



1    **A discrete wavelet spectrum approach to identifying non-monotonic trend**  
2    **pattern of hydroclimate data**

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23 **Abstract:** Hydroclimate system is changing non-monotonically and identifying its trend  
24 pattern is a great challenge. Building on the discrete wavelet transform theory, we develop a  
25 discrete wavelet spectrum (DWS) approach for identifying non-monotonic trend patterns in  
26 hydroclimate time series and evaluating their statistical significance. After validating the  
27 DWS approach using two typical synthetic time series, we examined the temperature and  
28 potential evaporation over China from 1961-2013, and found that the DWS approach detected  
29 both the “warming” and the “warming hiatus” in temperature, and the reversed changes in  
30 potential evaporation. Interestingly, the identified trend patterns showed stable significance  
31 when the time series was longer than 30 years or so (i.e., the widely defined “climate”  
32 timescale). Our results suggest that non-monotonic trend patterns of hydroclimate time series  
33 and their significance should be carefully identified, and the DWS approach has the potential  
34 for wide use in hydrological and climate sciences.

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36 **Key words:** trend identification; discrete wavelet spectrum; decadal variability; statistical  
37 significance; Mann-Kendall test

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45 **1. Introduction**

46 Climate and hydrological system are exhibiting great variability (Allen and Ingram, 2002;  
47 Trenberth et al., 2014). Quantitatively, identifying human-induced climate change signals in  
48 the usually changing hydroclimate system is of great socioeconomic significance  
49 (Diffenbaugh et al., 2008; IPCC, 2013), and remains a big challenge to both scientific and  
50 social communities. The simplest way to identify changes in the hydroclimate system would  
51 be to fit a monotonic (e.g., linear) trend at long time scales, at which a significance level  
52 would be assigned by a statistical test. Among the methods used for detection of trends, the  
53 Mann-Kendall non-parametric test is most widely used and has been successfully applied in  
54 climate change and its impact studies, when the time series is almost monotonic as required  
55 (Burn and Hag Elnur, 2002; Yue et al., 2002). However, due to its nonlinear and  
56 nonstationary nature, the hydroclimate system is changing and developing in a more  
57 complicated way rather than a monotonic trend way (Cohn and McMahon, 2005; Milly et al.,  
58 2008). For example, a debate on the recent change of global air temperature is receiving  
59 enormous public and scientific attention that the global air temperature increased during  
60 1980-1998 passing most statistical significance tests and then stabilized afterwards till now,  
61 widely called “global warming hiatus” (Kosaka and Xie, 2013; Roberts et al., 2015). Another  
62 known example is “evaporation paradox” (Brutsaert and Parlange, 1998; Roderick and  
63 Farquhar, 2002) that potential evaporation has worldwide declined from the 1960s, again  
64 passing most statistical significance tests, but then reversed after the 1990s. In practice, for  
65 the hydroclimate time series, the non-monotonicity is more the rule rather than the exception  
66 (Dixon et al., 2006; Adam and Lettenmaier, 2008; Gong et al., 2010). Therefore,



67 identifying the non-monotonic trend pattern hidden in those time series and assessing its  
68 statistical significance presents a significant research task.

69 Among those methods presently used in time series analysis, the wavelet method has the  
70 superior capability of handling the nonstationary characteristics of time series (Percival and  
71 Walden, 2000; Labat, 2005), so it may be more suitable for identifying non-monotonic trend  
72 patterns in hydroclimate time series. In a seminal work, Torrence and Compo (1998) placed  
73 the continuous wavelet transform in the framework of statistical analysis by formulating a  
74 significance test. Since then, the continuous wavelet method has become more applicable and  
75 rapidly developed to estimate the significance of variability in climate and hydrological  
76 studies. However, in the continuous wavelet results of time series, a known technical issue is  
77 the “data redundancy” (Gauchere, 2002; Nourani et al., 2014), which is the redundant  
78 information across timescales leading to more uncertainty.

79 On the contrary, the other type of wavelet transform, i.e., the discrete wavelet method,  
80 has the potential to overcome that problem of data redundancy, in that those wavelets used for  
81 discrete wavelet transform must meet the orthogonal properties. Therefore, the discrete  
82 wavelet method can be more effective to identify the non-monotonic trend pattern in time  
83 series. The discrete wavelet-aided identification of trend is usually influenced by some factors,  
84 such as choice of wavelet and decomposition level, and the uncertainty evaluation of results  
85 should also be considered. To overcome these problems, Sang et al. (2013) discussed the  
86 definition of trend, and further proposed a discrete wavelet energy function-based method for  
87 the identification of trend by comparing the difference of wavelet results between  
88 hydrological data and noise. The method used proper confidence interval to assess the



89 statistical significance of the identified trend, in which the key equation for quantifying  
90 trend's significance is based on the concept of quadratic sum. However, it disobeys the  
91 common practice of spectral analysis, and sometimes cannot reasonably assess the  
92 significance of non-monotonic trend, because it neglects the big influence of trend's mean  
93 value. For instance, for those trends with small variation but big mean value, the quadratic  
94 sums are big values, based on which the statistical significance of trends would inevitably be  
95 over-assessed. Therefore, the evaluation of statistical significance of a non-monotonic trend in  
96 a time series should be based on its own variability but not other factors.

97 By combining the advantages of the discrete wavelet method and successful practice in  
98 the spectral analysis methods, this study aims at developing a practical but reliable discrete  
99 wavelet spectrum approach for identifying non-monotonic trend patterns in hydroclimate time  
100 series and quantifying their statistical significance, and further improving the understanding  
101 of non-monotonic trends by investigating their variation with data length increase. To do that,  
102 Section 2 presents the details of the newly developed approach building on the wavelet theory  
103 and spectrum analysis. In Section 3, we use both synthetic time series and annual time series  
104 of air temperature and potential evaporation over China as examples to investigate the  
105 applicability of the approach, which is followed by the discussion and conclusion in Section  
106 4.

## 107 **2. A discrete wavelet spectrum approach**

108 Here we develop an approach, termed as “discrete wavelet spectrum approach,” for  
109 identifying non-monotonic trend pattern in hydroclimate time series, in which the discrete  
110 wavelet transform (DWT) is used first to separate the trend pattern, and its statistical



111 significance is then evaluated by using the discrete wavelet spectrum, whose confidence  
112 interval is described through the Monte-Carlo test.

113 Following the wavelet analysis theory (Percival and Walden, 2000), the discrete wavelet  
114 transform of a time series can be expressed as:

$$115 \quad W_f(j, k) = \int_{-\infty}^{+\infty} f(t) \psi_{j,k}^*(t) dt \quad \text{with} \quad \psi_{j,k}(t) = a_0^{-j/2} \psi(a_0^{-j} t - b_0 k) \quad (1)$$

116 where  $f(t)$  is the series to be analyzed with a time order  $t$ , and  $\psi^*(t)$  is the complex conjugate of  
117 mother wavelet  $\psi(t)$ ;  $a_0$  and  $b_0$  are constants, and integer  $k$  is a time translation factor;  $W_f(j, k)$  is  
118 the discrete wavelet coefficient under the decomposition level  $j$ . In practice, the dyadic DWT is  
119 used widely by assigning  $a_0=2$  and  $b_0=1$ :

$$120 \quad W_f(j, k) = \int_{-\infty}^{+\infty} f(t) \psi_{j,k}^*(t) dt \quad \text{with} \quad \psi_{j,k}(t) = 2^{-j/2} \psi(2^{-j} t - k) \quad (2)$$

121 The highest decomposition level  $M$  can be calculated as  $\log_2(L)$ , where  $L$  is the length of  
122 series. The sub-signal  $f_j(t)$  in the original series  $f(t)$  under each level  $j$  ( $j = 1, 2, \dots, M$ ) can be  
123 reconstructed as:

$$124 \quad f_j(t) = \sum_k W_f(j, k) \psi^*(2^{-j} t - k) \quad (3)$$

125 where the sub-signal  $f_j(t)$  at the highest decomposition level (when  $j = M$ ) defines the  
126 non-monotonic trend pattern of the series  $f(t)$ . Sang (2012) discussed the influence of wavelet  
127 and decomposition level choice and noise type on the discrete wavelet decomposition of  
128 series, and further proposed some methods to solve them. By doing Monte-Carlo experiments,  
129 he found that those seven wavelet families used for DWT can be divided into three types, and  
130 recommended the first type, by which wavelet energy functions of various types of noise are  
131 stable and thus have little influence on the wavelet decomposition of time series. Specifically,  
132 one chooses an appropriate wavelet, according to the relationship of statistical characteristics



133 among the original series, de-noised series and removed noise, chooses a proper  
134 decomposition level by analyzing the difference between energy function of the analyzed  
135 series and that of noise, and then identifies the deterministic components (including trend) by  
136 conducting significance testing of DWT. These methods are based on the hydroclimate time  
137 series itself, and thus are reliable and reasonable. They were used here to accurately identify  
138 the non-monotonic trend pattern in a time series and assess its statistical significance.

139 Further, to establish the discrete wavelet spectrum (DWS) of time series, we need to  
140 specify a spectrum value  $E(j)$  for each sub-signal  $f_j(t)$  (in Eq. 3). Here we define  $E(j)$  at the  $j$ th  
141 level by taking the variance of  $f_j(t)$  following the general practice in conventional spectral  
142 analysis methods (Fourier transform, maximum entropy spectral analysis, *etc.*):

$$143 \quad E(j) = \text{var}(f_j(t)) \quad (4)$$

144 It can accurately quantify the intensity of variation of sub-signals (including trend) by  
145 eliminating the influence of their mean value, which is obviously different from the quadratic  
146 sum-based method proposed in Sang et al. (2013). For hydroclimate time series, deterministic  
147 components generally have different characteristics from purely noise components (Sang et  
148 al., 2012; Rajaram et al., 2015). Interestingly, due to the grid of dyadic DWT (Partal and  
149 Cigizoglu, 2008), discrete wavelet spectra  $E_r(j)$  of various noise types strictly follow an  
150 exponentially decreasing rule with base 2 (Sang 2012):

$$151 \quad E_r(j) = 2^{-j} \quad (5)$$

152 The discrete wavelet spectra of deterministic components and that of noise are different.  
153 Hence, we define the DWS of noise as the “reference discrete wavelet spectrum (RDWS),”



154 based on which we evaluate the statistical significance of the non-monotonic trend pattern of  
155 a time series.

156 To be specific, we design a technical flowchart to show how we develop the DWS  
157 approach for identifying the non-monotonic trend pattern of time series, and also for  
158 evaluating the statistical significance of that trend pattern (see the detail in Figure 1):

159 (1) For the series  $f(t)$  normalized with length  $L$ , we analyze it using the DWT in Eq. (2),  
160 and calculate its discrete wavelet spectrum by Eq. (4);

161 (2) For the comparison purpose, we then use the Monte-Carlo method to generate the  
162 normalized noise data  $N$  with the same length as the series  $f(t)$ , and determine its  
163 RDWS by Eq. (4). Considering that discrete wavelet spectra of various types of noise  
164 just consistently follow Eq. (5), here we generate the noise data following the standard  
165 normal distribution;

166 (3) We repeat the above steps for 5000 times, and calculate the mean value and variance of  
167 the spectrum values (in Eq. 4) of the normalized noise data  $N$  at each decomposition  
168 level  $j$ , based on which we can estimate an appropriate confidence interval of RDWS at  
169 the concerned confidence level. In this study we mainly considered the 95%  
170 confidence level;

171 (4) In comparing the discrete wavelet spectrum of the series  $f(t)$  and the confidence  
172 interval generated by that of the noise (i.e., the RDWS), we can easily identify the  
173 deterministic components under the highest level as the non-monotonic trend pattern of  
174 the series, and determine whether it is significant. Specifically, if the spectrum value of  
175 the analyzed series' sub-signal under the highest level is above the confidence interval



176 of RDWS, it is thought that the non-monotonic trend pattern is statistically significant;  
177 otherwise, if the spectrum value of the sub-signal under the highest level is in the  
178 confidence interval of RDWS, it is not statistically significant.

179 Because the DWS approach fits well the common idea of spectral analysis, and its  
180 superiority compared to the method in Sang et al. (2013) can be clearly understood, but they  
181 are not compared here. In the following section, we mainly investigate the applicability of the  
182 DWS approach for identifying non-monotonic trend and its significance, and further  
183 investigate the variation of non-monotonic trend with data length increase to improve our  
184 understanding of trend.

185 < **Figure 1** >

### 186 **3. Results**

#### 187 **3.1 Synthetic series analysis**

188 To test and verify the applicability of the developed discrete wavelet spectrum (DWS)  
189 approach for identifying non-monotonic trend pattern of a time series, we consider the general  
190 hydrological situations and use two synthetic data, generated with known signals and noise a  
191 priori. For investigating the variation of non-monotonic trend with data length increase, we  
192 set the length of the two series as 200, and the noise in them follows a standard normal  
193 distribution. The first synthetic series S1 consists of an exponentially increasing line and a  
194 periodic curve (the periodicity is 200) with some noise content (Figure 2, left panel); and the  
195 second synthetic series S2 is generated by including a hemi-sine curve, a periodic curve (the  
196 periodicity is 50) and some noise content (Figure 2, right panel). Using the MK test, series S1  
197 shows a significant increase but the trend of series S2 is not significant.



198 In the DWS approach (Figure 1), we decompose series S1 into seven (i.e.,  $\langle \log_2 200 \rangle$ )  
199 sub-signals using Eq. (2) and Eq. (3), and take the sub-signals under the seventh level as the  
200 defined non-monotonic trend pattern. As shown in Figure 2 (left panel), the identified  
201 non-monotonic trend pattern in series S1 is similar to the true trend pattern. Interestingly, the  
202 linear fitting curve (a monotonic curve) could not capture the detail of the trend pattern. The  
203 same approach applies to series S2 in Figure 2 (right panel) and the conclusion is not changed.  
204 Moreover, for series S2 with large variation at long time scales, the linear fitting curve or  
205 other monotonic curves may not be physically meaningful.

206 < **Figure 2** >

207 We compute the discrete wavelet spectra of the two synthetic series using Eq. (4), and  
208 use the reference discrete wavelet spectrum with 95% confidence interval to evaluate the  
209 statistical significance of their trend patterns. Using our DWS approach, the trend pattern of  
210 series S1, which is quasi-monotonic, is found significant (Figure 3a) as in the MK test (Figure  
211 3c), but the non-monotonic series S2 shows a significant trend pattern (Figure 3b), which is  
212 different from the MK test (Figure 3d).

213 In Figure 3, we also present the significance of the identified trend patterns of the two  
214 series using both our DWS approach and the MK test and we change the length of the series  
215 to investigate the stability of the statistical significance. Generally, it would have more  
216 uncertainty when evaluating the statistical significance of trend pattern with a shorter length,  
217 corresponding to a bigger 95% confidence interval. Using our DWS approach, the 95%  
218 confidence interval for evaluating the statistical significance of trend pattern generally



219 decreases with the length of data, as expected. However, in the MK test, the significance is  
220 always determined by the constant thresholds of  $\pm 1.96$ , regardless of the data length.

221 One would expect that if the trend pattern of a series at a certain length is identified  
222 statistically significant, the significance may be more stable with a larger length of data  
223 considered. Using our DWS approach, the trend pattern of series S1 is significant when the  
224 data length is larger than 55 (Figure 3a), being similar to the result of the MK test (Figure 3c).  
225 Interestingly, using our DWS approach, the trend pattern of the series S2 is statistically  
226 significant when the data length is larger than 75 (Figure 3b), but using the MK test, the  
227 monotonic trend pattern of series S2 is significant only when the data length is between 40  
228 and 185 (Figure 3d). In summary, the significance of trend pattern identified by our DWS  
229 approach is more stable than that detected by the MK test, demonstrating the advantage of the  
230 DWS approach in dealing with non-monotonic hydroclimate time series.

231 < Figure 3 >

### 232 3.2 Observed data analysis

233 We use the annual time series of mean air temperature (denoted as TEM) and potential  
234 evaporation (denoted as PET) over China to further verify the applicability of our developed  
235 DWS approach for identifying non-monotonic trend pattern of a time series. The two series  
236 were obtained from the hydroclimate data measured at 740 meteorological stations over China,  
237 with the same measurement years from 1961 to 2013. The data have been quality-checked to  
238 ensure their reliability for scientific studies. The PET series was calculated from the  
239 Penman-Monteith approach (Chen et al., 2005).



240        Given the general nonstationary nature of observed hydroclimate time series, linear  
241 trends or more generally monotonic curves could not capture the trend pattern with large  
242 decadal variations and therefore are not particularly physically meaningful. In Figure 4 (left  
243 panel), we present the annual TEM time series visually showing nonstationary characteristics  
244 and non-monotonic variation. The TEM series decreases till the 1980s with fluctuations and  
245 then sharply rises till the 2000s, followed by a decreasing tendency. The large fluctuation of  
246 the mean air temperature after late 1990s is the known phenomenon of the “global warming  
247 hiatus” (Roberts et al., 2015). The linear fitting curve obviously missed out the more  
248 complicated trend pattern of the observed time series. Using our DWS approach, we  
249 decompose the TEM series into five (i.e.,  $\langle \log_2 53 \rangle$ ) sub-signals using Eq. (2) and Eq. (3), and  
250 take the sub-signals under the fifth level as the trend pattern, which realistically presents the  
251 nonstationary variability of temperature over China (Figure 4, left panel).

252        We also apply this DWS approach to the annual PET time series. In the time series of  
253 PET (Figure 4, right panel), there was a decreasing trend for the period from 1961 to the  
254 1990s, which is the known “evaporation paradox” leading to controversial interpretations  
255 continuing over the last decade on hydrological cycles (Brutsaert and Parlange, 1998;  
256 Roderick and Farquhar, 2002). That decreasing trend was then followed by an abrupt increase  
257 in around the 1990s, almost the same time when solar radiation was observed to be reversing  
258 its trend, widely termed as “global dimming to brightening” (Wild, 2009). Interestingly, after  
259 the mid-2000s, PET starts to decrease again (Figure 4, right panel). Sometimes, one would  
260 propose to fit linear curves for separate periods. Again, linear curves could not capture the  
261 overall trend pattern of the PET series. Using the same DWS approach, we identify



262 non-monotonic trend pattern of the PET series (Figure 4, right panel), which captures the two  
263 turning points of the changing trends in the 1990s and the 2000s.

264 < **Figure 4** >

265 The changes of trends in terms of magnitudes and signs for different periods lead to the  
266 difficulty in assessing and interpreting the significance of trends. For example, the PET time  
267 series shows a significant decrease using the MK test ( $-3.76 < -1.96$ ) during 1961-1992  
268 (Figure 5d). At that moment before the reversed trend reported, the significant decrease could  
269 be literally interpreted as that PET has significantly declined and might be declining in the  
270 future. However, the PET time series reversed after the 1990s and again in the 2000s, coming  
271 with an insignificant overall trend for the whole period of 1961-2013. For the more or less  
272 monotonic time series of the TEM series (1961-2013), the MK test detects a significant  
273 increase ( $6.00 > 1.96$ ) (Figure 5c), which leads to the surprise when air temperature was  
274 reported to have stopped increasing after late the 1990s. In summary, it becomes vital to  
275 develop an approach for testing the significance of trend pattern, which is suitable for  
276 non-monotonic time series.

277 In this study, building on the discrete wavelet transform theory, we propose an  
278 operational approach, i.e., the DWS, for evaluating the significance of non-monotonic trend  
279 pattern in the TEM (Figure 5a) and PET (Figure 5b) series. For comparison purpose, we also  
280 conduct the significance test for the two time series using the MK test (Figure 5c and 5d).  
281 Similar to Figure 3, we change the data length to investigate the stability of statistical  
282 significance (Figure 5). Again, the result indicates that the 95% confidence interval for  
283 evaluating the statistical significance generally decreases with the data length, which is



284 different from the constants  $\pm 1.96$  adopted in the MK test. The significance test using our  
285 DWS approach appears to be more stable with the data length than the MK test (Figure 5).  
286 Interestingly, using our DWS approach the trend pattern in the TEM series becomes  
287 significant when the data length is 30 and the significance is more stable when it is greater  
288 than 35 (Figure 5a). For the case of the PET series, the trend pattern becomes statistically  
289 significant when the data length is larger than 25 (Figure 5b). The findings here have  
290 important implications for non-monotonic hydroclimate time series analysis, in that the  
291 timescale of defining *climate* and *climate change* by the World Meteorological Organization  
292 is usually 30 years (Arguez and Vose, 2011) and in hydrological practice it is between 25-30  
293 years.

294 For the whole time series investigated here, whose length is larger than 30 years, we are  
295 able to examine the significance using the developed DWS approach. Combining the trend  
296 pattern in Figure 4 (left panel) and the significance test in Figure 5a, we can confirm that the  
297 trend pattern of the TEM time series from 1961-2013 identified in this study is significant at  
298 the 95% confidence interval. Similarly, the trend pattern in PET is also significant (Figure 4  
299 right panel and Figure 5b). The significance test results suggest that the three main stages of  
300 the series (red lines, Figure 4) are detectable as the overall trend pattern from the variability of  
301 the series and are vital to understanding how the temperature and the PET series are changing.  
302 In particular, the reversed changes in PET and its significance can be revealed by our DWS  
303 approach, which can provide more useful and physically meaningful information. Our results  
304 suggest that the non-monotonic trend pattern of hydroclimate time series and its significance  
305 should be carefully identified and evaluated.



306 < **Figure 5** >

307 **4. Summary and Conclusion**

308 Climate and hydrological system are changing non-monotonically. Identification of  
309 linear (or monotonic) trends in hydroclimate time series, as a common practice, cannot  
310 capture the detail of the trend pattern in the time series at long time scales, and then can lead  
311 to misinterpreting climatic and hydrological changes. Therefore, revealing the trend pattern of  
312 the time series and assessing its significance from the usually varying hydroclimate system  
313 remains a great challenge. To that end, we develop the discrete wavelet spectrum (DWS)  
314 approach for identifying the non-monotonic trend in hydroclimate time series, in which the  
315 discrete wavelet transform is used first to separate the trend pattern, and its statistical  
316 significance is then evaluated by using the discrete wavelet spectrum (Figure 1). Using two  
317 typical synthetic time series, we examine the developed DWS approach, and find that it can  
318 precisely identify non-monotonic trend pattern in the synthetic time series (Figure 2) and has  
319 the advantage in significance testing (Figure 3).

320 Using our DWS approach, we identify the trend pattern in the annual time series of  
321 average temperature and potential evaporation over China from 1961-2013 (Figure 4). The  
322 identified non-monotonic trend patterns precisely describe how temperature and PET are  
323 changing. Of particularly interest here is that the DWS approach can help detect both the  
324 “warming” and the “warming hiatus” in the temperature time series, and reveal the reversed  
325 changes and the latest decrease in the PET time series. The DWS approach can provide other  
326 aspects of information on the trend pattern in the time series, i.e., the significance test. Results  
327 show that the trend pattern becomes more significant and the significance test becomes more



328 stable when the time series is longer than a certain period like 30 years or so, the widely  
329 defined “climate” time scale (Figure 5). Using the DWS approach, in both time series of mean  
330 air temperature and potential evaporation, the identified trend patterns are found significant  
331 (Figure 5).

332 In summary, our results suggest that the non-monotonic trend pattern of hydroclimate  
333 time series and its statistical significance should be carefully identified and evaluated, and the  
334 DWS approach developed in this study has the potential for wide use in hydrological and  
335 climate sciences.

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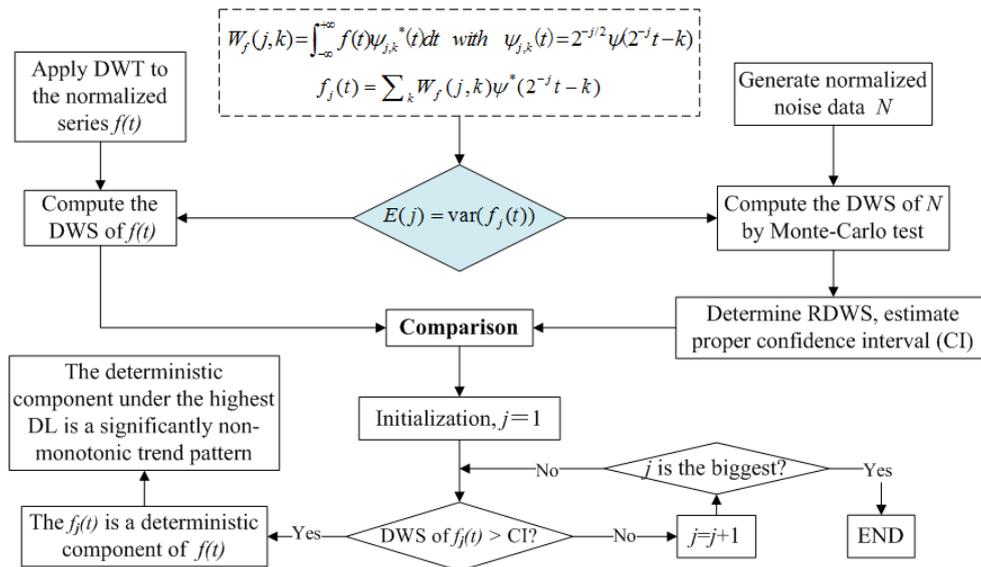
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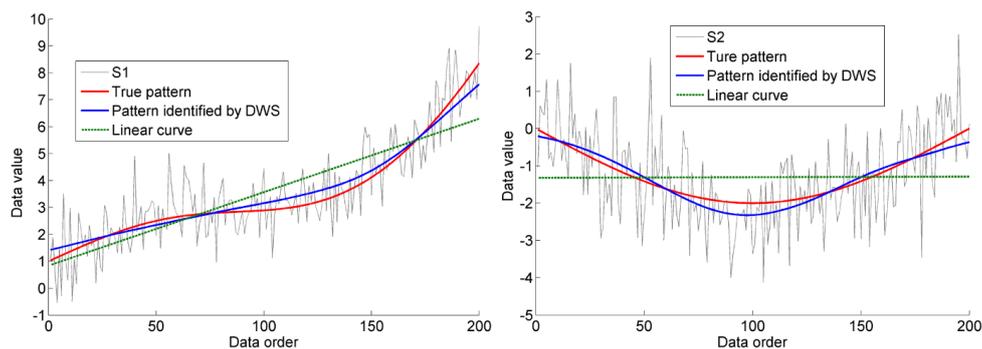
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**Figures**



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 421 **Figure 1.** Technical flowchart for identification of the non-monotonic trend pattern in a time  
 422 series using the discrete wavelet spectrum approach developed. In the figure, “DWT” is  
 423 the discrete wavelet transform, “DWS” is the discrete wavelet spectrum, “RDWS” is the  
 424 reference discrete wavelet spectrum, “DL” is the decomposition level, and “CI” is the  
 425 confidence interval.

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**Figure 2.** Non-monotonic trend patterns in the synthetic series S1 and S2 identified by the

430

discrete wavelet spectrum (DWS) approach, and the linear trends in the two series.

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Synthetic series S1 is generated as:  $S1=1.112^{0.1t}+0.8\times\sin(0.01\pi t)+\alpha$ ; and synthetic series

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S2 is generated as:  $S2=\sin(0.04\pi t)+2\times\sin(\pi+0.005\pi t)+\alpha$ , where  $\alpha$  is a random process

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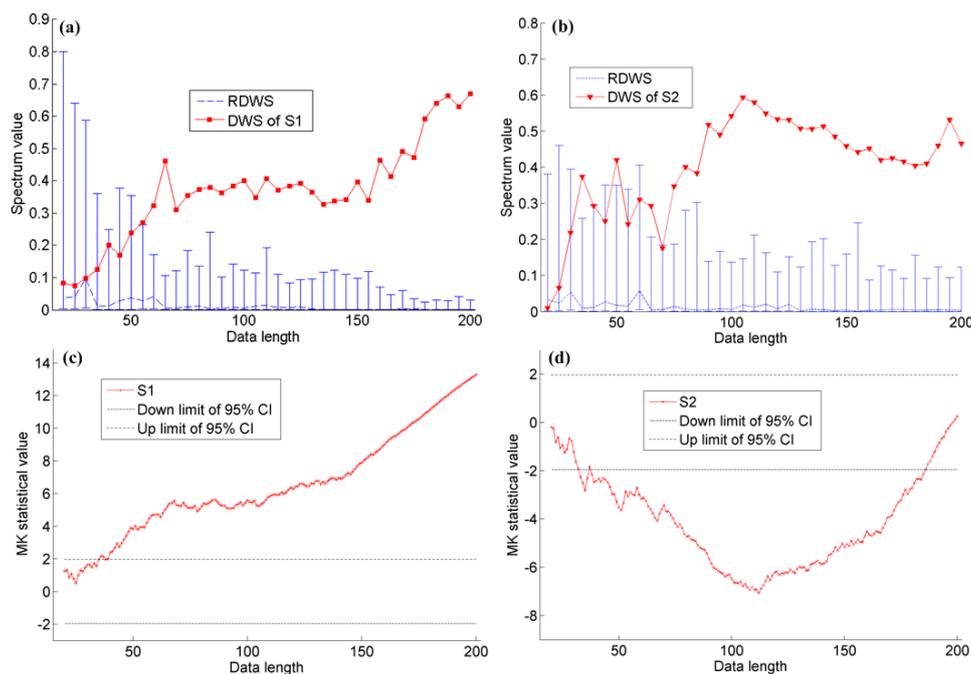
following the standard normal distribution.

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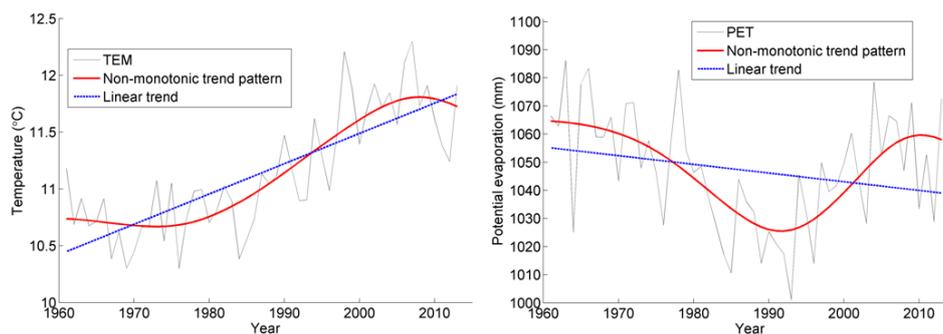
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439 **Figure 3.** Evaluation of statistical significance of non-monotonic trend patterns in the  
440 synthetic series S1 (a) and S2 (b) with different data length by the discrete wavelet  
441 spectrum (DWS) approach, and the results by the Mann-Kendall (MK) test (c and d). In  
442 figure a and b, the blue line is the reference discrete wavelet spectrum (RDWS) with 95%  
443 confidence interval under each data length; and in figure c and d, the two black dash lines  
444 indicate 95% confidence interval (CI) with the thresholds of  $\pm 1.96$  in the MK test.

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447 **Figure 4.** Non-monotonic trend patterns in the annual time series of the mean air temperature

448 (TEM) and the potential evaporation (PET) over China from 1961-2013 identified by the

449 discrete wavelet spectrum (DWS) approach, and the linear trends in the two series.

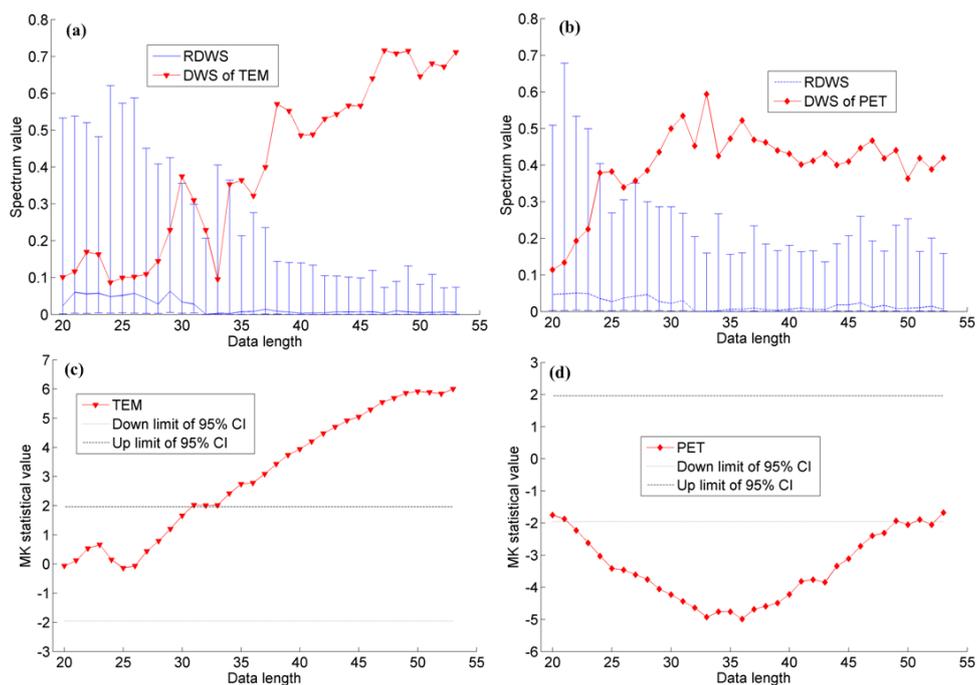
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456 **Figure 5.** Evaluation of statistical significance of non-monotonic trend patterns in the annual  
457 time series of the mean air temperature (TEM, a) and the potential evaporation (PET, b)  
458 over China with different data length by the discrete wavelet spectrum (DWS) approach,  
459 and the results by the Mann-Kendall (MK) test (c and d). In figure a and b, The blue line  
460 is the reference discrete wavelet spectrum (RDWS) with 95% confidence interval under  
461 each data length; and in figure c and d, the two black dash lines indicate 95% confidence  
462 interval (CI) with the thresholds of  $\pm 1.96$  in the MK test.