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**Abstract.** The density of new snow is operationally monitored by meteorological or hydrological services at daily time intervals, or occasionally measured in local field studies. However, meteorological conditions and thus settling of the freshly deposited snow rapidly alter the new snow density until measurement. Physically based snow models and now-casting applications make use of hourly weather data to determine the water equivalent of the snowfall and snow depth. In previous studies, a number of empirical parameterizations were developed to approximate the new snow density by meteorological parameters. These parameterizations are largely based on new snow measurements derived from local in-situ measurements. In this study a data set of automated snow measurements at four stations located in the European Alps is analysed for several winter seasons. Hourly new snow densities are calculated from the height of new snow and the water equivalent of snowfall. Considering the settling of the new snow and the old snowpack, the average hourly new snow density is 68 kg m<sup>-3</sup> with a standard deviation of 9 kg m<sup>-3</sup>. Seven existing parameterizations for estimating new snow densities were tested against these data, and most calculations overestimate the hourly automated measurements. Two of the tested parameterizations were capable of simulating low new snow densities observed at sheltered inner-alpine stations. The observed variability in new snow density from the automated measurements could not be described with satisfactory statistical significance by any of the investigated parameterizations. Applying simple linear regressions between new snow density and wet bulb temperature based on the measurements data resulted in significant relationships ( $r^2 > 0.5$  and  $p \leq 0.05$ ) for single periods at individual stations only. Higher new snow density was calculated for the highest elevated and most wind exposed station location. Whereas snow measurements using ultrasonic devices and snow pillows are appropriate for calculating station mean new snow densities, we recommend instruments with higher accuracy e.g. optical devices for more reliable investigations of the variability of new snow densities on sub daily intervals.

## 1 Introduction

In mountain regions there is an increasing demand for high-quality analysis, now-casting and short range forecasts of the spatial distribution of snowfall. Operational services, concerning avalanche warning, road maintenance and hydrology, as

well as hydropower companies and ski resorts need reliable information on the depth of new snow (HN) and the water equivalent (HNW) of snowfall. Therefore the new snow density ( $\rho_{\text{HN}}$ ) is needed to convert HN into HNW and vice versa. Information on HN is especially relevant for cold and windy conditions, when measuring HNW is a difficult task because conventional rain gauge measurements are prone to large errors (e.g. Goodison et al., 1998). Recent results of the Solid Precipitation Intercomparison Experiment (SPICE; Nitu et al., 2012) reveal that these errors still exist in standard meteorological measurements (e.g. Buisan et al., 2016; Pan et al., 2016). Many snow cover models calculate HN from HNW on subdaily time intervals, although reliable HNW input data are difficult to obtain (Egli et al., 2009), and thus the new snow density is needed in equal temporal resolution to convert between HNW and HN (e.g. Lehning et al., 2002; Roebber et al., 2003; Olefs et al., 2013). Additionally,  $\rho_{\text{HN}}$  has a considerable effect on the snow bulk density of the total snowpack (e.g. Schöber et al., 2016).

Since the 1960s ultrasonic rangiers have become more common for observing snow depth changes automatically even on sub-hourly time intervals (e.g. Gubler, 1981; Goodison et al., 1984; Lundberg et al., 2010). They have the advantage of a more objective method compared to subjective manual measurements of snow depth (Ryan et al. 2008). Beside snow depth (HS), the water equivalent of the snowpack (SWE) is observed operationally using weighing devices such as lysimetric snow pillows (e.g. Serreze et al. 2009; Egli et al., 2009; Lundberg et al., 2010; Krajci et al., 2017) and snow scales (e.g. <http://www.sommer.at/en/products/snow-ice/snow-scales-ssg>). Upward looking GPR (e.g. Heilig et al., 2009), GPS techniques (e.g. Koch et al., 2014; McCreight et al., 2014) and the combination of both (Schmidt et al., 2015) have been applied in scientific studies to monitor the depth, SWE and liquid water content of the snowpack. However, these techniques are rather expensive or not yet in use for long-term observations by operational services. In general, automatic measurements of SWE are prone to a high relative uncertainty and require a certain degree of maintenance, which makes them complex and labour-intensive (Smith et al., 2017). Due to such constraints, SWE measurement instrumentation is installed at considerably fewer stations compared to HS instruments, and only at sites with easy access for appropriate maintenance. Recent studies present the performance of cosmic ray neutron sensors (e.g. Schattan et al., 2017), which partly are used for long-term observations such as e.g. from Col de Porte (Morin et al., 2012).

The density of new snow is influenced by the shape and size of the snow crystals (e.g. Nayaka, 1951). Relationships between predominant snow crystal type, riming properties and snowfall density were already reported by Power et al. (1964) from snowstorm observations in Canada. Once the snow crystals have accumulated at the snow surface, the density of the fresh snow starts to increase depending on prevailing weather conditions and compaction caused by overlaying of snow. A common mean  $\rho_{\text{HN}}$  used to convert between HN and HNW is  $100 \text{ kg m}^{-3}$ . Many studies analysed  $\rho_{\text{HN}}$  values on a daily basis and confirmed this 10:1-rule as applicable for a first estimate (e.g. Roebber et al., 2003; Egli et al., 2009; Teutsch, 2009). However,  $\rho_{\text{HN}}$  span a wide range and values from 10 to  $350 \text{ kg m}^{-3}$  have been reported from American and European mountain ranges, with mean values between 70 and  $110 \text{ kg m}^{-3}$  (e.g. Diamond and Lowry, 1954; LaChapelle, 1962; Power et

al., 1964; Judson, 1965; McKay et al., 1981; Meister, 1985; Judson and Doesken, 2000; Valt et al., 2014). Most of the  $\rho_{\text{HN}}$  data analysed in these studies were observed using readings on a snow board. The density is calculated from HN measured with a ruler and HNW is derived from an external precipitation device or from weighing the new snow either in solid or melted form (Fierz et al., 2009).

5 Several studies have shown that measured  $\rho_{\text{HN}}$  can be related to meteorological parameters, although with different time intervals and different degrees of determination. In 1952, Gold and Power showed that the crystal type is related to its estimated formation temperature. Diamond and Lowry (1954) and Simeral et al. (2005) built an empirical calculation that ascertained relationships between  $\rho_{\text{HN}}$  and air temperature at the 700-mb level. Teutsch (2009) also concluded that  $\rho_{\text{HN}}$  of 12

10 hour intervals at valley stations is best correlated to the wet bulb temperature at mountain stations in close vicinity ( $r^2 = 0.86$ ). Judson and Doesken (2000) found that near-surface air temperature and new snow density at mountain stations could explain 52 % of the variance in snow density. Wetzel et al. (2004) presented a similar degree of correlation of  $\rho_{\text{HN}}$  to temperature at three high-elevation sites. Alcott and Steenburg (2009) showed that  $\rho_{\text{HN}}$  is correlated with near-crest-level temperature and wind speed particularly for high-SWE events. Wright et al. (2016) presented a statistical analysis of data from 42 seasons of manual daily snow density measurements along with air temperature and wind speed to derive

15 parameterizations to estimate new snow density. However, they end up with a low coefficient of determination. On the basis of data from 7 stations in Switzerland located between 1250 and 1800 m a.s.l., Meister (1985) concluded that  $\rho_{\text{HN}}$  does not correlate with the amount of new snow (HN), that it does not depend on altitude, and that air temperature does not accurately determine  $\rho_{\text{HN}}$ . Nevertheless, binning the data into temperature classes results in a statistical equation with a correlation coefficient of 0.85. Further, he recommended considering wind speed in addition to air temperature, at least for

20 stations higher than 1800 m a.s.l. On the basis of data sets from Schmidt and Gluns (1991) and the US Army Corps of Engineers (1956), Hedstrom and Pomeroy (1998) developed a power function using the air temperature for which they found a coefficient of determination of 0.84 and a standard error of estimate of  $9.3 \text{ kg m}^{-3}$ . Jordan et al. (1999) introduced an algorithm for assigning  $\rho_{\text{HN}}$  within the SNTHERM snow cover model. They added wind dependence to the temperature parameterization of Meister (1985). This achieved a reduction of the error, but a significant scatter remained between

25 observed and parameterized  $\rho_{\text{HN}}$  values. Lehning et al. (2002) built an empirical calculation for  $\rho_{\text{HN}}$  valid for a time interval of 30 to 60 minutes in the framework of the snow model SNOWPACK. They used air temperature, surface temperature, relative humidity and wind speed for the regression analysis and achieved an approximate multiple coefficient of determination of 0.83. Schmucki et al. (2014) used another empirical power relation, including air temperature, wind speed and relative humidity, to calculate the  $\rho_{\text{HN}}$  using SNOWPACK simulations for three contrasting sites in Switzerland.  $\rho_{\text{HN}}$

30 were analysed in short time intervals of one to two hours by Ishizaka et al. (2015). They measured even lower densities in comparison to  $\rho_{\text{HN}}$  estimates obtained using the SNOWPACK density model, especially for aggregated snow crystal types. On the basis of data from Col de Porte (1325m altitude, French Alps), Pahaut et al. (1976) developed a statistical relationship

including the melting point of water, air temperature and wind speed. This parameterization is used to calculate the density of new snow in the snow cover model CROCUS (Vionnet et al., 2012).

Settling of the new snow by its weight and destructive metamorphism may reduce HN and hence increase  $\rho_{\text{HN}}$  between snowfall and the HN reading and has to be considered when computing new snow density (e.g. Anderson 1976; Lehning et al., 2002; Steinkogler, 2009; Vionnet et al., 2012). The contribution of settling to snow depth changes is highest in the first hours after a snowfall. Wind drift and radiation input to the snow surface after the snowfall may increase  $\rho_{\text{HN}}$  in comparison to  $\rho_{\text{HN}}$  at the time of snowfall. However, direct measurements of  $\rho_{\text{HN}}$  at the time of snowfall are laborious and difficult to align with the hours of peak snowfall rates.

Whereas most of the studies have analysed daily and sub-daily, manual  $\rho_{\text{HN}}$  measurements, to our knowledge no extensive analysis of automated  $\rho_{\text{HN}}$  measurements in hourly intervals over several winter seasons exists. The aim of this study is to assess the value of automated measurements of hourly HN and HNW for the calculation of  $\rho_{\text{HN}}$  at different stations and in hourly time interval. Therefore we examine the following questions:

(1) Are automated measurements of HN and HNW suitable for the calculation of  $\rho_{\text{HN}}$  at hourly interval?

(2) How do the mean and the variability of observed  $\rho_{\text{HN}}$  differ between distinct study sites?

(3) How well do established density parametrisations represent observed hourly  $\rho_{\text{HN}}$  values?

To this end, we calculated  $\rho_{\text{HN}}$  from hourly snow depth changes (HN) and hourly SWE changes (HNW). The mean values and the variability of hourly  $\rho_{\text{HN}}$  are discussed for observations at four different meteorological stations and compared to calculations using established  $\rho_{\text{HN}}$  parameterizations. A critical assessment with outlook on next generation measurements techniques is given in the discussion.

## 2 Data and Methods

Data from four automatic weather stations (AWS) were used in this study (Fig. 1, Table 1). A prerequisite for the station selection was the combined measurement of HS and SWE at each station in addition to the standard meteorological measurements of air temperature, relative humidity, precipitation, wind speed and global radiation. HS data are measured using ultrasonic rangiers. SWE data are recorded using snow pillows. Details regarding the instruments at and the exact location of each AWS, as well as the start and end dates of the available data coverage are presented in Table 1.

The K uhroint station (Germany) is operated by the Bavarian Avalanche Warning Service. It is a well-equipped and maintained station for snow climate at the northern fringe of the Eastern Alps. It is located in a meadow below treeline.

The Kühtai station (Austria) is operated by the Tiroler Wasserkraft AG (TIWAG). It is located south of the Inntal valley, but north of the Alpine main ridge, and it is situated at a wind-sheltered location.

The station at Wattener Lizum (Austria) is operated by the Austrian Research Centre for Forests (BFW) of the Federal Ministry of Agriculture, Forestry, Environment and Water Management. This station is situated in a south-north oriented high alpine valley above the treeline near to the Alpine main ridge.

The station at Weissfluhjoch (Switzerland) is operated by the Institute for Snow and Avalanche Research (SLF), which is part of the Swiss Federal Institute for Forest, Snow and Landscape Research (WSL). Weissfluhjoch is the highest elevated station considered in this study.

On the basis of coinciding data availability we consider four time periods as presented in Table 1. Data outputs of the AWS are logged at time intervals ranging from 2 to 30 minutes. Hourly values were computed for global radiation, relative humidity, air temperature and wind speed. The hourly value is the mean of the previous hour. For precipitation it is the sum of the previous hour. To account for noise in the ultrasonic signal, HS and SWE were smoothed using a centred moving average over 3 values in the original data resolution. The hourly values for HS and SWE are the instantaneous values on the full hour from the smoothed time series.

The thermodynamic wet bulb temperature ( $T_w$ ) was computed applying the psychrometric equation (Sonntag, 1990) and an exact iterative approach presented by Olefs et al. (2010). A standard barometric equation was used to determine air pressure based on the station elevation. Air pressure dependency of  $T_w$  is generally minor and only relevant for air temperatures larger than +2°C (Olefs et al., 2010).

A necessary condition for all further analysis of the time series was the presence of a precipitation signal at the heated precipitation gauges in combination with positive snow depth changes. Then, the hourly height of new snow (HN) and the water equivalent of snowfall (HNW) were computed as the change in HS and SWE. Within the next filtering step, only HN and HNW values with  $T_w$  less than 0°C and a wind speed ( $u$ ) of less than 5 m s<sup>-1</sup> were considered.

Constraints have to be made in order to avoid low values of HNW and HN, which are prone to large relative errors due to random and systemic measurement uncertainties in HN and SWE, but a minimum of approx. 100 remaining samples for statistical analysis must be ensured.

To investigate the influence of different minimum HNW and HN limits, a distribution matrix was calculated by varying the minimum HNW and HN limits in steps of 0.5 mm for HNW and 0.5 cm for HN, respectively. To account for settling during ongoing snowfall, the compaction correction described in Anderson (1976) was applied. The approach was simplified with respect to HS, SWE and snow density by considering only two layers of the snowpack: the new snow and the total snowpack of the previous time step. Destructive settling ( $S$ ) of HN is considered for each time step where the snow depth increases (Eq. 1). The destructive settling of the new snow ( $S_{HN}$ ) for each time step is calculated by

$$S_{\text{HN}} = -0.000002777 \cdot e^{(0.04 \cdot T)} \quad \{\rho_{\text{HN}} \leq 150 \text{ kg m}^{-3}\} \quad (1a)$$

$$S_{\text{HN}} = S_{\text{HN}} \cdot e^{(0.046 \cdot T \cdot (\rho_{\text{HN}} - 150))} \quad \{\rho_{\text{HN}} \geq 150 \text{ kg m}^{-3}\}, \quad (1b)$$

where  $T$  is the air temperature. Settling of the new snow layer caused by the weight of the ongoing snow accumulation is not taken into account.

Settling within the old snowpack is computed considering the total snow depth (HS). The destructive settling within the old snow layer ( $S_{\text{HS}}$ ) is calculated using Eq. (1), substituting HS for HN and using the bulk density of the old snowpack ( $\rho_{\text{HS}}$ ) calculated from HS and total SWE of the previous time step. Settling within the old snowpack caused by the weight of the snowpack ( $S_{\text{wHS}}$ ) is given as:

$$S_{\text{wHS}} = -248.976 \cdot \frac{HN}{3600000} \cdot e^{0.8 \cdot T} \cdot e^{-0.021 \cdot \rho_{\text{HS}}} . \quad (2)$$

The resulting settling factors of  $S_{\text{HN}}$ ,  $S_{\text{HS}}$  and  $S_{\text{wHS}}$  are multiplied with HS and HN to adjust HN accordingly.

New snow density ( $\rho_{\text{HN}}$ ) was obtained from the ratio of HN to HNW. Outliers below the 5 % percentile and higher than the 95 % percentile were excluded. The  $\rho_{\text{HN}}$  data were grouped by wet bulb temperature and wind speed, using bins of 1°C and 0.5 ms<sup>-1</sup> respectively. A least squares regression was carried out using both the ungrouped data and the median of the grouped data to quantify possible correlations of  $\rho_{\text{HN}}$  with  $T_w$  and  $u$ .

The  $\rho_{\text{HN}}$  were compared to the following parameterizations developed in previous studies. In these parametrisations,  $\rho_{\text{HN}}$  is a function of meteorological parameters such as air temperature ( $T$ ), wind speed ( $u$ ) and relative humidity ( $rH$ ). The time interval for  $\rho_{\text{HN}}$  readings of the respective study is given in the brackets.

$$\rho_{\text{HP}} = 67.92 + 51.52 \cdot e^{\frac{T}{2.59}} \quad (\text{Hedstrom and Pomeroy 1998, event/daily}) \quad (3)$$

$$\rho_{\text{D}} = 119 + 6.48 T \quad (\text{Diamond and Lowry 1954, frequent interval during event}) \quad (4)$$

$$\rho_{\text{LC}} = 50 + 1.7 \cdot (T + 15)^{1.5} \quad (\text{LaChapelle 1962, event}) \quad (5)$$

$$\rho_{\text{J}} = 500 \cdot (1 - 0.951 \cdot e^{-1.4 \cdot (5 - T)^{-1.15} - 0.008 \cdot u^{1.7}}) \quad \{-13^\circ \text{ C} < T \leq 2.5^\circ \text{ C}\} \quad (6a)$$

$$\rho_{\text{J}} = 500 \cdot (1 - 0.904 \cdot e^{-0.008 \cdot u^{1.7}}) \quad \{T \leq 13^\circ \text{ C}\} \quad (\text{Jordan et al., 1999, event/daily}) \quad (6b)$$

$$\rho_{\text{V}} = 109 + 6 \cdot (T - T_f) + 26 u^{0.5} \quad (\text{Vionnet et al., 2012, event/daily}) \quad (7)$$

$$\rho_{\text{S}} = 10^{3.28 + 0.03T - 0.36 - 0.75 \cdot \arcsin(\sqrt{0.01 \cdot rH} + 0.03 \cdot \log_{10} u)} \quad \{T \geq -14^\circ \text{ C}\} \quad (8a)$$

$$\rho_S = 10^{3.28 + 0.03T - 0.75 \cdot \arcsin(\sqrt{0.01 rH} + 0.03 \cdot \log_{10} u)} \quad \{T < -14^\circ \text{ C}\} \quad (\text{Schmucki et al., 2014, event/hourly}) \quad (8b)$$

$$\rho_L = 70 + 6.5 T + 7.5 T_s + 0.26 rH + 13 u - 4.5 T T_s - 0.65 T u - 0.17 rH u + 0.06 T T_s rH \quad (9)$$

(Lehning et al., 2002, event/hourly)

5 The melting point of snow ( $T_f$ ) in Eq. (7) was approximated as  $0^\circ\text{C}$  (Vionnet et al., 2012). Following Schmucki et al. (2014), we limited the parameter range and set  $rH$  to a constant value of 0.8 (80 %) during snowfall and the lower boundary for the wind speed to  $2 \text{ ms}^{-1}$ .

The temperature of the snow surface ( $T_s$ ) is required in Eq. (9). As this was not available for each station, we used the approximation  $T_s = T$ . We argue that  $T_s$  could not considerably exceed  $0^\circ$ , because of the maximum  $T_w$  of  $0^\circ\text{C}$ . Since only

10 precipitation events are considered,  $rH$  can be expected to be high, and thus difference between  $T_w$  and  $T$  is small.

The uncertainty of ultrasonic measurements on snow can be assumed to be in the range of  $\pm 1 \text{ cm}$ , which partly is a consequence of changes in signal velocity due to meteorological conditions. However, we used the original HS data logged in mm-resolution to avoid the effects caused by rounding to full cm when calculating HN. Likewise, we used the tenths mm SWE data logged at the pillows. Another documented error source of the HS measurement is signal blocking by e.g. dense

15 snowfall or drifting snow, which causes peaks of the HS. However, with the filtering procedure applied in this study, no such spikes were left in the analysis.

A source of uncertainty is the spatial offset between the HS measurements and the SWE measurements. HS is measured directly above the SWE measurement at Kühtai station, Kühroint station and Wattener Lizum station (Fig. 1). However, the footprint of the snow depth sensor may be smaller than the surface area of the pillow, and it is decreasing with increasing

20 HS. A spatial variability of HS on the pillow can be caused by snow drift and differing snow settling or snow melt.

For the calculations within this study we used the changes in HS and SWE over the time period of snowfall only. Errors due to spatial variability in HS and SWE caused by spatial differences in energy consumption and snow drift between precipitation events are reduced. This is especially valid for the HS and SWE measurements at the stations with matching HS and SWE measurements. The snow depth sensor and the snow pillow of Weissfluhjoch station are separated by 9 meters.

25 Schmid et al. (2014) suggest a small-scale variability in HS of  $\pm 4.3 \%$  at the Weissfluhjoch station. Again, the error may be smaller due to using temporally limited changes of HS, but an additional uncertainty of  $\pm 5 \%$  can be assumed here.

A well-known issue with snow pillows are bridging effects (e.g. Serreze et al., 1999; Johnson and Schaefer, 2002). Dense snow layers and crusts within the snowpack sustain the weight of the new snow so that HNw, and thus  $\rho_{HN}$ , are

30 underestimated. We cannot exclude such data explicitly. However, all filtering conditions have to be fulfilled for including values in the analysis, so that data without or with lagged HN increase were not considered. Additionally, the chosen snow stations are well maintained in case of implausible data due to their overall good accessibility. E.g. trenches are dug out

around the base area of the snow pillow at Kühtai station to cut off the measured part of the snowpack to avoid bridging effects.

Nevertheless, the measurement uncertainty is  $\pm 1$  cm for HN and 0.1 cm for HNW. Considering mean HN (Table 2) and HNW values, the uncertainty is  $\pm 25 \text{ kg m}^{-3}$  or 37 % of the mean density. This value is lower considering higher HN, but increases to 80 % for the combination of minimum HN and minimum HNW of 1.6 cm and 0.2 mm respectively.

### 3 Results and discussion

#### 3.1 Data filtering, correction of settling and evaluation

Figure (2) presents the median new snow density ( $\rho_{\text{HN}}$ ) data calculated from all filtered HN and HNW exceeding the respective minimum HN and HNW limits. This presentation highlights the variability of  $\rho_{\text{HN}}$  by using different minimum limits with respect to the high relative uncertainty of low HN and HNW values. Changing the minimum limits for HN and HNW affects the resulting  $\rho_{\text{HN}}$  considerably. However, increasing the minimum limits for HN and HNW results in a distinct lowering of the number of data remaining for the subsequent analysis (Fig. 2). There are certain differences between the stations for high minimum HNW limits. Calculated  $\rho_{\text{HN}}$  decrease when low minimum HN and high minimum HNW limits are applied at Kühtai and Wattener Lizum station. In contrast,  $\rho_{\text{HN}}$  increase for equal minimum limits at Kühroint and Weissfluhjoch station. At Kühtai and Wattener Lizum station, high HNW values of more than 3 mm HNW are accompanied by rather high HN (Fig. 2). In contrast, low HN occurring with high HNW at Kühroint and Weissfluhjoch cause high  $\rho_{\text{HN}}$ . However, these results are based on a small number of values only. In general, the calculated median  $\rho_{\text{HN}}$  are rather constant following the 1:1 line of minimum HNW and HN limits (Fig. 2).

In order to avoid low values of HNW and HN, but ensuring an appropriate number of approx. 100 samples and with respect to the results of the Figures 2, we decided to use a minimum limit of 1.5 mm in HNW and 2.0 cm in HN. This leads to the exclusion of on average 94 % of all data points that have a precipitation signal and positive snow depth changes (Table 2). Frequency distributions for HN, HNW,  $T_w$  and  $u$  of the unfiltered and filtered data are presented for each station and for each time period in the supplement Figures (S01) to (S09).

The exclusion of high wind speeds has only a small effect at the lower stations and is more noticeable at the more wind exposed stations of Wattener Lizum and Weissfluhjoch. Considering period 1 comprising all stations, the filtering process causes the highest filtering rate for Weissfluhjoch station, with 6 % of data remaining after applying the filtering. The overall highest amount of data reduction is found at Kühtai station, with 5 % of the data remaining after filtering of the longer



periods 3 and 4 (Table 2). There was a considerable fraction of data with positive HS changes, a precipitation signal and positive  $T_w$ . Most of these data seem to be paired with very small HS changes and are eliminated for the final data set applying the HN minimum limit.

The correction of the HN underestimation caused by settling of the snowpack during snowfall leads to an average reduction of mean  $\rho_{HN}$  of  $10.2 \text{ kg m}^{-3}$  with a standard deviation ( $\sigma$ ) of  $2.6 \text{ kg m}^{-3}$  (Table 2). This corresponds to 13.5 % with a  $\sigma$  of 3.7 %. The compaction correction causes noticeably less change in  $\rho_{HN}$  at Weissfluhjoch in period 1 (5 % reduction of mean  $\rho_{HN}$ ) than at the other time periods and stations. The next closest is K hroint, also in period 1, with a reduction in  $\rho_{HN}$  of 7 %. Unless otherwise stated in the text,  $\rho_{HN}$  always refers to the corrected densities hereafter. Based on a 15 year data set of Weissfluhjoch (WSL Institute for Snow and Avalanche Research SLF, doi:10.16904/1.) from 1 September 1999 to 31 December 2015, the contribution of settling relative to HN was calculated using the multi-layer SNOWPACK model (e.g. Lehning et al., 2002) and the approach from Anderson (1976) to compare the results of this study to a more physically based estimate. Results are presented in Fig. 3. While a median relative contribution of settling to HN by 19 % was calculated with SNOWPACK, the approach of Anderson (1976) resulted in lower values of 5 % in median and 9 % in mean. Thus, the settling considered for the presented data can be assumed to be lower estimation. However, higher contributions of settling would result in lower  $\rho_{HN}$  with increased HN assuming a fixed HNW.

Figure 4 shows the distribution of  $\rho_{HN}$  values obtained from the filtered data at K hroint station as a representative for all stations and periods (Fig. S10 to S17 in the supplement). In general, the  $\rho_{HN}$  show high variability at all stations. Nevertheless,  $\rho_{HN}$  values are within a reasonable range of less than  $200 \text{ kg m}^{-3}$ . The histograms of  $\rho_{HN}$  show one-tailed distributions towards higher  $\rho_{HN}$ . Median  $\rho_{HN}$  of the different stations and for different periods range between 66 and 86  $\text{kg m}^{-3}$  for uncorrected values and between 54 and 83  $\text{kg m}^{-3}$  for  $\rho_{HN}$  corrected for settling (Table 2).

### 3.2 Station dependent differences

The distributions of  $\rho_{HN}$ ,  $T_w$  and  $u$  during all filtered snowfall data are presented in Fig. 5 and 6 and in Table 2. The lowest  $T_w$  and highest wind speeds were observed during snowfall at Weissfluhjoch station. However, the range and distribution of  $T_w$  at Weissfluhjoch station result in higher median  $T_w$  during snowfall compared to  $T_w$  at Wattener Lizum station. With respect to wind speeds, Wattener Lizum is second. Lowest wind speeds at K htai station occur together with lowest  $\rho_{HN}$ . Weissfluhjoch station has the highest median  $\rho_{HN}$  by a large margin with  $83 \text{ kg m}^{-3}$  in period 1 compared to, respectively, 67, 61 and  $66 \text{ kg m}^{-3}$  at K hroint, K htai and Wattener Lizum station.

Wind influence may be the reason for higher  $\rho_{HN}$  at Weissfluhjoch station. Snow grains are fragmented by snow drift (e.g. Sato et al., 2008), and thus more packed into the layer of new snow during windy conditions even over the course of only one hour. The K htai station shows lowest  $\rho_{HN}$  and the difference of mean  $\rho_{HN}$  is  $17 \text{ kg m}^{-3}$  between Weissfluhjoch and

Kühtai station for period 1. Median  $\rho_{\text{HN}}$  and median  $T_w$  of the different periods show a relationship between the periods at Kühtai station, with higher  $\rho_{\text{HN}}$  for higher  $T_w$  (Fig. 6, Table 2).

The overall mean hourly  $\rho_{\text{HN}}$  of all stations and time periods is  $68 \text{ kg m}^{-3}$  with a standard deviation of  $9 \text{ kg m}^{-3}$ . In general, this is considerably lower than new snow densities from daily measurements (e.g. Roebber et al., 2003; Egli et al., 2009; Teutsch, 2009). Meister (1985) measured  $\rho_{\text{HN}}$  lower than  $100 \text{ kg m}^{-3}$  on a daily basis analysing data with a HN of more than 0.1 m. In contrast, the presented  $\rho_{\text{HN}}$  are closer to the time of the snowfall event, and density changes over several hours due to e.g. energy exchanges and wind drift at the uppermost snow layer can be excluded. On the basis of  $\rho_{\text{HN}}$  in-situ measurements in hourly resolution Lehning et al. (2002) emphasized that at sub-daily time intervals lower densities in comparison to daily new snow densities have to be applied. Comparatively low  $\rho_{\text{HN}}$  values close to  $50 \text{ kg m}^{-3}$  were also presented by Ishizaka et al. (2016), with an average  $\rho_{\text{HN}}$  of  $52 \text{ kg m}^{-3}$  for aggregated snowflakes and  $55 \text{ kg m}^{-3}$  for small hydrometeors. They further found a mean  $\rho_{\text{HN}}$  of  $72 \text{ kg m}^{-3}$  for a second group of smaller crystals and  $99.4 \text{ kg m}^{-3}$  for graupel type hydrometeors.

The observed inter-station variability shows the importance of differing  $\rho_{\text{HN}}$  between more windy mountain stations and less windy stations in the valleys.

### 3.3 Density parametrisations

A simple linear regression analysis showed that the short term variability of  $\rho_{\text{HN}}$  cannot be explained with corresponding changes in  $T_w$  or  $u$  (Table 3, Fig. 7 and S18 to S26). An increase of  $\rho_{\text{HN}}$  with increasing  $T_w$  can be identified in the figures, and the slopes of the least squares regressions show an increase of  $\rho_{\text{HN}}$  with an increase of wet bulb temperature for all stations (Table 3). However, no consistent relationship between  $\rho_{\text{HN}}$  and  $u$  could be found neither for single stations nor for different periods at one station. The binned analysis based on  $T_w$  showed a considerable  $r^2$  of more than 0.5 on a 0.01 significance level at Kühroint and Kühtai station, with intercepts of 70 to  $80 \text{ kg m}^{-3}$  and gradients of about 3 to  $4 \text{ kg m}^{-3}$  per  $1^\circ\text{C}$ .

Although the regressions generally show the expected trends, it must be noted that the variability of  $\rho_{\text{HN}}$  remains unexplained. This could partly be attributed to the measurement uncertainties. However, the variability caused by measurement uncertainties is assumed to be equalized considering mean and median of  $\rho_{\text{HN}}$  values only for total time periods. Relationships between  $\rho_{\text{HN}}$  and  $T_w$  were recognized for distinct periods and stations only, but with a similar coefficients of determination in comparison to the results of e.g. Judson and Doesken (2000), Wetzel et al. (2004) or Wright et al. (2016).

Testing multiple regressions using additional meteorological parameters didn't increase the statistical significance. Instead a comparison to existing parameterizations of  $\rho_{\text{HN}}$  was performed for all stations and periods.

Considering the various parameterizations, which use meteorological parameters to approximate new snow density (Eq. 3 to 9), it is evident that the observed variability of  $\rho_{\text{HN}}$  is not correlated to the variability of parameterized new snow densities (Table 4). Most of the seven parametrizations overestimate the median of the observed  $\rho_{\text{HN}}$  values (Fig. 3, 7 and 8, Table 4). However, some parameterizations produce considerably better results than others for median  $\rho_{\text{HN}}$  values. The parameterizations of LaChapelle (1962), Diamond and Lowry (1954) and Vionnet et al., (2012) consistently overestimate  $\rho_{\text{HN}}$ . The parameterization of Hedstrom and Pomeroy (1998) overestimates  $\rho_{\text{HN}}$  at K hroint, K htai and Wattener Lizum station (Fig. 7 and 8), but converges with the median  $\rho_{\text{HN}}$  at Weissfluhjoch station for period 1 (Fig. 7, Table 4). In general, the  $\rho_{\text{HN}}$  simulated using the parameterization of Jordan et al. (1999) are closer to calculated  $\rho_{\text{HN}}$ , but median  $\rho_{\text{HN}}$  are underestimated for Weissfluhjoch station. Median  $\rho_{\text{HN}}$  and the range of  $\rho_{\text{HN}}$  at Weissfluhjoch are well simulated using the parameterization of Schmucki et al. (2014), but it overestimates median  $\rho_{\text{HN}}$  of K hroint, K htai and Wattener Lizum station (Fig. 3 and 7, Table 4). However, this parameterization was fitted to original density data from Weissfluhjoch.

The lowest root mean squared error (R) was achieved for Weissfluhjoch station with the parameterization of Diamond and Lowry (1954). The parameterizations of Lehning et al. (2002) and Jordan et al. (1999) result in lowest R (Table 4) compared to  $\rho_{\text{HN}}$  at K hroint, K htai and Wattener Lizum station, with slightly lower density values using the parameterization of Lehning et al. (2002) fitting best to the low median  $\rho_{\text{HN}}$  values of the K htai station.

Thus, the parameterization of Lehning et al. (2002) appears to be the first choice regarding the calculation of hourly new snow densities for high elevations and inner alpine regions. This parameterization requires multiple input parameters. Where such data is not available, the parameterization of Jordan et al. (1999), requiring temperature and wind data only, might be a good alternative. Even though, correlations are low in general, some of the highest Pearson correlation values (r, Table 4) were achieved by applying the simpler, linear equations by Diamond and Lowry (1954), LaChapelle (1962) and Vionnet et al. (2012). In addition to the regressions presented in Table 3, this shows again the identifiable relation between snow density and temperature.

Mair et al. (2015) evaluated some of the parameterizations also considered in this study. Using a distinctly larger time window for smoothing their HS data (5-hour-average), they calculated median  $\rho_{\text{HN}}$  between 75 and 100 kg m<sup>-3</sup> using the parameterizations of Jordan et al. (1999) and Hedstrom and Pomeroy (1998), which is close to the results presented in this study. They also found that using the parameterization of LaChapelle (1962) results in mean  $\rho_{\text{HN}}$  higher than 100 kg m<sup>-3</sup>. In general they concluded that using a constant  $\rho_{\text{HN}}$  of 100 kg m<sup>-3</sup> caused an overestimation of seasonal precipitation by up to 30 %. Conversely, a mean  $\rho_{\text{HN}}$  of 70 kg m<sup>-3</sup> will result in better SWE estimations. This is in accordance with the resulting average  $\rho_{\text{HN}}$  of 68 kg m<sup>-3</sup> calculated from automated measurements within our study.

## 4 Conclusion

The aim of this study was to assess the value of automated measurements of snow depth (HS) and snow water equivalent (SWE) to compute new snow density ( $\rho_{\text{HN}}$ ) on an hourly time interval. Complementary data sets of HS and SWE measurements using ultrasonic devices and snow pillows from four mountain stations were used to calculate the height of new snow (HN) and the water equivalent of snowfall (HNW). Subsequently,  $\rho_{\text{HN}}$  was calculated from HN and HNW considering potential underestimation of HN by settling of the snowpack.

The snow measurements using ultrasonic devices and snow pillows were found to be appropriate for the calculation of station average hourly  $\rho_{\text{HN}}$  values. However, the observed variability in  $\rho_{\text{HN}}$  from the automated measurements could not be described with appropriate statistical significance by any of the investigated algorithms. An average  $\rho_{\text{HN}}$  of  $68 \text{ kg m}^{-3}$  with a standard deviation of  $9 \text{ kg m}^{-3}$  was calculated considering all stations and time periods. Average  $\rho_{\text{HN}}$  for individual stations in a common period ranged from  $61 \text{ kg m}^{-3}$  to  $83 \text{ kg m}^{-3}$ , with higher  $\rho_{\text{HN}}$  at more windy locations. Thus, wind speed is a crucial parameter for the inter-station variability of  $\rho_{\text{HN}}$ .

Seven existing parameterizations for estimating new snow densities were tested, and most calculations overestimate  $\rho_{\text{HN}}$  in comparison to the results from the hourly automated measurements. Two of the tested parameterizations were capable of simulating low  $\rho_{\text{HN}}$  at sheltered inner-alpine stations. This reveals that it has to be carefully considered which parameterization should be used for which application and environment.

Nevertheless, the natural variability of  $\rho_{\text{HN}}$  is masked using the combination of ultrasonic ranging and snow pillow data for  $\rho_{\text{HN}}$  calculation, because of the limited accuracy of the sensors and snow depth changes due to settling of the snowpack and wind drift. We conclude that the value of the analysed data is given by the mean and median  $\rho_{\text{HN}}$  and its variation between different stations and time periods, and the considerably lower  $\rho_{\text{HN}}$  values in contrast to  $\rho_{\text{HN}}$  calculated on daily or event-based measurements.

The study shows the potential of collocated measurements of HS and SWE for determining  $\rho_{\text{HN}}$  automatically. However, recent developments in optical distance sensors and weighing devices increase the accuracy of such snow measurements and hence decrease the uncertainty of subsequent calculations. We therefore recommend the use of high accuracy sensors for the determination of  $\rho_{\text{HN}}$  on sub daily intervals.

## Data availability

A processed set of SNOWPACK input data from Weissfluhjoch station is available at: *WSL Institute for Snow and Avalanche Research SLF (2015): WFJ\_MOD: Meteorological and snowpack measurements from Weissfluhjoch, Davos, Switzerland; WSL Institute for Snow and Avalanche Research SLF; doi:10.16904/1.*

- 5 Detailed information about the Weissfluhjoch data set can be found in WSL Institute for Snow and Avalanche Research SLF (2015) and in Marty and Meister (2012). Data of Kühtai station are published by Krajčič et al. (2017) and are available from the Zenodo repository doi:10.5281/zenodo.556110.

Data of Kühroint station are available on request from the Bavarian avalanche service (<https://www.lawinenwarndienst-bayern.de/organisation/lawinenwarnzentrale/kontakt.php>).

- 10 Data of Wattener Lizum station are available on request from the Austrian Federal Research and Training Centre for Forests, Natural Hazards and Landscap (BFW, <https://bfw.ac.at/rz/bfwcms.web?dok=6057>).

## Author contribution

Kay Helfricht is the main investigator of this study. Lea Hartl performed snow density analysis within the pluSnow project.

- 15 Roland Koch performed initial quality control, provision and setup of project database for all station and meta data. Christoph Marty prepared the data of Weissfluhjoch station, contributed fruitful discussions and helped to focus the analysis and the manuscript. Marc Olefs contributed significantly to analysis and discussions as the main project partner within the framework of the pluSnow project.

## 20 Competing Interests

The authors declare that they have no conflict of interest.

## Acknowledgements

- 25 The pluSnow project is financed by the Gottfried and Vera Weiss Science Foundation (WWW). The project funding is managed in trust by the Austrian Science Fund (FWF): P 28099-N34. Project duration 10/2015 - 09/2018. The authors want to thank the colleagues of the Tiroler Wasserkraft AG (TIWAG), of the Federal Research and Training Centre for Forests

(BFW) and of the Bavarian avalanche service for data provision. In particular we are grateful for the close collaboration by Johannes Schöber (TIWAG) and Reinhard Fromm (BFW). We also want to thank Michael Lehning, Charles Fierz and the two reviewers for their helpful comments and fruitful discussion of the results.

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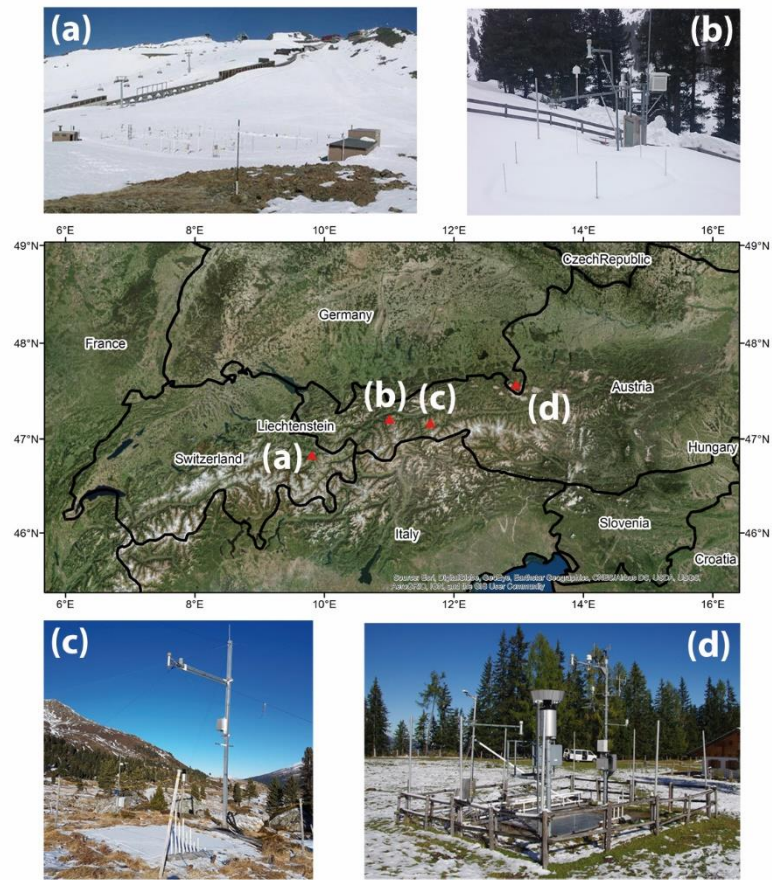
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**Figure 1: Map of the station locations. Pictures are given for (a) Weissfluhjoch station, (b) Kühtai station, (c) Wattener Lizum station and (d) Kühroint station.**

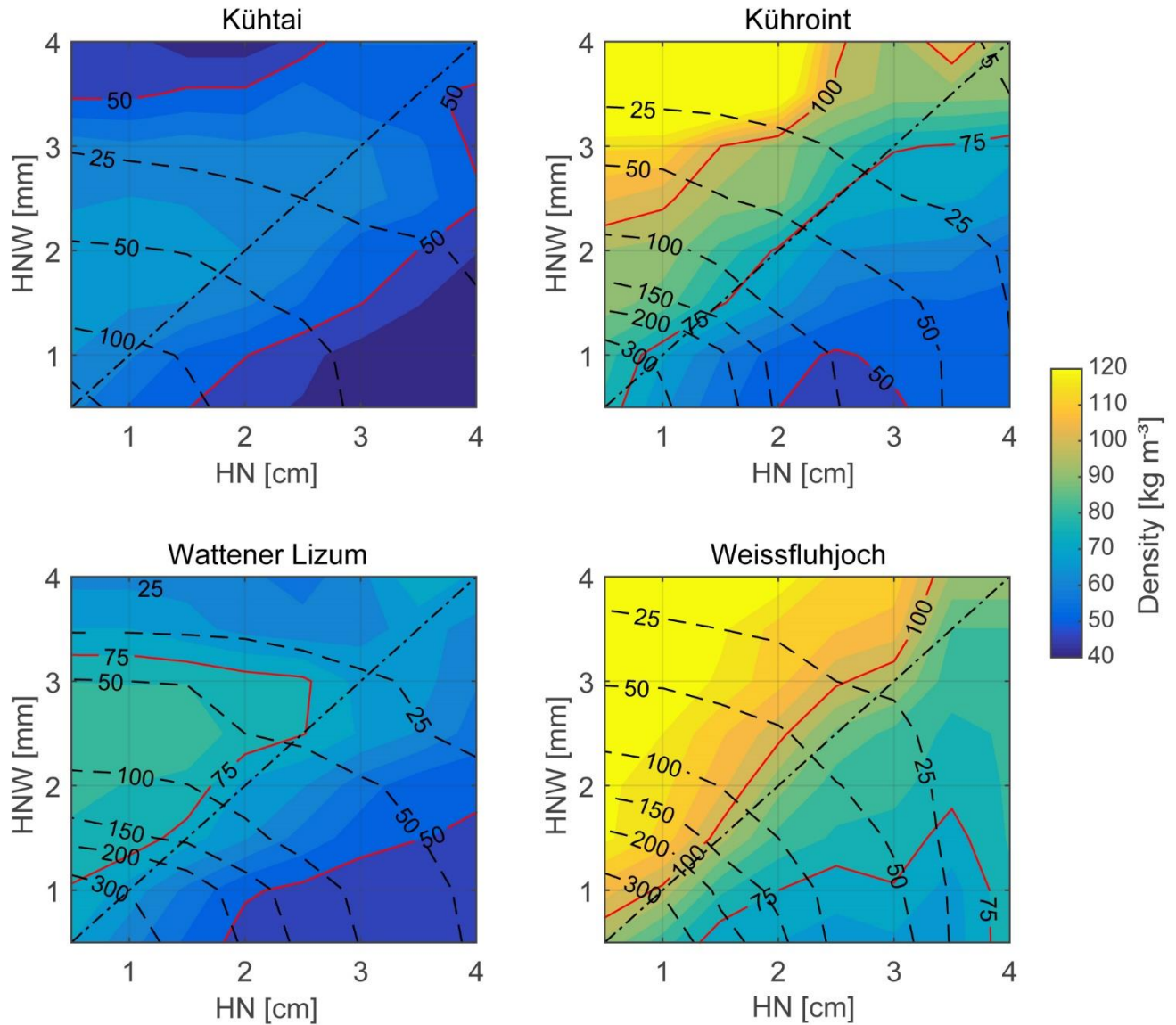
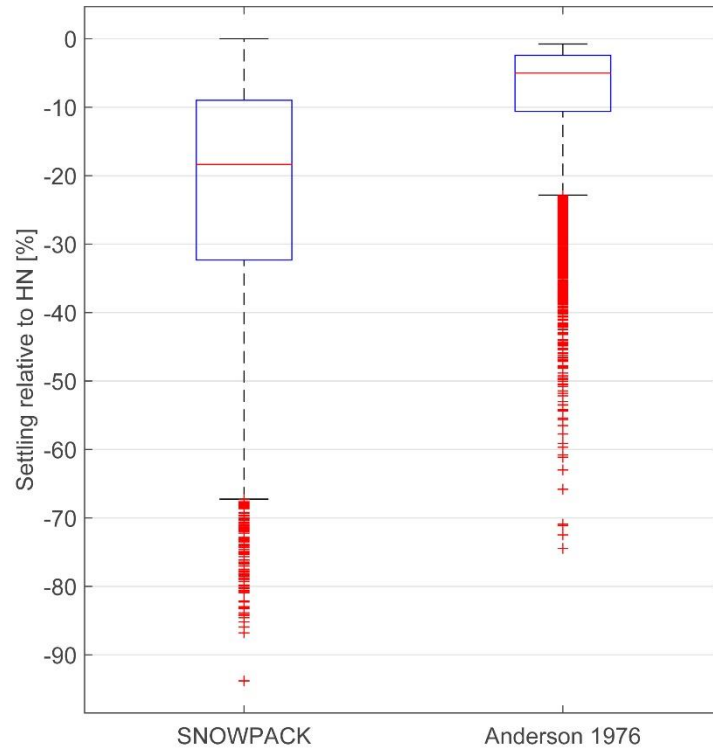
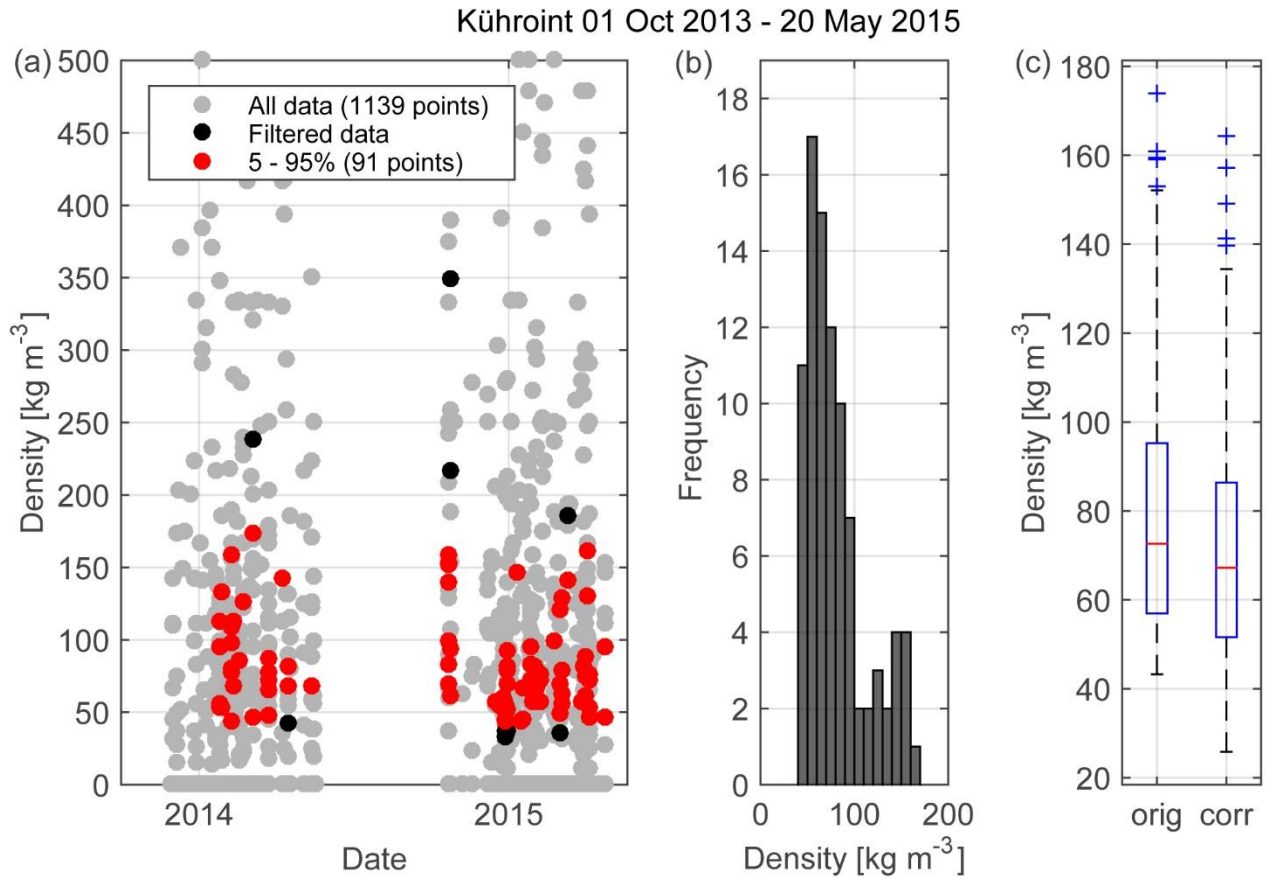


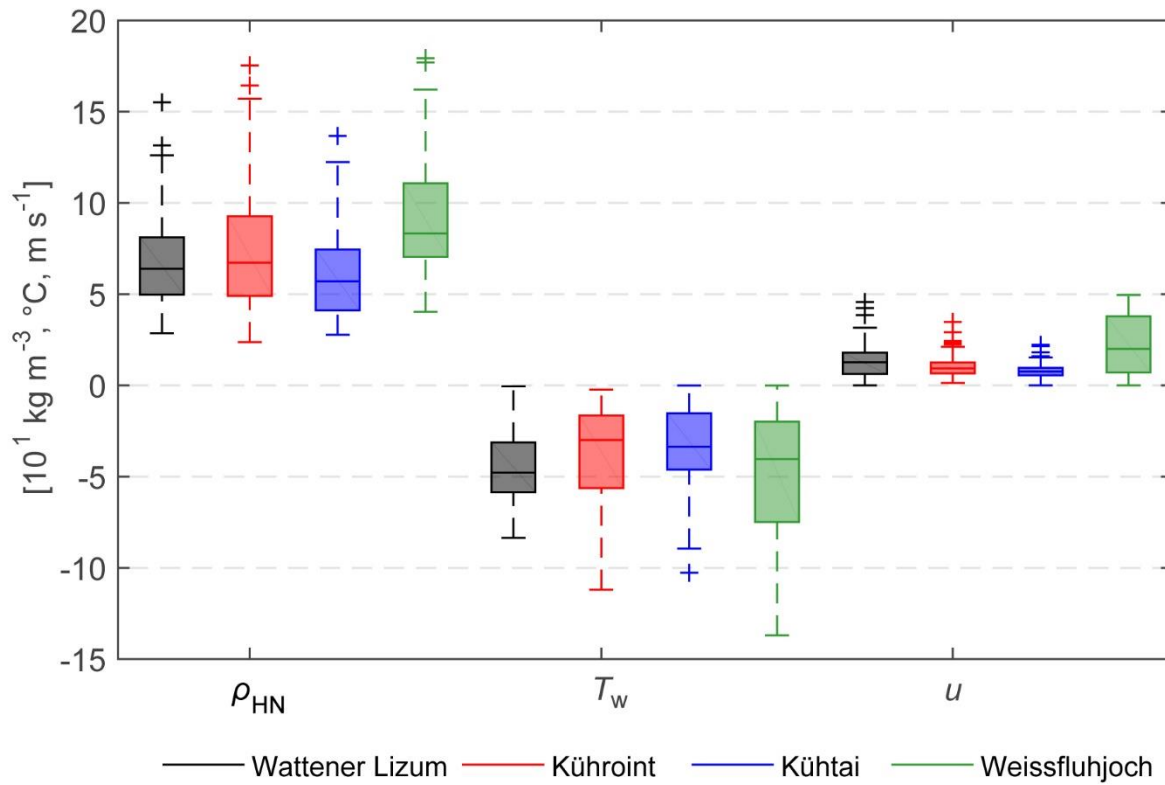
Figure 2: Median new snow densities (colour scale) calculated using all data exceeding different minimum limits of the height of new snow (HN) and the water equivalent of snowfall (HNW) for the period 1 (1 Oct 2013 - 20 May 2015). Note that multiples of 25  $\text{kg m}^{-3}$  are highlighted with red contour lines. The labelled black dashed lines give the count of the hourly data remaining after filtering. The straight dot-dashed lines show results for equal minimum limits of HN [cm] and HNW [mm].



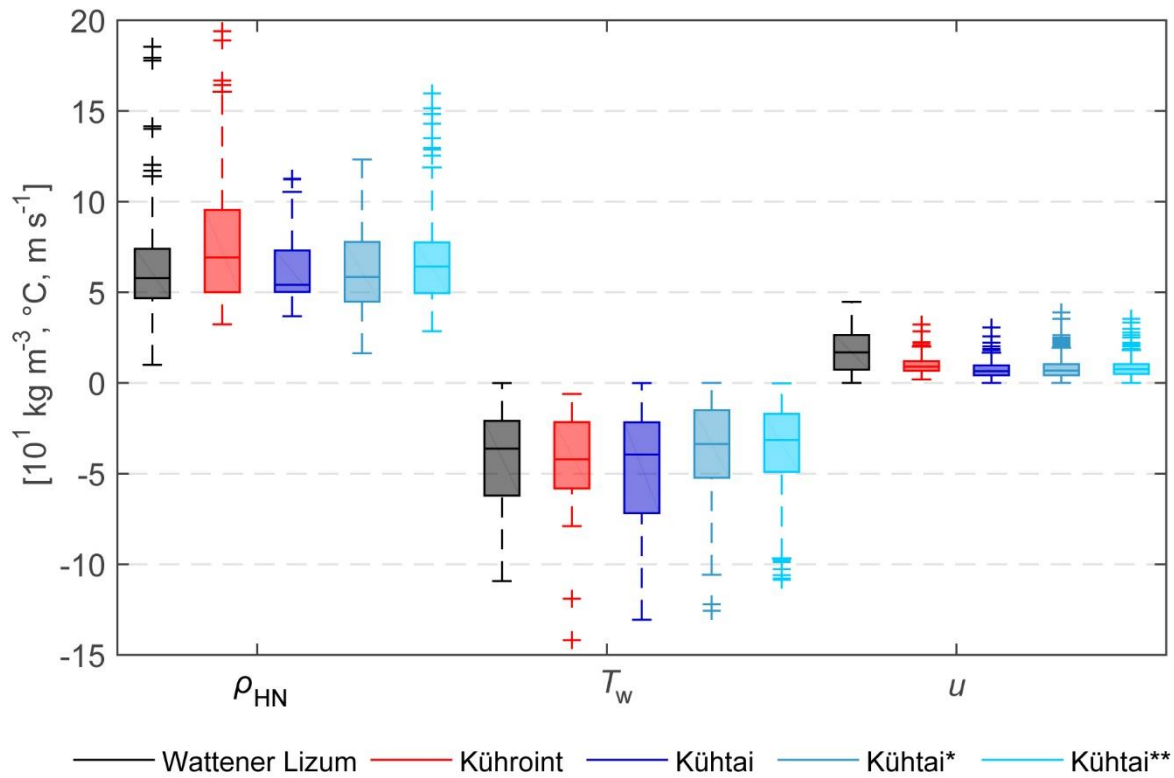
**Figure 3: Boxplots (Median, 25% and 75% percentiles, 1.5 x interquartile ranges, outliers) of settling relative to hourly new snow heights (HN) modelled with SNOWPACK and using the approach presented by Anderson (1976).**



**Figure 4:** Distribution of calculated new snow densities at Kühroint station for the period 1 (1 Oct 2013 - 20 May 2015). (a) All data with precipitation signal and positive HS change, all data filtered with  $\text{HN} > 2 \text{ cm}$ ,  $\text{HNW} > 1.5 \text{ mm}$ ,  $T_w < 0^\circ\text{C}$  and  $u < 5 \text{ ms}^{-1}$ , and filtered data reduced by cutting off at 5 % and 95 % percentiles. (b) Histogram of all filtered densities. (c) The boxplot showing median, 25 % and 75 % interquartile range of uncorrected densities and densities corrected for settling of the snowpack. Note that similar figures are available in the supplement (Fig. S10 – S17) for all stations and all time periods considered in this study.

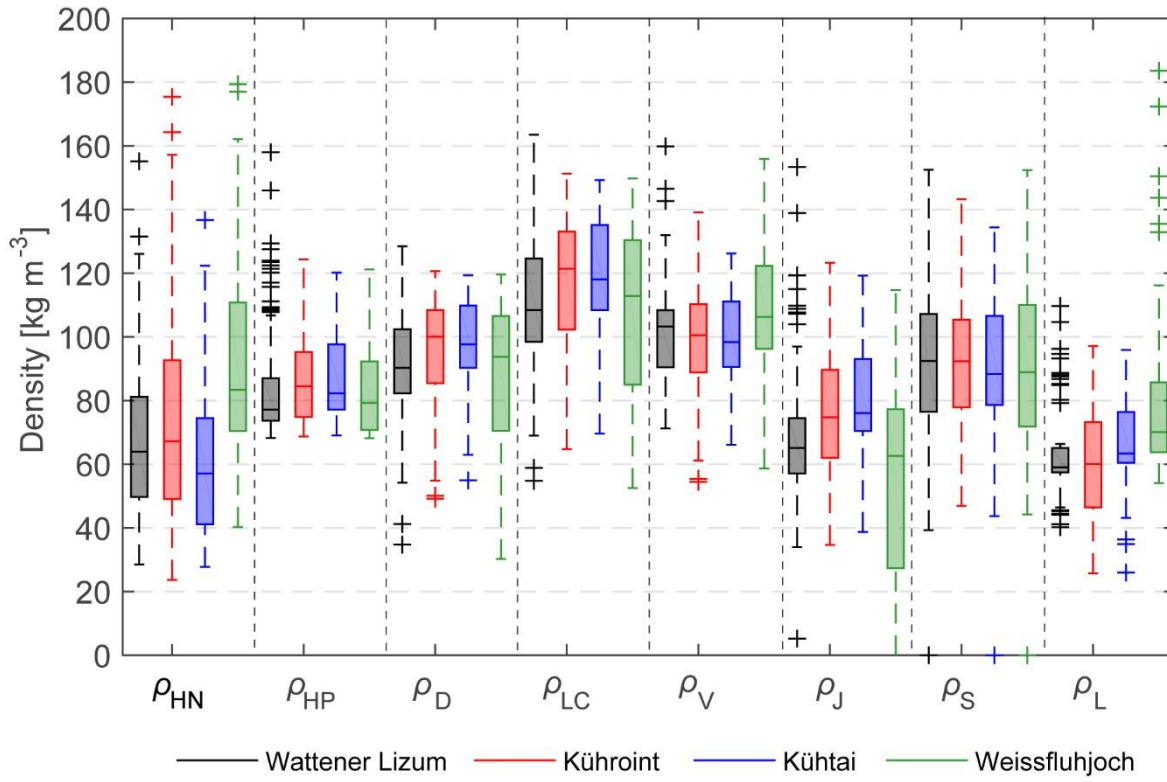


**Figure 5: Boxplot (Median, 25% and 75% percentiles, 1.5 x interquartile ranges, outliers) of calculated new snow densities ( $\rho_{HN}$ ) based on observations, wet bulb temperature ( $T_w$ ) and wind speed ( $u$ ) for filtered snowfall events (Table 2) at all four stations within period 1 (1 Oct 2013 - 20 May 2015).**



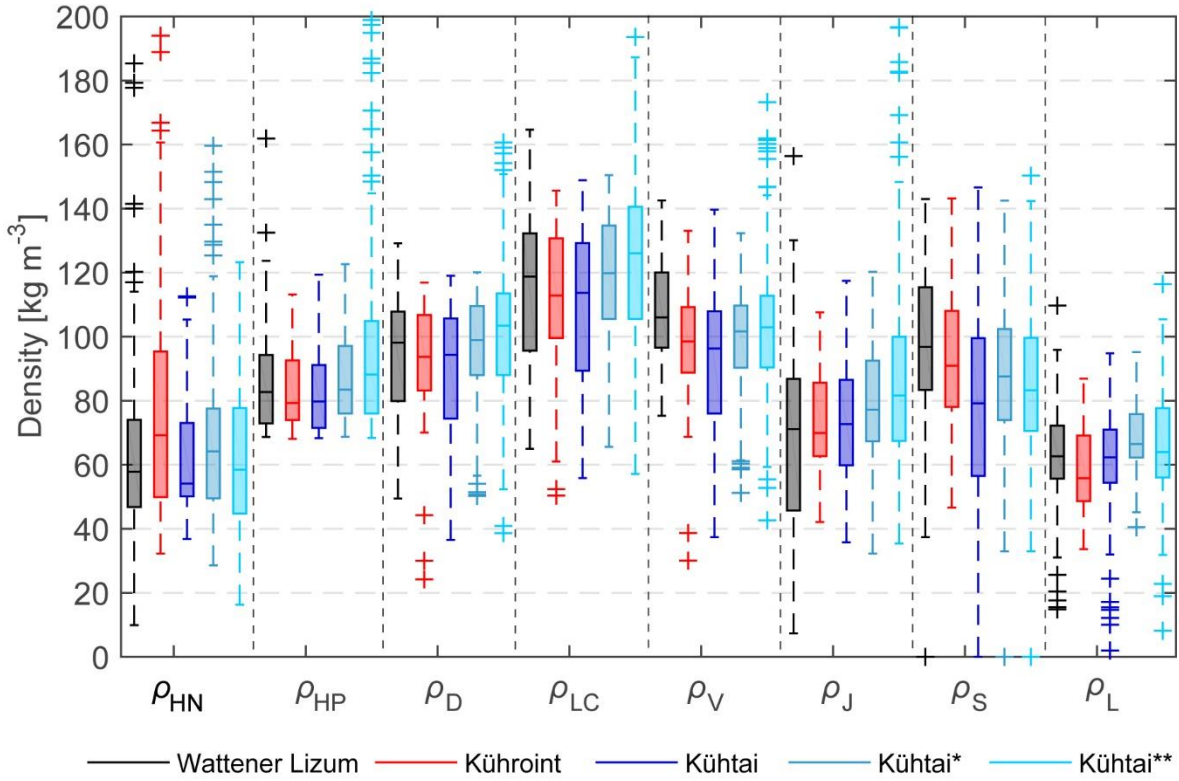
**Figure 6:** Boxplots (Median, 25% and 75% percentiles, 1.5 x interquartile ranges, outliers) of calculated new snow densities ( $\rho_{HN}$ ) based on observations, wet bulb temperature ( $T_w$ ) and wind speed ( $u$ ) for filtered snowfall events (Table 2) at three stations within period 2 (1 Oct 2011 - 01 Oct 2013) and at Kühtai station within period 3 (index \*, 01 Oct 1999 - 30 Sep 2011) and period 4 (index \*\*, 27 Feb 1987 – 30 Sep 1999) .





**Figure 7: Boxplots (Median, 25% and 75% percentiles, 1.5 x interquartile ranges, outliers) of calculated new snow densities ( $\rho_{HN}$ ) based on observations and densities calculated using parameterizations developed in previous studies (Eq. (3) – (9)) at all four stations within period 1 (1 Oct 2013 - 20 May 2015).**

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**Figure 8: Boxplots (Median, 25% and 75% percentiles, 1.5 x interquartile ranges, outliers) of calculated new snow densities ( $\rho_{HN}$ ) based on observations and densities calculated using parameterizations developed in previous studies (Eq. (3) – (9)) at three stations within period 2 (1 Oct 2011 - 01 Oct 2013) and at Kühtai station within period 3 (index \*, 01 Oct 1999 - 30 Sep 2011) and period 4 (index \*\*, 27 Feb 1987 – 30 Sep 1999).**

**Table 1: Coordinates and data availability of the four snow stations are given. The instrumentations for measuring snow depth (HS), snow water equivalent (SWE), temperature ( $T$ ), relative humidity ( $rH$ ), precipitation ( $P$ ), wind speed ( $u$ ) and global radiation ( $r$ ) are listed.**

Station Abbreviation		Kühroint KRO	Kühtai KTA	Wattener Lizum WAL	Weissfluhjoch WFJ
Location	East	12°57'35.5"	11°00'21.6"	11°38'18.6"	9°48'35.7"
	North	47°34'12.4"	47°12'25.6"	47°10'05.5"	46°49'46.4"
	$z$ (m a.s.l.)	1420	1970	1994	2540
Data		01 Jan 2011 - 02 Dec 2015	27 Feb 1987 - 20 May 2015	01 Oct 2010 – 30 Dec 2016	01 Oct 2013 - 29 Sep 2015
Instruments	HS	Sommer USH 8	Sommer USH 8	Sommer USH 8	Campbell Scientific SR50A
	SWE	Sommer Snow Scale SSG	OTT Thalimedes Shaft Encoder, Endress+Hauser Deltapilot M	Sommer Snowpillow	Sommer Snowpillow
	$T$	Rotronic MP408	Kroneis NTC	Vaisala HMP45C	Rotronic Hydroclip S3
	$rH$	Rotronic MP408	Pernix hair hygrometer	Vaisala HMP45C	Rotronic Hydroclip S3
	$P$	Sommer NIWA/Med-K505	Ott Pluvio since 2001, custom built tipping bucket before	Sommer NIWA/Med-K505	Lambrecht Pluvio 1518 H3
	$u$	Young 05103	Kroneis cup anemometer + vane	YOUNG Wind Monitor	Young 05103
	$r$	Schenk 8101	Schenk 8101	Kipp&Zonen CM21	Kipp&Zonen CM21
Comments			data gap winter 2012/13, wind regionalized from 1999	Meteorological measurements at 2041 m a.s.l.	

**Table 2: Time periods analysed in this study with mean and the median of hourly values for the height of new snow (HN), wet bulb temperature ( $T_w$ ), wind speed ( $u$ ), calculated densities from observed values ( $\rho$ ) and calculated densities corrected for settling of the snowpack ( $\rho_{HN}$ ). The results are valid for the filtered data values ( $n_{th}$ ) with  $HN > 2$  cm,  $HNW > 1.5$  mm,  $T_w < 0^\circ$  C and  $u < 5$  ms<sup>-1</sup> as a subset of all data consisting precipitation signal and positive HS change ( $n_p$ ).**

Station	#	Period	count data		HN [cm]		$T_w$ [°C]		$u$ [m s <sup>-1</sup> ]		$\rho$ [kg m <sup>-3</sup> ]		$\rho_{HN}$ [kg m <sup>-3</sup> ]	
			$n_p$	$n_{th}$	mean	median	mean	median	mean	median	mean	median	mean	median
KRO	1	1 Oct 2013 - 20 May 2015	1139	91	3.2	3.1	-3.9	-3.0	1.1	0.9	82	73	73	67
	2	1 Oct 2011 - 30 Sep 2013	1576	118	3.4	3.1	-4.2	-4.2	1.0	0.9	87	77	74	69
KTA	1	1 Oct 2013 - 20 May 2015	579	53	3.8	3.3	-3.4	-3.4	0.8	0.8	70	69	61	61
	2	1 Oct 2011 - 30 Sep 2013	506	36	3.3	2.8	-4.8	-4.0	0.8	0.7	75	66	60	54
	3	1 Oct 1999 - 30 Sep 2011	5293	252	3.5	3.2	-3.5	-3.2	0.8	0.8	74	74	64	64
	4	27 Feb 1987 - 30 Sep 1999	7958	387	3.7	3.3	-3.6	-3.4	0.8	0.7	74	75	61	59
WAL	1	1 Oct 2013 - 20 May 2015	1248	111	3.6	3.4	-4.3	-4.8	1.3	1.3	76	72	68	66
	2	1 Oct 2011 - 30 Sep 2013	1588	126	3.9	3.5	-4.3	-3.6	1.7	1.7	71	69	62	58
WFJ	1	1 Oct 2013 - 20 May 2015	1619	100	3.0	2.7	-4.9	-4.0	2.2	2.0	95	86	91	83

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**Table 3: Results of a single linear regression between the corrected densities ( $\rho_{HN}$ ) as dependent variable and wet bulb temperature ( $T_w$ ) as well as wind speed ( $u$ ) as explanatory variables for the class median values based on all filtered data points binned into  $0.5^\circ$  K classes and classes of  $0.5\text{ ms}^{-1}$ , respectively. The corresponding coefficient of determination ( $r^2$ ) and the p – value are presented.**

Station	Period #	$T_w$				$u$			
		Intercept	$\delta\rho/\delta T_w$	$r^2$	p	Intercept	$\delta\rho/\delta u$	$r^2$	p
KRO	1	82.07	4.00	0.65	0.00	45.12	19.10	0.35	0.16
	2	76.54	0.99	0.11	0.35	64.84	1.29	0.00	0.90
KTA	1	66.37	1.84	0.12	0.44	72.59	-14.44	0.41	0.36
	2	55.15	-0.37	0.02	0.75	54.25	3.37	0.53	0.17
	3	68.18	1.51	0.56	0.01	64.81	-3.82	0.39	0.10
	4	72.41	3.75	0.82	0.00	49.31	9.41	0.30	0.26
WAL	1	78.84	2.88	0.47	0.06	65.28	1.32	0.02	0.71
	2	64.58	0.97	0.17	0.21	59.43	1.50	0.05	0.57
WFJ	1	92.68	0.71	0.04	0.53	92.88	-2.91	0.18	0.23

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**Table 4: Comparison of corrected density values ( $\rho_{HN}$ , [kg m<sup>-3</sup>]) and parameterizations applying the Eq. (3) to (9) presented in section 2. Median values (m, [kg m<sup>-3</sup>]) are shown together with the Pearson correlation coefficient (r) and the root mean squared error (R, [kg m<sup>-3</sup>]) between the respective calculations and  $\rho_{HN}$ . Best values of the performance measures are highlighted for each station and time period using underlined bold numbers.**

Stat	Period	$\rho_{HN}$	$\rho_{HP}$			$\rho_D$			$\rho_{LC}$			$\rho_V$			$\rho_I$			$\rho_S$			$\rho_L$		
		m	m	r	R	m	r	R	m	r	R	m	r	R	m	r	R	m	r	R	m	r	R
KRO	1	67	85	0.28	14.4	100	0.45	23.3	121	0.44	44.4	101	<u><b>0.47</b></u>	25.3	75	0.29	<u><b>0.5</b></u>	92	0.36	18.9	60	0.40	15.8
	2	69	79	0.18	8.7	94	0.18	18.8	113	0.19	38.8	99	0.13	25.7	70	<u><b>0.20</b></u>	<u><b>1.0</b></u>	91	0.10	19.8	56	<u><b>0.20</b></u>	14.5
KTA	1	61	82	0.35	28.7	98	<u><b>0.38</b></u>	40.1	118	<u><b>0.38</b></u>	62.3	98	0.33	41.7	76	0.37	23.0	88	0.26	32.2	63	0.35	<u><b>8.9</b></u>
	2	54	80	0.14	22.4	94	0.21	33.2	114	0.21	53.2	96	0.27	35.6	73	0.12	14.3	79	<u><b>0.36</b></u>	22.1	62	0.05	<u><b>4.9</b></u>
	3	64	83	0.21	22.0	99	<u><b>0.25</b></u>	32.1	120	0.24	53.2	102	0.19	34.5	77	0.24	14.5	88	0.09	20.3	67	0.06	<u><b>4.8</b></u>
	4	59	88	0.25	26.7	103	<u><b>0.32</b></u>	35.7	126	0.31	57.1	103	<u><b>0.32</b></u>	37.6	82	0.25	19.5	83	0.10	24.2	64	0.26	<u><b>5.4</b></u>
WAL	1	66	77	0.26	16.1	90	<u><b>0.33</b></u>	23.9	108	0.32	43.9	103	0.25	32.8	65	0.24	<u><b>0.7</b></u>	92	0.04	17.9	59	0.10	5.9
	2	58	83	0.08	24.0	98	0.14	31.8	119	0.13	52.5	106	<u><b>0.15</b></u>	45.5	71	0.06	6.9	97	-0.09	28.9	63	0.08	<u><b>1.7</b></u>
WFJ	1	83	79	0.08	8.0	94	<u><b>0.10</b></u>	<u><b>2.1</b></u>	113	<u><b>0.10</b></u>	17.6	106	0.00	19.1	63	0.09	34.6	89	0.01	2.7	70	-0.03	14.6

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