long-term average ET fluxes at the 1° by 1° grid scale 226 (Fig. 3e) highlights regions around the globe where ET fluxes are likely to be systematically overestimated. The 227 spatial distribution of the heterogeneity bias calculated using Eq. 4 (Fig. 3e) closely coincides with locations where 228 the aridity index is highly variable (Fig. 3c), which is driven in turn by topographic variability (Fig. 3d). Strongly 229 heterogeneous landscapes exhibit significant-larger estimated heterogeneity biases in long-term average ET fluxes. 230 Although the global average of our Budyko-based heterogeneity bias estimates is small (<1%), physically based ET 231 calculations may exhibit larger heterogeneity biases than the modest values we calculate here, because the Budyko 232 approach already subsumes spatial heterogeneity effects at the catchment scale (and also temporal heterogeneity 233 effects due to its steady-state assumptions). The heterogeneity biases in ET estimates shown in Fig. 3e correspond 234 to long-term average ET estimates. Given the fact that P and PET can vary temporally (i.e., seasonality), the actual 235 bias could be much larger, particularly where P and PET are inversely correlated (see the last term of Eq. 4).

- 236
- 237 Our results show that the topographic gradient, and hence the variability in the aridity index across a given grid
- 238 scale, drives consistent, predictable patterns of heterogeneity bias in evapotranspiration estimates at that scale.
- 239 Equation 4 shows that this bias is equally sensitive to fractional variability in P and PET (standard deviation divided
- by mean). However, because P is typically more variable (in percentage terms) than PET across landscapes, the
- 241 variability in P will usually make a larger contribution to the <u>estimated</u> heterogeneity bias.

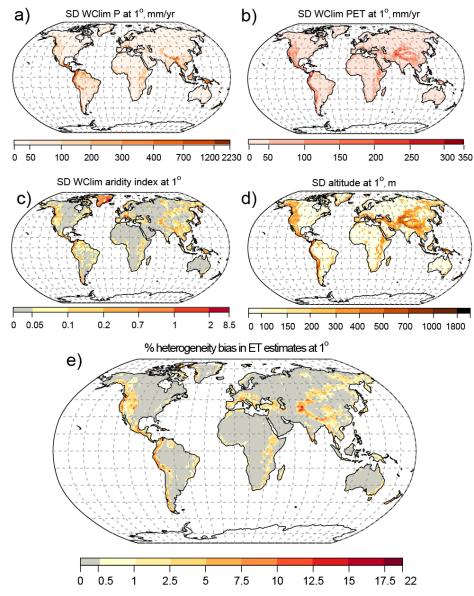


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243050100200300500750100015003300244Figure 2. Global distribution of one-kilometer resolution annual mean precipitation (a: P; WorldClim; Hijmans et al.,2452005), potential evapotranspiration (b: PET; WorldClim; Hijmans et al., 2005), aridity index (c: AI=P/PET; WorldClim;

Hijmans et al., 2005), and topography (d: SRTM; Jarvis et al., 2008), along with (e) evapotranspiration (ET) at 1° by

247 1° scale by averaging 1km values of ET calculated using the Budyko function (Eq. 1).



250 Figure 3. Global spatial distribution of variability (standard deviation) of one-kilometer values of a) precipitation (P),

b) potential evapotranspiration (PET), c) aridity index (AI=P/PET), and d) altitude at 1° by 1° grid cell. The



and aridity index have larger heterogeneity bias.

254 4. Variation in heterogeneity bias across climate zones, data sources, and grid scales

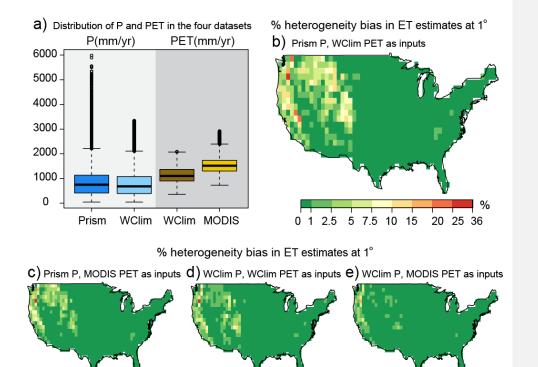
255 With increased availability of spatial data, it is becoming standard practice to assess input data uncertainties and 256 their propagated impacts on water and energy flux estimates in land surface models. To quantify how choices 257 among alternative input data products could affect the heterogeneity bias in ET estimates, we calculated the 258 heterogeneity bias at 1° by 1° grid cell resolution across the contiguous US using four different pairs of P and PET 259 data products. Two precipitation data sets, Prism (http://prism.oregonstate.edu) and WorldClim (Hijmans et al., 260 2005), along with two PET data sets, MODIS (Mu et al., 2007) and WorldClim (Hijmans et al., 2005). As Prism precipitation data is available at 4 km resolution, all other data sets were aggregated to 4 km. Two P products and 261 262 two PET products were combined in all possible pairs. The WorldClim PET dataset (Hijmans et al., 2005) is based on 263 the Hargreaves method (Hargreaves and Samani 1985) while the MODIS PET product (Mu et al, 2007) is based on 264 the Penman–Monteith equation (Monteith, 1965). The heterogeneity bias in ET estimates (Eq. 4), as outlined in 265 Sect. 2, was evaluated from 4km values of P, PET, and the estimated average ET using the Budyko relationship (Eq. 266 1) for each of the four input data pairs. Figure 4a-e compares the spatial distributions of heterogeneity bias across 267 the contiguous US for the four pairs of P and PET data products. The heterogeneity bias in ET estimates reached as 268 high as 36 % in the western US using Prism P and WorldClim PET as input to the ET model (Fig. 4b). A visual 269 comparison of Figs. 4b and Fig. 4d shows that the choice of P data source (Prism vs. WorldClim) had a bigger effect 270 on the heterogeneity bias than the choice of PET data source (MODIS vs. WorldClim), meaning that the fractional 271 variability in P is the dominant variable. In all cases, data sources that were more variable in relation to their means 272 (Prism for P and WorldClim for PET; Fig. 4b) led to larger estimates of heterogeneity biases, as expected from Eq. 273 (4). Thus we infer that we would have obtained larger heterogeneity biases if we had conducted our global analysis 274 (Fig. 3) with Prism P and either WorldClim or MODIS PET, but we cannot show that result explicitly at global scale 275 because Prism P is not freely available globally. 276

277 If we separate the heterogeneity biases shown in Fig. 4 according to Köppen-Geiger climate zones (Peel et al., 2007; 278 Fig. 5a), we see that they are distinctly higher in particular climate-terrain combinations. Estimated hHeterogeneity 279 biases are higher in regions with temperate climates and dry summers (climate zone Cs) and in regions with cold, 280 dry summers (climate zone Ds), most likely due to the sharp spatial gradient in their water and energy sources for 281 evapotranspiration (Fig. 5b). These areas typically have high topographic relief, combined with seasonal climate. 282 The heterogeneity effects on ET estimates in these regions are expected to be even larger when a mechanistic 283 model of ET is used. We expect that averaging over temporal variations of drivers of ET, especially in places with 284 strong seasonality, could substantially bias the ET estimates, but this cannot be quantified in the Budyko framework 285 due to its underlying steady-state assumptions. Figure 5b also illustrates the relative magnitudes of the 286 heterogeneity biases obtained with the four pairs of P and PET data sources. The estimated heterogeneity bias is 287 the highest when the Prism P and WorldClim PET datasets are used, followed by the combination of Prism P and 288 MODIS PET, which resulted in the second-highest heterogeneity bias across different climate zones. Wilcoxon 289 signed-rank tests was performed to evaluate the statistical significance of the differences between heterogeneity

- 290 bias in ET estimates using all pairs of climate zones and data sources that are shown in Fig. 5b (Table S1). These
- 291 analysis show that while the difference between heterogeneity biases estimated in Cs and Ds climate zones are not
- 292 statistically significant across all four combinations of datasets, the difference between estimated heterogeneity
- 293 bias in Cs versus Cf, Ds versus Cf, as well as Cs versus Bs climate zones are significant across all four data
- 294 combinations (highlighted in Table S1 of the supplementary material).

295

296 Equation 4 shows that heterogeneity biases in Budyko estimates of ET are equally sensitive to the same percentage 297 variability in P and PET. Thus the degree of sensitivity, per se, to P and PET variations expressed in percentage terms 298 is the same. Although Figs. 5c and 5d give the visual impression that PET is more variable than P across climate 299 zones and between data sources, Fig. 5e shows that the fractional variability in P is systematically higher than PET, 300 and it also varies more across the climate zones and between the two data sets. Because P is typically more 301 variable than PET (in percentage terms) across landscapes, the variability in P will make a larger contribution to the 302 heterogeneity bias (Fig. 5e) estimated usingin the Budyko approach. Whether this is true for more physically based 303 ET estimates remains to be seen. Analysis of percent variability of P and PET products shows that percent 304 variabilities of precipitation products are in general larger than PET products and hence contribute more to 305 heterogeneity (Fig 5e). While the percent variabilities of the two PET products are in the same range, the percent 306 variability in Prism precipitation is slightly larger than in WorldClim precipitation, in regions with dry summers (Cs and Ds climate zones in Fig. 5a).



309

310 Figure 4. The distribution of P and PET in the four datasets is shown in a). Estimated heterogeneity bias (Eq. 4)

311 across the contiguous US using four-kilometer values of b) Prism P and WorldClim PET c) Prism P and MODIS PET d)

0 1 2.5 5 7.5 10 15 20 25 36

%

0 1 2.5 5 7.5 10 15

312 WorldClim P and WorldClim PET, and e) WorldClim P and MODIS PET as inputs.

%

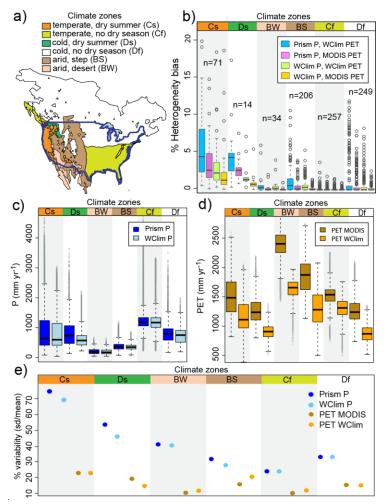
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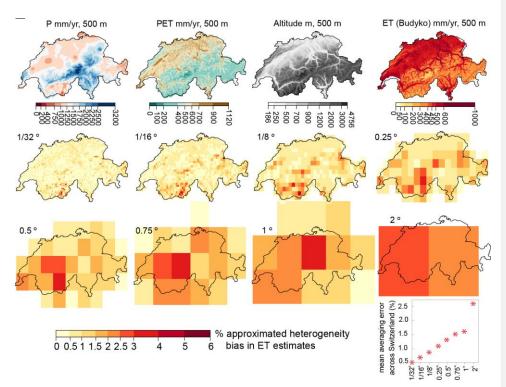
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315 Figure 5. a) Köppen-Geiger climate classification (Peel et al., 2007 in Beck et al. 2013) across the contiguous US, b) 316 the distribution of calculated heterogeneity bias in ET estimates (Eq. 4) at 1° by 1° grid cell in individual climate 317 zones, shown by boxplot (three data points with heterogeneity biases of over 20% are off-scale). The significance of 318 differences between the pairs are presented in Table S1. Panels c and d show the distribution of precipitation 319 products (Prism and WorldClim) and potential evaporation products (MODIS and WorldClim) at individual climate 320 zones, respectively. The color-coded climate zones at the tops of panels b, c, and d correspond to the climate zones 321 mapped in panel a. Panel e compares the percentage variability of the two P and PET data products across climate 322 zones, showing that the percentage variability in P is markedly higher than in PET, and the percentage variability in 323 Prism P is somewhat higher than in WorldClim P, particularly in climate zones with dry summers.

324 Because future increases in computing power will lead to ESMs with smaller grid cells, it is useful to ask how 325 changes in grid resolution affect the heterogeneity biases that we have estimated in this paper. To quantify the 326 heterogeneity bias in ET estimates as a function of grid scale, we repeated our analysis at various grid resolutions 327 using Switzerland as a test case. We started with high-resolution (500m) maps of long-term average annual 328 precipitation and PET across the Swiss landscape (Fig. 6), and then used Eq. 4 to estimate the heterogeneity bias at 329 grid scales ranging from 1/32° to 2° (~3 km to ~200 km). As Fig. 6 shows, aggregating P and PET over larger scales 330 leads to larger, and more widespread, overestimates in ET. Conversely, at finer grid resolutions, the average 331 heterogeneity bias is smaller, and the locations with large biases are more localized. On average, the heterogeneity 332 bias across Switzerland as a whole grows exponentially as the inputs are averaged over larger grids (as shown in the 333 lower-right panel in Fig. 6). 334 335



³³⁶

- 337 Figure 6. Heterogeneity bias in ET estimates at various scales across Switzerland, estimated from 500m climate
- data. ET is calculated using the Budyko relationship (Eq. 1). Heterogeneity bias was estimated from 500m
- 339 precipitation (P) and potential evapotranspiration (PET), and their variances at each grid scale, using Eq. 4. At finer
- 340 grid resolutions, the heterogeneity bias is more localized, and smaller on average.
- 341

342 5. Summary and discussion

343 Because evapotranspiration (ET) processes are inherently bounded by water and energy constraints, over the long 344 term, ET is always a nonlinear function of available water and PET, whether this function is expressed as a Budyko 345 curve or another ET model. These nonlinearities imply that spatial heterogeneity will not simply average out in 346 predictions of land surface water and energy fluxes in ESMs. Overlooking sub-grid spatial heterogeneity in large-347 scale land surface models ESMs could lead to biases in estimating estimated water and energy fluxes these fluxes 348 (e.g., ET rates). Here we have shown that, across several scales, averaging over spatially heterogeneous land 349 surface properties and processes leads to biases in evapotranspiration estimates. We examined the global 350 distribution of this bias, its scale dependence, and its sensitivity to variations in P versus PET, and showed under 351 what conditions, this heterogeneity bias is likely to be most important. Our analysis does not quantify the 352 heterogeneity biases in ESMs, owing to the many differences between these mechanistic models and the simple 353 empirical Budyko curve. But if the heterogeneity biases in ESMs can be quantified, they can be used as correction 354 factors to improve ESM estimates of surface-atmosphere water and energy fluxes across landscapes. Our paper 355 highlights a general methodology that can be used to estimate heterogeneity biases and to map their spatial 356 patterns, but not to calculate their absolute magnitudes because those will change significantly depending on the 357 ET formulation that is used.

358

359 In this study, we used Budyko curves as simple models of ET, in which long-term average ET rates are functionally 360 related to long-term averages of P and PET. We used an approach outlined by Rouholahnejad Freund and Kirchner 361 (2017) to estimate the heterogeneity bias in modeled ET at 1-degree grid scale across the globe (Fig. 3), and also at 362 multiple grid scales across Switzerland (Fig. 6), using finer-resolution P and PET values as drivers of ET. We showed 363 how the heterogeneity effects on ET estimates vary with the nonlinearity in the governing equations and with the 364 variability in land surface properties. Our analysis shows that heterogeneity effects on ET fluxes matter the most in 365 areas with sharp gradients in the aridity index, which are in turn controlled by topographic gradients, and not 366 merely in areas that are either arid or humid (e.g., compare Fig. 3e with Fig. 2c).

367

According to our analysis, regions within the U.S. that have temperate climates and dry summers exhibit greater heterogeneity bias in ET estimates (Fig. 5). We show that the <u>estimated</u> heterogeneity bias in ET estimates at each grid scale depends on the variance in the drivers of ET at that scale (Fig. 4), and on the choice of data sources used to estimate ET. Heterogeneity bias <u>estimates were</u> was significantly larger across the contiguous United States when P and PET data sources with larger variances were used (Fig. 4).

373

374 We also explored how heterogeneity biases and their spatial distribution vary with the magnitude and spatial

375 distribution of heterogeneity bias in ET estimates as a function of the scale at which the climatic drivers of ET are

- averaged. We found that as heterogeneous climatic variables are aggregated to larger scales, the heterogeneity
- 377 biases in ET estimates become greater on average, and extend over larger areas (Fig. 6). At smaller grid scales,

378 estimated the heterogeneity biasheterogeneity biases do does not completely disappear, but instead becomes 379 more localized around areas with sharp topographic gradients. Finding an effective scale at which one can average 380 over the heterogeneity of land surface properties and processes has been a longstanding problem in Earth science. 381 Our analysis shows that at smaller resolutions the average heterogeneity bias as seen from the atmosphere 382 becomes smaller, but there is no characteristic scale at which it vanishes entirely (Fig. 6). The magnitude and spatial 383 distribution of this bias depend strongly on the scale of the averaging and degree of the nonlinearity in the 384 underlying processes. The heterogeneity bias concept is general and extendable to any convex or concave function (Rouholahnejad Freund and Kirchner 2017), meaning that in any nonlinear process, averaging over spatial and 385 386 temporal heterogeneity can potentially lead to bias.

387

388 In the analysis presented here, we have assumed a value of 2 for the Budyko parameter n, which approximates the 389 variation of ET/PET with respect to P/PET in MODIS and WorldClim data across continental Europe (Mu et al. 2007; 390 Hijmans et al. 2005; Rouholahnejad Freund & Kirchner, 2017). Although there are suggestions in the literature that 391 n can vary with land use and other landscape properties (e.g., Teuling et al., 2019), here we have assumed that n is 392 spatially and temporally constant in order to focus on the effects of heterogeneity in P and PET. In the supplement 393 we present a sensitivity analysis with values of p ranging from 2 to 5 (Fig. S1). That analysis shows that, as expected 394 from Eqs. 3 and 4, higher values of *p* lead to larger heterogeneity biases, because higher values of *p* localize the 395 curvature of the Budyko function more strongly at the transition between the energy and water limits (Fig. 1b), 396 increasing the heterogeneity bias for P/PET values near this transition. Nonetheless, but the spatial pattern shown 397 in Fig. 3e remains largely unchanged over the full range of n values that we analyzed, and t. The Taylor 398 approximation in Eqs. 3 and 4 yields realistic estimates of the heterogeneity bias for all values of n that were tested

(Fig. S2). Thus while our numerical estimates of heterogeneity bias depend somewhat on the value of *p*, our
conclusions do not.

402 One should keep in mind that the true mechanistic equations that determine point-scale ET as a function of point-403 scale water availability and PET (if such data were available) may be much more nonlinear than Budyko's empirical 404 curves, because these curves already average over significant spatial and temporal heterogeneity. Thus, we expect 405 that the real-world effects of sub-grid heterogeneity are probably larger than those we have estimated in Sects. 3 406 and 4 of this study. In addition, the 1km P and PET values that are used in our global analysis might be still too 407 coarse to represent small-scale heterogeneity that is important to evapotranspiration processes.

408

401

409 Budyko curves are empirical relationships that functionally relate evaporation processes to the supply of water and

- 410 energy under steady-state conditions in closed catchments with no changes in storage. Our analysis likewise
- 411 assumes no changes in storage, nor any lateral transfer between the model grid cells, although both lateral
- transfers and changes in storage may be important, both in the real world and in models. Unlike the Budyko
- 413 framework, ET fluxes in most ESMs are often physically based (not merely functions of P and PET) and are

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414	calculated at much smaller time steps (seconds to minutes). These models often represent more processes that are
415	important to evapotranspiration (such as storage variations) and include their dynamics to the extent that is
416	computationally feasible. Because these relationships may be much more nonlinear than Budyko curves, much
417	larger there may also be significant heterogeneity biases could result when complex physically based models are
418	used to estimate ET from spatially aggregated data. Therefore, we are now working to quantify heterogeneity bias
419	in ET fluxes using a more mechanistic land surface model.
420	
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424	resolution data that enabled the analysis shown in Fig. 6.
425	
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