Response to reviewers:

Review #1

5 General comments

The paper addresses the important issue of calibrating outputs from a General Circulation Model (GCM) for 12 Australian catchments. The authors implement two relevant methodological choices: (i) calibration of daily precipitation and (ii) use of a calibration method that accounts for the correlation between forecasts and observations. Furthermore, the authors evaluate the added value (value as a gain in skill) that the proposed method has over the use of raw forecasts and the simpler and popular

- 10 Quantile Mapping (QM). The paper builds upon the work of the research group which fits nicely within the general scope of HESS and within the objectives of the "Sub-seasonal to seasonal hydrological forecasting" special edition. I believe that post-processing of seasonal forecasts is a subject still in its infancy. Any effort to improve the forecasts is a relevant addition to the field of forecasting due to the negative societal effects of issuing forecasts that exhibit biases. My recommendation is that the paper should be accepted after minor revisions which are mainly clarifications needed to ease the reading. Detailed comments
- 15 are given below.

Response: Thank you for the positive comments. We hope that this work encourages new and innovative approaches to calibration of seasonal forecasts and delivers better forecasts for end users.

Specific comments — Major

Page 3, Lines 30-32 and Sect. 3.4: As I understand it, the set up for the QM uses a sliding window of 11-days in order to estimate the predictive and the observed distribution functions, while RPP-S uses 14-days for week 1 and week 2 up to 56 days for weeks 13-16. I think the implementation of both methods should be consistent in order to achieve robust conclusions. Perhaps something in the lines of what is discussed in Page 12, Lines 17-21. Could the authors explain how the difference in implementations between QM and RPP-S may impact the conclusions?

Response: It is entirely reasonable to suggest that the RPP-S and QM models should make use of similar data windows to give
more confidence in the conclusions. We have modified the manuscript and now show results with RPP-S set up to use an 11-day sliding window. That is, a separate RPP-S model is used to forecast each day, so the results are more comparable with

QM. Our results and conclusions do not change materially. In fact, we see improvements in the results for some cases like the dry ORO catchment, presumably due to the inclusion of more information in the early forecasts.

Page 5, Line 18: To my understanding, the number of ensembles after post-processing is 1000, is that correct? If this is the

30 case, then I would consider the comparison to the raw and quantile mapping corrected forecasts unfair. Perhaps the authors should make an effort to debias the CRPS as in Ferro, et al., (2008) to add quantitative evidence of the claim in Page 13 Lines 11-14. On the other hand, if the number of ensembles is 11, then it should be stated more clearly on the text. Response: The RPP-S forecasts are made up of 1000 ensemble members; correct. We considered carefully the impact of the ensemble sizes during the course of the study and indeed did consider adjusting the CRPS following the work of Ferro et al.; however, the adjustment of Ferro et al. applies in an idealised situation and from our point of view isn't relatable if forecasts from different models exhibit differences in reliability. We have added to the existing discussion on this matter in section 5 on

5 page 12, from line 24.

Specific comments — Minor

Page 3, Lines 30-32: These lines can be removed as you are also explaining this on Sect. 3.4.

Response: Agreed, the lines were removed.

Page 4-7: Merge Sect. 3.1 with Sect. 3.2.2 and 3.2.3 to avoid repetition and to make the manuscript shorter.

10 Response: Section 3.1 has been streamlined by combining subsections and reducing repetition.

Page 7, Lines 14-20. Perhaps change this paragraph to the result section.

Response: During the manuscript preparation we considered including the forecast example in the results but we feel it sits better in the methods section. We have left the example in the method section.

Page 21 and 22, Fig. 4 and Fig. 5: Change label "Bias (mm)" to "AB (mm/day)".

15 Response: We have defined bias in equation (6) and therefore the y-axis label is correct as Bias (mm). Average absolute bias (AB) is calculated and added as text in the top left corner of the subplots.

Technical Corrections

Page 4, Line 14. "are mapped in Figure 1" should be "are mapped in

Fig. 1". In general, mentions to figures should be changed throughout the

20 manuscript to comply with HESS guidelines (https://www.hydrology-and-earth-systemsciences.

net/for_authors/manuscript_preparation.html).

Page 5, Line 1. Missing parenthesis at the end of Eq. (1).

Page 8, Line 9. Change "streamflow" to "rainfall".

Page 8, Line 14. Missing parenthesis after "... (or simply PIT plot) ... ".

25 Page 11, Line 9. I am not sure the word "bleeding" is the correct one.

Page 13, Line 13. Change "(Figure 5)" to "(Fig. 6)".

Response: Thank you for reading our manuscript so closely – all the problems have now been fixed.

General comments

This manuscript reports the development of a rainfall postprocessor for GCM forecasts in the sub-seasonal to seasonal period (RPP-S). The proposed method is surely an important contribution as it attempts to advance in methods for post-processing

- 5 rainfall forecasts in this time scale. The method elaborates on authors' previous work and makes use of the Bayesian joint probability (BJP) modelling approach to account for predictor-predictand skill relationships. The post-processor generates daily amounts which are then aggregated to in-season totals using the Schaake Shuffle. The method is applied to rainfall forecasts from the ACCESS-S model for a set of catchments in Australia, and is found more skillful than ACCESS-S forecasts post-processed using quantile mapping (QM).
- 10 I find the paper well written and the experimental setting well described, although in some instances additional clarifications would be desirable. I have only minor comments on some methodological assumptions which need more justification to ease the readability and warrant reproducibility of the proposed method. Apart from that, I found this manuscript suitable to be published on this special issue. Below, I elaborate these minor/specific comments:

Response: Thank you for the positive comments and understanding of our work.

15 Specific comments

In section 3.2.1, pooling of multiple GCM runs and grouping forecast days are key steps in the proposed method. The proposed RPP-S follows a particular configuration and authors argue that this is a practical measure to enable post-processing of rainfall forecasts across a range of perennial and ephemeral catchments, but there is no restriction for the RPP-S configuration. P6 L11-L13. The size of the day groups expands farther from the initialization day. Smaller day groups in early periods are

- 20 intended to extract skill from initial conditions. Larger day groups in later periods are intended to better approximate the climatological distributions." Could the authors elaborate this statement in a more generic way? The point being, how can one estimate the size of day groups, without knowledge of the rainfall forecasts distribution properties? I see this issue is extensively discussed later in section 5, e.g. lines: L22-L27 in P12. Perhaps, some elements of the discussion should come earlier in the paper, e.g. in section 3.2.1
- 25 P12 L4-L5: "RPP-S forecast outperform QM forecasts, primarily because QM does not take into account the correlation between forecast and observations.". I also share the first reviewer's concerns on how differences in the implementation of the QM (as described in section 3.4) and the RPP-S method could impact the results from the benchmarking experiment. It would be worth to discuss at length those differences and their implications on the conclusions.

Response: We respond to the above two comments together since they are related. It is right that we intend for the RPP-S

30 approach to be flexible in terms of how forecasts and observations are paired and how days are grouped together. The best way to do it will depend on the available GCM data as well as catchment characteristics.

Regarding the concern about differences in model set up, we wish to give confidence in our method and results. We have modified the manuscript to show results with RPP-S set up to use an 11-day sliding window. That is, a separate RPP-S model

will is used to forecast each day, and the configuration is comparable to the QM set up. Our conclusions don't change materially. In fact, we see improvements in the results for some cases like the dry ORO catchment, presumably due to the inclusion of more information in the early forecasts.

- 5 P7 L10-L13: "BJP forecast ensemble members are randomized and are not linked across days by default. To deal with the problem, we apply the Shaake Shuffle." The use of the Shaake Shuffle approach is an important component of the proposed method, which is of major relevance when looking at time windows beyond the weather scale. Given the reliance on such technique to create realistic temporal patterns from BJP forecasts, a few lines describing details of the Shaake Shuffle rationale and its operational implementation are needed.
- 10 Response: We agree that a fuller description of the Schaake Shuffle is warranted. We have included a more detailed description in section 3.3. from page 6.

Technical corrections

P3 L25: Reference Hudson et al. (2017) is missing in the list of references

P5 L11: Symbols in Eq. (5) are no described

P8 L9: "the forecast probability integral transforms (PITs) of streamflow observations: ::" Is it not rainfall observations?
 P8 L14: "probability plot (or simply PIT plot." Missing parenthesis

P12 L14: "yet QM and RPP-S forecasts do not exhibit any obvious differences in the magnitude of biases (Figure 3)". Is it not Figure 4?

P13 L13: "reliable than RPP-S forecasts (Figure 5)". Is it not Figure 6?

20 Response: Thank you for reading our manuscript so closely – these problems have all now been fixed.

A new<u>A Bayesian modelling</u> method for post-processing daily subseasonal to seasonal rainfall forecasts from GCMs and evaluation for 12 Australian catchments

5 Andrew Schepen¹, Tongtiegang Zhao², Quan. J. Wang³Wang² and David E. Robertson²Robertson³

¹CSIRO Land and Water, Dutton Park, 4102, Australia
 ²CSIRO Land and Water, Clayton, 3168, Australia
 ³Department²Department of Infrastructure Engineering, The University of Melbourne, Parkville, 3010, Australia
 ³CSIRO Land and Water, Clayton, 3168, Australia

10

Correspondence to: Andrew Schepen (Andrew.Schepen@csiro.au)

Abstract. Rainfall forecasts are an integral part of hydrological forecasting systems at sub-seasonal to seasonal time scales. In seasonal forecasting, global climate models (GCMs) are now the go-to source for rainfall forecasts. However, for hydrological applications, GCM forecasts are often biased and unreliable in uncertainty spread, and therefore calibration is required before

- 15 use. There are sophisticated statistical techniques for calibrating monthly and seasonal aggregations of the forecasts. However, calibration of seasonal forecasts at the daily time step typically uses very simple statistical methods or climate analogue methods. These methods generally lack the sophistication to achieve unbiased, reliable and coherent forecasts of daily amounts and seasonal accumulated totals. In this study, we propose and evaluate a Rainfall Post-Processing method for Seasonal forecasts (RPP-S) based on the Bayesian joint probability modelling approach for calibrating daily forecasts and the Schaake
- 20 Shuffle approach for connecting the daily ensemble members of different lead times. We apply the method to post-process ACCESS-S forecasts for 12 perennial and ephemeral catchments across Australia and for 12 initialisation dates. RPP-S significantly reduces bias in raw forecasts and improves both skill and reliability. RPP-S forecasts are more skilful and reliable than forecasts derived from ACCESS-S forecasts that have been post-processed using quantile mapping, especially for monthly and seasonal accumulations. Several opportunities to improve the robustness and skill of RPP-S are identified. The new RPP-
- 25 S post-processed forecasts will be used in ensemble sub-seasonal to seasonal streamflow applications.

1 Introduction

Rainfall forecasts are an integral part of hydrological forecasting systems at sub-seasonal to seasonal time scales (Bennett et al., 2016; Crochemore et al., 2017; Wang et al., 2011). Inclusion of climate information in seasonal streamflow forecasts enhances streamflow predictability (Wood et al., 2016). One strategy for integrating climate information into hydrological models is to conditionally resample historical rainfall (e.g. Beckers et al., 2016; Wang et al., 2011). An alternative approach is

30 models is to conditionally resample historical rainfall (e.g. Beckers et al., 2016; Wang et al., 2011). An alternative approach is to use rainfall forecasts from dynamical climate models.

Ensemble rainfall forecasts from GCMs (global climate models) are attractive for hydrological prediction in that they forecast multiple seasons ahead and have a well-established spatial and temporal forecast structure. On the other hand, a major issue with GCM forecasts at sub-seasonal to seasonal time scales is that the forecasts are often biased and lacking in predictability of local climate (e.g.Kim et al., 2012; Tian et al., 2017). It is therefore necessary to post-process GCM rainfall forecasts using statistical or dynamical methods before they can be used in hydrological models (Yuan et al., 2015).

Several conceptually simple statistical correction methods are used for directly post-processing daily GCM rainfall forecasts including: additive bias correction, multiplicative bias correction and quantile mapping (Ines and Hansen, 2006). For example, Crochemore et al. (2016) recently evaluated linear scaling and quantile mapping for post-processing ECMWF System4 rainfall forecasts in France. Quantile mapping adjusts forecast means and ensemble spread but it is not a full calibration method

5

10 because it does not account for the correlation between forecasts and observations (Zhao et al., 2017). It is useful for bias correction of climate change projections where a full statistical calibration is inappropriate (Teutschbein and Seibert, 2012). Since additive and multiplicative bias correction and quantile mapping methods do not account for intrinsic GCM skill they are ineffective to use as post-processing tools when GCM forecasts are unskilful.

Post-processing methods that take into account model skill typically fall under the model output statistics (MOS; Glahn and

- 15 Lowry, 1972) banner. MOS type approaches are well established in the weather forecasting community for short term forecasting and modern variants are normally probabilistic. For example Wilks and Hamill (2007) studied ensemble MOS approaches for post-processing global forecast system (GFS) forecasts of rainfall and temperature up to 14 days ahead. MOS methods can also be thought of as full calibration methods. In this regard, several Bayesian calibration approaches are known to be effective at post-processing GCM rainfall totals aggregated to monthly and seasonal time scales (Hawthorne et al.)
- al., 2013; Luo et al., 2007; Schepen et al., 2014). However, it is apparent that full calibration methods aren't normally applied to the post-processing of daily GCM forecasts of rainfall in the sub-seasonal to seasonal period.
 That said, some studies have explored more sophisticated methods for post-processing daily rainfall forecasts from GCMs.
 For example, Pineda and Willems (2016) applied a non-homogenous hidden Markov model (NHHMM) to forecasts in the northwestern region of South America. Their method extracted information from GCM forecasts of rainfall fields and SSTs.
- 25 Their study built on the NHHMM method originally developed by Robertson et al. (2004) for prediction of rainfall occurrence in Brazil. In Australia, Shao and Li (2013) and Charles et al. (2013) applied an analogue downscaling method (Timbal and McAvaney, 2001) to produce downscaled daily rainfall forecasts from POAMA (the Predictive Ocean Atmosphere Model for Australia). The NHHMM and analogue methods are not straightforward to apply operationally, as they require the identification of optimal climate predictors in different climatic regions. The methods do not by design lead to forecasts that
- 30 are always reliable in ensemble spread and at least not worse than climatology forecasts. Statistical post-processing of daily rainfall forecasts is a formidable challenge, perhaps explaining the lack of sophisticated methods. Barriers include short GCM hindcast records, a high prevalence of zero rainfall amounts, seasonal variations in rainfall patterns, and intrinsically low GCM skill. Amplifying the challenge is that GCM skill decays rapidly as the age of ensemble memberslead time increases. For example, Lavers et al. (2009) examined temperature and rainfall forecasts using

DEMETER and CFS GCMs, and found that, in an idealized scenario, skill in the first 30 days is primarily attributable to skill in the first 15 days and much less to skill over the next 15 days. Post-processing methods should therefore be designed to capture as much skill as possible in the first fortnight and take into account GCM skill when post-processing forecasts further ahead.

- 5 In this study, we seek to develop a new, more direct daily rainfall post-processing method that operates solely on rainfall output from GCMs and provides a full forecast calibration taking into account GCM skill. We build a new sub-seasonal to seasonal rainfall post-processor, which has some similarities to the rainfall forecast post-processor (RPP) developed by Robertson et al. (2013b) and Shrestha et al. (2015) for post-processing numerical weather prediction (NWP) forecasts for short-term streamflow forecasting. The new system is hereafter called RPP-S, which stands for rainfall forecast post-processor seasonal.
- 10 The proposed RPP-S method applies the Bayesian joint probability (BJP) modelling approach to post-process daily GCM forecasts of rainfall. BJP has never before been used in this situation and it is therefore important to fully evaluate the merits of BJP as a component of a daily forecast post-processing system. As GCMs can produce ensemble forecasts for over 100 days ahead, RPP-S is developed to generate daily rainfall amounts that aggregate to intra-seasonal and seasonal totals. To this end, the Schaake Shuffle (Clark et al., 2004) is also included as a component of RPP-S.
- 15 We apply RPP-S to post-process raw catchment rainfall forecasts from ACCESS-S, Australia's new seasonal forecasting GCM. We <u>comprehensively assessevaluate</u> bias, <u>skill and</u>-reliability and <u>skill</u> for 12 ephemeral and perennial catchments across Australia. RPP-S catchment rainfall forecasts are compared to catchment rainfall forecasts derived from the Bureau of Meteorology's bias-corrected product for ACCESS-S. Opportunities to develop the method further are discussed.

2 Data and catchments

30

20 2.1 ACCESS-S rainfall forecasts

Raw, gridded GCM rainfall forecasts are obtained from the BoM's new ACCESS-S forecasting system. Raw catchment rainfall forecasts are derived through a process of area-weighted averaging of the ensemble mean.

ACCESS-S is a customised version of the UK Met Office's seasonal forecasting system. It contains a fully coupled model representing the interactions among the Earth's atmosphere, oceans, and land surface and <u>including</u> sea ice. The current

25 horizontal spatial resolution of the ACCESS-S atmospheric model is approximately 60 km in the mid-latitudes. Full details of the current system are provided by (Hudson et al., 2017)Hudson et al. (2017).

<u>Available</u> ACCESS-S hindcasts are initialised at midnight UTC on days 1 and 25 of each month₋ (additional start dates will become available in the future). A burst ensemble comprising 11 ensemble members is generated. Hindcasts are available for the period 1990–2012 (23 years). A longer hindcast set is not possible because of a lack of suitable initial conditions. The next version of ACCESS-S will be delivered with more sets of initial conditions and longer hindcasts.

The BoM post-processes ACCESS-S forecasts using quantile mapping (QM) at the 60 km grid scale. We use QM forecasts as a benchmark for RPP-S forecasts. QM catchment rainfall forecasts are derived in the same way as for raw catchment rainfall forecasts.

2.2 Observed rainfall

5 Observed rainfall is derived from the Australian Bureau of Meteorology's 5km daily rainfall analysis (AWAP). We make use of AWAP data from 1950 onwards (further details in section 3.2.3). Catchment rainfall is derived through a process of areaweighted averaging.

An important note on time zones: Australian rainfall is recorded as 24 hour totals to 9 am local time. Consequently, AWAP data are not perfectly synchronised with the ACCESS-S forecasts, which are initialised at midnight UTC. The asynchronism

is compounded by time zone differences and daylight savings.
 We align the GCM data and observed data as best we can. For east coast locations, an ACCESS-S forecast overlaps with the following day's AWAP rainfall analysis with a discrepancy of 1-2 hours. The time discrepancy for west coast locations is

approximately 2 hours more.

2.3 Catchments

15 Twelve perennial and ephemeral catchments spread over Australia are selected for application and evaluation. Catchment information including name, gauge ID, regional location and size are shown in Table 1. Catchment locations and climate zones are mapped in FigureFig. 1. The catchments reside in highly distinct climate zones and vary markedly in size from 100 km² to 119034 km². Evaluation across climate zones and for varying catchment sizes helps draw more general conclusions about post-processing effectiveness. However, relating forecast performance to catchment characteristics is outside the scope of this

20 studyto comprehensively evaluate the effectiveness of the post-processing methods.

3 Methods

3.1 Bayesian joint probability models

Our post-processing method embeds the Bayesian joint probability (BJP) modelling approach, which was originally designed for forecasting seasonal streamflow totals (Robertson et al., 2013a; Wang and Robertson, 2011; Wang et al., 2009). BJP has

25 since been applied to calibrate hourly rainfall forecasts (Robertson et al., 2013b; Shrestha et al., 2015) and seasonal rainfall forecasts (Hawthorne et al., 2013; Khan et al., 2015; Peng et al., 2014; Schepen and Wang, 2014; Schepen et al., 2014). BJP was most recently adapted for sub-seasonal to seasonal streamflow forecasting (Schepen et al., 2016; Zhao et al., 2016). BJP has never been applied to post process daily rainfall forecasts from GCMs.

Salient details of the BJP modelling approach are provided here. Full details can be found by reference to the above studies.

30 How the models are used in RPP S is described in the next section.

BJP formulates a joint probability distribution to characterize the relationship between forecast ensemble means (predictors) and corresponding observations (predictands). The joint distribution is modelled as a bivariate normal distribution after transformation of the marginal distributions. Data transformation is handled using the flexible log-sinh transformation (Wang et al., 2012). The log-sinh transformation transforms y by

5
$$f(y) = \beta^{-1} \ln(\sinh(\alpha + \beta y))$$
 (1)

where α and β are transformation parameters. A predictor x is transformed to g. Likewise, a predictand y is transformed to h. The relationship between g and h Therefore is formulated by a bivariate normal distribution:

$$\begin{bmatrix} g \\ h \end{bmatrix} \sim N(\mathbf{\mu}, \mathbf{\Sigma}) \tag{2}$$

where

10
$$\boldsymbol{\mu} = \begin{bmatrix} \boldsymbol{\mu}_g \\ \boldsymbol{\mu}_h \end{bmatrix}$$
(3)

is the vector of means and

$$\boldsymbol{\Sigma} = \begin{bmatrix} \sigma_g^2 & \rho_{gh} \sigma_h \sigma_g \\ \rho_{gh} \sigma_h \sigma_g & \sigma_h^2 \end{bmatrix}$$
(4)

is the covariance matrix.

In fitting mode, BJP model parameters are inferred from data. The transformation parameters is estimated by maximum a

15 posteriori probability (MAP) solution. The parameters of the bivariate normal distribution are inferred from a Bayesian formulation, using a Markov chain Monte Carlo (MCMC) algorithm to sample 1000 sets of parameters to represent parameter uncertainty.

In forecasting mode, a BJP model is conditioned on new datapredictor values. For a particular parameter set, θ , and new transformed predictor value g_{new} ,

20
$$h_{new} \mid g_{new}, \theta \sim N\left(\mu_h + \rho_{gh} \frac{\sigma_h}{\sigma_g} (g_{new} - \mu_g), \sigma_h^2 (1 - \rho_{gh}^2)\right)$$
 (5)

A sample of is obtained by drawing a random sample of from Eq. (5) and then back transforming using the inverse of Eq. (1). A complete forecast ensemble is built up by sampling a single for each parameter set.

3.2 Daily rainfall post-processing

We note some specifics of the BJP implementation used in this study. <u>Values</u><u>To allow the use of the continuous bivariate</u> normal distribution, values of zero rainfall are treated as censored values, meaning the true value is assumed to be less than or equal to zero. The predictor and predictand censor thresholds are and , respectively. Data are resealed to lie in the range [0,5]

5 prior to transformation toTo simplify the inference of transformationmodel parameters, data rescaling is used within the modelling process. To prevent undesirable extrapolation effects in forecasting mode, predictor values are limited to twice the maximum predictor value used to fit the model. The strategy of limiting predictor values has been used successfully in earlier studies. (future research will address this limitation).

3.2 The seasonal rainfall forecast post-processor (RPP-S)

10 3.2.1 Pooling of data

Model parameters may be poorly estimated when there are insufficient nonzero data values used in data inference. In We find that in very dry catchments, certain days of the year may have forecasts or observations with no non-zero values at all. In such cases, inference is not possible. To overcome this problem, predictor and predictand data sets are created by pooling data for consecutivenearby days and for multiple GCM initialisation dates. Figure 2 illustrates our RPP S configuration using 6 BJP

- 15 modelsWe choose to cover 112pool data within an 11-day sliding window to be consistent with the Bureau of Meteorology's bias correction scheme (see section 3.4); for the first 5 days. The the window is fixed to the first 11 days. However, the RPP-S method places no restriction on the configuration, and alternative configurations are discussed in section 5. We pool daily data for six "day groups": week 1, week 2, weeks 3–4, and subsequent 4 week periods. The data in day group 1 are used to fit model BJP1, the data in day group 2 are used to fit model BJP2, and so on. The size of the day groups expands farther from
- 20 the initialisation date. Smaller day groups in early periods are intended to extract skill from the initial conditions. Larger day groups in later periods are intended to better approximate the climatological distribution. We choose to pool data for ACCESS-S hindcasts initialised on days 1 of each month and day 25 of the month prior. Predictor and predictand data are subsequently paired according to the number of days since forecast initialisation. Consider, for example, post-processing of forecasts initialised on 1 February. Days 5–15 from the 1 February run and days 5–15 from the 25 January run are used to fit a model.
- 25 that is used to post-process forecasts for 10 February. We also pool data for multiple GCM initialisation dates. As noted in section 2.1, ACCESS S hindcasts are initialised on days 1 and 25 of each month. We assume it is better to estimate the correlation for a day group rather than for each day separately (see section 5 for further discussion). Predictor and predictand data are paired according to the day since forecast initialisation. For example, in Figure 2, days 1–7 from the 1 February and 25 January runs are pooled and used to fit model BJP1.
- 30 <u>PooledWe note that pooled</u> data are not completely independent meaning there is potential to underestimate parameter uncertainty using BJP. Nevertheless, we expect the effect to be limited by the weak predictor-predictand relationships and

weak persistence in daily rainfall. Additionally, we have not dealt with overlapping data in any special way (see section 5 for further discussion).

3.2.2 Fitting BJP models

- 5 <u>We now describe the model fitting and forecasting steps.</u> After pooling the data, model fitting for each BJP model proceeds as follows:
 - (1) Rescale predictor and predictand data so that each series ranges within [0,5]
 - (2) Estimate the log-sinh transformation parameters for the predictors using <u>the</u> maximum a posteriori (MAP) <u>solutionestimation</u>
 - (3) Estimate the log-sinh transformation parameters for the predictands in the same way
 - (4) Apply the transformations to normalise the predictor and predictand data
 - (5) Apply the transformations to transform the predictor and predictand censoring thresholds
 - (6) Sample parameter sets representing the posterior distribution of the bivariate normal distribution parameters using MCMC sampling, the transformed data and transformed censoring thresholds.

15 **3.2.3** Forecasting

10

20

A BJP model is used to forecast all of the days in the group. For example, model BJP1 is used to forecast days 1–7. Forecasting a single day using BJP proceeds as follows:

- (1) Transform the predictor value using the log-sinh transformation for predictors
- (2) Sample one h_{new} for each parameter set
- (3) Back-transform the ensemble members using the transformation for predictands
 - (4) Rescale the ensemble members to the original space (opposite of step (1) in $\frac{3.2.2 \text{ model fitting}}{3.2.2 \text{ model fitting}}$).
- (5) Set negative values to zero

3.3 Forming ensemble time-series forecasts

BJP forecast ensemble members are randomisedinitially random and are not linked across days-by default. To deal with the

25 problem, we apply the Schaake Shuffle (Clark et al., 2004), which uses historical data to link ensemble members and create sequences with realistic temporal patterns. The Schaake Shuffle works by imposing the rank correlation of observations on the forecast ensembles by using the trajectories of historical observations as a dependence template. Several variations on the Schaake Shuffle exist and choices can be made about how to construct a temporal dependence template, For example, Schefzik (2016)Fifty years of historical data from 1950 is used to apply the Schaake Shuffle, thus the forecast ensemble shuffled in

blocks of 50. propose selecting historical observations using a similarity criterion. The steps of the Schaake Shuffle as implemented in this study are as follows:

- (1) Select a large number of years from the past with the same month and day as the forecast initialisation date (excluding the date of the current forecast)
- 5 (2) Construct a two-dimensional dependence template by incrementing the dates by one day for each day ahead to be forecast
 - (3) Extract the observed data for each date in the dependence template
 - (4) For each day, reorder the ensemble according to the rank of the historical data for that day. For example, if the second year selected has the smallest observed value, place the smallest ensemble member in second position.
- 10 (5) Repeat steps (2)–(4) for blocks of ensemble members until the full ensemble has been shuffled.
 After applying the Schaake Shuffle to all ensemble members across all lead times, the forecasts can be considered ensemble time-series forecasts.

An example RPP-S forecast for the BRS catchment is shown in Figure 3Fig. 2. The top panel is the forecast of daily amounts and the bottom panel is the forecast of accumulated totals. Hereafter, a daily amount is taken to mean a 24-hour rainfall total

15 on any given day. An accumulated total is the sum of daily rainfall amounts over a number of days. In this example, the forecast correctly predicts a dry beginning to the forecast period. The quantile ranges of the daily <u>amountsforecasts</u> are reasonably consistent with the observed values. The accumulated forecast is somewhat narrower than the climatology reference forecast at monthly and seasonal time scales and the forecast is predicting a seasonal rainfall total less than the climatological median. It is clear that RPP S is able to produce ensemble forecasts of daily rainfall amounts.

20 3.3. Verification

We use RPP-S to post-process all available ACCESS-S forecasts initialised on day 1 of each month for 100 days ahead (the sub-seasonal to seasonal forecast period). For water resources management, whilst it is important that daily amounts are realistic, it is vital that the accumulated totals are as reliable and as skilful as possible. Forecasts are verified against both daily amounts and accumulated totals.

25 RPP-S forecasts are generated using leave-one-year-out cross-validation. RPP-S forecasts are compared with the BoM's QM forecasts (see section 2.13.4), which are also generated using leave-one-year-out cross-validation. Bias is recognised as the correspondence between the mean of forecasts and the mean of observations. Bias is visually assessed by plotting the bias for a set of events against the average forecast for the same set of events. We calculate bias as the mean error

$$BIAS \coloneqq \frac{1}{T} \sum_{t=1}^{T} \left(\overline{y}_{fcst,t} - y_{obs,t} \right)$$
(6)

where $\overline{y}_{fcst,t}$ is the forecast ensemble mean for event *t* and $y_{obs,t}$ is the corresponding observation. Bias is calculated separately for each catchment, initialisation date and day. The bias is calculated across 23 events. For a given <u>daylead time</u>, we <u>also</u> calculate the average absolute bias (AB) across all 12 catchments and 12 initialisation dates. Notwithstanding that the bias is

- 5 scale dependent, the average absolute bias is used to compare the magnitude of biases for different model forecasts and for different days ahead. We do not compare relative biases because of the occasional divide by zero problem. lead times. Reliability is the statistical consistency of forecasts and observations — a reliable forecasting system will accurately estimate the likelihood of an event. Reliability is checked by analysing the forecast probability integral transforms (PITs) of streamflowrainfall observations. The PIT for a forecast-observation pair is defined by $\pi = F(y_{obs})$ where F is the forecast
- cumulative distribution function (CDF). In the case that y_{obs} = 0, a pseudo-PIT value is sampled from a uniform distribution with a range [0, π] (Wang and Robertson, 2011). If a forecasting system is reliable, π follows a standard uniform distribution. Reliability can be visually examined by plotting the set of π_t (t=1,2,...,T) with the corresponding theoretical quantile of the uniform distribution using the PIT uniform probability plot (or simply PIT plot-). A perfectly reliable forecast follows the 1:1 line. In this study, we do not plot individual PIT diagrams. Instead, reliability is summarised using the α-index (Renard et al., 2010)

$$\alpha = 1.0 - \frac{2}{n} \sum_{t=1}^{n} \left| \pi_t^* - \frac{t}{n+1} \right| \tag{7}$$

where π_t^* is the sorted π_t in increasing order. The α -index represents the total deviation of π_t^* from the corresponding 20 uniform quantile (i.e. the tendency to deviate from the 1:1 line in PIT diagrams). The α -index ranges from 0 (worst reliability) to 1 (perfect reliability).

Forecast skill is evaluated using the continuous ranked probability score (CRPS; Matheson and Winkler, 1976). The CRPS for a given forecast and observation is defined as

$$\operatorname{CRPS} = \int \left[F(y) - H(y - y_{obs}) \right]^2 dy \tag{8}$$

25 where *y* is the forecast variable; y_{obs} is the observed value; *F* is the forecast CDF; and *H* is the Heaviside step function, which equals 0 if $y < y_{obs}$ and equals 1 otherwise. Model forecasts are compared to reference forecasts by calculating skill scores:

$$CRPS \text{ skill score} = \frac{\overline{CRPS_{ref}} - \overline{CRPS}}{\overline{CRPS_{ref}}} \times 100 \quad (\%)$$
(9)

where the overbar indicates averaging across a set of events. Reference forecasts <u>for each day</u> are taken as the day of year elimatology, produced using a BJP model fitted usingto observed data <u>5 days either side of the within an 11-day of year</u>, matching the window-size that the BoM uses to post process the OM. The Schaake Shuffle is also applied to instil correct

5 temporal characteristics into the reference forecasts (see section 3.4). The CRPS skill score is positively oriented (whereas CRPS is negatively oriented). As a percentage, a maximum score of 100 is indicative of perfectperfectly sharp and accurate forecasts. A score of 0 indicates no overall improvement compared to the reference forecast. A negative score indicates poor quality forecasts in the sense that athe naïve climatologyreference forecast is more skilful.

10 3.4 Comparison with forecasts post-processed by using quantile mapping

We compare RPP-S forecasts with forecasts that have been post-processed at the ACCESS-S grid scale using quantile mapping (QM). The gridded QM forecasts are supplied by the Bureau of Meteorology alongsideand we derive the rawcatchment forecasts through area-weighting. QM matches the statistical distribution of past forecasts to the distribution of observations to reduce errors in the forecast mean and improve forecast spread (Crochemore et al., 2016; Zhao et al., 2017). A post-processed

- 15 forecast value is obtained by first working out the quantile fraction (cumulative probability) of the new forecast using the CDF of past forecasts, then inverting the quantile fraction using the CDF of observations. The Bureau of Meteorology applied a separate quantile mapping model to each day. The CDF of the past forecasts is formed using 11 ensemble members in an 11-day sliding window and 22 years of data in leave-one-year-out cross-validation. The statistical distribution of the observations is formed using the observations in an 11-day sliding window and 22 years of data. For the first 5 days the window is fixed to
- 20 the first 11 days. If the raw forecast ensemble member is above the previously known maximum forecast value, then the forecast value is instead linearly rescaled by $o_{\text{max}} / f_{\text{max}}$ where o_{max} is the maximum observed value and f_{max} is the maximum past forecast value.

4 Results

4.1 Bias in forecasts of daily amounts

25 Biases in forecasts of daily amounts are analysed for selected days using Figure 4Fig. 3. Each circle represents the bias for a catchment and initialisation date. The bias is plotted against the average forecast (averaged over all events). As expected, raw forecasts are more biased than post-processed forecasts. The AB for raw forecasts ranges, for the examples given, from about 1.3mm to 1.5mm. The bias for raw forecasts tends to be negative, indicating that ACCESS-S has a propensity to underestimate daily rainfall amounts.

QM and RPP-S are similarly effective at reducing biases in daily amounts and both reduce bias <u>significantly-substantially</u>. After post-processing, <u>some</u> residual bias remains for any single day since the bias is corrected using pooled observations. The AB for QM forecasts ranges from about 0.8mm to 1.1mm. The AB for RPP-S forecasts ranges from about 0.6mm7mm to 1.1mm. We do not think the The differences in bias between QM and RPP-S for daily amounts are <u>significant and the results</u>.

5 vary for different selections of days (not shown).evidently insignificant. Visual examination of the QM and RPP-S scatter plots shows no wholesale bias in either the QM or RPP-S forecasts of daily amounts.

4.2. Bias in forecasts of accumulated totals

Biases in forecasts of accumulated totals are analysed using Figure 5Fig. 4 after rescaling to mm/day. As with the results for daily amounts, the raw forecasts are more biased than post-processed forecasts and the bias tends to be negative. Visual

10 examination of the scatterplots for the raw forecasts reveals, particularly for days 30, 60 and 90, that for some catchments and initialisation times, the raw forecasts are unbiased for sub-seasonal to seasonal rainfall totals; which suggests that the GCM performs well in some regions. Nevertheless, it is evident that post-processing is still very-necessary before using the ACCESS-S forecasts in hydrological forecasting.

QM and RPP S both reduce bias significantly. In contrast to the results for daily amounts, QM and RPP-S have differing

15 efficacy for reducing biases in accumulated totals-, although both still reduce bias substantially. Visual examination of the QM and RPP-S scatter plots shows that the RPP-S points fit more tightly around the zero line than the QM points. The better ability of the RPP-S forecasts to reduce biases is reflected in the AB; for example, for day 90 the AB for QM forecasts is 0.28 mm/day and for RPP-S forecasts the AB is 0.4213 mm/day (c.f. the AB for raw forecasts of 0.93 mm/day).

4.3 Reliability

20 Reliability is analysed using Figure 6Fig. 5, which presents boxplots of α -index for forecasts of daily rainfall amounts (left panel) and accumulated totals (right panel). The boxplots describe the distribution of α -index values for the same cases as we evaluated bias in section 4.2 except that we omit the day 1 result from the accumulated totals analysis. Results are presented for RPP-S before and after the Schaake Shuffle has been applied.

Raw forecasts have the poorest reliability for both daily amounts and accumulated totals. RPP-S has marginally better

25 reliability <u>overall</u> than QM for daily amounts. It is noted that the Schaake Shuffle has no effect on forecast reliability for daily amounts since it is plainly a reordering of already randomised ensemble members <u>over lead times</u>. RPP-S forecasts are more reliable than raw <u>forecast_forecasts</u> and QM <u>forecast_forecasts</u> for accumulated totals. The RPP-S forecasts become significantly more reliable after applying the Schaake Shuffle and linking the ensemble members <u>into aso</u> that they have realistic temporal patternpatterns (discussed further in section 5).

4.4 Skill scores – overall performance

Skill scores for QM forecasts are plotted against skill scores for RPP-S forecasts in a scatterplot (Figure 7Fig. 6). Skill scores for daily amounts are plotted in the left panel; skill scores for accumulated totals are plotted in the right panel. Accumulated totals are for two days or more. RPP-S forecasts and OM forecasts tend to be positively skill for the same cases, although there

- 5 is considerable variation in the magnitude of the skill scores. A striking feature of the scatterplots is that when QM forecasts are negatively skilful, the RPP-S forecasts tend to be neutrally skilful, particularly for the accumulated totals. It is evident that skill for daily amounts can be sharply negative. The skill for daily amounts can be difficult to estimate because of the small sample size and the inability to accurately forecast daily amounts beyond about 10 days. In contrast, the skill for accumulated totals is easier to estimate as it benefits from temporal averaging and the bleeding through accumulation of skill from earlier
- 10 periods.

4.5 Skill scores - detailed evaluation

Skill scores are partitioned according to catchment, days aheadlead time and initialisation date in Figures Figs. 7, 8, 9 and 109, respectively. Since forecasts of accumulated totals are more informative for water resources management and forecast skill is generally known to be limited beyond the first week or two, we focus the remainder of the results on accumulated totals.

- 15 CRPS skill scores are plotted for each catchment in Figure 8Fig. 7. Skill scores for RPP-S forecasts are vastly positive except for, most noticeably, some cases in ORO, HRDNMN and HLG. The negative Negative skill scores for RPP S forecasts in ORO, when skill scores from QM forecasts are positive, coincide with the very dry July August period in that catchment (see ahead to Figure 10). It highlights an extreme case where there is can be caused by insufficient information to fit thea stable **RPP-S** model despite the pooling and can also be an artefact of data.cross-validation in the presence of extreme events. In
- 20 several catchments, e.g., WLC and HRDDRT, RPP-S forecasts are seen to significantly outperform QM forecasts by virtue of QM forecasts frequently being significantly negatively skilful and RPP-S forecasts rarely becoming being negatively skilful. In several catchments, for example, BRP and CTG, both QM and RPP-S skill scores are predominantly positive, demonstrating that simple bias correction techniques willcan appear sufficient in somelocalised cases.
- CRPS skill scores are plotted for groups of dayslead times in Figure 9Fig. 8. For days 2–10, the RPP-S and QM forecasts are 25 similarly skilful and vastly positively skilful. There are some instances of negative skill, which are likely to be artefacts of cross-validation. For days 11–19 and 20–28, the positive relationship between OM and RPP-S forecasts gradually weakens. From day 29 onwards, the OM forecasts become negatively skilful in many instances whereas the RPP-S skill scores tend to level out around zero and are rarely negative to the same degree. For days far ahead of model initialisation For very long lead times, QM skill scores can be higher than RPP-S skill scores although overall the skill scores for accumulated seasonal totals
- 30 are quite low (< 30%). Both these factors signal room for further improvement is possible in the RPP-S forecasts.

CRPS skill scores are plotted for each initialisation date in Figure 10Fig. 9. Forecasts initialised from 1 September to 1 November have the highest proportion of cases where both QM and RPP-S forecasts are positively skilful, suggesting the ACCESS-S produces it's best sub-seasonal to seasonal forecasts during the Austral spring and summer.

5 Discussion

- 5 We demonstrate that RPP-S is able to improve daily GCM rainfall forecasts by: reducing bias (Figs. <u>3</u>-4-5); improving reliability (Fig 65); and ensuring that forecasts are typically at least as skilful as a climatological reference forecast (Figs 7-106-9). RPP-S forecasts are comprehensively compared with quantile mapping (QM) forecasts. RPP-S forecasts outperform QM forecasts, primarily because QM does not take into account the correlation between forecasts and observations-<u>nor</u> <u>corrects for autocorrelation problems</u>. Since RPP-S is built upon the Bayesian joint probability (BJP) modelling approach, it
- 10 explicitly models the correlation between the forecasts and observations, and thus takes into account model skill in the calibration. Our results add to the findings of Zhao et al. (2017) who studied the post-processing of monthly rainfall forecasts from POAMA (Australia's GCM preceding ACCESS-S). While Zhao et al. (2017) demonstrated that QM is very effective for bias correction, they did not consider accumulated totals. In our study, we find that the RPP-S forecasts are less biased and more reliable than QM forecasts for accumulated totals.
- 15 FigureFig. 5 illustrates the importance of the Schaake Shuffle for producing reliable forecasts using RPP-S. If the forecasts are not shuffled, then the forecasts of accumulated totals will tend to be too narrow in terms of ensemble spread, making the forecasts over-confident and reducing reliability.less reliable. Related to this, RPP-S forecasts are more reliable for daily amounts than QM forecasts, yet QM and RPP-S forecasts do not exhibit any obvious differences in the magnitude of biases (FigureFig. 3);) evidence that the RPP-S forecasts have a more appropriate ensemble spread, since the α-index integrates 20 information about forecast bias and ensemble spread.
- We establish six-We applied a consistent methodology to perennial and ephemeral catchments and established BJP models separately for each day, catchment and initialisation date to forecast 100+ days ahead. Alternative configurations are possible. For example, a separate BJPwe trialled establishing only six models by pooling data within week 1, week 2, weeks 3–4, and subsequent 4 week periods (not shown). The data in week 1 were used to fit model can be established for each of the 100+
- 25 days where the data areBJP1, the data in week 2 were used to fit model BJP2, and so on. Fewer days were pooled within a sliding window. The disadvantage of having a separate model for each day isclose to the initialisation date in an attempt to extract skill from the added computational cost.initial conditions and more days were pooled for longer lead times to better approximate the climatological distribution. We are not suggesting that it is infeasible (RPP S is very fast), merely pointing out that the found significant efficiency gains with little loss in performance-benefits would need to be weighed up against the
- 30 computational costs and , except for dry catchments, where a small increase in the number of parameters in future investigations.

We applied a consistent methodology to perennial and ephemeral catchments, which is unlikely to be optimal. Our results for the ORO catchment (Figures 8–10) demonstrate that pooling seven days of data and two ACCESS S runs is probably a minimum requirement in very dry catchments. In perennial frequency of negative skill scores occurred. Therefore, performance and efficiency need to be considered when establishing RPP-S models in new catchments and, more specifically, where there

5 are strong predictor predictand relationships, higher skill scores may be achieved by establishing additional models in the earlier periods. Conversely, in very dry periods it may be better to simply approximate the climatology well by pooling data over a longer period.

The RPP-S method is sophisticated in that it is a full calibration approach. However, there are opportunities to improve the methodology. For example, by pooling the data for many days in model parameter inference, it is assumed that rainfall from

- 10 one day to the next is independent, an assumption that is almost certainly false.which is an oversimplification. New inference methods that treat the rainfall data as conditionally independent ought to be investigated. Another area to improve is to handle overlaps in data, which in our study arise as we pool a large number of days to approximate climatological distributions well. A simple strategy to avoid overlaps in data would be to limit day groups to the size of the gap between GCM forecast initialisation dates. Future studies will seek to address the issues Future studies will seek to address the issues Future studies will seek to address the matters of independence
- 15 and overlapping data in more sophisticated ways.

We make use of ACCESS-S runs initialised on day 1 of the month and day 25 of the previous month. These initialisation dates are only 4–7 days apart and therefore the climatology of daily rainfall is unlikely to change significantly over that period of time. It is technically possible to establish an RPP-S model using initialisation dates spanning several months. If far apart initialisation dates are included in model parameter inference, new strategies may be needed to ensure that the effects of

20 seasonality are minimised. One possible approach is to standardise the forecasts and observations prior to fitting the BJP model. In this way, the model transformation and climatological parameters will be allowed to vary by day of year. Such strategies for building more robust RPP-S models and coincidentally minimising the effect of seasonality will be investigated in follow up work.

RPP-S and QM CRPS scores are calculated using ensembles of different sizes- and we consider the effect of ensemble size on

- 25 <u>our results.</u> When <u>competinga</u> forecasting <u>systems aresystem is</u> perfectly reliable, <u>a larger ensemblesensemble</u> should yield <u>a</u> better CRPS <u>scoresscore</u> (Ferro et al., 2008). <u>However, CRPS has a weakness in that it discourages forecasting extremes</u> (Fricker et al., 2013); <u>however QM forecasts tend to be less reliable than RPP S forecasts (Figure 5). It's; indeed it is</u> our experience that <u>ensembles that are unbiased but too</u> narrow <u>yet unreliable(like the QM</u> forecasts) can score overly well in terms of CRPS. <u>Despite these factors, our majorOur position is that because the RPP-S and QM forecasts are not similarly reliable</u>.
- 30 we are unable to make meaningful adjustments to CRPS scores to allow for the effect of ensemble size. In any case, we understand that QM forecasts can be significantly negatively skilful and unreliable, and small adjustments to the CRPS of QM forecasts will not affect those conclusions are unlikely to change due to ensemble size effects.

RPP-S is designed to post-process forecasts of daily amounts. An alternative strategy is to post-process accumulated totals and subsequently disaggregate them-to daily amounts. BJP models may be applied to post-process monthly and seasonal totals,

which are then disaggregated them to daily time scales. Future work will investigate the relative merits of direct daily postprocessing versus a seasonal calibration and disaggregation approach. This follow_up work is of particular interest as our study has shown daily post-processing skill is limited beyond 10-15 days. It is not clear how much of the seasonal forecasting skill is attributable to skill in the initial period and how much of the skill is attributable to seasonal climate signals in the GCM.

5 While our study has focused on rainfall, other variables are important in hydrology. For example, temperature forecasts are required in areas of snowmelt. Therefore, future research will investigate the extension of the RPP-S ideas to other meteorological variables.

An alternative to statistical post-processing of GCM outputs is to run a regional climate model (RCM) to provide much more localised information than a global GCM. A review study by (Xue et al., 2014) found that RCMs have limited downscaling

10 ability for sub-seasonal to seasonal forecasts. In that regard, RCM outputs may be statistically post-processed also, which may lead to better forecasts in some regions. RCMs are suited to specialised studies and less suited to post-processing operational GCM forecasts in support of national scale hydrological forecasting services.

6 Conclusion

We have developed a novel method for post-processing daily rainfall amounts from seasonal forecasting GCMs. RPP-S is a

15 full calibration approach that makes use of <u>the</u> Bayesian joint probability (BJP) modelling <u>approach</u> to account for predictorpredictand skill relationships in the post-processing. Reliable forecasts of sub-seasonal and seasonal accumulated totals are produced by linking ensemble members together using the Schaake Shuffle.

We applied RPP-S to 12 catchments across Australia in diverse climate zones and to 12 ACCESS-S initialisation dates. The method is robust in terms of being capable of post-processing forecasts in all cases, even in very dry catchments.

- 20 Compared to raw forecasts and quantile mapping (QM) post-processing (which does not account for predictor-predictand skill relationships), RPP-S performs significantly better in terms of correcting bias, reliability and skill. The only exception to this conclusion is that QM and RPP-S are similarly effective for correcting biases in daily amounts. RPP-S is particularly effective at delivering reliable, skilful, monthly and seasonal rainfall forecasts. Thus RPP-S forecasts are highly suitable for feeding into hydrological models for seasonal streamflow forecasting and other water resources management applications.
- 25 RPP S is efficient and makes use only six calibration models to produce forecasts over 100 days ahead. Pooling multiple GCM runs and grouping forecast<u>data for adjacent</u> days in model parameter inference are<u>is a</u> practical measures applied to<u>measure</u> that can enable statistical post-processing across a range of perennial and ephemeral streams. There are many avenues of research that could significantly improve the robustness and performance of RPP-S forecasts. In the future, more effort can be devoted to applying RPP-S more widely and relating forecast performance to catchment characteristics and to yield better
- 30 understandings of the forecast bias, reliability and skill.

Acknowledgements

This research was supported by the Water Information Research and Development Alliance (WIRADA), a partnership between CSIRO and the Bureau of Meteorology. We thank the Bureau of Meteorology for providing the ACCESS-S data, AWAP data and catchment information used in this study. We thank Durga Lal Shrestha for helpful comments on the manuscript-<u>and two</u>

5 anonymous reviewers for their insightful comments that helped improve the manuscript.

References

40

45

Beckers, J. V. L., Weerts, A. H., Tijdeman, E., and Welles, E.: ENSO-conditioned weather resampling method for seasonal ensemble streamflow prediction, Hydrol. Earth Syst. Sci., 20, 3277-3287, 2016.

5 Bennett, J. C., Wang, Q. J., Li, M., Robertson, D. E., and Schepen, A.: Reliable long-range ensemble streamflow forecasts: Combining calibrated climate forecasts with a conceptual runoff model and a staged error model, Water Resources Research, 52, 8238-8259, 2016.

Charles, A., Timbal, B., Fernandez, E., and Hendon, H.: Analog downscaling of seasonal rainfall forecasts in the Murray darling basin, Monthly Weather Review, 141, 1099-1117, 2013.

- 10 Clark, M., Gangopadhyay, S., Hay, L., Rajagopalan, B., and Wilby, R.: The Schaake shuffle: A method for reconstructing space-time variability in forecasted precipitation and temperature fields, Journal of Hydrometeorology, 5, 243-262, 2004. Crochemore, L., Ramos, M. H., and Pappenberger, F.: Bias correcting precipitation forecasts to improve the skill of seasonal streamflow forecasts, Hydrol. Earth Syst. Sci., 20, 3601-3618, 2016.
- Crochemore, L., Ramos, M. H., Pappenberger, F., and Perrin, C.: Seasonal streamflow forecasting by conditioning climatology with precipitation indices, Hydrol. Earth Syst. Sci., 21, 1573-1591, 2017.

Ferro, C. A. T., Richardson, D. S., and Weigel, A. P.: On the effect of ensemble size on the discrete and continuous ranked probability scores, Meteorological Applications, 15, 19-24, 2008.
 Fricker, T. E., Ferro, C. A., and Stephenson, D. B.: Three recommendations for evaluating climate predictions, Meteorological

Applications, 20, 246-255, 2013.

Glahn, H. R. and Lowry, D. A.: The Use of Model Output Statistics (MOS) in Objective Weather Forecasting, Journal of Applied Meteorology, 11, 1203-1211, 1972.
 Hawthorne, S., Wang, Q., Schepen, A., and Robertson, D.: Effective use of general circulation model outputs for forecasting monthly rainfalls to long lead times, Water Resources Research, 49, 5427-5436, 2013.

Hudson, D., Shi, L., Alves, O., Zhao, M., Hendon, H. H., and Young, G.: Performance of ACCESS-S1 for key horticultural regions, Bureau of Meteorology, Melbourne, 39 pp., 2017.

Ines, A. V. and Hansen, J. W.: Bias correction of daily GCM rainfall for crop simulation studies, Agricultural and forest meteorology, 138, 44-53, 2006.

Khan, M. Z. K., Sharma, A., Mehrotra, R., Schepen, A., and Wang, Q.: Does improved SSTA prediction ensure better seasonal rainfall forecasts?, Water Resources Research, 2015. 2015.

30 Kim, H.-M., Webster, P. J., and Curry, J. A.: Seasonal prediction skill of ECMWF System 4 and NCEP CFSv2 retrospective forecast for the Northern Hemisphere Winter, Climate Dynamics, 39, 2957-2973, 2012. Lavers, D., Luo, L., and Wood, E. F.: A multiple model assessment of seasonal climate forecast skill for applications, Geophysical Research Letters, 36, L23711, 2009.

Luo, L., Wood, E. F., and Pan, M.: Bayesian merging of multiple climate model forecasts for seasonal hydrological predictions,
Journal of Geophysical Research: Atmospheres, 112, 2007.

Matheson, J. E. and Winkler, R. L.: Scoring rules for continuous probability distributions, Management science, 22, 1087-1096, 1976.

Peng, Z., Wang, Q., Bennett, J. C., Schepen, A., Pappenberger, F., Pokhrel, P., and Wang, Z.: Statistical calibration and bridging of ECMWF System4 outputs for forecasting seasonal precipitation over China, Journal of Geophysical Research: Atmospheres, 119, 7116-7135, 2014.

Pineda, L. E. and Willems, P.: Multisite Downscaling of Seasonal Predictions to Daily Rainfall Characteristics over Pacific– Andean River Basins in Ecuador and Peru Using a Nonhomogeneous Hidden Markov Model, Journal of Hydrometeorology, 17, 481-498, 2016.

Renard, B., Kavetski, D., Kuczera, G., Thyer, M., and Franks, S. W.: Understanding predictive uncertainty in hydrologic modeling: The challenge of identifying input and structural errors, Water Resources Research, 46, W05521, 2010.

Robertson, A. W., Kirshner, S., and Smyth, P.: Downscaling of Daily Rainfall Occurrence over Northeast Brazil Using a Hidden Markov Model, Journal of Climate, 17, 4407-4424, 2004.

Robertson, D., Pokhrel, P., and Wang, Q.: Improving statistical forecasts of seasonal streamflows using hydrological model output, Hydrology and Earth System Sciences, 17, 579-593, 2013a.

Robertson, D., Shrestha, D., and Wang, Q.: Post-processing rainfall forecasts from numerical weather prediction models for short-term streamflow forecasting, Hydrology and Earth System Sciences, 17, 3587-3603, 2013b.

- 5 Schefzik, R.: A similarity-based implementation of the Schaake shuffle, Monthly Weather Review, 144, 1909-1921, 2016. Schepen, A. and Wang, Q.: Ensemble forecasts of monthly catchment rainfall out to long lead times by post-processing coupled general circulation model output, Journal of Hydrology, 519, 2920-2931, 2014. Schepen, A., Wang, Q., and Robertson, D. E.: Seasonal forecasts of Australian rainfall through calibration and bridging of coupled GCM outputs, Monthly Weather Review, 142, 1758-1770, 2014.
- 10 Schepen, A., Zhao, T., Wang, Q. J., Zhou, S., and Feikema, P.: Optimising seasonal streamflow forecast lead time for operational decision making in Australia, Hydrol. Earth Syst. Sci., 20, 4117-4128, 2016. Shao, Q. and Li, M.: An improved statistical analogue downscaling procedure for seasonal precipitation forecast, Stochastic environmental research and risk assessment, 27, 819-830, 2013.
 Sharethe, D.L., Backerten, D.E., Barnett, L.C., Burger, C.J., Burger, C.J., Barnett, B.C., Shao, C., Shao, C.,
- Shrestha, D. L., Robertson, D. E., Bennett, J. C., and Wang, Q.: Improving Precipitation Forecasts by Generating Ensembles
 through Postprocessing, Monthly Weather Review, 143, 3642-3663, 2015.
- Teutschbein, C. and Seibert, J.: Bias correction of regional climate model simulations for hydrological climate-change impact studies: Review and evaluation of different methods, Journal of Hydrology, 456, 12-29, 2012.

Tian, D., Wood, E. F., and Yuan, X.: CFSv2-based sub-seasonal precipitation and temperature forecast skill over the contiguous United States, Hydrol. Earth Syst. Sci., 21, 1477-1490, 2017.

20 Timbal, B. and McAvaney, B.: An analogue-based method to downscale surface air temperature: application for Australia, Climate Dynamics, 17, 947-963, 2001. Wang, E., Zhang, Y., Luo, J., Chiew, F. H. S., and Wang, Q. J.: Monthly and seasonal streamflow forecasts using rainfall-runoff modeling and historical weather data, Water Resources Research, 47, W05516, 2011.

Wang, Q. and Robertson, D.: Multisite probabilistic forecasting of seasonal flows for streams with zero value occurrences, Water Resources Research, 47, 2011.

Wang, Q., Robertson, D., and Chiew, F.: A Bayesian joint probability modeling approach for seasonal forecasting of streamflows at multiple sites, Water Resources Research, 45, 2009.

Wang, Q. J., Shrestha, D. L., Robertson, D. E., and Pokhrel, P.: A log-sinh transformation for data normalization and variance stabilization, Water Resources Research, 48, W05514, 2012.

Wilks, D. S. and Hamill, T. M.: Comparison of Ensemble-MOS Methods Using GFS Reforecasts, Monthly Weather Review, 135, 2379-2390, 2007.
 Wood, A. W., Hopson, T., Newman, A., Brekke, L., Arnold, J., and Clark, M.: Quantifying Streamflow Forecast Skill Elasticity

to Initial Condition and Climate Prediction Skill, Journal of Hydrometeorology, 17, 651-668, 2016.

Xue, Y., Janjic, Z., Dudhia, J., Vasic, R., and De Sales, F.: A review on regional dynamical downscaling in intraseasonal to seasonal simulation/prediction and major factors that affect downscaling ability, Atmospheric Research, 147, 68-85, 2014.

Yuan, X., Wood, E. F., and Ma, Z.: A review on climate-model-based seasonal hydrologic forecasting: physical understanding and system development, Wiley Interdisciplinary Reviews: Water, 2, 523-536, 2015. Zhao, T., Bennett, J., Wang, O. J., Schepen, A., Wood, A., Robertson, D., and Ramos, M.-H.; How suitable is quantile mapping

for post-processing GCM precipitation forecasts?, Journal of Climate, doi: 10.1175/JCLI-D-16-0652.1, 2017. 2017.

40 Zhao, T., Schepen, A., and Wang, Q.: Ensemble forecasting of sub-seasonal to seasonal streamflow by a Bayesian joint probability modelling approach, Journal of Hydrology, 2016. 2016.

ID	Catchment name	Gauge id ID	Region	Area (km ²)
BRP	Barron River above Picnic Crossing	110003A	Queensland	228
BRS	Burdekin River above Sellheim	120002	Queensland	36260
DIB	Diamantina River at Birdsville	A0020101	Queensland	119034
NMN	Namoi River above North Cuerindi	419005	New South Wales	2532
WLC	Wollomombi River above Coninside	206014	New South Wales	377
CTG	Cotter River above Gingera	410730	Murray-Darling Basin	148
MRB	Murray River above Biggara	401012	Murray-Darling Basin	1165
DVC	Davey River D/S Crossing River	473	Tasmania	698
HLG	Hellyer River above Guilford Junction	61	Tasmania	100
DRT	Deep River above Teds Pool	606001	Western Australia	474
HRD	Harvey River above Dingo Road	613002	Western Australia	148
ORO	Ord River at Old Ord Homestead	809316	Northern Australia	19513

Table 1 Catchment ID, catchment name (river and gauging location), gauge ID, region and catchment size

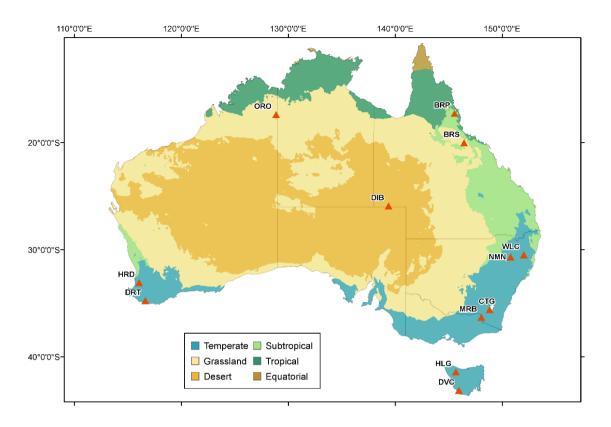


Figure 1. Map of Australian climate zones overlaid by gauging locations plotted as red triangles and labelled with catchment ID. Details of the catchments, including size, are presented in Table 1.



Figure 2. Schematic of data used to establish six BJP models for calibrating daily rainfall forecasts initialised on 1 February. Data are pooled for contiguous numbers of days since forecast. Each BJP model is used to calibrate daily forecasts within the time period it covers.

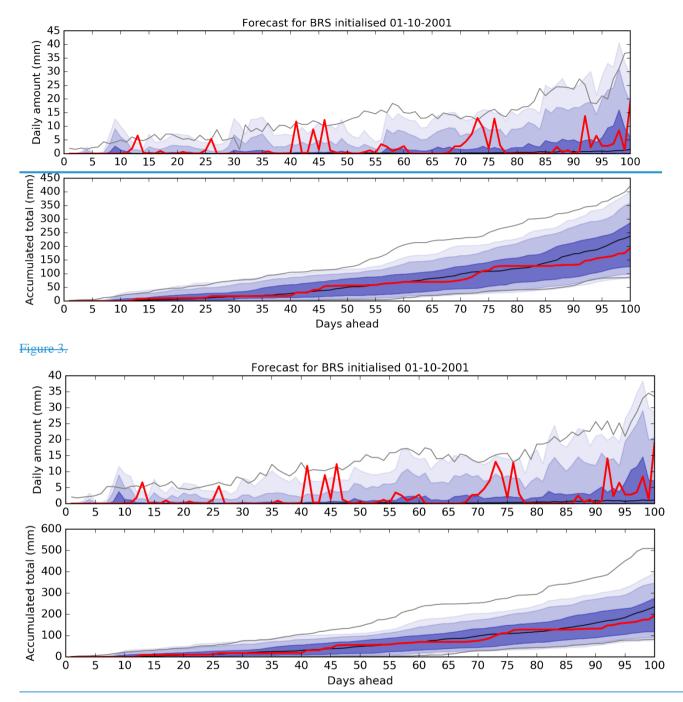
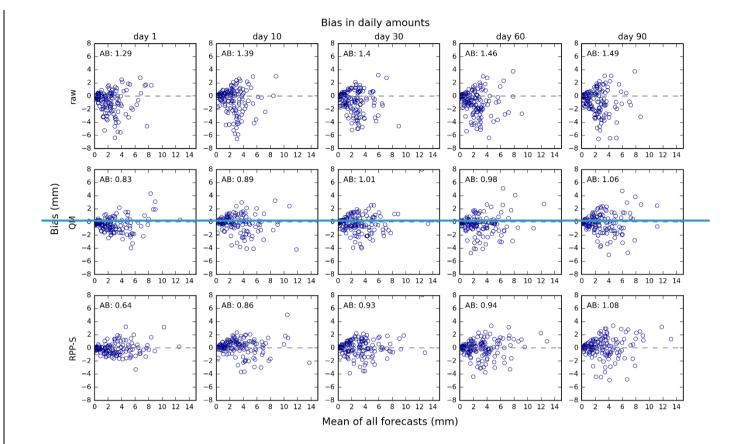


Figure 2. Example RPP-S rainfall forecast for the Burdekin River at Sellheim, initialised on the 1 October 2001 and forecasting 100 days ahead. Forecasts of daily amounts are shown in the top panel and forecasts of accumulated totals are shown in the bottom panel. Dark blue is the forecast [0.25,0.75] quantile range, medium blue is the forecast [0.10,0.90] quantile range and light blue is the forecast [0.05, 0.95] quantile range. Grey lines are the climatological reference forecast [0.05, 0.95] quantile range. The black line is the climatological reference forecast median. The red line is the observation.



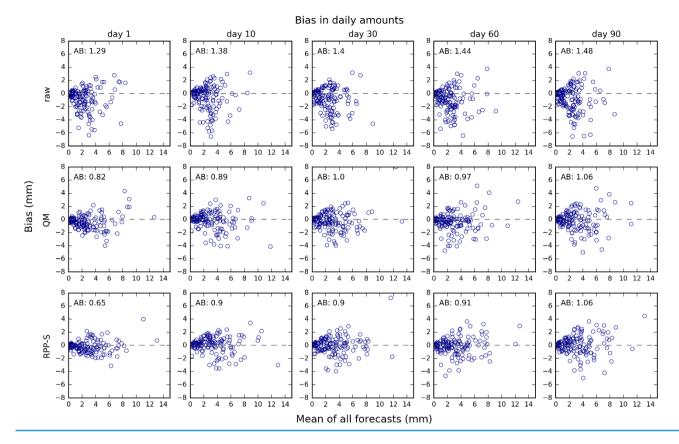
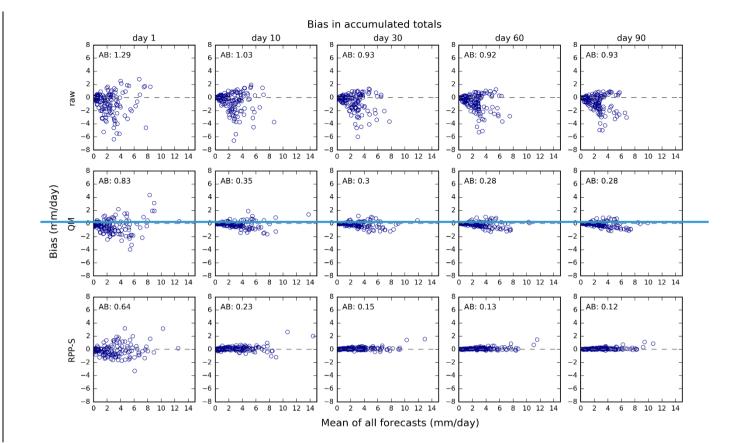


Figure 4<u>3</u>. Bias in daily rainfall forecasts for raw, QM and RPP forecasts (rows) and selected days ahead <u>/ lead times</u> (columns). The scatterplots are of forecast mean (x-axis) versus observation (y-axis). Both axes are in mm. There is one blue circle for each catchment and forecast initialisation time. The average absolute bias (AB) is printed in the top left corner.



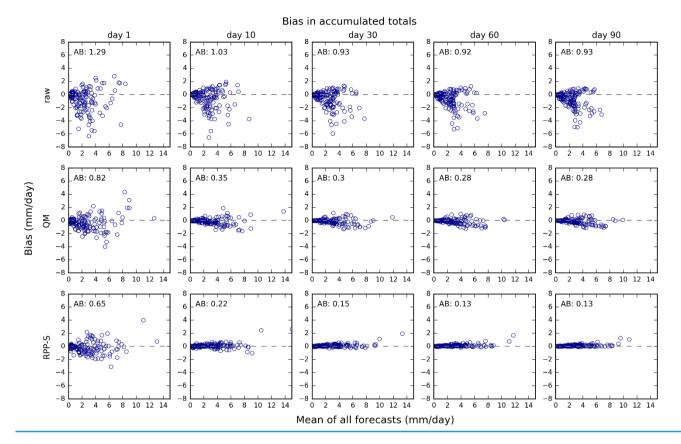


Figure 54. As for Figure 4, except for accumulated totals

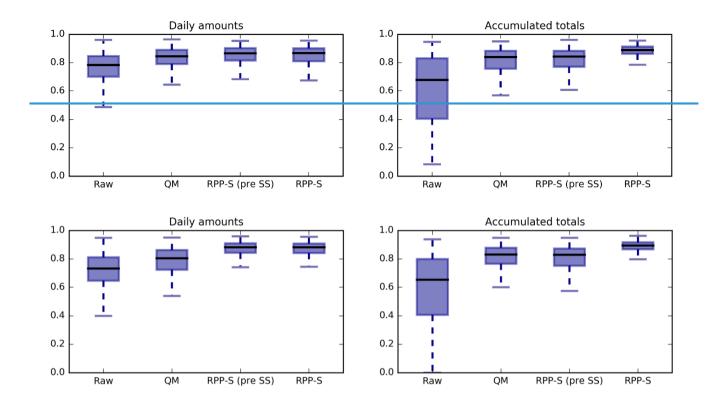


Figure 65. α-index of reliability for forecast daily rainfall amounts (left panel) and forecast accumulated rainfall totals (right panel).
Results for four types of forecasts are presented: raw, QM, RPP-S before Schaake Shuffle (pre SS) and RPP-S forecasts after the Schaake Shuffle. Higher α-index indicates better reliability. The boxplots display the median as a black line. The box spans the interquartile range and the whiskers span the [0.1,0.9] quantile range.

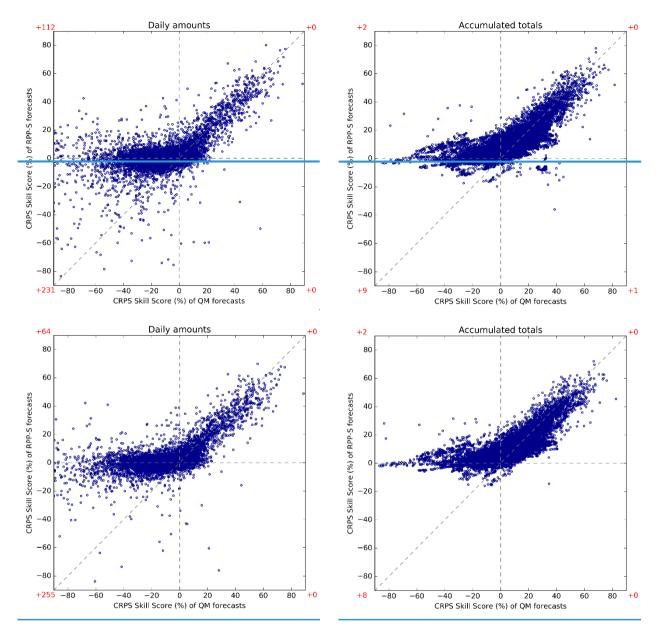
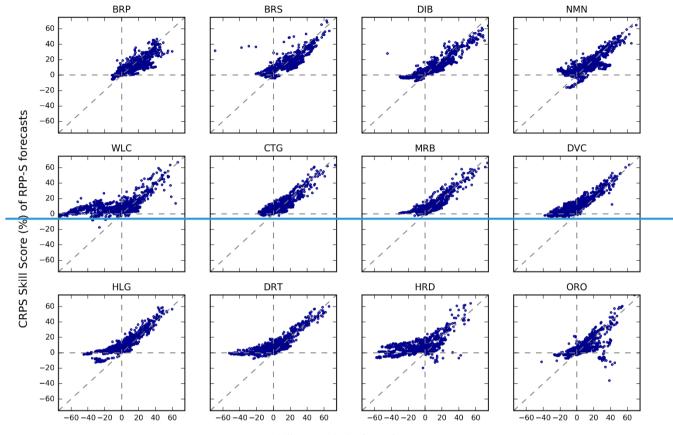


Figure 76. Scatterplots of CRPS skill scores for daily amounts (left panel) and accumulated totals (right panel). Results for QM are on the horizontal axis and results for RPP-S are on the vertical axis. Higher CRPS skills scores reflect better forecast performance. Red text preceded by a "+" symbol indicates the number of points plotted outside the axis limits in the quadrant nearest the text.



CRPS Skill Score (%) of QM forecasts

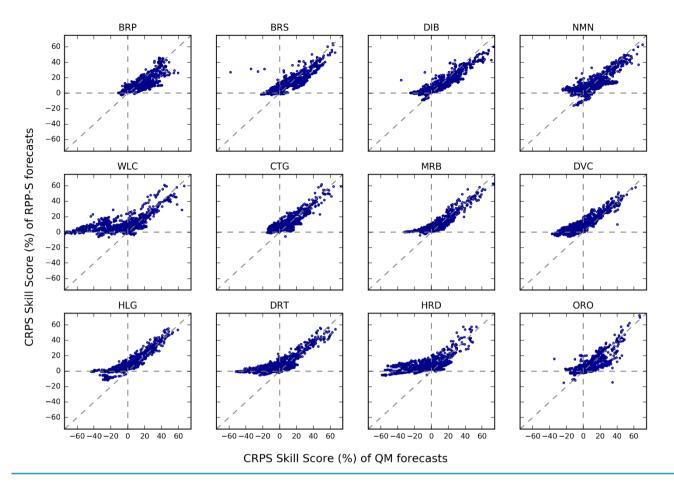


Figure <u>87</u>. Scatterplots of CRPS skill scores for accumulated totals for each catchment. The scatterplots plot QM skill scores (horizontal axis) against RPP-S skill scores (vertical axis). <u>Each scatterplot combines results for all initialisation dates and all lead times.</u>

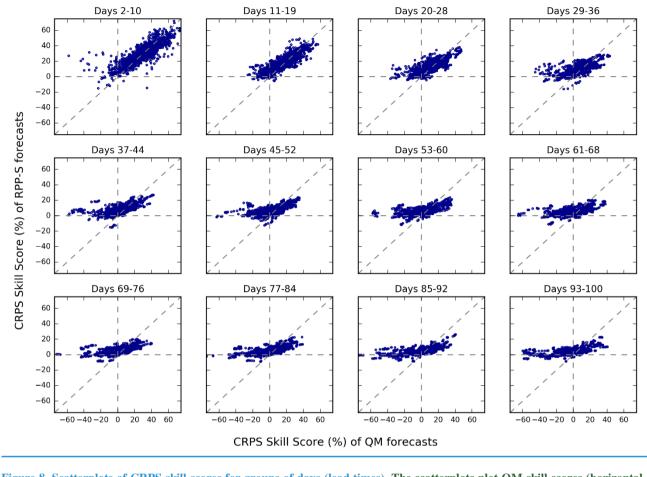


Figure 8. Scatterplots of CRPS skill scores for groups of days (lead times). The scatterplots plot QM skill scores (horizontal axis) against RPP-S skill scores (vertical axis). Each scatterplot poolscombines results for all initialisation dates and all days sheadcatchments.

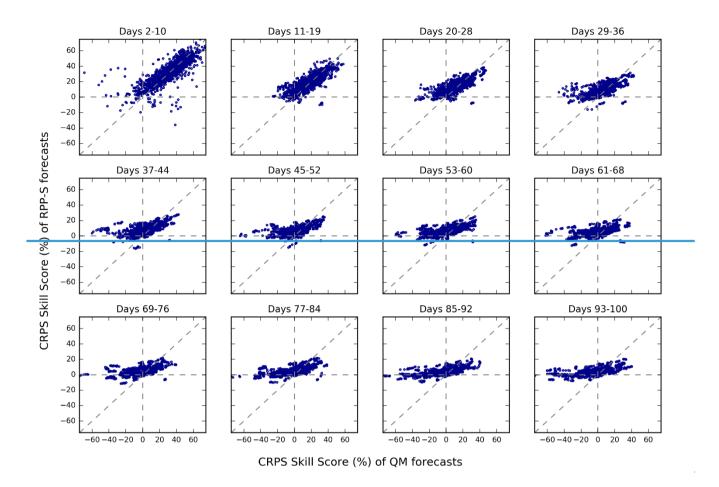


Figure 9. Scatterplots of CRPS skill scores for groups of days.

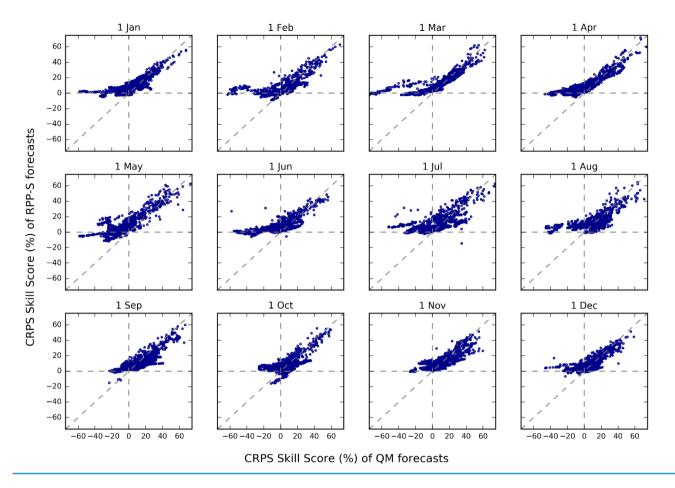
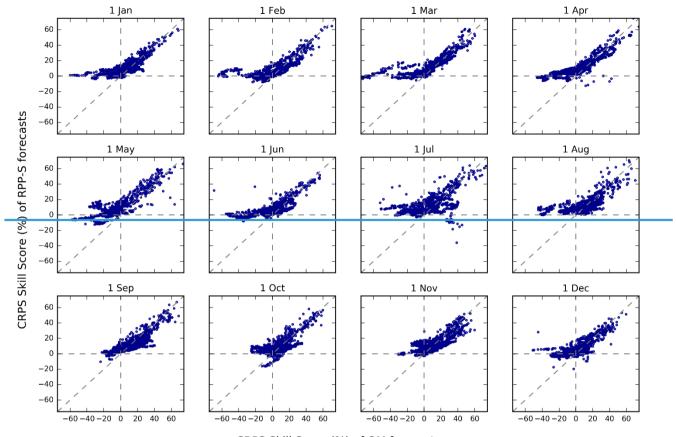


Figure 9. The scatterplots plot QM skill scores (horizontal axis) against RPP-S skill scores (vertical axis). Each scatterplot pools results for all initialisation dates and all catchments.



CRPS Skill Score (%) of QM forecasts

Figure 10. Scatterplots of CRPS skill scores for accumulated totals for each initialisation date. The scatterplots plot QM skill scores (horizontal axis) against RPP-S skill scores (vertical axis). Each scatterplot <u>poolscombines</u> results for all catchments and all <u>days</u> aheadlead times.