## To: Editor, Hydrology and Earth System Sciences

Subject: Revised manuscript (#hess-2017-6)

The Authors: Sang Y.F., et al.

**Title:** A discrete wavelet spectrum approach to identifying non-monotonic trend pattern of hydroclimate data

## **Response:**

The authors appreciate the Editor and Reviewers for helpful and constructive comments that improved our original manuscript submitted to HESS.

## **Response to Editor's comments:**

While Reviewer #1 felt that the paper could be published "as is", Reviewer #2 did not feel that revision to address the novelty of this paper's contribution was addressed in a sufficient way. After looking over the revised track-changes manuscript, the authors' proposed responses, and the reviewer comments, I concur with Reviewer #2. For example, the authors noted in their proposed responses that they would, "carefully rewrite and add many contents in the revised manuscript, and add some new references about the use of DWT for trend identification, mainly to more clearly explain the DWS approach proposed and emphasize the advantage of the approach." In examining the track changes version of the revised discussion paper, the only change I was able to observe is that a few more references have been added to the introduction. Since neither reviewer recommended rejection, I will not overrule this recommendation. However, I do hope to see the comments of Reviewer #2 fully addressed in this next revision. The discussion paper will once again be sent back out for review before making a final decision on publication.

**Reply:** Thank you very much for giving these helpful comments. We have revised the manuscript following your comments and suggestions. As you can find in the revised manuscript, we added some new contents to clarify the novelty of the study. To be specific, we added those contents in lines 81-83 to describe the wide use of the continuous wavelet transform and the continuous wavelet spectrum in hydrology studies, and further added those contents in lines 91-94 to emphasize the lack of an effective discrete wavelet spectrum approach in the wavelet methodology: "However, there lacked an effective discrete wavelet spectrum in the wavelet methodology. Without it, uncertainty in the discrete wavelet-aided identification of trend cannot be accurately estimated, and the significance level of the discrete wavelet spectrum (DWS) approach for detecting non-monotonic trends in hydroclimate time series, as an important basis of understanding the variability of hydroclimate process at large time scales.

Moreover, we added those contents in lines 339-370 and the new Figure 6 to describe the spatial distribution of significance of trends in potential evaporation over China, and found the big different results gotten from the DWS approach proposed and the Mann-Kendall (MK)

test. We found that the MK test underestimated the significant of those trends (at 150 stations among total 520 stations) with non-monotonic variations, which is unfavorable for accurately understanding the temporal and spatial variability of potential evaporation and hydroclimate process in China. Therefore, we think that the DWS approach performs better for detection of non-monotonic trends in hydroclimate time series, and it is the novelty of this study.

More details can be found in the following point-to-point response.

## **Response to Reviwer#2's comments:**

Comment 1. DWS is a well established approach and has been widely applied, especially in signal analysis. I simply cannot see the novelty despite the authors stated they developed a new DWS approach. The novelty should be further elaborated and highlighted should the authors consider to revise and resubmit to another journal.

**Reply:** Thanks for giving this helpful comment. We know that the wavelet methodology, including both continuous and discrete wavelet transform, has been widely used for hydrology studies. And the continuous wavelet spectrum (i.e., continuous wavelet variance) was also established to detect those significant variabilities in the hydroclimate process. However, there is a "data redundancy" problem in continuous wavelet transform. Comparatively, the discrete wavelet transform can overcome the problem, and can also describe the trend pattern using those sub-signals of the original series at large time scales, so it can be more suitable for trend identification. However, there lacked an effective discrete wavelet spectrum in the wavelet methodology. Without it, uncertainty in the discrete wavelet-aided identification of trend cannot be accurately estimated, and the significance level of the identified trend cannot be quantitatively evaluated, either. Therefore, we proposed the DWS approach in the manuscript, and it is the main novelty of the study.

Following the helpful comments, we added those contents in lines 81-83 to describe the wide use of the continuous wavelet transform and the continuous wavelet spectrum in hydrology studies, and then added those contents in lines 88-94 to emphasize the lack of an effective discrete wavelet spectrum approach in the wavelet methodology. It is the incentive of proposing the DWS approach in the study. Moreover, we added those contents in lines 339-370 (new Figure 6) to more clearly verify the better performance of the DWS approach compared with the MK test used widely for trend identification, and also to describe the necessity of the study for detecting non-monotonic trends.

Comment 2. Application of DWS is limited on time series trend identification. Data interpretation is indeed important to understand the hydro-climate system. However, it would be much practically useful if the application can be extended to trend/data forecasting.

**Reply:** Thanks for giving this favorable comment. We agree the opinion, that is, detection of variability (including trend) in hydroclimate process is an important basis of hydrological simulation and prediction at large time scales, as a practical guide for water resources

planning and management. Considering that hydrological prediction is another big issue and is related to many other issues, we didn't discuss it too much here. However, following the favorable comment, we added some contents in the manuscript (in lines 49-50, 308-309 and 367-368) to briefly clarify the importance of the issue and its relationship with this study.

<u>Comment 3. The analyzed hydro-climate data are averaged time series over 740</u> <u>meteorological stations over China, if I understand correctly. By averaging, the features</u> <u>related to different climatic regimes. geological characteristics and geographical locations,</u> <u>etc. will be filtered out. To analyze the time series with different features would be of more</u> <u>interest and revealing than just to analyze the averaged data. Also, I don't think a time series</u> <u>with 53 annual value is long enough to detect the reliable trend</u>

**Reply:** Thanks very much for giving the valuable comment. Following the comment, we added those contents in lines 339-370 and the new Figure 6 to describe the spatial distribution of significance of trends in potential evaporation (PET) in China, based on which we further verify the better performance of the DWS approach proposed.

Interestingly, we found that when using the MK test, the monotonic trends were detected as significant in those annual PET time series measured at 230 stations (in lines 344-351), however, the significant non-monotonic trends in PET time series can be detected at 380 stations throughout China. That means, those annual PET time series measured at 150 stations (28.8% of the total stations and mainly in the south part of China) mainly indicated non-monotonic variations rather than monotonic trends at interdecadal scales, with similar phenomena as shown in Figure 4 (right panel), and their significance was underestimated by the MK test (in lines 352-357). Following previous studies, we know that potential evaporation process was influenced by more physical factors (precipitation, air temperature, wind speed, relative humidity, etc) in the south part of China rather than the north part; thus, potential evaporation process in South China presented more complex variability, and was more difficult to detect and attribute its physical causes (in lines 357-362).

Therefore, from the results in Figure 6 we can further verify the better performance and effectiveness of the DWS approach proposed for the detection of non-monotonic trends in hydroclimate time series, and suggest that the non-monotonic trend pattern of hydroclimate time series and its significance should be carefully identified and evaluated.

Besides, we think that detection of trend is closely related to the time scales concerned. In this study, we mainly used the observed hydroclimate data with 53 years to investigate the variability of TMP and PET process at interdecadal scales.

Thank you very much!

Best Regards! Yan-Fang Sang

1	A	discrete wavelet spectrum approach <u>for <del>to</del>-</u> identifying non-monotonic	
2	tr	end pattern <mark>s</mark> of hydroclimate data	
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25	Abstract: <u>The h</u> Hydroclimate system process is changing non-monotonically and
26	identifying its trend pattern is a great challenge. Building on the discrete wavelet transform
27	theory, we develop <u>ed</u> a discrete wavelet spectrum (DWS) approach for identifying
28	non-monotonic trend patterns in hydroclimate time series and evaluating their statistical
29	significance. After validating the DWS approach using two typical synthetic time series, we
30	examined the-annual temperature and potential evaporation over China from 1961-2013, and
31	found that the DWS approach detected both the "warming" and the "warming hiatus" in
32	temperature, and the reversed changes in potential evaporation. Interestingly, tFurther, the
33	identified <u>non-monotonic</u> trend patterns showed stable significance when the time series was
34	longer than 30 years or so (i.e., the widely defined "climate" timescale). The significance of
35	trends in potential evaporation measured at 150 stations in China, with an obvious
36	non-monotonic pattern, was underestimated and was not detected by the Mann-Kendall test.
37	Comparatively, the DWS approach can-overcaome the problem and detected those significant
38	non-monotonic trends at 380 stations, which is favorable for understanding and interpreting
39	the- spatiotemporal variability of the hydroclimatice process. Our results suggest that
40	non-monotonic trend patterns of hydroclimate time series and their significance should be
41	carefully identified, and the DWS approach proposed has the potential for wide use in
42	hydrological and climate sciences.

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# Key words: trend identification; discrete wavelet spectrum; decadal variability; statistical significance; Mann-Kendall test **1. Introduction** Climate and hydrological systemprocesses are exhibiting great variability (Allen and Ingram, 2002; Trenberth et al., 2014). Quantitatively, identifying human-induced those changing climate change signals in the usually changing the hydroclimate system process is of great socioeconomic significance (Diffenbaugh et al., 2008; IPCC, 2013), ), as an important basis for <del>of</del>-hydrological modelling, understanding the future hydroclimate regimes, and water resources planning and management. However, and it remains a big challenge to both scientific and social communities. The simplest and the most straightforward way to identify changes in the hydroclimate system process would be to fit a monotonic (e.g., linear) trend at <u>a long time scalescertain time period</u>, at which a significance level would be assigned by a statistical test. Among the methods used for detection of trends, the Mann-Kendall non-parametric test is most widely used and has been successfully applied in studies on climate change and its impact studies, when the time series is almost monotonic as required,

66	and a statistical threshold of $\pm 1.96$ is set to judge the significance of trends at 95% confidence
67	level(Burn and Hag Elnur, 2002; Yue et al., 2002)). However, due to its nonlinear and
68	nonstationary nature, the hydroclimate system process is changing and developing in a more
69	complicated way rather than a monotonic trend way at large time scales (Cohn and McMahon,
70	2005; Milly et al., 2008). For example, a debate on the recent change of global air temperature
71	is receiving enormous public and scientific attention that the global air temperature increased
72	during 1980-1998 passing most statistical significance tests and then stabilized afterwards till
73	now, widely called "global warming hiatus" (Kosaka and Xie, 2013; Roberts et al., 2015;
74	Medhaug et al., 2017). Another known example is "evaporation paradox" (Brutsaert and
75	Parlange, 1998; Roderick and Farquhar, 2002) that potential evaporation has worldwide
76	declined from the 1960s, again passing most statistical significance tests, but then reversed
77	after the 1990s. In practice, for the hydroclimate time series, the non-monotonicity is more the
78	rule rather than the exception (Dixon et al., 2006; Adam and Lettenmaier, 2008; Gong et al.,
79	2010). Therefore, identifying the non-monotonic trend pattern hidden in those hydroclimate
80	time series and assessing its statistical significance presents a significant research task for
81	understanding hydroclimatic variability and changes at large time scales.
82	Among those methods presently used in time series analysis, the wavelet method,
83	including both continuous and discrete wavelet transforms, has the superior capability of
84	handling the nonstationary characteristics of time series at multi-time scales (Percival and
85	Walden, 2000; Labat, 2005), so it may be more suitable for identifying non-monotonic trend
86	patterns in hydroclimate time series at large time scales. In a seminal work, Torrence and
87	Compo (1998) placed the continuous wavelet transform in the framework of statistical

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88	analysis by formulating a significance test. Since then, the continuous wavelet method has
89	become more applicable and rapidly developed to estimate the significance of variability in
90	climate and hydrological studies. Especially, the continuous wavelet spectrum (i.e.,
91	continuous wavelet variance) was established to detect those significant variabilities in the
92	hydroclimate process (Labat et al., 2000). However, in the continuous wavelet results of a
93	time series, a known technical issue is the "data redundancy" (Gaucherel, 2002; Nourani et al.,
94	2014), which is the redundant information across timescales leading to more uncertainty.
95	On the contrary, the other type of wavelet transform, i.e., the discrete wavelet
96	method <u>transform</u> , has the potential to overcome that problem of data redundancy, in that
97	those wavelets used for discrete wavelet transform must meet the orthogonal properties.
98	Therefore, the discrete wavelet method can be more effective to identify and describe the
99	non-monotonic trend pattern in <u>a</u> time series (Almasri et al., 2008; de Artigas et al., 2006;
100	Kallache et al., 2005; Partal and Kucuk, 2006; Nalley et al., 2012). However, there lacked an
101	effective discrete wavelet spectrum in the wavelet methodology Wwithout which it,
102	uncertainty in the discrete wavelet-aided identification of a trend cannot be accurately
103	estimated, and the significance level of the identified trend cannot be quantitatively evaluated,
104	either. The discrete wavelet aided identification of trend is usually influenced by some factors,
105	such as choice of wavelet and decomposition level, and; moreover, the uncertainty evaluation
106	of results should also be <u>carefully</u> considered. ToFor overcomeovercoming these problems,
107	Sang et al. (2013) discussed the definition of trend, and further proposed tried to proposed a
108	discrete wavelet energy function-based method for the identification of trends, with the a
109	basic idea of by comparing the difference of discrete wavelet results between hydrological

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110	data and noise. The method used <u>a proper confidence interval</u> to assess the statistical
111	significance of the identified trend, in which the key equation for quantifying trend's
112	significance iwas based on the concept of quadratic sum. However, it the practice computation
113	of quadratic sum disobeys the common practice customary practice of computing variance in
114	of-spectral analysis, and By using the quadratic sum, the significance of a non-monotonic
115	trend cannot be reasonably assessed, because it neglects the big influence of trend's mean
116	value. sometimes cannot reasonably assess the significance of non monotonic trend, because
117	it neglects the big influence of trend's mean value. For instance, for those trends with small
118	variations but big mean values, the quadratic sums are big values, based on which the
119	statistical significance of trends would inevitably be over-assessed. Therefore, the evaluation
120	of statistical significance of a non-monotonic trend in a time series should be based on its own
121	variability, and the influence of other factors should also be eliminated but not other factors.
122	By combining the advantages of the discrete wavelet method transform and successful
123	practice in the spectral analysis methods, this study aimeds at developing a practical but
124	
	reliable discrete wavelet spectrum approach approach for identifying non-monotonic trend
125	reliable discrete wavelet spectrum <u>approach approach</u> for identifying non-monotonic trend patterns in hydroclimate time series and quantifying their statistical significance, and further
125 126	
	patterns in hydroclimate time series and quantifying their statistical significance, and further
126	patterns in hydroclimate time series and quantifying their statistical significance, and further improving the understanding of non-monotonic trends by investigating their variation with
126 127	patterns in hydroclimate time series and quantifying their statistical significance, and further improving the understanding of non-monotonic trends by investigating their variation with data length increase. To do that, Section 2 presents the details of the newly developed
126 127 128	patterns in hydroclimate time series and quantifying their statistical significance, and further improving the understanding of non-monotonic trends by investigating their variation with data length increase. To do that, Section 2 presents the details of the newly developed approach building on the wavelet theory and spectrum analysis. In Section 3, we use both

#### 132 **2. A discrete wavelet spectrum approach**

Here we develop an approach, termed as "discrete wavelet spectrum approach," for identifying non-monotonic trend patterns in hydroclimate time series, in which the discrete wavelet transform (DWT) is used first to separate the trend pattern at large time scales, and its statistical significance is then evaluated by using the discrete wavelet spectrum, whose confidence interval is quantified and described through the-Monte-Carlo test.

138 Following the wavelet analysis theory (Percival and Walden, 2000), the discrete wavelet

139 140

transform of a time series $f(t)$ with a time order t can be expressed as:

$$W_{f}(j,k) = \int_{-\infty}^{+\infty} f(t)\psi_{j,k}^{*}(t)dt \quad \text{with} \quad \psi_{j,k}(t) = a_{0}^{-j/2}\psi(a_{0}^{-j}t - b_{0}k)$$
(1)

141 where where f(t)t is the series to be analyzed with a time order t, and  $\psi^*(t)$  is the complex 142 conjugate of mother wavelet  $\psi(t)$ ;  $a_0$  and  $b_0$  are constants, and integer k is a time translation 143 factor;  $W_f(j,k)$  is the discrete wavelet coefficient under the decomposition level j (i.e., time

144 <u>scale  $a_0^{j}$ </u>. In practice, the dyadic DWT is used widely by assigning  $a_0=2$  and  $b_0=1$ :

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$$W_{f}(j,k) = \int_{-\infty}^{+\infty} f(t)\psi_{j,k}^{*}(t)dt \quad with \quad \psi_{j,k}(t) = 2^{-j/2}\psi(2^{-j}t-k)$$

The highest decomposition level *M* is determined by the length *L* of series f(t), and can be calculated as  $log_2(L)$  (Fourfoula-Georgiou and Kumar, 2014),). where *L* is the length of series f(t). The sub-signal  $f_j(t)$  in the original series f(t) under each level j (j = 1, 2, ..., M) can be reconstructed as:

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$$f_{j}(t) = \sum_{k} W_{f}(j,k) \psi^{*}(2^{-j}t - k)$$
(3)

where the sub-signal  $f_j(t)$  at the highest decomposition level (when j=-M)-\_defines and describes the non-monotonic trend pattern of the series f(t), as generally understood. However, it should be noted that a meaningful trend closely depends on the temporal time scale

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154concerned. If the variability of series f(t) on a certain smaller time scale K (K < L) is concerned,155the proper decomposition level can be determined as  $log_2(K)$ , j-then, the sum of all those156sub-signals at the time scales equal to and bigger than MK can be the non-monotonic trend

157 pattern identified.-

158Sang (2012) discussed the influence of the choice of mother wavelet and decomposition 159level, as well as choice and noise types, on the discrete wavelet decomposition of time series, 160 and further proposed some methods to solve them. By doing conducting Monte-Carlo 161 experiments, he found that those seven wavelet families (126 mother wavelets) used for DWT 162 can be divided into three types, and recommended the first type, by which wavelet energy 163 functions of various types diverse types of noise data are keep stable and thus have little 164 influence on the wavelet decomposition of time series. Specifically, one chooses an 165 appropriate wavelet, according to the relationship of statistical characteristics among the 166 original series, de-noised series and removed noise, chooses a proper decomposition level 167 level by analyzing the difference between energy function of the analyzed series and that of 168 noise, and then identifies the deterministic components (including trend) by conducting 169 significance testing of DWT. These methods are closely-based built on the composition and variability of hydroclimate time series itselfat different time scales., and thus are reliable and 170 171 reasonable. They were used here to accurately identify and describe the non-monotonic trend 172 pattern in a time series, and assess its statistical significance. 173 Further, to establish the <u>a reliable</u> discrete wavelet spectrum (DWS) of time series, we 174 need to specify a spectrum value E(j) for each sub-signal  $f_i(t)$  (in Eq. 3), based on which we

175 <u>can quantitatively evaluate its importance and statistical significance</u>. Following the general

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practice in conventional spectral analysis methods (Fourier transform, maximum entropy

177 spectral analysis, *etc.*), Hhere we define E(j) at the *j*th level by taking the variance of  $f_j(t)$ 178 following the general practice in conventional spectral analysis methods (Fourier transform,

179 maximum entropy spectral analysis, *etc.*):

$$E(j) = \operatorname{var}(f_i(t)) \tag{4}$$

181 It can accurately quantify the intensity of variation of sub-signals (including trend) by eliminating the influence of their mean values, which is obviously different from the 182 183 quadratic sum-based method proposed by in-Sang et al. (2013). For hydroclimate time series, 184 different both stochastic and deterministic components generally have 185 characteristicsdistinctive characteristics from purely noise components (Sang et al., 2012; 186 Rajaram et al., 2015). Interestingly, dDue to the grid of dyadic DWT (Partal and Cigizoglu, 187 2008), discrete wavelet spectra  $E_r(j)$  of various noise types strictly follow an exponentially 188 decreasing rule with a base 2 (Sang 2012):

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 $E_r(j) = 2^{-j} \tag{5}$ 

The discrete wavelet spectra of deterministic components and that of noise are <u>obviously</u> different. Hence, we define the DWS of noise<u>data</u> as the "reference discrete wavelet spectrum (RDWS);", based on which we evaluate the statistical significance of the non-monotonic trend pattern of a time series.

To be specific, we design a technical flowchart to show how we develop the DWS approach for identifying the non-monotonic trend pattern of time series, and also for evaluating the statistical significance of that trend pattern (see the detail in Figure 1): ( 带格式的: 字体: 非倾斜( 带格式的: 字体: 非倾斜

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197	(1) For the series $f(t)$ normalized with length $L_{to be analyzed}$ , we normalize it, and	带格式的	1: 字体:	非倾斜
198	decompose it analyze it using the DWT method in Eq. (2) and (3), );			
199	(1)(2) We and calculate its the discrete wavelet spectrum of the series $f(t)$ by Eq. (4);	带格式的 带格式的		
200	(2)(3) For the comparison purpose, we then use the Monte-Carlo method to generate the			
201	normalized noise data N with the same length as the series $f(t)$ , and determine compute			
202	its RDWS by Eq. (4). Considering that discrete wavelet spectra of various typesdiverse			
203	types of noise data just-consistently follow Eq. (5), here we generate the-noise data			
204	following the standard normal probabilisty ie-distribution;			
205	(3)(4) We repeat the above <u>steps step</u> for 5000 times, and calculate the mean value and			
206	variance of the spectrum values (in Eq. 4) of the normalized noise data $N$ at each			
207	decomposition level <u><i>j</i>. Based on it</u> , based on which we ean estimate $\underline{d}$ an appropriate			
208	confidence interval of RDWS at the concerned confidence level. In this studystudy, we			
209	mainly considered the 95% confidence level;			
210	(5) In comparing the discrete wavelet spectrum <u>DWS</u> of the series $f(t)$ and the confidence			
211	interval generated by that of the-noise (i.e., the RDWS), we can easily-identifiedy the			
212	deterministic components under the highest decomposition level as the non-monotonic			
213	trend pattern of the series, and determine judged whether it wais significant.			
214	Specifically, if the spectrum value of the analyzed series' sub-signal under the highest			
215	level <u>wa</u> is above the confidence interval of RDWS, it <u>wa</u> is <u>considered</u> thought that the			
216	non-monotonic trend pattern was is-statistically significant; otherwise, if the spectrum			
217	value of the sub-signal under the highest level is infiall s-into the confidence interval of			
218	RDWS, it <u>wai</u> s not statistically significant:			

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219	(4)(6) If a smaller time scale <u>K</u> is concerned, we can use the decomposition level <u>log<sub>2</sub>(K)</u> ,	带格式 带格式
220	instead of $M$ , and then repeat the steps (1-45) to identify the non-monotonic trend	带格式 带格式
221	pattern at that time scale. <del>.</del>	带格式
222	Because the DWS approach fits well the common idea of spectral analysis, and its	带格式
223	superiority compared to the method in Sang et al. (2013) can be clearly understood, but they	
224	are not compared here. In the following section, we mainly investigate the applicability and	
225	reliability of the DWS approach for identifying the non-monotonic trend and assessing its	
226	significance, and further investigate the variation of the non-monotonic trend with data length	
227	increase to improve our understanding of trend at large time scales.	
228	< Figure 1>	
229	3. Results	
230	3.1 Synthetic series analysis	
230 231	<b>3.1 Synthetic series analysis</b> To test and verify the applicability reliability of the developed discrete wavelet spectrum	
231	To test and verify the applicability reliability of the developed discrete wavelet spectrum	
231 232	To test and verify the <u>applicability-reliability</u> of the developed discrete wavelet spectrum (DWS) approach for identifying <u>the</u> non-monotonic trend pattern of a time series, we	
231 232 233	To test and verify the <u>applicability-reliability</u> of the developed discrete wavelet spectrum (DWS) approach for identifying <u>the</u> _non-monotonic trend pattern of a time series, we consider <u>ed</u> the general hydrological situations and <u>use-generated</u> two synthetic <u>series</u> data,	
231 232 233 234	To test and verify the <u>applicability-reliability</u> of the developed discrete wavelet spectrum (DWS) approach for identifying <u>the</u> _non-monotonic trend pattern of a time series, we consider <u>ed</u> the general hydrological situations and <u>use-generated</u> two synthetic <u>series</u> data, <u>generated</u> with known signals and noise a priori. For investigating the variation of	
<ul> <li>231</li> <li>232</li> <li>233</li> <li>234</li> <li>235</li> </ul>	To test and verify the <u>applicability-reliability</u> of the developed discrete wavelet spectrum (DWS) approach for identifying <u>the</u> _non-monotonic trend pattern of a time series, we consider <u>ed</u> the general hydrological situations and <u>use-generated</u> two synthetic <u>series</u> data, <u>generated</u> _with known signals and noise a priori. For investigating the variation of non-monotonic trend with data length increase, we set the length of the two <u>synthetic</u> series as	
<ul> <li>231</li> <li>232</li> <li>233</li> <li>234</li> <li>235</li> <li>236</li> </ul>	To test and verify the <u>applicability-reliability</u> of the developed discrete wavelet spectrum (DWS) approach for identifying <u>the</u> _non-monotonic trend pattern of a time series, we consider <u>ed</u> the general hydrological situations and <u>use-generated</u> two synthetic <u>series</u> data, <u>generated</u> —with known signals and noise a priori. For investigating the variation of non-monotonic trend with data length increase, we set the length of the two <u>synthetic</u> series as 200, and the noise in them follow <u>eds</u> a standard normal <u>probability stie-</u> distribution. The first	
<ul> <li>231</li> <li>232</li> <li>233</li> <li>234</li> <li>235</li> <li>236</li> <li>237</li> </ul>	To test and verify the <u>applicability-reliability</u> of the developed discrete wavelet spectrum (DWS) approach for identifying <u>the</u> _non-monotonic trend pattern of a time series, we considered the general hydrological situations and <u>use-generated</u> two synthetic <u>series</u> data, <u>generated</u> with known signals and noise a priori. For investigating the variation of non-monotonic trend with data length increase, we set the length of the two <u>synthetic series</u> as 200, and the noise in them followeds a standard normal <u>probability stie-</u> distribution. The first synthetic series S1 consisteds of an exponentially increasing line and a periodic curve (the	

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241 <u>considering monotonic trends</u>, series S1 show<u>eds</u> a significant increase but the trend of series
242 S2 <u>wais</u> not significant.

243 In-When using the DWS approach (Figure 1), we concerned the time scale as data length, 244 and used the Daubechies (db8) wavelet to decompose series S1 into seven (i.e.,  $< log_2 200$ ) 245 sub-signals using Eq. (2) and Eq. (3), and. Then, we took take the sub-signals under the 246 seventh level as the defined non-monotonic trend pattern. As shown in Figure 2 (left panel), 247 the identified non-monotonic trend pattern in series S1 wais similar to the true trend pattern. 248 InterestinglyHowever, the linear fitting curve (a monotonic curve) could not capture the detail 249 of the <u>non-monotonic</u> trend pattern. The same approach applieds to series S2 in Figure 2 250 (right panel) and the conclusion did is not changed. Moreover, for series S2 with large 251 variation variability at long large time scales, the linear fitting curve or other monotonic 252curves may not be physically meaningful.

253 < Figure 2>

254 We computed the discrete wavelet spectra of the two synthetic series using Eq. (4), and 255 used the reference discrete wavelet spectrum with 95% confidence interval to evaluate the 256 statistical significance of their non-monotonic trend patterns. That is, if the red point at a 257 certain data length wais above the 95% confidence bar, described by the blue line in Figure 3, 258 it wais considered thought that the trend pattern wais significant at 95% confidence level. 259 Using our the DWS approach, the trend pattern of series S1, which wais quasi-monotonic, 260 wais found significant (Figure 3a) as in the MK test (Figure 3c), but the non-monotonic series 261 S2 showeds a significant trend pattern (Figure 3b), which wais greatly different from the MK 262 test (Figure 3d).

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263	In Figure 3, we also presented the significance of the identified trend patterns of the two
264	series using both our DWS approach and the MK test, and we changed the data length of the
265	series to investigate the stability of the statistical significance of the non-monotonic trend
266	pattern. Generally, it would have more uncertainty when evaluating the statistical significance
267	of trend pattern with a shorter length, corresponding to a bigger 95% confidence interval.
268	Using our DWS approach, the 95% confidence interval (i.e., the height of blue bars in Figure
269	<u>3)</u> for evaluating the statistical significance of trend pattern generally decrease $ds$ with the
270	increase of data length of data, as expected. However, in the MK test, the significance wais
271	always determined by the constant thresholds of +/-1.96, regardless of the data length.
272	In the DWS results in Figure 3, the significance levels of the non-monotonic trend
273	patterns did o-not consistently decrease with data length, but showed some fluctuation, One
274	would expect that as the proportions of different components (including trend) in the original
275	series variedy with data length. Furthermore, one would expect that if the trend pattern of a
276	series at a certain length wais identified statistically significant, the trend pattern would
277	extend with the increase of data length, thus its the significance- may be more stable with a
278	larger length of data considered. Using our DWS approach, the trend pattern of series S1 wais
279	significant when the data length wais larger than 55 (Figure 3a), being similar to the result of
280	the MK test (Figure 3c). Interestingly, using our DWS approach;;. the The trend pattern of the
281	series S2 <u>wa</u> is statistically significant when the data length <u>wais</u> larger than 75 (Figure 3b), ).
282	However, but using the MK test, the monotonic trend pattern of series S2 wais significant only
283	when the data length wais between 40 and 185 (Figure 3d). In summary, the significance of
284	trend pattern identified by our DWS approach wais more stable than that detected by the MK

test, demonstrating the advantage of the DWS approach in dealing with non-monotonic
variation of hydroclimate time series.

287 < Figure 3>

288 3.2 Observed data analysis

289 We used the annual time series time series of mean-air temperature (denoted as TEM) and 290 potential evaporation (denoted as PET) over China to further verify the applicability of our 291 developed DWS approach for identifying non-monotonic trend patterns of a time series. The 292 twoThese series time series were obtained from the hydroclimate data measured at 520 293 meteorological stations over China, with the same measurement years from 1961 to 2013. The 294 data have been quality-checked to ensure their reliability for scientific studiesresearches. The 295 PET-series data were was calculated from the Penman-Monteith approach (Chen et al., 2005). 296 The average time series of TEM and PET measured at 520 stations weare first 297 considered. Given the general nonstationary nature of observed hydroclimate time series, 298 linear trends or more generally monotonic curves could not capture the trend pattern with 299 large interdecadal variations and therefore weare not particularly physically meaningful. In 300 Figure 4 (left panel), we presented the average annual TEM time series visually showing 301 nonstationary characteristics and non-monotonic variation. The TEM series decreaseds till the 302 1980s with fluctuations and then sharply roises till the 2000s, followed by a decreasing 303 tendency. The large fluctuation of the mean-average air temperature after the late 1990s is the 304 well-known phenomenon of the "global warming hiatus" (Roberts et al., 2015). The linear 305 fitting curve obviously missed out the more complicated trend pattern of the observed 306 temperature time series. Using our DWS approach, we decomposed the TEM series into five 307 (i.e.,  $<log_253$ ) sub-signals using Eq. (2) and Eq. (3), and t<u>ook\_ake</u>-the sub-signals under the 308 fifth level as the trend pattern, which realistically present<u>eds</u> the nonstationary variability of 309 temperature-over <u>China\_at large time scales</u> (Figure 4, left panel).

310 We also applied y this the DWS approach to the average annual PET time series. In the 311 time series of PET (Figure 4, right panel), there was a decreasing trend for the period from 312 1961 to the 1990s, which is the well-known "evaporation paradox" leading to controversial 313 interpretations continuing over the last decade of hydrological cycles (Brutsaert and 314 Parlange, 1998; Roderick and Farquhar, 2002). That decreasing trend was then followed by an 315 abrupt increase in around the 1990s, almost the same time when solar radiation was observed 316 to be reversing its trend, widely termed as "global dimming to brightening" (Wild, 2009). 317 Interestingly Surprisingly, after the mid-2000s, PET starteds to decrease again (Figure 4, right 318 panel). Sometimes, one would propose to fit linear curves for separate time periods. Again, 319 linear curves could not capture the overall <u>non-monotonic</u> trend pattern of the PET series. 320 Using the same DWS approach, we identified y-the non-monotonic trend pattern of the PET 321 time series (Figure 4, right panel), which captureds the two turning points of the changing 322 trends in the 1990s and the 2000s.

323 < Figure 4>

The changes of trends in terms of magnitudes and signs for different periods lead to the difficulty in assessing and interpreting the significance of trends. For example, the PET time series show<u>eds</u> a significant decrease using the MK test (-3.76 < -1.96) during 1961-1992 (Figure 5d). At that moment before the reversed trend reported, the significant decrease could be literally interpreted as that PET ha<u>ds</u> significantly declined and might be declining in the

329	future. However, the PET time series reversed after the 1990s and again in the 2000s, coming
330	with an insignificant overall trend for the whole period of 1961-2013. For the more or less
331	monotonic time series of the TEM series (1961-2013), the MK test detecteds a significant
332	increase $(6.00 > 1.96)$ (Figure 5c), which leads to the surprise when air temperature was
333	reported to have stopped increasing after late the 1990s. In summary, it becomes vital to
334	develop an approach for testing the significance of trend pattern, which is suitable for
335	non-monotonic time series, as it is an important basis and prerequisite for hydrological
336	simulation and prediction at decadal scales.
337	In this study, building on the discrete wavelet transform theory, we proposed an
338	operational approach, i.e., the DWS, for evaluating the significance of non-monotonic trend
339	pattern in the TEM (Figure 5a) and PET (Figure 5b) series. For comparison purpose, we also
340	conducted the significance test for the two time series using the MK test (Figure 5c and 5d).
341	Similar to Figure 3, we changed the data length to investigate the stability of statistical
342	significance (Figure 5). Again, the result indicateds that the 95% confidence interval for
343	evaluating the statistical significance of non-monotonic trend pattern generally decreaseds
344	with the data length, which wais different from the constant thresholds +/-1.96 adopted in the
345	MK test. The significance test using our DWS approach appeareds to be more stable with the
346	data length than the MK test (Figure 5). Interestingly, uUsing our DWS approach, the trend
347	pattern in the TEM series becaomes significant when the data length is-increaseds to 30 and
348	the significance <u>wa</u> is more stable when it <u>wa</u> is greater than 35 (Figure 5a). For the case of the
349	PET series, the trend pattern becaemes statistically significant when the data length wais
350	larger than 25 (Figure 5b). The findings here have important implications for non-monotonic

hydroclimate time series analysis, in that the timescale of defining *climate* and *climate change*by the World Meteorological Organization is usually 30 years (Arguez and Vose, 2011) and
in hydrological practice it is between 25-30 years.

354 For the whole time series investigated here, whose length wais larger than 30 years, we 355 weare able to examine the significance using the developed DWS approach. Combining the 356 trend pattern in Figure 4 (left panel) and the significance test in Figure 5a, we can confirmed 357 that the trend pattern of the TEM time series from 1961-2013 identified in this study wais 358 significant at the 95% confidence interval. Similarly, the trend pattern in PET wais also 359 significant (Figure 4 right panel and Figure 5b). The significance test results suggested that 360 the three main stages of the series (red lines, Figure 4) weare detectable as the overall trend 361 pattern from the variability of the series and weare vital to understanding how the temperature 362 and the PET series weare changing at interdecadal scales. In particularly, the reversed changes 363 in PET and its significance can be revealed by our DWS approach, which can provide more 364 useful and physically meaningful information. Our results suggest that the non ignificanco should be oorofully identifi 365 f hydroelimate d patter 366 <mark>and evaluated.</mark> 367 < Figure 5>

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368	We further detected and evaluated the significance of non-monotonic trends of the PET	<b>带格式的:</b> 缩i 符	进: 首行缩进:	4
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369	time series measured at 520 stations for investigating their spatial difference. Because the			
370	trends in annual TEM time series weare quasi-monotonic, and they weare statistically			
371	significant at most of the all-stations, no matter using our DWS approach or the MK test,			
372	more details of TEM data were not repeated here. As for the trend patterns in the PET data,			

373	the results gotten from our DWS approach (Figure 6, left panel) and those in the MK test
374	presented substantial differences. When conducting the statistical significance test using the
375	MK test, the monotonic trends were detected as significant in those annual PET time series
376	measured at 230 stations. Significant downward monotonic trends were mainly found in the
377	southern part of the Songliao River basin, the Haihe River basin, the Huaihe River basin,
378	some regions in South China, and the Northwest China. Significant upward monotonic trends
379	were mainly found in the northern part of the Songliao River basin, the upper reach of the
380	Yellow River basin, the southwest corner of China, and some regions in the Yangtze River
381	Delta.
382	Comparatively, the significant non-monotonic trends in the PET time series were ean be
383	detected at 380 stations throughout China. That means that - those annual PET time series
384	measured at 150 stations (28.8% of the total stations and mainly in the south part of China)
385	mainly indicated non-monotonic variations rather than monotonic trends at interdecadal scales,
386	with similar phenomena as shown in Figure 4 (right panel), and their significance was
387	underestimated by the MK test, which can only handle monotonic trends. Previous studies
388	(Zhang et al., 2016; Jiang et al., 2007) indicated that potential evaporation process-was
389	influenced by more physical factors (precipitation, air temperature, wind speed, relative
390	humidity, etc.) in the southern part of China rather than the northern part; thus, potential
391	evaporation process in South China presented a more complex variability, and was more
392	difficult to detect and attribute its physical causes. As a result, it is known here that annual
393	potential evaporation process in most part of China indicated significance variability at
394	interdecadal scales, but it was underestimated by the conventional MK test; moreover, only

395	considering monotonic trends would cause a great difficulty in accurately understanding the	
396	temporal and spatial variability of potential evaporation and hydroclimate process in China,	
397	and also would be is-unfavorable for hydrological prediction at interdecadal scales. Our	
398	results suggest that the non-monotonic trend pattern of hydroclimate time series and its	
399	significance should be carefully identified and evaluated.	
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400 < Figure 6>

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### 401 **4. Summary and Conclusion**

402 Climate and hydrological system processes are changing non-monotonically. 403 Identification of linear (or monotonic) trends in hydroclimate time series, as a common 404 practice, cannot capture the detail of the non-monotonic trend pattern in the time series at long 405 large time scales, and then can lead to misinterpreting climatic and hydrological changes. 406 Therefore, revealing the trend pattern of the time series and assessing its significance from the 407 usually varying hydroclimate system process remains a great challenge. To that end, we 408 develop the discrete wavelet spectrum (DWS) approach for identifying the non-monotonic 409 trend in hydroclimate time series, in which the discrete wavelet transform is used first to 410 separate the trend pattern, and its statistical significance is then evaluated by using the 411 discrete wavelet spectrum (Figure 1). Using two typical synthetic time series, we examine the 412 developed DWS approach, and find that it can precisely identify non-monotonic trend pattern 413 in the synthetic time series (Figure 2) and has an the advantage in significance testing (Figure 414 3).

Using our DWS approach, we identify the trend pattern in the annual time series of average temperature and potential evaporation over China from 1961-2013 (Figure 4). The

417	identified non-monotonic trend patterns precisely describe how temperature and PET are
418	changing at interdecadal scales. Of particularly interest here is that the DWS approach can
419	help detect both the "warming" and the "warming hiatus" in the temperature time series, and
420	reveal the reversed changes and the latest decrease in the PET time series. The DWS approach
421	can provide other aspects of information on the trend pattern in the time series, i.e., the
422	significance test. Results show that the trend pattern becomes more significant and the
423	significance test becomes more stable when the time series is longer than a certain period like
424	30 years or so, the widely defined "climate" time scale (Figure 5). Using the DWS approach,
425	in both time series of mean air temperature and potential evaporation, the identified trend
426	patterns are found significant (Figure 5). Moreover, significance of trend patterns in the PET
427	time series obtained gotten-from the DWS approach and the MK test has obviously different
428	spatial distributions (Figure 6). Variability of hydroclimate process at large time scales,
429	especially for non-monotonic trend patterns, would be underestimated by the MK test, which
430	causes a great difficulty in understanding and interpreting the spatiotemporal variability of
431	hydroclimate process. Comparatively, the developed DWS approach can quantitatively assess
432	the statistical significance of non-monotonic trend pattern in the hydroclimate process, and so
433	can meet practical needs much better.
434	In summary our results suggest that the non-monotonic trend pattern of hydroclimate

In summary, our results suggest that the non-monotonic trend pattern of hydroclimate time series and its statistical significance should be carefully identified and evaluated, and the DWS approach developed in this study has the potential for wide<u>r</u> use in hydrological and climate sciences.

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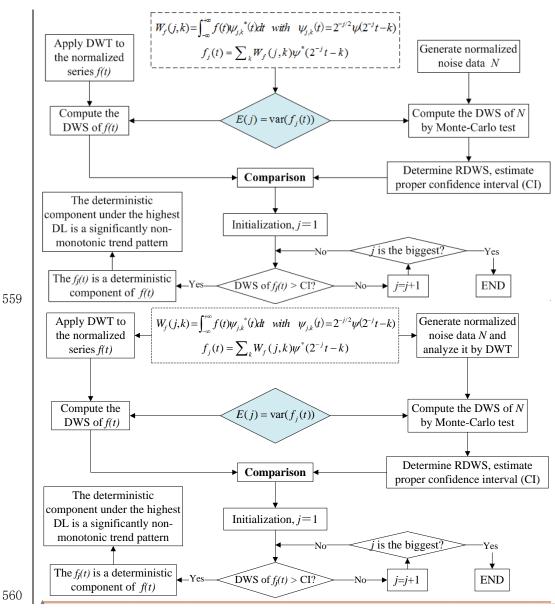
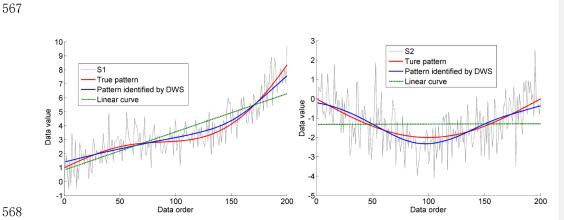


Figure 1. Technical flowchart for identification of the non-monotonic trend pattern in a time series using the discrete wavelet spectrum approach developed. In the figure, "DWT" is the discrete wavelet transform, "DWS" is the discrete wavelet spectrum, "RDWS" is the reference discrete wavelet spectrum, "DL" is the decomposition level, and "CI" is the confidence interval.

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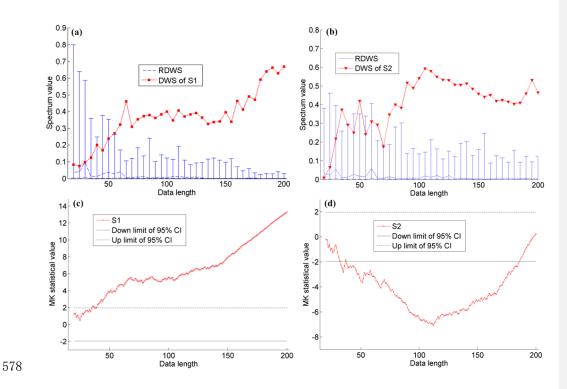


**Figure 2.** Non-monotonic trend patterns in the synthetic series S1 and S2 identified by the discrete wavelet spectrum (DWS) approach, and the linear trends in the two series. Synthetic series S1 is generated as:  $S1=1.112^{0.1t}+0.8 \times \sin(0.01\pi t)+\alpha$ ; and synthetic series S2 is generated as:  $S2=\sin(0.04\pi t)+2 \times \sin(\pi+0.005\pi t)+\alpha$ , where  $\alpha$  is a random process following the standard normal distribution.

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579 Figure 3. Evaluation of statistical significance of non-monotonic trend patterns in the 580synthetic series S1 (a) and S2 (b) with different data length by the discrete wavelet spectrum (DWS) approach, and the results by the Mann-Kendall (MK) test (c and d). In 581 582figure a and b, the blue line is the reference discrete wavelet spectrum (RDWS) with 95% 583 confidence interval under each data length; if the red point at certain data length is above 584 the blue bar, it is thought that the trend pattern is significant at 95% confidence level.; and 585i In figure c and d, the two black dash lines indicate 95% confidence interval (CI) with the 586 thresholds of +/- 1.96 in the MK test.

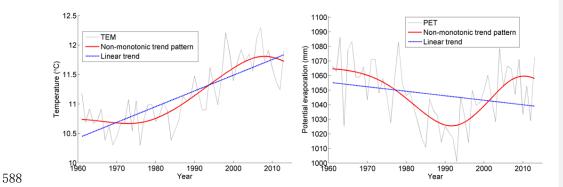
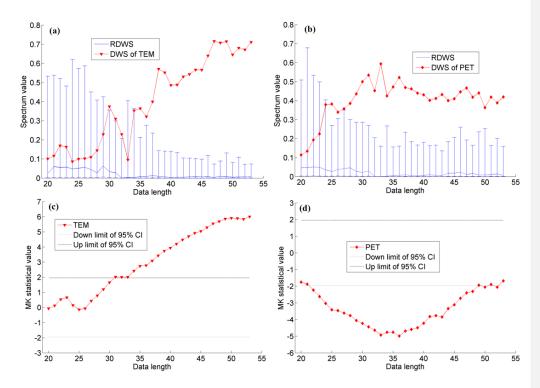


Figure 4. Non-monotonic trend patterns in the annual time series of the mean air temperature (TEM) and the potential evaporation (PET) over China from 1961-2013 identified by the discrete wavelet spectrum (DWS) approach, and the linear trends in the two series.

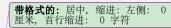


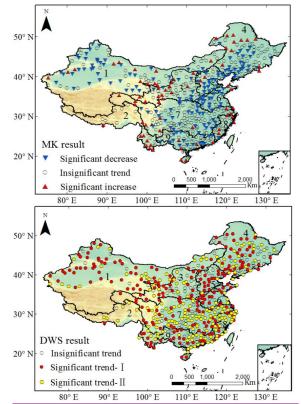
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Figure 5. Evaluation of statistical significance of non-monotonic trend patterns in the annual time series of the mean air temperature (TEM, a) and the potential evaporation (PET, b) over China with different data length by the discrete wavelet spectrum (DWS) approach, and the results by the Mann-Kendall (MK) test (c and d). In figure a and b, The blue line is the reference discrete wavelet spectrum (RDWS) with 95% confidence interval under each data length; and in figure c and d, the two black dash lines indicate 95% confidence interval (CI) with the thresholds of +/- 1.96 in the MK test.

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610	80° E 90° E 100° E 110° E 120° E 130° E
611	Figure 6. Spatial distribution of the significance of trends in the annual potential evaporation
612	data during 1961-2013 and measured at 520 weather stations over China. The result
613	above was gotten from the Mann-Kendall (MK) test. The result below was gotten from
614	the discrete wavelet spectrum (DWS) approach developed, in which significant trend-I
615	means those significant trends (at 230 stations) can be identified by both the DWS
616	approach and the MK test, but significant trend-II means those significant trends (at 150
617	stations) can only be identified by the DWS approach but not the MK test. 1, the
618	Northwest Inland River basin; 2, the Southwest River basin; 3, the Yellow River basin; 4,
619	the Songliao River basin; 5, the Haihe River basin; 6, the Huaihe River basin; 7, the
620	Yangtze River basin; 8, the Southeast River basin; and 9, the Pearl River basin.