

1 **Comparing soil moisture anomalies from multiple independent sources over**
2 **different regions across the globe**

3
4 Carmelo Cammalleri, Jürgen V. Vogt, Bernard Bisselink, Ad de Roo

5 European Commission, Joint Research Centre (JRC), Ispra, Italy.

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7 *Correspondence to:* C. Cammalleri, European Commission - Joint Research Centre, via E. Fermi 2749,
8 I-21027 Ispra (VA), Italy. Bldg. 26b, Room 140, TP 267. Phone: +39 (0)332.78.9869, e-mail:
9 carmelo.cammalleri@ec.europa.eu.

10
11 **Abstract:** Agricultural drought events can affect large regions across the World, implying the urge for a
12 suitable global tool for an accurate monitoring of this phenomenon. Soil moisture anomalies are
13 considered a good metric to capture the occurrence of agricultural drought events, and they have
14 become an important component of several operational drought monitoring systems. In the framework
15 of the JRC Global Drought Observatory (GDO, <http://edo.jrc.ec.europa.eu/gdo/>) the suitability of three
16 datasets as possible representation of root zone soil moisture anomalies has been evaluated: (1) soil
17 moisture from the Lisflood distributed hydrological model (namely LIS), (2) remotely sensed Land
18 Surface Temperature data from the MODIS satellite (namely LST), and (3) the ESA Climate Change
19 Initiative combined passive/active microwave skin soil moisture dataset (namely CCI). Due to the
20 independency of these three datasets, the Triple Collocation (TC) technique has been applied, aiming at
21 quantifying the likely error associated to each dataset in comparison to the unknown true status of the
22 system. TC analysis was performed on five macro-regions (namely North America, Europe, India,
23 Southern Africa and Australia) detected as suitable for the experiment, providing insight into the mutual
24 relationship between these datasets as well as an assessment of the accuracy of each method. Even if no
25 definitive statement on the spatial distribution of errors can be provided, a clear outcome of the TC
26 analysis is the good performance of remote sensing datasets, especially CCI, over dry regions such as
27 Australia and Southern Africa, whereas the outputs of LIS seem to be more reliable over areas that are

28 well monitored through meteorological ground station networks, such as North America and Europe. In
29 a global drought monitoring system, the results of the error analysis are used to design a weighted-
30 average ensemble system that exploits the advantages of each dataset.

31

32 **1. Introduction**

33

34 Drought is a recurring natural extreme, triggered by lower than normal rainfall, often exacerbated
35 by a strong evaporative demand due to high temperatures and strong winds. Drought events may occur
36 in all climates and in most parts of the world, since drought is defined as a temporary deviation from the
37 local normal condition. Due to the usually wide extension of the interested area, drought affects millions
38 of people across the Globe each year (Wilhite, 2000).

39 On the basis of the economic and natural sectors impacted by this phenomenon, a drought event is
40 usually classified in meteorological, agricultural and hydrological drought, depending on the persistence
41 of the water deficit within the hydrological cycle. Of particular interest for this study are the agricultural
42 (or ecosystem) drought events, defined as prolonged periods with drier than usual soils that negatively
43 affect vegetation growth and crop production, and, as a consequence, human welfare (Dai, 2011).

44 Soil moisture is commonly seen as one of the most suitable variables to monitor and quantify the
45 impact of water shortage on vegetated lands due to its effects on the terrestrial biosphere and the
46 feedback into the atmospheric system, as highlighted by the inclusion of time-aggregated soil moisture
47 anomalies (e.g., monthly) in numerous drought monitoring systems at regional to continental scales
48 (i.e., European Drought Observatory, <http://edo.jrc.ec.europa.eu>; United States Drought Monitor,
49 <http://droughtmonitor.unl.edu>; African Flood and Drought Monitor,
50 <http://hydrology.princeton.edu/adm/>; among others).

51 In the context of drought monitoring, the soil moisture dynamic over large areas is usually
52 modelled through either distributed hydrological models or land-surface schemes of climate models
53 (Crow et al., 2012; Sheffield et al., 2004), as well as by thermal or passive/active microwave remote
54 sensing-derived quantities (see e.g., Anderson et al., 2007; Houborg et al., 2012; Mo et al., 2010).

55 Particularly, with regard to a global-scale monitoring, remote sensing-based approaches have the
56 advantage of an intrinsic worldwide coverage, but the drawbacks, in the case of microwave sensors, of
57 exploring only the first few centimeters of soil and a decreasing sensitivity with the increase of
58 vegetation coverage (Jackson, 2006). In the case of thermal data, the lack of coverage during cloudy
59 conditions and the nontrivial connection between thermal and soil moisture signals (Price, 1980) are
60 other limitations. On the contrary, diagnostic models allow for a continuous monitoring of soil moisture
61 at the desired soil depths, but the accuracy of the data is constrained by uncertainties in the
62 parameterization of soil hydrological characteristics, as well as by the actual availability of near-real
63 time reliable meteorological forcing data. Generally, the use of in-situ observations for large area
64 monitoring is limited, mainly due to the lack of long records, the sparseness of recording stations and
65 the high spatial heterogeneity of soil moisture fields.

66 It follows that both satellite measurements and model predictions are subject to errors and
67 uncertainties that need to be accounted for in their interpretation and application (Gruber et al., 2016).
68 This also suggests that a monitoring system based on a single model is rarely capable to provide global
69 reliable estimates, and a combination of different data sources is desirable in order to minimize the
70 errors in the detection of drought events. Recently, Cammalleri et al. (2015) demonstrated the value of
71 an ensemble of modelled soil moisture anomalies for drought monitoring over Europe, similarly to the
72 findings of the U.S. National Land Data Assimilation System (NLDAS) (Dirmeyer et al., 2006).
73 However, a key point in combining different modelled data is the need to estimate the affinity and
74 divergence between the models across the modelling domain.

75 In the most recent years, the Triple Collocation (TC) technique (Stoffelen, 1998) has been
76 established as a practical approach to evaluate the unknown error variance (with respect to the truth) of
77 three mutually independent measurement systems without knowing the “true” status of the system
78 (Yilmaz and Crow, 2014). This technique has been widely applied in hydrology to estimate errors in
79 soil moisture, as well as to evaluate precipitation and vegetation property indicators (Dorigo et al.,
80 2010; McColl et al., 2014). One key requirement in TC is the existence of linearity between the three
81 estimates and the truth, which can fail in the case of strongly seasonal geophysical variables such as soil

82 moisture (Su et al., 2014). Luckily, drought monitoring systems are usually based on soil moisture
83 anomalies rather than actual values, hence providing a partial remedy to this problem and making soil
84 moisture anomalies directly suitable for this methodology (Miralles et al., 2010). However, since most
85 of TC studies focused on soil moisture dynamics rather than standardized anomalies, specific analyses
86 are required to evaluate the accuracy of each dataset across the spatial domain.

87 In the frame of an operational monitoring of agriculture and ecosystem drought, the availability of
88 soil moisture, or proxy datasets available in near-real time is crucial; within the Global Drought
89 Observatory (GDO, <http://edo.jrc.ec.europa.eu/gdo/>), developed by the Joint Research Centre (JRC) of
90 the European Commission, the soil moisture outputs of the Lisflood hydrological model and the land
91 surface temperature (LST) anomalies derived from the Moderate-Resolution Imaging Spectroradiometer
92 (MODIS) onboard the Terra satellite have been detected as suitable datasets for a near-real time
93 monitoring. In particular, Cammalleri and Vogt (2016) have highlighted how monthly-average LST
94 anomalies represent the best proxy of soil moisture variations across different climates in Europe when
95 compared to other LST-derived quantities.

96 As a third dataset for the TC analysis, the combined active/passive microwave soil moisture
97 dataset produced by the European Space Agency (ESA) in the context of the Climate Change Initiative
98 (CCI) is used; even if this dataset is not currently updated in near-real time, it represents a valuable
99 reference dataset for a global consistent time-series of microwave-based soil moisture maps (also, near-
100 real time updating is foreseen in the framework of the Copernicus Climate Change Services).

101 The agreement between anomaly time-series derived from these three products has not been fully
102 investigated in the literature, especially at global scale; hence, given the independency of the three
103 sources of data (hydrological model, thermal and microwave remote sensing) and the likely fulfilling of
104 the main TC key hypothesis (i.e., independency between the errors of the three datasets), the TC
105 approach seems suitable for quantifying the spatial distribution of the errors associated to each dataset.

106 Following these considerations, the overall goal of this study is twofold. First, the agreement
107 between the monthly anomalies of the three datasets is evaluated, in order to identify the macro-areas
108 where a reliable monitoring of soil moisture extreme conditions can be performed according to these

109 three datasets available globally and suitable for use in a near-real time monitoring system. Second, the
110 TC analysis is performed over those macro-areas to quantify the spatial distribution of the expected
111 random errors for each model compared to the unknown true status. Ultimate objective of the error
112 analysis reported in this study is to provide information on the accuracy of the datasets that can be
113 injected into a weighted-average ensemble procedure for a near-real time detection of the occurrence of
114 ecosystem drought events, contributing to the future development of a robust agricultural drought
115 monitoring index within the GDO system.

116

117 2. Methods

118

119 Drought events are commonly defined as prolonged periods during which a given drought
120 indicator significantly deviates from the usual condition for the specific site and period (e.g., soil
121 moisture content is lower than the climatology). Following this definition, this study will focus on
122 standardized z-score values in order to make the different datasets directly comparable (i.e., minimizing
123 the differences related to seasonality, soil depth, etc.). Specifically, monthly z-score values, or
124 anomalies, are evaluated as:

$$125 \quad Z_{x,i,k} = \frac{x_{i,k} - \mu_{x,i}}{\sigma_{x,i}} \quad (1)$$

126 where $x_{i,k}$ is the monthly average variable for the i -th month at the k -th year, $\mu_{x,i}$ and $\sigma_{x,i}$ are the long-
127 term average and standard deviation of the variable x for the i -th month, respectively. The baseline
128 period adopted to compute the twelve μ and σ monthly reference values should be of 15-30 years in
129 order to ensure a stable benchmark. The three datasets used here, as described in the next section, are
130 the root zone soil moisture data from the Lisflood model ($x = LIS$), the ESA Climate Change Initiative
131 skin soil moisture microwave combined product ($x = CCI$) and the thermal remote sensing derived Land
132 Surface Temperature ($x = LST$); in the case of LST data, the sign of the anomalies is reversed due to the
133 expected inverse relationship between soil moisture and LST.

134 The monthly aggregation period is chosen to ensure a statistical robustness of the computed
135 anomalies, as well as to minimize the presence of missing data in the remote sensing datasets due to
136 sub-optimal acquisition conditions (e.g., cloudy days for LST). The transition from daily data to
137 monthly aggregated values also ensures a reduction in the likely discrepancies among the three datasets
138 introduced by the differences in the explored soil depth, since the phase shift in time-aggregated
139 quantities is usually less marked (Campbell and Norman, 1998). Additionally, the anomalies computed
140 according to Eq. (1), characterized by a null average and a unitary standard deviation, allow for a direct
141 comparison of the different datasets thanks to the removal of potential biases. In the particular case of a
142 regression analysis between two standardized anomaly quantities, the Pearson correlation coefficient, R ,
143 represents not only a measure of the linear dependency of the two random variables but also the slope of
144 the linear relationship and a proxy of the difference and biases of the two datasets. In this respect, R can
145 be seen as a good synthetic descriptor of the relationship between two standardized z-score datasets.
146 The statistical significance of the existence of a positive correlation can be evaluated by means of the
147 Student's t-test (2 sided) by computing the R value corresponding to a significance level $p = 0.05$.

148 Analysis of the correlation among the datasets is interesting in the framework of the triple
149 collocation (TC) technique and its basic hypotheses. In TC, a first key hypothesis is the existence of
150 linearity between the 'true' status of the system and the three models; this is formally expressed as:

$$151 \quad z_x = \alpha_x + \beta_x z_\Theta + \varepsilon_x \quad (2)$$

152 where z_Θ is the unknown true dataset of soil moisture anomalies, α_x and β_x are the systematic slope and
153 bias parameters for the dataset x with respect to the truth, and ε_x is the additive zero-mean random noise.
154 It follows that the absence of a statistical significant linear relation between all three models openly
155 violates this hypothesis.

156 Other key underling hypotheses of TC are the stationarity of both signals and errors, the
157 independency between the errors and the signal (error orthogonality) and the independence between the
158 errors of the three datasets (zero-cross correlation) (Gruber et al., 2016). Finally, operational limitations
159 regard the minimum sample size of each dataset, which is commonly assumed equal to 100 values

160 (Scipal et al., 2008; Dorigo et al., 2010), even if some other authors suggest much larger sample sizes
161 for a lower relative uncertainty (Zwieback et al., 2012).

162 Under these assumptions, Stoffelen (1998) proposed a formulation to estimate each model error
163 variance, $\sigma_{\varepsilon_x}^2$, based on a combination of the covariance between the datasets. In this approach, known
164 as the covariance notation (Gruber et al., 2016), the error variance values are computed without a
165 common (arbitrary) reference dataset as:

$$\begin{aligned} \sigma_{\varepsilon_1}^2 &= \sigma_1^2 - \frac{\sigma_{12}\sigma_{13}}{\sigma_{23}} \\ \sigma_{\varepsilon_2}^2 &= \sigma_2^2 - \frac{\sigma_{21}\sigma_{23}}{\sigma_{13}} \\ \sigma_{\varepsilon_3}^2 &= \sigma_3^2 - \frac{\sigma_{31}\sigma_{32}}{\sigma_{12}} \end{aligned} \quad (3)$$

167 where, for the sake of simplicity, LIS, LST and CCI were renamed 1, 2, 3, respectively. The first term
168 on the right side of Eqs. (3) represents the single model data variance, whereas the second term
169 represents the so-called sensitivity of the model to variations in the true status, which is a function of the
170 covariance terms between the three models. The advantage of this formulation is to directly estimate the
171 unscaled error variances, which can (eventually) be scaled to a common data space, if needed.

172 In the case of the application of the covariance notation to standardized quantities (with zero mean
173 and unitary standard deviation), the error variance values computed through Eqs. (3) are expressed as
174 dimensionless multiples of standard deviation, and a transformation to a common data space is not
175 needed.

176 Different performance metrics can be derived from the covariance notation, including relative
177 error variance metrics such as the fractional root-mean-squared-error (fRMSE, Draper et al., 2013) and
178 the correlation coefficient of each model with the underlying true signal (McColl et al., 2014).
179 However, these metrics can be derived from each other by means of simple relationships (see Gruber et
180 al., 2016) and they are analogous to the absolute error variance values in the case of z-scores that have
181 known unitary dataset variance.

182

183 3. Data and Materials

184

185 *3.1 Lisflood model soil moisture*

186

187 Root zone soil moisture dynamics are simulated by means of the Lisflood model (de Roo et al.,
188 2000), a GIS-based distributed hydrological rainfall-runoff-routing model designed to reproduce the
189 main hydrological processes that occur in large and trans-national European river catchments. The
190 model simulates all the main hydrological processes occurring in the land-atmosphere system, including
191 infiltration, actual evapotranspiration, soil water redistribution in three sub-layers (surface, root zone
192 and sub-soil), surface runoff routing to channel, and groundwater storage and transport (Burek et al.,
193 2013).

194 Static maps used by the model are related to topography (i.e., digital elevation model, local drain
195 direction, slope gradient, elevation range), land use (i.e., land use classes, forest fraction, fraction of
196 urban area), soil (i.e., soil texture classes, soil depth), and channel geometry (i.e., channel gradient,
197 Manning's roughness, bankfull channel depth, channel length, bottom width and side slope). Root zone
198 depth is defined for each modelling cell on the basis of soil type and land use, where the soil-related
199 hydraulic properties are obtained from the ISRIC 1-km SoilGrids database (Hengl et al., 2014), whereas
200 topography data are obtained from the Hydrosheds database (Lehner et al., 2008).

201 Daily meteorological forcing maps are derived from the European Centre for Medium-range
202 Weather Forecasts (ECMWF) data as spatially resampled and harmonized by the JRC Monitoring
203 Agricultural ResourceS (MARS) group. The dataset includes daily average air temperature, potential
204 evapotranspiration (for soil, water and reference surfaces) and total rainfall at 0.25 degree spatial
205 resolution, which were resampled on the model grid using the nearest neighbors algorithm.

206 The model run used in this study includes daily maps at 0.1 degree resolution between 1989 and
207 2015; the grid domain of this dataset is used as reference for the other two, whereas the baseline for the
208 anomalies computation is defined by the period 2001-2015 in order to match the LST data availability.
209 Monthly data to be used in Eq. (1) are computed as a simple average of all the data available for each
210 month, given that no gaps can be found in this dataset due to its continuous nature as hydrological

211 model. However, some areas were masked out due to the minimum or null temporal dynamic of soil
212 moisture, such as Greenland and the Sahara desert.

213

214 *3.2 Land Surface Temperature dataset*

215

216 The use of the land surface temperature (LST) anomalies as a proxy of soil moisture anomalies is
217 based on the well-known role of LST in the surface energy budget as a control factor for the partitioning
218 between latent and sensible heat fluxes. In recent years, the existence of a connection between soil
219 moisture and LST has been analyzed, mainly through the thermal inertia and the triangle methods (e.g.,
220 Carlson 2007; Verstraeten et al., 2006), as well as a direct proxy (see e.g., Park et al., 2014; Srivastava
221 et al., 2016). In a study over the pan-European domain, Cammalleri and Vogt (2016) have demonstrated
222 the good agreement between monthly LST and LIS-based root zone soil moisture z-score values during
223 summer time, where LST outperforms other LST-based indicators such as the day-night difference and
224 the surface-air gradient.

225 Following these findings, this study adopts the dataset collected by the Moderate-Resolution
226 Imaging Spectroradiometer (MODIS) sensor on board of the Terra satellite
227 (<http://terra.nasa.gov/about/terra-instruments/modis>) as a source of monthly-scale long records of LST
228 maps. In particular, the MOD11C3 Monthly CMG (Climate Modelling Grid) LST product is used in
229 this study, which is constituted by monthly composited and averaged temperature and emissivity maps
230 at a spatial resolution of 0.05 degrees over a regular latitude/longitude grid; data for the period 2001–
231 2015 are used, as the only fully completed years at the time of the analysis.

232 This monthly composite product is obtained as an average of the clear-sky data in the MOD11C1
233 products on the calendar days of the specific month, which are derived after a re-projecting and a re-
234 sampling of the MOD11B1 product. Details on the algorithms used to obtain the daily MOD11B1 maps
235 can be found in Wan et al. (2002); in summary, a double screening procedure is applied, based on: i) the
236 difference between the two independent LST estimates of the day/night algorithm (Wan and Li, 1997)

237 and the generalized split-window algorithm (Wan and Dozier, 1996), and ii) the histogram of the
238 difference between daytime and nighttime LSTs.

239 LST monthly maps were spatially co-registered to the Lisflood 0.1 degree regular
240 latitude/longitude grid by means of a simple average of the values within each cell, and anomaly maps
241 were computed according to Eq. (1) by using only the data for which $LST > 1\text{ }^{\circ}\text{C}$; this threshold value
242 (commonly used in snowmelt and snow/rainfall discrimination procedures; WMO, 1986) allows
243 removing the data that are likely affected by snow/frost from the analysis.

244

245 *3.3 Microwave combined dataset*

246

247 The ESA Climate Change Initiative (CCI) aims at developing a multi-satellite soil moisture
248 dataset by combining data collected in both past and present by passive and active microwave
249 instruments (Liu et al., 2012; Wagner et al., 2012). The current version of the dataset (v03.2) combines
250 data from nine different sensors (SMMR, ERS-1/2, TMI, SSM/I, AMSR-E, ASCAT, WindSat, AMSR2
251 and SMOS) between 1978 and 2015.

252 Satellite-based microwave estimates of soil moisture are usually related to the first few
253 centimeters of soil column (i.e., skin layer), which is quite closely related to the soil moisture content in
254 the root zone (Paulik et al., 2014), except for very dry conditions in sandy soils. Additionally, numerous
255 validations against land surface models have highlighted good performance across the globe, with
256 notable exceptions over densely vegetated areas (e.g., Loew et al., 2013).

257 The algorithm adopted to merge the different data sources is the one developed by Liu et al.
258 (2012), which is a three-step procedure that: i) merges the original passive microwave products, ii)
259 merges the original active microwave products, and iii) blends the two merged products into a single
260 final dataset. The merging procedure of passive datasets includes pixel-scale separation between
261 seasonality and anomalies, rescaling of the data based on the piece-wise cumulative distribution
262 function (CDF) and merging of the dataset using a common reference seasonality. For the active
263 microwave instruments, the CDFs are directly used to rescale the data under the assumption that active

264 datasets have an identical dynamic range, this mainly due to the limited overlap between datasets. The
265 final blending of the two merged datasets is obtained by adopting a common resolution of
266 approximately 25 km and daily frequency, as well as by using the GLDAS-1-Noah model
267 (<ftp://agdisc.gsfc.nasa.gov/data/s4pa/>) as a reference dataset for the CDF matching.

268 In this study, the daily blended dataset is spatially resampled to a 0.1 degree regular
269 latitude/longitude grid (the same used in Lisflood simulations) by means of the nearest neighbor
270 algorithm, and successively aggregated to monthly time scale by simply averaging the data (only if at
271 least 8 daily values were available in the specific month). Monthly average maps were converted into z-
272 score maps by using the baseline period 2001-2015 (the timeframe available for the LST dataset).
273 Monthly aggregated z-score values of skin soil moisture are analyzed, jointly with the other two
274 datasets, under the assumption that time-aggregation and normalization procedures minimize some of
275 the discrepancies that are likely present between skin and root zone daily time-series.

276

277 4. Results and Discussion

278

279 Considering the assumption of linearity between each one of the datasets and the unknown true
280 status of the system in TC, a preliminary analysis on the linear correlation between the three anomaly
281 products has been performed in order to detect the macro-areas where the TC procedure can be applied
282 without violating this basic hypothesis. The correlation analysis was performed by using only the
283 monthly anomalies that were available for all three datasets, with at least a sample size of 100 values
284 (max sample size = 12 months \times 15 years = 180), and by defining a minimum correlation threshold
285 ($R_{0.05}$) that ensures a statistical significance of the linear relationship on the basis of the Student's t-test
286 (at $p = 0.05$).

287 The map in Fig. 1 reports in grey the areas where all three datasets are significantly linearly
288 correlated according to the described criteria, representing the areas where the first basic hypothesis of
289 the TC is not clearly violated. It is worth to point out that some areas are excluded from the analysis by
290 the lack of data in LIS (low temporal variability, as over Greenland and the Sahara desert), LST (due to

291 the minimum temperature threshold or low temporal variability) or CCI (densely vegetated areas, such
292 as the Amazon forest and the Congo basin). These results suggest to focus the successive detailed
293 analysis on five macro-regions (demarked by the boxes in Fig. 1) that have consistent positive
294 correlation values for all the three datasets; these areas are named, from now on, as: 1) NA (including
295 the contiguous U.S. and Mexico), 2) EU (Southern and Central Europe), 3) SA (Southern countries of
296 the African continent and Madagascar), 4) IN (Indian subcontinent), and 5) AU (Australia)*.

297 The correlation coefficient maps over those regions, obtained by inter-comparing the three
298 datasets, are reported in Figs. 2 to 4, where the cells in red and yellow are the ones with negative or not-
299 significant correlation, respectively, whereas the blue scale represents the cells with increasing
300 significant linear correlation (from light to dark tones). The comparison between LIS and LST (Fig. 2)
301 shows an overall good agreement between the two datasets, with only minor areas characterized by
302 negative/not-significant correlation values; notably, low correlation can be observed over the Great
303 Lakes and Rocky mountain areas in the U.S., over the Alps in Europe, North Angola and Western
304 Himalaya. Similar results can be observed in Fig. 3, where LIS and CCI datasets are compared; this
305 comparison shows an increasing number of negative values in the Western U.S., the Alps, and Southern
306 Turkey, but overall high correlation values across most of the five regions. Finally, the comparison
307 between LST and CCI reported in Fig. 4 shows an increase of areas with low/not-significant correlation
308 in the Eastern and Western U.S. and both North- and South-Eastern Europe and the Alps, whereas a
309 high correlation can be observed all over the other regions.

310 On average, the data in Table 1 summarize the results obtained for all the regions together, as well
311 as for each region independently, showing how CCI and LST are the two datasets best correlated to
312 each other overall, even if this result is mainly driven by the results over AU, SA and IN macro-areas.
313 The data of the LIS model are similarly correlated to the ones of LST and CCI, with a more uniform
314 distribution of the results across the various sub-regions. Another outcome of this analysis is that the
315 area with the lowest average correlation between the three datasets is the EU, probably due to the high
316 heterogeneity of this region at the 0.1 degree spatial scale.

* Consider the countries and boundaries reported here only as indicative of the interested areas, and they may not in any circumstances be regarded as stating an official position of the European Commission.

317 Some of the discrepancies observed in Figs. 2 to 4 can be explained by the differences in both
318 horizontal and vertical resolution of the three raw datasets. LIS is characterized by an higher spatial
319 resolution (5-km) compared to CCI (25-km) and a vertical resolution that encompasses the full root
320 zone against the skin soil moisture of the latter; LST has the same spatial resolution of LIS but a vertical
321 resolution that varies as function of the vegetation coverage between skin (for bare soil) to root zone
322 (for full vegetation coverage). The impact of such differences is partially reflected in the observed
323 results, with CCI-LST better related over shallow soil in homogeneous areas, and LIS-LST better in
324 agreement over sparse agricultural areas in Europe. Overall, it seems that the adopted expedients (i.e.,
325 monthly average, standardization) successfully minimized these issues, given that the results in Table 1
326 shows a substantial and similar agreement of the three datasets in the main areas.

327 Additionally, the obtained results seem to suggest that it is reliable to adopt LST anomalies as
328 proxy of soil moisture anomalies, since there is a clear consistency of LST anomalies with the other two
329 datasets. Similar results were obtained by Fang et al. (2016) over the continental United States, where
330 the outputs of the thermal-based ALEXI (Atmosphere Land EXchange Inverse) model compare well
331 with soil moisture anomalies from CCI and Noah land-surface model. This consideration allows
332 applying the TC analysis to the LST dataset as well, whereas most of the studies in the literature focus
333 on land modelled and microwave soil moisture datasets (i.e., Dorigo et al., 2010; Gruber et al., 2016; Su
334 et al., 2014) with only few notable exceptions including thermal data (e.g., Hain et al., 2011; Yilmaz et
335 al., 2012).

336 The outputs of the correlation analysis were used to detect the cells suitable for the TC technique;
337 since a key hypothesis of the technique is the existence of a linear relation between each model and the
338 (unknown) truth, a necessary condition (even if not sufficient) is the existence of linear relationships
339 among the three datasets. As outcome of the correlation analysis, around 10% of the five macro-areas
340 were removed from the TC analysis due to the absence of this basic condition.

341 The maps in Figs. 5 to 7 show the main outcome of the TC analysis, which is the spatial
342 distribution of the error variance (dimensionless, showing the multiple of model standard deviation) for
343 each model, as detailed by Eqs. (3). The blank areas in those maps correspond to the cells where no

344 significant linear correlation was observed between all three datasets. The results for LIS (Fig. 5) show
345 how the highest errors are observed over the Western U.S., Northern Cape in South Africa and
346 Western/Southern Australia, whereas low errors are observed over the Eastern U.S. On the opposite, the
347 LST dataset displays the highest errors over the latter area (Fig. 6), whereas the lowest errors are
348 observed over Queensland in Australia, Eastern Cape in South Africa and Lesotho. The maps in Fig. 7
349 show that the CCI dataset has consistent patterns of low error variance values over most of Australia,
350 Western India and Central U.S.

351 Overall, on the one hand, it seems evident how CCI tends to outperform the other two methods
352 over dry areas such as Australia and South Africa, but on the other hand, a region like the U.S. is almost
353 equally subdivided among the three datasets, where LIS performs better in the East, LST in the West
354 and CCI in the center. Differences among products can be partially explained by the differences in the
355 soil layer monitored by each dataset, i.e., microwave system capturing skin soil moisture whereas
356 Lisflood models the full root zone; indeed, even if the use of monthly anomalies allows minimizing
357 some of the discrepancies, skin soil moisture remains more reliable for dry/bare areas (Das et al., 2015).
358 Even if these considerations partially explain the agreement/disagreement of the three datasets, it is not
359 straightforward to pinpoint in detail climate and/or vegetation derived patterns in the spatial distribution
360 of the TC outputs.

361 These findings are summarized in the data reported in Table 2, where the average error variance
362 for each model and macro-area is reported aside its spatial standard deviation. The data in Table 2
363 confirm that CCI has an overall better performance (lower errors) than LIS and LST, which perform
364 quite closely, mainly thanks to the very low error variance observed over Australia and, to a minor
365 extend, Southern Africa. The LIS model shows to perform better over NA and EU regions, likely due to
366 the better meteorological forcing datasets available over those regions compared to the other macro-
367 areas (due to denser ground networks). The LST dataset seems to perform moderately well over all five
368 macro-regions, with the only notable exception of EU; however, it rarely outperforms the other two
369 datasets, constituting a “second-best” option in most of the cases. It is also worth to point out that the
370 CCI dataset is often masked-out over those regions where the error of microwave techniques are likely

371 high, whereas the data of the other two datasets are mostly produced globally; hence, a possible
372 explanation of the better performance of CCI compared to LIS and LST may be linked to this
373 preliminary screening of the data.

374 The outcome that LIS slightly outperforms the other two datasets over NA is in agreement with
375 the results reported by Hain et al. (2011), where the Noah land-surface model slightly outperforms (on
376 average) the microwave and thermal datasets over Contiguous U.S. However, it should be pointed out
377 how the spatial distribution of the error estimates for LIS differs from the ones reported for Noah, likely
378 due to the differences in both meteorological forcing and modelling approaches. Some qualitative
379 analogies can also be observed with the results reported in Pierdicca et al. (2015), which show smaller
380 average errors at daily time scale over Europe for the ERA-LAND modelled datasets compared to two
381 microwave-based datasets, even if both the temporal scale and the adopted methodology of the latter
382 differ from the ones used in our study. This previous study seems to confirm that land modelling
383 approaches are more reliable, on average, over these regions, likely due to the reliability of
384 meteorological forcing and model parameterizations, even if there can be significant differences among
385 the performances of different land-surface models.

386 Over AU sub-region, the spatial distribution of the errors in CCI are quite in agreement with the
387 results reported in Su et al. (2014) for two microwave datasets, with larger errors along the South-East
388 Australian coast. This result supports the assumption that microwave data are more reliable over dry
389 bare soil areas, which is further highlighted by the results obtained in SA and IN sub-regions. The
390 subdivision of the NA domain in three main regions is similar to the one observed by Gruber et al.
391 (2016) in comparing ASCAT and AMSR-E microwave datasets, suggesting key differences in the soil
392 moisture behavior over these three sub-regions. Overall, the spatial patterns of microwave and land
393 model errors show similarities with the ones observed by Dorigo et al. (2010), even if no thermal data
394 were included in their analysis.

395 The error variance values can also be interpreted as the correlation coefficient of each dataset with
396 the underlying true signal, following the definition of McColl et al. (2014). In fact, for the special case
397 of anomalies with unitary variance ($\sigma_x^2 = 1$), the TC-derived R_x of each dataset is simply equal to

398 $\sqrt{1 - \sigma_{\varepsilon_x}^2}$, which ranges on average over all five regions (not shown) between 0.91 (for CCI in AU) to
399 0.66 (for LST over EU); these values show a good capability of the datasets to capture, on average,
400 temporal variations in soil moisture anomalies.

401 In order to provide a simple synthetic representation of the likely best model for each area, the
402 map in Fig. 8 depicts for each cell the dataset with the lowest error variance by associating different
403 colors to the three datasets (red for LIS, blue for LST and green for CCI). Even if this approach is rather
404 simplistic, as it cannot account for two products performing really close over some areas, the major
405 relevant features, like the predominance of the CCI model over Australia, are made evident by these
406 maps.

407 The maps in Fig. 8 confirm CCI as the dataset with the lowest error variance values over most of
408 AU, SA and IN, whereas the three datasets almost equally split the other two macro-areas; this is even
409 more evident in the data reported in Table 3, where the percentage of sub-areas where each model is the
410 best is reported. These data confirm the good performance of CCI over AU, SA and IN macro-regions,
411 whereas the NA territory is almost equally divided among the three datasets and LIS outperforms both
412 LST and CCI over 50% of EU domain. In the latter, the areas where the LIS dataset outperforms the
413 other two datasets partially resemble the results obtained by Pierdicca et al. (2011) for the ERA-LAND
414 model; however, the present study includes also remote sensing thermal data and not only microwave-
415 derived datasets. Overall, the CCI dataset outperforms the other two datasets in about 50% of cells, with
416 the remaining almost equally split between LIS and LST.

417 Finally, the spatial distribution of the weighting factor of each dataset, computed according to the
418 least square theory (Yilmaz et al., 2012), is represented in Figs. 9 to 11. The color scale of the figures
419 was designed to represent in a neutral color the cells that have a weighing factor close to the one for a
420 simple-average (1/3), in green scale the weights greater than a simple-average (larger contribute) and in
421 orange the weights lower than the simple-average (smaller contribute). The visual intercomparison of
422 the three maps further emphasizes the good performance of the CCI product over AU and SA, the best
423 performance of LIS over the Eastern US and EU, and the good results obtained for LST in Western US

424 and Northern AU. It is worth noticing that the use of a weighted average based on the TC error analysis
425 does not seem to bring advantages over large areas of central US, EU and Eastern IN where the
426 weighting factors are close to the ones for a simple arithmetic average. The behavior of the weighting
427 factors over the five macro-areas can be synthesized by the frequency diagram in Fig. 12. This plot
428 shows the high fraction of weighting factors > 0.4 for the CCI dataset, representing a predominant
429 contribution on the ensemble mean of this product over the others, whereas LST has a peak of
430 frequency center around $1/3$ (arithmetic average) and LIS has a hint of bi-modal distribution. These
431 data, together with the map in Fig. 8, confirm the fact that CCI outperforms the other two datasets in
432 50% of the domain, whereas LST is often the second-best option behind either CCI or LIS.

433

434 **5. Summary and Conclusions**

435

436 Three datasets have been compared as proxy of the unknown true status of soil moisture
437 anomalies in the context of the global drought monitoring system under development by the JRC of the
438 European Commission. Key assumption of the study is the inability of a single dataset to accurately
439 capture the soil moisture dynamic over the large range of variability of conditions that can be observed
440 at continental to global scale.

441 The inter-comparison between the three datasets, namely the outputs of the Lisflood hydrological
442 model (LIS), the MODIS-based land surface temperature (LST) and the combined active/passive
443 satellite microwave (CCI) data, confirms some inconsistencies between the three datasets over certain
444 areas, as well as the difficulties in comparing the three datasets over peculiar areas (e.g., Sahara desert,
445 Amazon rainforest) due to the lack of coverage from one or more datasets. Generally, the three datasets
446 seem comparable over most of the globe, thanks to the use of time-aggregation and standardization
447 procedures that remove temporal inconsistencies and biases among the series. Focusing the analysis
448 only on the areas where the three datasets are substantially in agreement (following a linear regression
449 analysis), five macro-regions were detected as suitable for further investigations according to the Triple
450 Collocation (TC) technique. This analysis allows quantifying the likely random error associated with

451 each model (with regard to the true status) even in absence of an observation of the “truth”, under the
452 hypothesis that certain criteria are met.

453 The main outcome of the TC analysis further confirms the need of a multi-source approach for a
454 reliable assessment of soil moisture anomalies over those five regions, given that no model outperforms
455 the others (in terms of expected error variance) for the entire study domain. Emblematic are the results
456 over North America, where each model outperforms the others in one sub-region, like the LIS approach
457 in Eastern U.S., LST in the Southern-Western domain and CCI in Central U.S. Even if no clear insight
458 on the general patterns of the errors can be provided as outcome of the study, overall, the obtained
459 results seem suggesting that remote sensing datasets perform better over dry areas and sparsely
460 monitored areas (e.g., Australia and Southern Africa), whereas the LIS dataset seems more reliable over
461 NA and EU where dense networks of meteorological ground stations are deployed.

462 It has been highlighted how some differences among the datasets can also be related to the depth
463 of the soil layer monitored by each dataset, i.e., the microwave system capturing skin soil moisture
464 whereas Lisflood models the full root zone; indeed, even if the use of monthly anomalies allows
465 minimizing some of the discrepancies and biases, our results confirm that skin soil moisture remains
466 more reliable for areas where the effects of vegetation coverage is minimal (Das et al., 2015), whereas
467 hydrological models are more suited for agricultural and densely vegetated regions. However, the three
468 datasets seems to be overall comparable in terms of average performances, supporting the success of the
469 adopted homogenization procedures. Some analogies between the obtained results and the ones already
470 available in the literature have been found, but the inclusion of thermal data into the analysis enlarges
471 the understanding of the mutual relationship between the different datasets.

472 The results of this study represent a robust starting point for the development of a global drought
473 monitoring system based on such anomaly datasets, which can exploit the main findings of the TC
474 analysis in order to develop a suitable ensemble product over the investigated regions. The error
475 characterization derived from TC was used to estimate the weighing factors of an ensemble mean
476 procedure, based on the least squares framework reported in Yilmaz et al. (2012). Currently, an

477 operational implementation of such ensemble product is foreseen for the GDO system as soon as the
478 CCI product becomes available in near-real time.

479 Further analyses are required to be able to extend the test to the areas currently not included in this
480 study, especially the ones where the three datasets are available but provide inconsistent or contrasting
481 results. In this context, the analysis of further global datasets may help in unveil the reasons behind such
482 discrepancies.

483

484 **References**

485

486 Anderson, M.C., Norman, J.M., Mecikalski, J.R., Otkin, J.P., Kustas, W.P., 2007. A climatological
487 study of evapotranspiration and moisture stress across the continental U.S. based on thermal
488 remote sensing: II. Surface moisture climatology. *J. Geophys. Res.* 112, D11112,
489 doi:10.1029/2006JD007507.

490 Burek, P., van der Knijff, J.M., de Roo, A., 2013. LISFLOOD: Distributed Water Balance and Flood
491 Simulation Model. JRC Scientific and Technical Reports, EUR 26162 EN, 142 pp.
492 doi:10.2788/24719.

493 Cammalleri, C., Vogt, J.V., 2016. On the role of Land Surface Temperature as proxy of soil moisture
494 status for drought monitoring in Europe. *Remote Sens.* 7, 16849-16864.

495 Cammalleri, C., Micale, F., Vogt, J.V., 2015. On the value of combining different modelled soil
496 moisture products for European drought monitoring. *J. Hydrol.* 525, 547-558.

497 Campbell, G.S., Norman, J.M., 1998. An introduction to environmental biophysics, Springer-Verlag.,
498 New York (NY), USA. doi: 10.1007/978-1-4612-1626-1.

499 Carlson, T., 2007. An overview of the “Triangle Method” for estimating surface evapotranspiration and
500 soil moisture from satellite imagery. *Sensors* 7(8), 1612-1629.

501 Crow, W.T., Kumar, S.V., Bolten, J.D., 2012. On the utility of land surface models for agricultural
502 drought monitoring. *Hydrol. Earth Syst. Sci.* 16, 3451-3460.

503 Dai, A., 2011. Drought under global warming: A review. *Wiley Interdiscip. Rev. Clim. Change* 2, 45-
504 65.

505 Das, K., Paul, P.K., Dobesova, Z., 2015. Present status of soil moisture estimation by microwave
506 remote sensing. *Cogent Geoscience* 1, 1084669.

507 de Roo, A., Wesseling, C., van Deusen, W., 2000. Physically based river basin modelling within a GIS:
508 The LISFLOOD model. *Hydrol. Process.* 14, 1981-1992.

509 Dirmeyer, P.A., Gao, X., Zhao, M., Guo, Z., Oki, T., Hanasaki, N., 2006. GSWP-2: multimodel analysis
510 and implications for our perception of the land surface. *Bull. Amer. Meteor. Soc.* 87, 1381–1397.

511 Dorigo, W.A., Scipal, K., Parinussa, R.M., Liu, Y.Y., Wagner, W., de Jeu, R.A.M., Naeimi, V., 2010.
512 Error characterisation of global active and passive microwave soil moisture datasets. *Hydrol.*
513 *Earth Syst. Sci.* 14, 2605-2616.

514 Draper, C., Reichle, R., de Jeu, R., Naeimi, V., Parinussa, R., Wagner, W., 2013. Estimating root mean
515 square errors in remotely sensed soil moisture over continental scale domains. *Remote Sens.*
516 *Environ.* 137, 288-298.

517 Fang, L., Hain, C.R., Zhan, X., Anderson, M.C., 2016. An inter-comparison of soil moisture data
518 products from satellite remote sensing and a land surface model. *Int. J. Appl. Earth Obs. Geoinf.*
519 48, 37-50.

520 Gruber, A., Su, C.-H., Zwieback, S., Crow, W., Dorigo, W., Wagner, W., 2016. Recent advances in
521 (soil moisture) triple collocation analysis. *Int. J. Appl. Earth Obs. Geoinf.* 45, 200-211.

522 Hain, C.R., Crow, W.T., Mecikalski, J.R., Anderson, M.C., Holmes, T., 2011. An intercomparison of
523 available soil moisture estimates from thermal infrared and passive microwave remote sensing
524 and land surface modeling. *J. Geophys. Res.* 116, D15107.

525 Hengl, T., de Jesus, J.M., MacMillan, R.A., Batjes, N.H., Heuvelink, G.B.M., Ribeiro, E., et al., 2014.
526 SoilGrids1km — Global Soil Information Based on Automated Mapping. *PLoS ONE* 9(8),
527 e105992.

528 Houborg, R., Rodell, M., Li, B., Reichle, R., Zaitchik, B., 2012. Drought indicators based on model
529 assimilated GRACE terrestrial water storage observations. *Wat. Resour. Res.* 48, W07525.
530 doi:10.1029/2011WR011291.

531 Jackson, T.J., 2006. Estimation of Surface Soil Moisture Using Microwave Sensors. *Encyclopedia of*
532 *Hydrological Sciences, Part 5: Remote Sensing.* doi: 10.1002/0470848944.hsa060.

533 Lehner, B., Verdin, K., Jarvis, A., 2008. New global hydrography derived from spaceborne elevation
534 data, *Eos* 89(10), 93–94.

535 Liu, Y.Y., Dorigo, W.A., Parinussa, R.M., de Jeu, R.A.M., Wagner, W., McCabe, M.F., Evans, J.P., van
536 Dijk, A.I.J.M., 2012. Trend-preserving blending of passive and active microwave soil moisture
537 retrievals. *Remote Sens. Environ.* 123, 280-297.

538 Loew, A., Stacke, T., Dorigo, W., de Jeu, R., Hagemann, S., 2013. Potential and limitations of
539 multidecadal satellite soil moisture observations for selected climate model evaluation studies.
540 Hydrol. Earth Syst. Sci. 17, 3523-3542.

541 McColl, K.A., Vogelzang, J., Konings, A.G., Entekhabi, D., Piles, M., Stoffelen, A., 2014. Extended
542 triple collocation: Estimating errors and correlation coefficients with respect to an unknown
543 target. Geophys. Res. Lett. 41, 6229-6236.

544 Miralles, D.G., Crow, W.T., Cosh, M.H., 2010. Estimating spatial sampling errors in coarse-scale soil
545 moisture estimates derived from point-scale observations. J. Hydrometeorol. 11, 1423-1429.

546 Mo, K.C., Long, L.N., Xia, Y., Yang, S.K., Schemm, J.E., Ek, M.B., 2010. Drought indices based on
547 the Climate Forecast System Reanalysis and ensemble NLDAS. J. Hydrometeorol. 12, 185-210.

548 Park, J.-Y., Ahn, S.-R., Hwang, S.-J., Jang, C.-H., Park, G.-A., Kim, S.-J., 2014. Paddy Water Environ.
549 12(1), 77-88.

550 Paulik, C., Dorigo, W., Wagner, W., Kidd, R., 2014. Validation of the ASCAT Soil Water Index using
551 in situ data from the International Soil Moisture Network. Int. J. Appl. Earth Obs. Geoinfo. 30, 1-
552 8.

553 Pierdicca, N., Fascetti, F., Pulvirenti, L., Crapolicchio, R., Munõz-Sabater, J., 2015. Analysis of
554 ASCAT, SMOS, in-situ and land model soil moisture as a regionalized variable over Europe and
555 North Africa. Remote Sens. Environ. 170, 280-289.

556 Price, J.C., 1980. The potential of remotely sensed thermal infrared data to infer surface soil moisture
557 and evaporation. Water Resour. Res. 16(4), 787-795.

558 Scipal, K., Holmes, T., de jeu, R., Naemi, V., Wagner, W., 2008. A possible solution for the problem of
559 estimating the error structure of global soil moisture datasets. Geophys. Res. Lett. 35, L24404.
560 doi:10.1029/2008GL035599.

561 Sheffield, J., Goteti, G., Wen, F., Wood, E.F., 2004. A simulated soil moisture based drought analysis
562 for the United States. J. Geophys. Res. 109, D24108. doi:10.1029/2004JD005182.

563 Stoffelen, A., 1998. Toward the true near-surface wind speed: Error modelling and calibration using
564 triple collocation. J. Geophys. Res. 103, 7755-7766.

565 Srivastava, P.K., Islam, T., Singh, S.K., Gupta, M., Petropoulos, G.P., Gupta, D.K., Wan Jaafar, W.Z.,
566 Prasad, R., 2016. Soil moisture deficit estimation through SMOS soil moisture and MODIS land
567 surface temperature. In: *Satellite Soil Moisture Retrieval: Techniques and Applications*, P.K.
568 Srivastava, G.P. Petropoulos, Y.H. Kerr (Eds.), Elsevier B.V.

569 Su, C.-H., Ryu, D., Crow, W.T., Western, A.W., 2014. Beyond triple collocation: Applications to soil
570 moisture monitoring. *J. Geophys. Res. Atmos.* 119, 6419-6439.

571 Verstraeten, W.W., Veroustraete, F., van der Sande, C.J., Grootaers, I., Feyen, J., 2006. Soil moisture
572 retrieval using thermal inertia, determined with visible and thermal spaceborne data, validated for
573 European forests. *Remote Sens. Environ.* 101(3), 299-314.

574 Wagner, W., Dorigo, W., de Jeu, R., Fernandez, D., Benveniste, J., Haas, E., Ertl, M., 2012. Fusion of
575 active and passive microwave observations to create an essential Climate Variable data record on
576 soil moisture. *ISPRS Annal of the Photogrammetry, Remote Sensing and Spatial Information*
577 *Sciences*, volume I-7. XXII ISPRS Congress, 25 August – 01 September 2012, Melbourne,
578 Australia.

579 Wan, Z., Zhang, Y., Zhang, Q., Li, Z.-L., 2002. Validation of the land-surface temperature products
580 retrieved from Terra Moderate Resolution Imaging Spectroradiometer data. *Remote Sens.*
581 *Environ.* 83, 163-180.

582 Wan, Z., Li, Z.-L., 1997. A physics-based algorithm for retrieving land-surface emissivity and
583 temperature from EOS/MODIS data. *IEEE Trans. Geosci. Remote Sens.* 35, 980-996.

584 Wan, Z., Dozier, J., 1996. A generalized split-window algorithm for retrieving land surface temperature
585 from space. *IEEE Trans. Geosci. Remote Sens.* 34, 892-905.

586 Wilhite, D.A., 2000. *Drought as a natural hazard: Concepts and definitions*. N: Disasters series.
587 Routledge Publishers, UK, 213-230.

588 World Meteorological Organization, 1986. *Intercomparison of models of snowmelt runoff*. Operational
589 Hydrological Report, 23.

590 Yilmaz, M.T., Crow, W.T., Anderson, M.C., Hain, C.R., 2012. An objective methodology for merging
591 satellite- and model-based soil moisture products. *Water Resour. Res.* 48(11), W11502.
592 doi:10.1029/2011WR011682, 2012.

593 Yilmaz, M.T., Crow, W.T., 2014. Evaluation of assumptions in soil moisture triple collocation analysis.
594 *J. Hydrometeorol.* 15, 1293-1302.

595 Zwieback, S., Scipal, K., Dorigo, W., Wagner, W., 2012. Structural and statistical properties of the
596 collocation technique for error characterization. *Nonlin. Processes Geophys.* 19, 69-80.

597 **Tables**

598

599 **Table 1.** Summary of the Pearson correlation coefficient values (average \pm standard deviation) observed
 600 for all the regions.

Comparison	ALL	NA	EU	SA	IN	AU
LIS vs. LST	0.44 ± 0.09	0.41 ± 0.08	0.39 ± 0.07	0.48 ± 0.09	0.44 ± 0.07	0.50 ± 0.10
LIS vs. CCI	0.49 ± 0.10	0.47 ± 0.09	0.42 ± 0.08	0.48 ± 0.10	0.48 ± 0.08	0.58 ± 0.11
CCI vs. LST	0.56 ± 0.13	0.49 ± 0.14	0.37 ± 0.09	0.63 ± 0.09	0.52 ± 0.10	0.68 ± 0.07

601

602

603 **Table 2.** Summary of the TC error variance analysis, reporting the spatial average (\pm standard
 604 deviation) values observed over each macro-region.

Model	ALL	NA	EU	SA	IN	AU
LIS	0.48 ± 0.13	0.42 ± 0.14	0.44 ± 0.12	0.54 ± 0.11	0.49 ± 0.10	0.54 ± 0.14
LST	0.44 ± 0.13	0.46 ± 0.15	0.56 ± 0.10	0.37 ± 0.10	0.48 ± 0.09	0.38 ± 0.11
CCI	0.36 ± 0.18	0.46 ± 0.16	0.54 ± 0.12	0.30 ± 0.14	0.38 ± 0.16	0.17 ± 0.10

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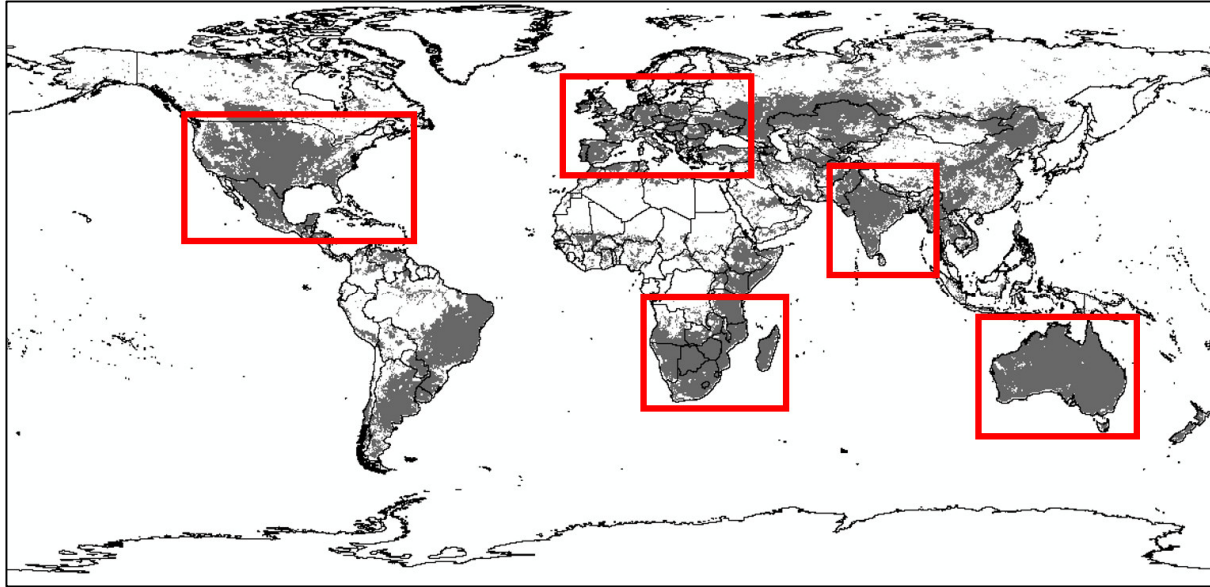
606

607 **Table 3.** Fraction of each macro-area (as percentage) where one model outperforms the other two.

Model	ALL	NA	EU	SA	IN	AU
LIS	25.5	39.2	50.0	10.6	28.2	4.3
LST	25.7	28.8	23.1	36.0	20.3	18.6
CCI	48.8	32.0	26.9	53.4	51.5	77.1

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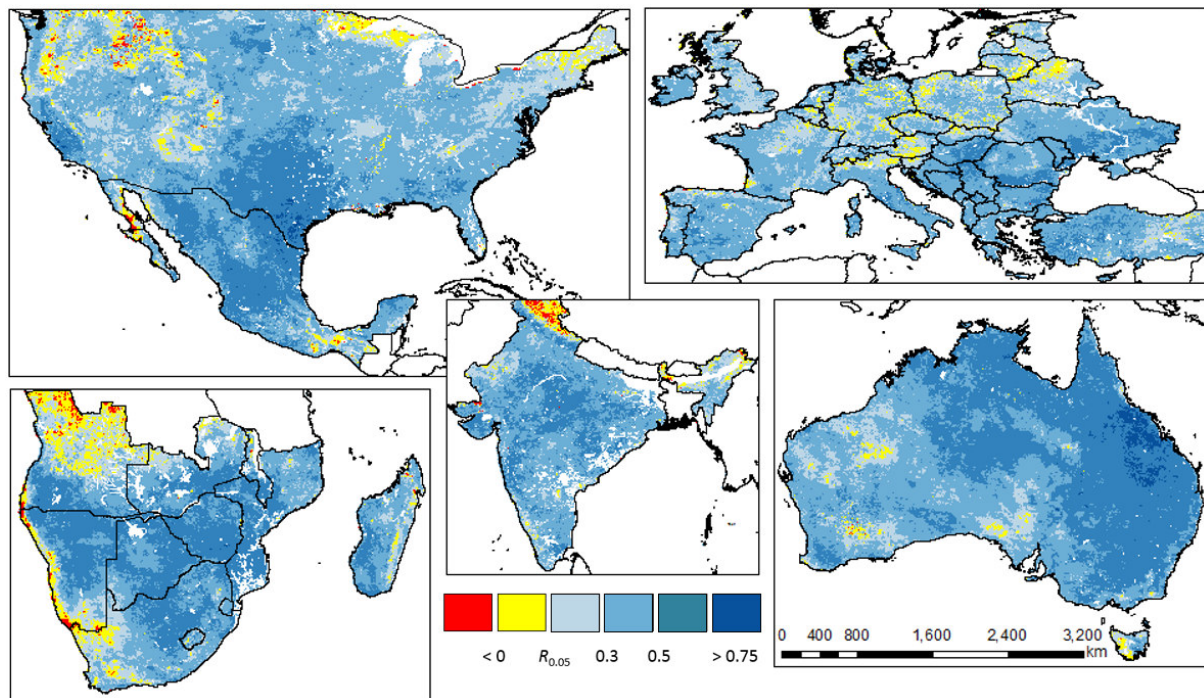
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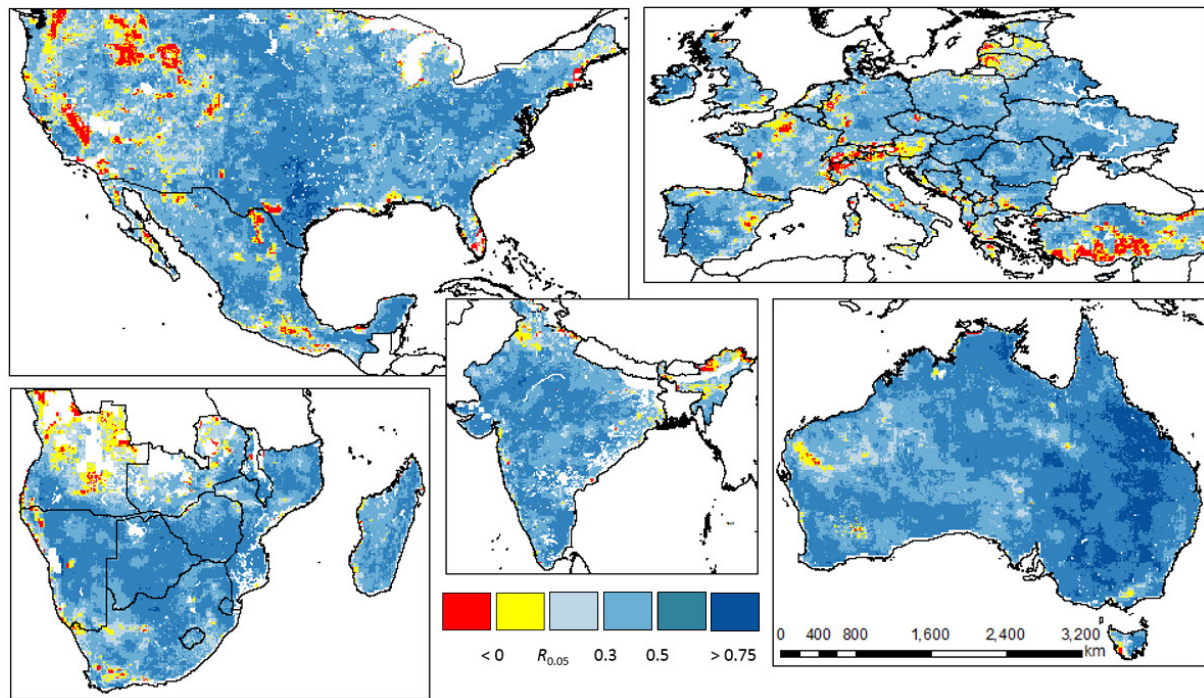
612 **Fig. 1.** Map of the areas where all the three models are positively significantly linearly correlated (cells
613 in grey) according to the Student's t-test at $p = 0.05$. The boxes delimitate the macro-regions selected
614 for the successive analyses.

615



616
 617 **Fig. 2.** Spatial distribution of the Pearson correlation coefficient (R) between Lisflood soil moisture
 618 anomalies (LIS) and land surface temperature anomalies (LST) over the five selected macro-regions.
 619 Values in red and yellow are negatively correlated or not significant at $p = 0.05$, respectively.

620



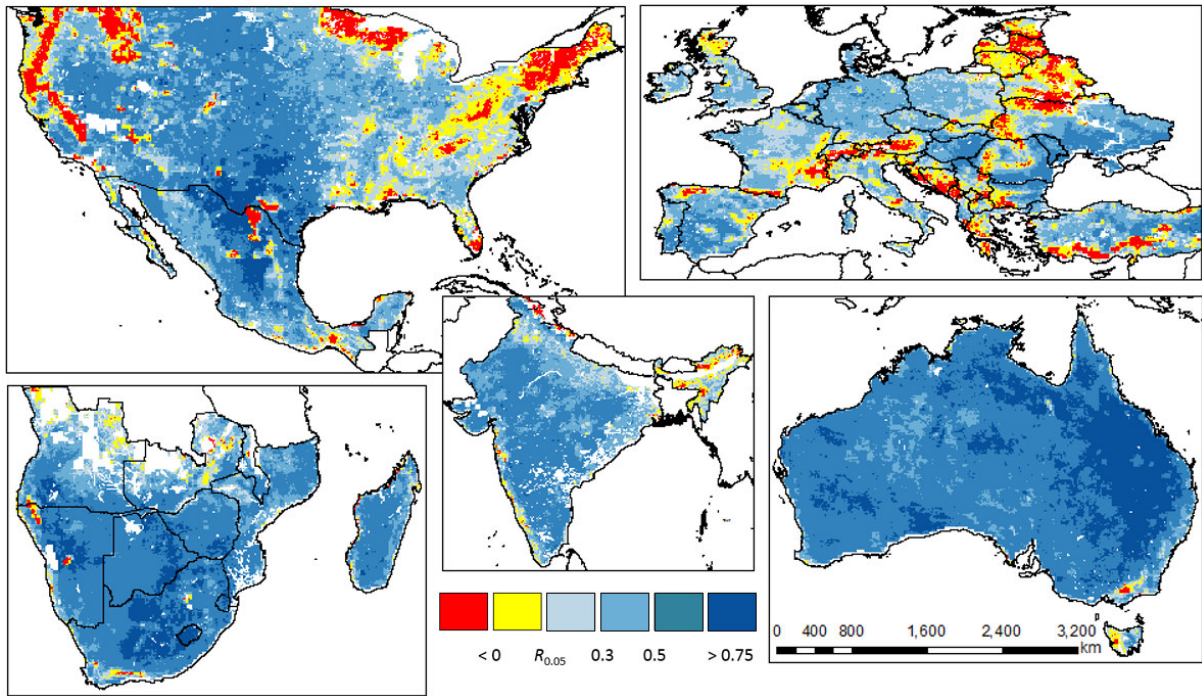
621

622 **Fig. 3.** Spatial distribution of the Pearson correlation coefficient (R) between Lisflood (LIS) and ESA

623 Climate Change Initiative (CCI) soil moisture anomalies over the five selected macro-regions. Values in

624 red and yellow are negatively correlated or not significant at $p = 0.05$, respectively.

625



626

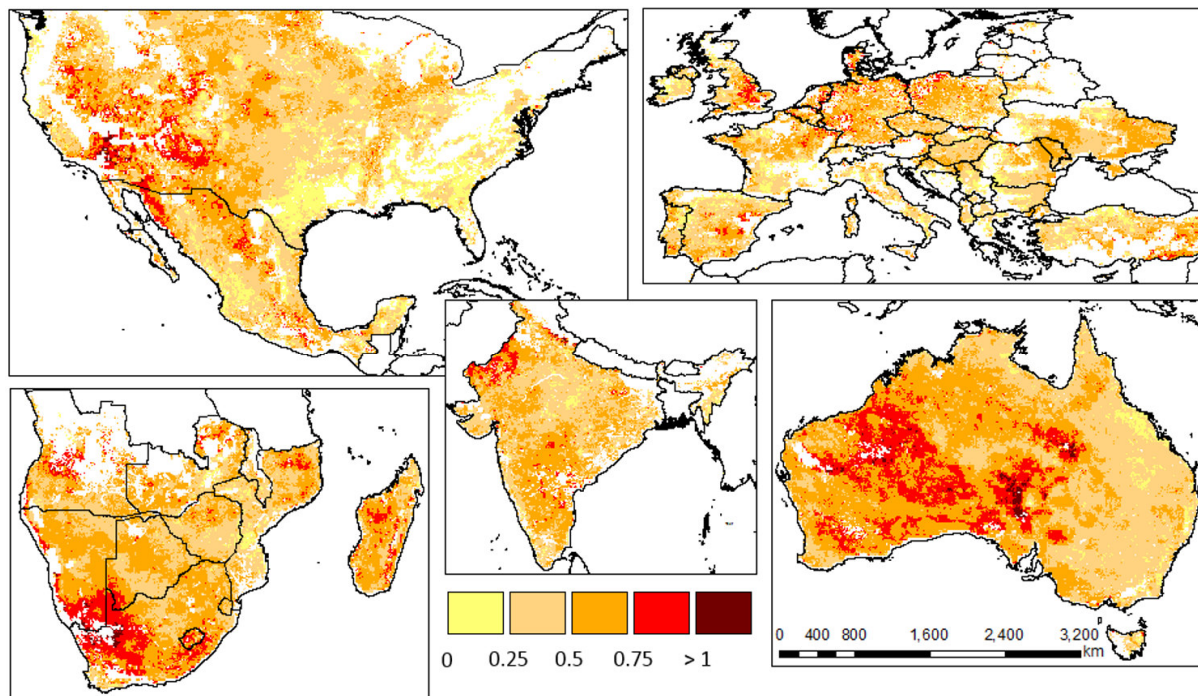
627 **Fig. 4.** Spatial distribution of the Pearson correlation coefficient (R) between ESA Climate Change

628 Initiative soil moisture anomalies (CCI) and land surface temperature anomalies (LST) over the five

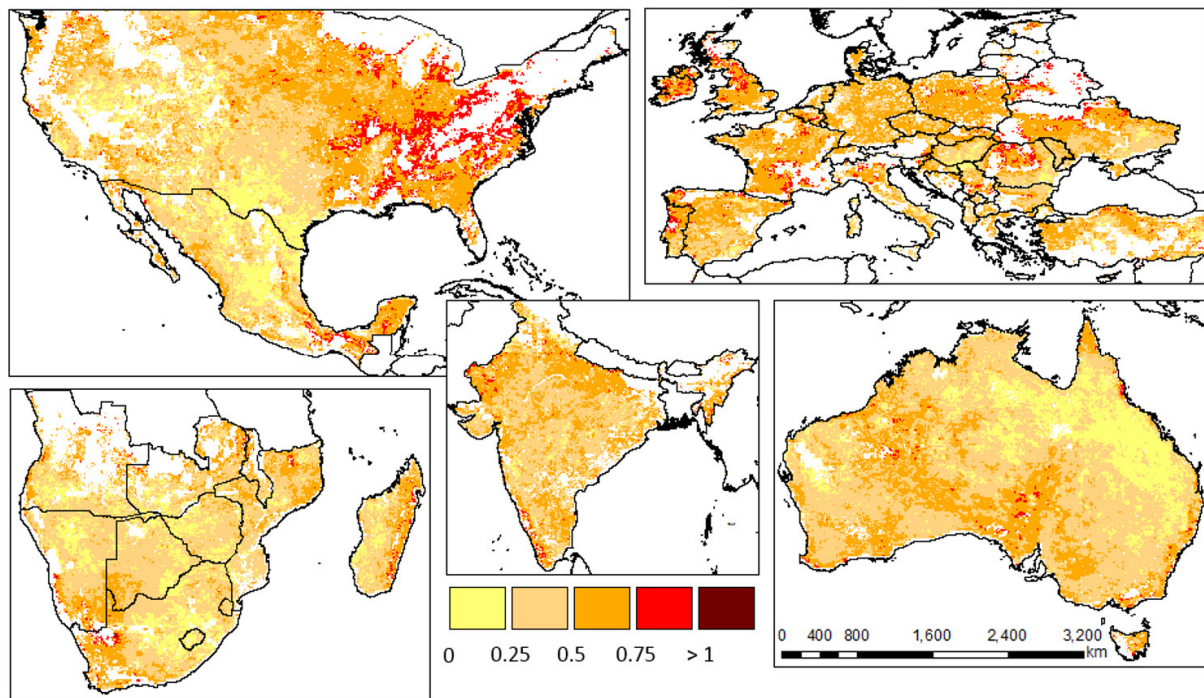
629 selected macro-regions. Values in red and yellow are negatively correlated or not significant at $p = 0.05$,

630 respectively.

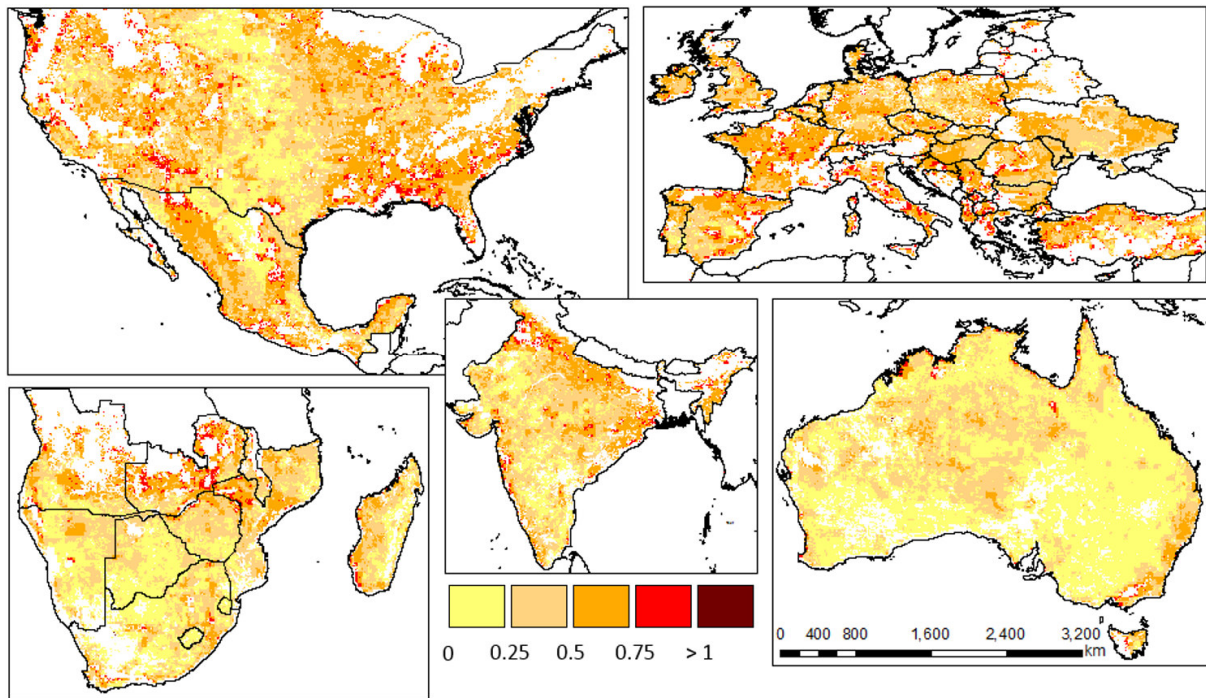
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632
633 **Fig. 5.** Spatial distribution of the error variance for the Lisflood (LIS) dataset over the five selected
634 macro-regions.
635



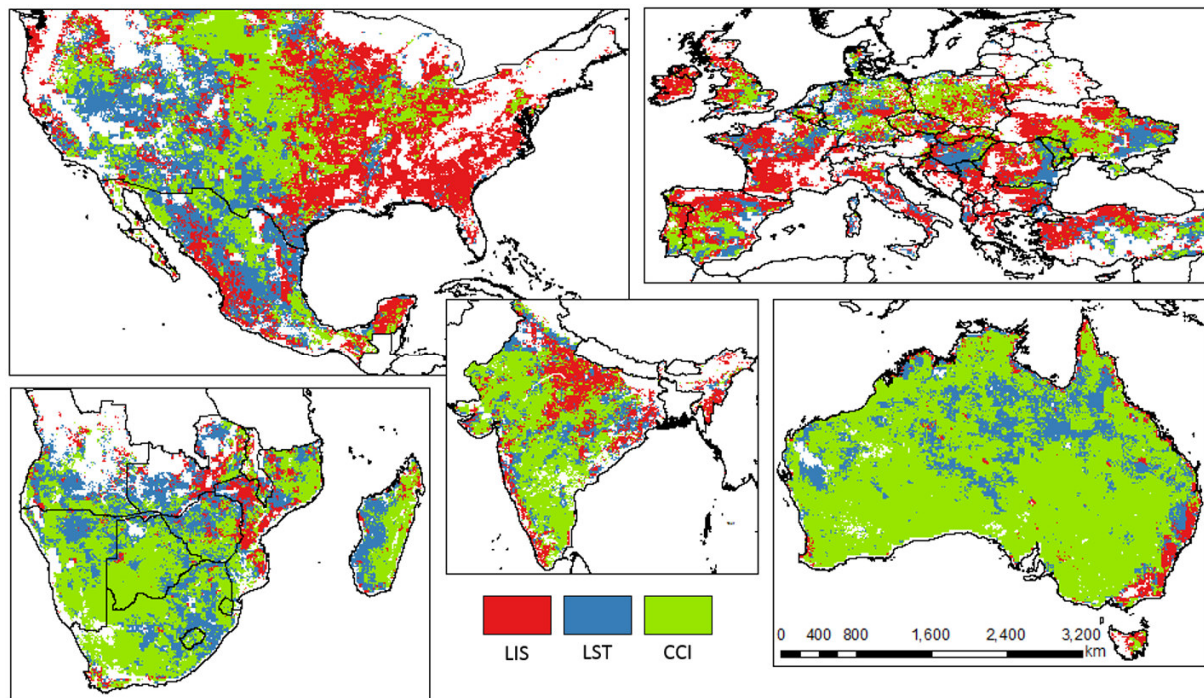
636
 637 **Fig. 6.** Spatial distribution of the error variance for the land surface temperature (LST) dataset over the
 638 five selected macro-regions.
 639



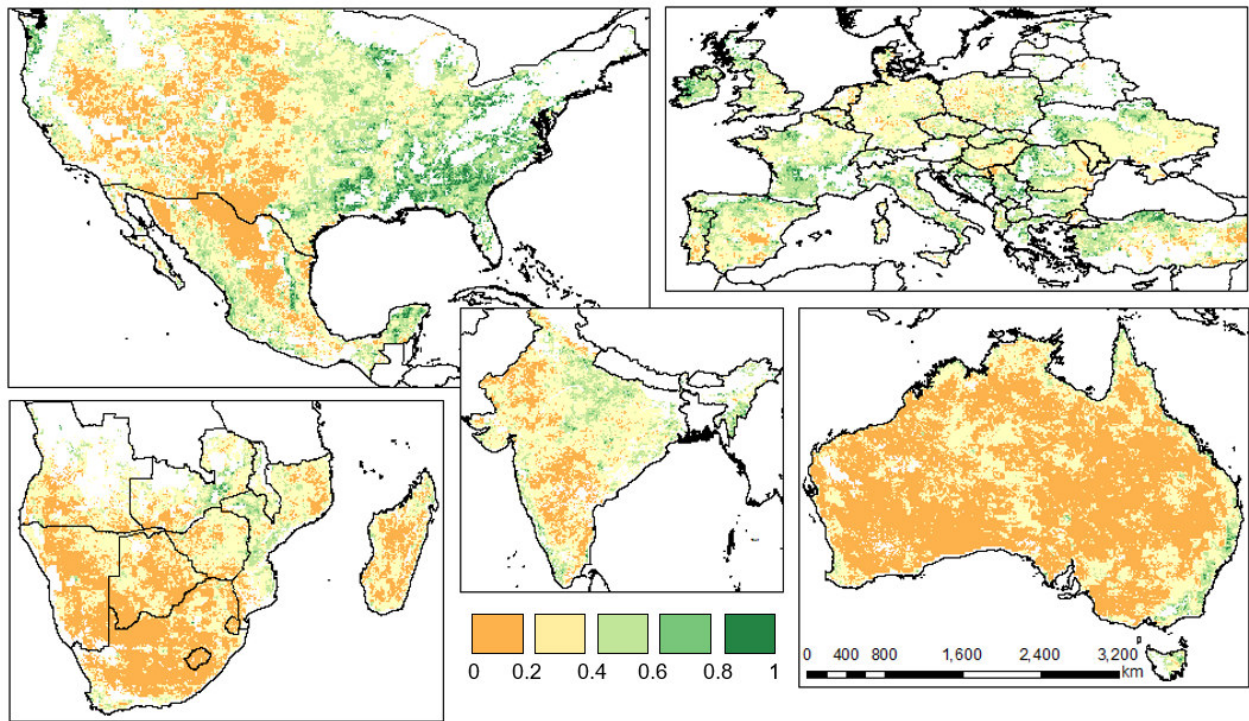
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641 **Fig. 7.** Spatial distribution of the error variance for the ESA Climate Change Initiative (CCI) dataset
642 over the five selected macro-regions.

643



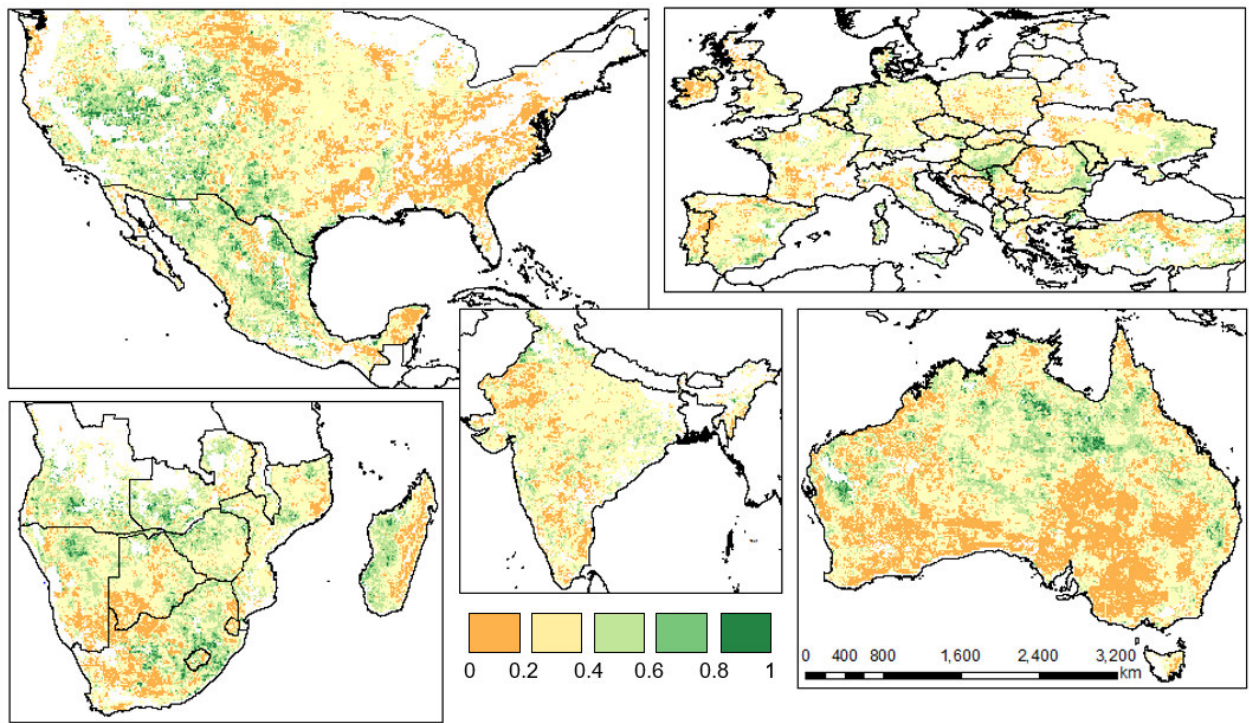
644
 645 **Fig. 8.** Maps representing the best performing (lowest error variance) dataset for each cell according to
 646 the TC analysis.
 647



648

649 **Fig. 9.** Maps representing the ensemble mean weighting factor for the LIS dataset according to the error
 650 maps derived from the TC analysis.

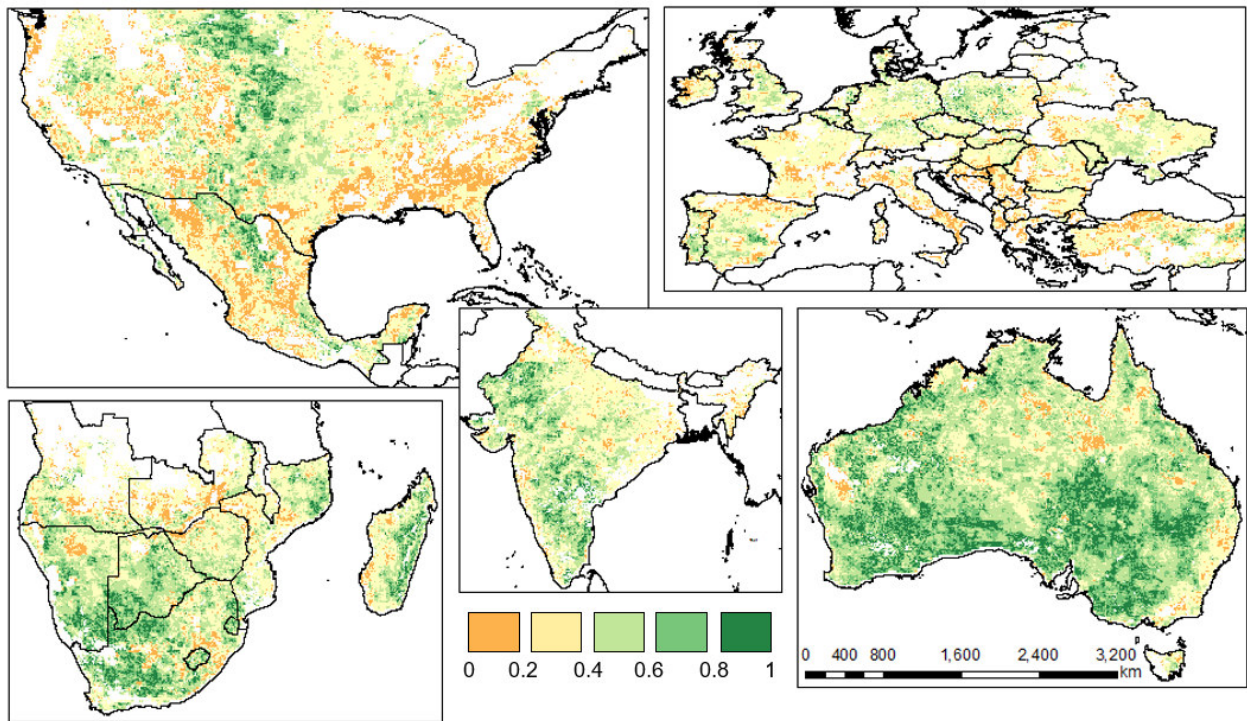
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652

653 **Fig. 10.** Maps representing the ensemble mean weighting factor for the LST dataset according to the
 654 error maps derived from the TC analysis.

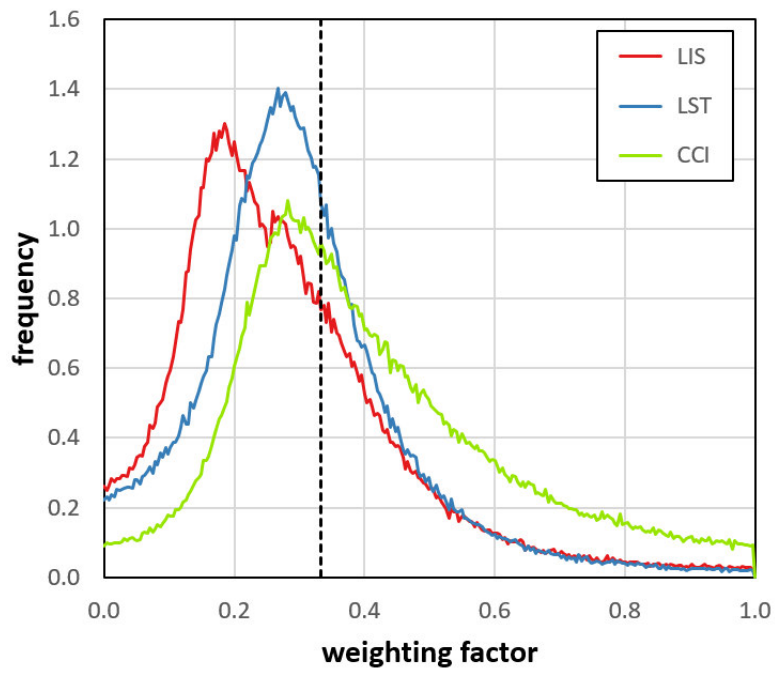
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656

657 **Fig. 11.** Maps representing the ensemble mean weighting factor for the CCI dataset according to the
 658 error maps derived from the TC analysis.

659



660

661 **Fig. 12.** Frequency distribution of the ensemble mean weighting factor for each dataset computed
 662 according to the TC analysis. The black dotted line represents the value corresponding to a simple
 663 arithmetic average (1/3).

664