Improvement of hydrological model calibration by selecting multiple parameter ranges

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Abstract. The parameters are usually calibrated to achieve good performance of hydrological models, owing to the highly non-linear problem of hydrology process modelling. However, parameter calibration efficiency has a direct relation with parameter range. Furthermore, parameter range selection is affected by probability distribution of parameter values, parameter sensitivity and correlation. A newly proposed method is employed to determine the optimal combination of multi-parameter ranges for improving the calibration of hydrological models. At first, single-parameter probability distributions were analysed based on 100 samples obtained from independent Genetic Algorithms (GA) calibration performed on a Xinganjiang model with a corresponding initial parameter range and, the distribution type (i.e. normal, exponential and uniform distributions) was specified for each parameter of the model. Then, the optimal range for each parameter was determined by comparing $E_{\rm NS}$ values calculated separately with the initial range, the minimum and maximum ranges of a given cumulative frequency of 50% (i.e. MINR and MAXR) and the extended range. Next, parameter correlation and sensibility were evaluated by quantifying two indexes R_{CYX} and S_E which can be used to coordinate with the negatively correlated parameters to specify the optimal combination of ranges of all parameters for calibrating models. It is shown from the investigation that the probability distribution of calibrated values of any particular parameter in a Xinanjiang model is closely approximated by a normal or exponential distribution. The multi-parameter optimal range selection method is superior to the single-parameter one for calibrating hydrological models with multiple parameters. The combination of optimal ranges of all parameters is not the optimum inasmuch as some parameters like Im have negative effects on other parameters. The application of the proposed methodology gives a rise to an increase of 0.01 in minimum $E_{\rm NS}$ compared with that of the pure GA method. The rising of minimum $E_{\rm NS}$ with little change of the maximum may shrink the range of the possible solutions, which can effectively reduce uncertainty of the model performance.

Key words: hydrological model, calibration, parameter ranges, probability distribution

1. Introduction

Hydrological process modelling is an important tool for research on water resources management, flood control and disaster mitigation, water conservancy project planning and design, hydrological response to climate change and so on (Zanon et al., 2010;Papathanasiou et al., 2015). The initial hydrological model was a black-box model in 1932 (Sherman, 1932) and conceptual & distributed models are subsequently put forward in 1960s (Freeze and Harlan, 1969). The three kinds of hydrological models have been significantly improved in recent years and their structures become more and more mature. Theoretically, distributed models have definite physical mechanism of the water cycle and all parameters can be measured insitu (Abbott et al., 1986;Huang et al., 2014). Conceptual models express hydrological processes in form of some abstract models which come from some physical phenomenon and experience. For example, the interflow and the base flow are simplified as the flow from linear reservoirs (Caviedes-Voullième et al., 2012;Lü et al., 2013). As a result, some parameters of conceptual models need calibrating. In general, conceptual models have better performance of modelling the streamflow at the catchment outlet than distributed models do, especially for catchments lacking sufficient data (Bao et al., 2010;Cullmann et al., 2011). Thus, many conceptual models such as HBV model, TOPMODEL, Tank model and Xinanjiang model are of strong vitality (Abebe et al., 2010;Vincendon et al., 2010;Hao et al., 2015; Xie et al., 2015). Additionally, the performance of distributed models can be improved after calibration of some parameters. Therefore, all of the hydrological models should be calibrated before engineering applications.

There are two kinds of calibration methods for hydrological models, the trial-error method and auto-calibration method. The trial-error method depends on plenty of trials for reducing the error of the objective. However, it is difficult to obtain an exact optimal solution due to limited enumeration (Boyle et al., 2000). The auto-calibration method is based on stochastic or mathematical calculations, having a wide application in the non-linear parameter optimization. Compared with the trial-error method, it is more efficient and effective, avoiding the interference of anthropogenic factors (Madsen, 2000; Getirana, 2010). The initial automatic optimization methods, such as the Rosenbrock Method (Rosenbrock, 1960) and the Simplex Method (Nelder and Mead, 1965), are classical and useful methods, but has its limitation of initial value ranges of parameters. Therefore, it can only be regarded as local optimization algorithms (Gupta and Sorooshian, 1985). Different from classical methods above, the Genetic Algorithm (GA) is of random search strategy that avoids problem of local search, being a global optimization algorithm in a real sense (Wang, 1991, 1997; Sedki et al., 2009; Chandwani et al., 2015). After that, many global optimization algorithms have been proposed inheriting the random search strategy. The Shuffled Complex Evolution (SCE-UA) method combines many advantages of Genetic Algorithm and Simplex Method, having a powerful capability of calibrating the rainfallrunoff model (Duan et al., 1994; Zhang and Shi, 2011). The Particle Swarm Optimization (PSO) based on random solution can directly obtain the identification parameters through the iterative search for an optimal solution (Kennedy, 1997; Zambrano-Bigiarini and Rojas, 2013). Though the auto-calibration method has been intensively employed to calibrate parameters in the field of hydrology, the most advanced algorithm inevitably falls into local solution because of the strong non-linear problem of a hydrological model and parameter correlation (Chu et al., 2010; Jiang et al., 2010; Jiang et al., 2015).

In general, parameter variables obey some special probability distributions within the given range after multiple independent calibrations (Viola et al., 2009; Jin et al., 2010; Li et al., 2010). Graziani et al. (2008) stated that the shape of a parameter probability distribution can be significantly affected by a parameter range. Ben et al. (2013) studied the effect of different probability distributions (e.g., Normal distribution and uniform distribution) of parameters values on parameter sensitivity, and found that the probability distribution can be provide a clue to realize parameter sensitivity. Although normal and uniform distributions are greatly studied in practice, other types of probability distributions seldom were investigated in previous researches (Kucherenko et al., 2012; Esmaeili et al., 2014).

Most hydrological models contain many parameters of different sensitive characteristics and correlation behaviour. Some researchers believe that the sensitive parameter should be calibrated, but the insensitive parameter can be set as a fixed value by experience (Beck, 1987; Cheng et al., 2006). Inappropriate parameter ranges or fixed values may result in the instability of calibrated results. Furthermore, the range setting of one parameter may influence the calibration of other related parameters (Song et al., 2015). The model parameter sensitivity analysis has been a growing concern in recent years. Parameters sensitivity varies with catchment characteristics, objective functions and parameter ranges (van Griensven et al., 2006). Wang et al. (2013) noted the different parameter ranges could lead to changes in parameter sensitivity. Shin et al. (2013) reported that reducing or extending ranges might make insensitive parameters become sensitive ones or vice versa. Thus, parameter ranges and correlation should be taken into considered when the calibration of multi-parameters models is performed.

Parameter ranges are generally given roughly due to lack of knowledge concerning physical settings of a local catchment (Song et al., 2013; Hao et al., 2015). The more deviation between an optimal range and a given range, the more uncertainty of the calibration result. The selection of appropriate parameter ranges is critical for calibrating the model efficiently. However, few literature covers information on how to select the appropriate parameter range for improving the calibration of hydrological models. Furthermore, the calibration of multiple parameters is more complex due to parameter sensitivity and correlation. Hence, it is necessary to find a way to coordinate the range settings of all parameters.

Considering the effect of parameter ranges on calibration efficiency of hydrological models, an approach of parameter ranges selection (PRS) is put forward to improve the calibration of hydrological models with multiple parameters. At first, probability distribution of each parameter was analysed based on a lot of independent calibrations by using a GA method. Then the optimal range of a single parameter was specified for calibration according to its probability distribution. Finally, parameter correlation and sensitivity were estimated to determine the optimal combination of multiple parameters ranges. The proposed method is expected to be helpful for an effective and efficient calibration of hydrological models with multiple parameters.

2. Study area and data collection

The Chaotianhe River catchment is located in the northeast of Guangxi Zhuang Autonomous Region in Southwest China (Fig. 1). The Chaotianhe River is the major tributary of the Lijiang River of well-known karst landscape. The total catchment area is 476.24 km². The annual precipitation is approximately 1704 mm and 78% precipitation concentrates in flood seasons (March

to August). The thickness of soil varies spatially in most karst areas. Limestone is exposed to air in some peak-cluster region. Clay soil with thickness ranging from 2 to 10 m is distributed in the depressions and valleies. In clastic rock mountain areas, the thickness of the soil is usually less than 0.5 m. Thus, the soil moisture storage capacity varies significantly with space. Moreover, the underground rivers are very well developed in the karst area, which makes the flood gather rapidly and recess slowly due to higher underground flow rate.

The daily data concerning precipitation, evaporation and streamflow were collected from national gauging stations for the 5-year period of 1996–2000. Four precipitation stations, one streamflow gauging station and one evaporation station are selected for the investigation. Areal precipitation was calculated using data from the four precipitation stations by using a Thiessen polygon method under GIS environment (Cai et al., 2014). The streamflow gauging station is at the catchment outlet. Some metro-hydrological statistical data of the studied catchment are summarized in Table 1. From 1996–2000, the maximum of daily streamflow is about 719 m³/s, the minimum 0.53 m³/s and the average is 13.31 m³/s at the outlet. The maximum areal daily precipitation varied with years in the studied catchment and reached the value of 235 mm/d in 1996.

3. Methodology

3.1 Hydrological model selection

The method of parameters ranges selection (PRS) is designed for most of hydrological models. At present, there have been many hydrological models for hydrological processes simulation. Considering the climate characteristics of the study area, the Xinanjiang model which is suitable for humid regions was chosen to serve as a hydrological model for the investigation. The Xinanjiang model mainly includes three evapotranspiration layers and three runoff components (i.e. surface-, subsurface runoff and groundwater) (Zhao et al., 1980;Zhao, 1992). The surface runoff is routed by the Unit Hydrograph (UH) which is derived from the observed streamflow and, other runoff components are simplified as linear reservoirs (Ju et al., 2009). With regard to the Xinanjiang model, there are 10 parameters that should be calibrated. The definitions of the parameters are given in Table 2 (Lin et al., 2014;Hao et al., 2015). The proposed PRS method is introduced as follows, when a Xinanjiang model is taken as an example.

3.2 Probability distribution analysis of calibrated parameter value

25 3.2.1 Sample collection of calibrated parameter value

In theory, the parameter values calibrated by using a stochastic-based auto-calibration method are not same to each other but obey a certain probability distribution under a reasonable convergence condition (Jiang et al., 2015). In order to analyse the probability distribution of parameter values, a stochastic-based auto-calibration is used to calibrate the model, and samples of calibrated parameters values are obtained. As far as the sample size is concerned, 100 samples are enough to estimate the probability distribution of calibrated parameter values in the investigation, which is deduced from the results of trial tests as shown in Fig. 2. It can be seen that both maximum and minimum E_{NS} keep stable when sampling size is greater than 100.

A Genetic Algorithm (GA) was selected as the auto-calibration method in the investigation, because GA is a common and widespread used global optimization algorithm based on stochastic and evolutionary optimization. Many studies show that evolutionary algorithms provide equal or better performance of a model than other algorithms do (Cooper et al., 1997; Jha et al., 2006; Zhang et al., 2009). The Nash–Sutcliffe efficiency (E_{NS}) was chosen as an objective function (Eq. (1)) for GA, which represents the agreement between observed and simulated data.

$$E_{NS} = 1 - \frac{\sum_{i=1}^{n} (Q_{\text{obs},i} - Q_{\text{sim},i})^2}{\sum_{i=1}^{n} (Q_{\text{obs},i} - Q_{\text{mean}})^2}$$
(1)

where E_{NS} is Nash–Sutcliffe efficiency, i serial number of the step; n total number of the observed streamflow data, $Q_{obs,i}$ the observed streamflow at step i, $Q_{sim,i}$ the simulated streamflow at step i, and Q_{mean} is the mean value of observed streamflow.

3.2.2 Determination of probability distribution types

The probability distributions of calibrated parameters values can be estimated roughly by using box-plot charts, cumulative frequency curves and frequency histograms. The symmetry of the box-plot chart (including one box and two whiskers) and the length ratio of the whisker to the box, the shape of the cumulative frequency curve and the frequency histogram are important indicators for the identification of the distribution type. Based on these indicators, three types of probability distributions are listed as follows: (1) Normal distributions, the box and whiskers are approximately symmetrical along the Y-axis direction, the length of either whisker is longer than half height of the box in a box-plot chart (Fig. 3a), the cumulative frequency curve is S shaped and the histogram bell shaped (Fig. 3b); (2) Exponential distributions, the whole chart is distinctly asymmetrical in the Y-axis direction, which means that the average value (marked with a small hollow square) deviates from the median value (marked with a centre line in box), the box is inclined to one side with the extreme shorter whisker (Fig. 3a), the cumulative frequency curve is parabola shaped, and the histogram tends to increase or decline gradually (Fig. 3c); (3) Uniform distribution, the box and whiskers are approximately symmetrical along the Y-axis direction, the length of two whiskers is close to that of the box (Fig. 3a), the cumulative frequency curve tends to a straight line and the histogram varies little along the X-axis (Fig. 3d).

A Kolmogorov-Simirnov test (K-S test) tries to examine whether a data set fit a reference probability distribution or not (Haktanir, 1991). In a K-S test, for any variable x_i in a data set, the empirical distribution function value (Fi) is calculated by using a plotting position formula, and the cumulative distribution function value (Fi*) is computed by using the reference probability distribution. The maximum deviation between the two values, Δ_{Max} , is expressed in Eq. (2).

$$\Delta_{Max} = |F_i^* - F_i| \tag{2}$$

According to the acceptable level of significance α (α =0.2) and the total number of values in a data set n, Δ_{table} can be obtained from the K-S table. If $\Delta_{Max} < \Delta_{table}$, the reference probability distribution is identified to fit to the data set.

3.3 Parameters ranges selections

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3.3.1 Single parameter range selection (S-PRS)

In order to improve E_{NS} , the initial range of a parameter requires adjusting properly. In consideration of the three probability distribution types mentioned above, the different ways to specify the optimal range for a single parameter are presented in the investigation. For the parameter of a uniform distribution, it is better to keep the initial range due to little influence of ranges on calibration results. For the parameter of a normal distribution, the cumulative frequency curve is employed to seek some reduced ranges with a given cumulative frequency (e.g. 50%), and the minimum and maximum ranges (namely MINR and MAXR) are obtained as depicted in Fig. 4. The MINR and MAXR represent the ranges of maximum and minimum probability density of parameter values under a given cumulative frequency, respectively. As for the parameter of an exponential distribution, the initial range can be extended appropriately towards one side of high probability density, if the parameter has reasonable meaning in the extended range. Then, the optimal range of the parameter can be specified by comparing different $E_{\rm NS}$ calculated separately by using the initial range, the MINR or MAXR of the initial range, the MINR or MAXR of the extended range. If the initial range cannot be extended, the MINR and MAXR are sought out according to the cumulative frequency curve. Figure 5 gives the variation curves of maximum and minimum $E_{\rm NS}$ of a single parameter with cumulative frequency values. It is found that the maximum $E_{\rm NS}$ keeps constant despite a cumulative frequency value varying, while the minimum $E_{\rm NS}$ approaches the peak value of 0.881 when the cumulative frequency value is equal to 50%. Considering that higher minimum E_{NS} contributes to more efficient calibration, the fixed cumulative frequency value of 50% was selected to determine the ranges of maximum and minimum probability density (i.e. MINR and MAXR) for each parameter. In short, the optimal range of a single parameter can be determined by properly extending or reducing the initial range to make calibrated parameter values distributed quite closely to a uniform distribution.

3.3.2 Multiple parameters ranges selection (M-PRS)

In general, there is more or less correlation between parameters for most hydrological models. As far as a Xinanjiang model is concerned, parameters WM and B refer to the water storage volume – area curve that represents the spatial variability of soil moisture storage. If the curve is fixed, a larger WM results in a smaller B (Zhao et al., 1980). As a result, the range change of parameter WM may affect the range setting and calibration of parameter B. The correlations among parameters, therefore, should be taken into account, if the ranges of the related parameters require adjusting. If the range change of one parameter has positive influence on calibration of other parameters, using the optimal range of the parameter instead of the initial one can contribute to better calibration results. On the contrary, the negative impact may result in a worse model calibration, although the optimal ranges of the parameters are used. Thus, some coordination measures should be taken to deal with such a contradiction. The index R_C (Eq. (3)) was quantified to analyse the influence degree of one-parameter range change on the calibration of other parameters. When $R_{C,Y,X}$ is closer to 1, the range change of parameter X has a greater positive influence on the calibration of parameter Y. If $R_{C,Y,X}$ is minus, it means a negative influence.

$$R_{\rm CY,X} = 1 - \frac{L_{\rm Y,X} - L_{\rm Y,Y}}{L_{\rm Y,Initial} - L_{\rm Y,Y}} \tag{3}$$

Where $R_{CY,X}$ is the influence degree of the range change of parameter X on the calibration of parameter Y; $L_{Y,X}$ the range of parameter Y calibrated with the optimal range of parameter X and initial ranges of other parameters, $L_{Y,Y}$ the range of parameter Y calibrated with the optimal range of parameter Y and initial ranges of other parameters, and $L_{Y,Initial}$ is the range of parameter Y calibrated with initial ranges of all parameters. The calibrated range of any parameter is calculated except extreme outliers. If there is a negative influence between two parameters, the optimal range of the parameter of higher sensitivity is used and the initial range of the other parameter kept for calibration generally to mitigate the negative impact. It is due to that sensitive parameters play more important roles than insensitive parameters do in a multi-parameter calibration. In order to assess the sensitivity of parameter range change to E_{NS} , index S_E as expressed in Eq. (4) is computed by performing an S-PRS method on each parameter. The larger value of R_E , the more concentrated distribution of E_{NS} which means more efficient parameter calibration. Thus, the parameter of higher S_E is given priority to use the optimal range when the R_C of two parameters is minus.

$$S_E = 1 - \frac{E'_{\text{NS Max}} - E'_{\text{NS Min}}}{E_{\text{NS Max}} - E_{\text{NS Min}}} \tag{4}$$

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where $S_{\rm E}$ is sensitivity of parameter range change to $E_{\rm NS}$, $E_{\rm NS~Max}$ and $E_{\rm NS~Min}$ maximum and minimum $E_{\rm NS}$ calibrated with an initial range, $E'_{\rm NS~Max}$ and $E'_{\rm NS~Min}$ are maximum and minimum $E_{\rm NS}$ calibrated with an optimal range. The statistical analysis of $E_{\rm NS}$ excludes extreme outliers.

Considering there are more than two parameters in most hydrological models, the accumulative influence and the coordination of range selection were investigated in the study. The mean value of $R_{\rm C}$ ($R_{\rm C}$ mean) is the index to judge the accumulative influence of one-parameter range change on the calibration of other parameters. Thus, for parameters of a negative $R_{\rm C}$ mean, the initial ranges instead of the optimal ones are adopted for the multi-parameters calibration.

The flow chart of the parameter range selection method is shown in Fig. 6. In stage 1, a set of initial ranges of parameters is given for a hydrological model and the probability distribution for each parameter analysed based on the 100 independent parameters values calibrated by an auto-calibration method. In stage 2, there are three range adjustment methods with response to a probability distribution of parameter values: for a normal distribution, the optimal range of a single parameter is obtained by reducing the initial range; for an exponential distribution, the initial range of a single parameter is extended to specify the optimal range, or the initial range is reduced to seek the optimal range for calibration when the extension of the parameter range is limited; for a uniform distribution, the initial range is kept. In stage 3, the single-parameter range selection (S-PRS) is performed on each parameter. Based on the indexes S_E and R_C estimated, the optimal combination of ranges is determined by coordinating the range selection of all parameters.

4. Results and discussion

4.1 Probability distribution characteristics of calibrated parameter values of the Xinanjiang model

A series of calibrated parameters values were obtained through 100 independent calibration runs by using a GA method. Trial tests were employed to determine the optimal GA control parameters: crossover probability of 0.5, mutation probability of 0.7 for the individual, mutation probability of 0.5 for each gene, population size of 21, maximum generation number of 500 and maximum iteration number of 50. These parameters were kept constant for GA calibrations in the investigation. The initial and calibrated ranges of parameters are presented in Table 3. The ratio of the calibrated range length to the initial one in Table 3 is less than 60% for most parameters (i.e. parameter CI, Kc, KI, SM, B, and WM), which implies that reducing the ranges can help calibrate most parameters efficiently. For any particular parameter, calibrated values were normalized by dividing a deviation between a calibrated value and the lower limit of the initial range by the length of the initial range. Based on 100 calibrated values after normalization, a box-plot for a parameter is depicted. It is obvious from Fig. 7 that the box and whiskers are approximately symmetrical and the length of whiskers is longer than that of half box along the direction of the Y axis for parameters CI, SM and Kc. But for other parameters, it is shown from the box-plot charts that the mean value deviates from the median one, which means an asymmetric chart. According to the characteristics of the box-plots, it is shown that the probability distributions of the calibrated values are normal for parameters CI, SM, and Kc, while those are exponential for other parameters. Furthermore, K-S tests were employed to determine the probability distributions of parameters and the corresponding results are listed in Table 3. It is shown that only a normal distribution is accepted for parameters CI & SM. Despite the fact that both normal and uniform distributions are accepted for parameter KC, the probability distribution of parameter KC is regarded as a normal distribution. It is because that the Δ_{Max} will become smaller if a normal distribution serves as a reference distribution instead of a uniform distribution. In addition, just an exponential distribution is accepted for the rest of the parameters. Thus, the three parameters follow normal distributions and the others exponential distributions in the Xinanjiang model. The ratio of the calibrated range length to the initial range length is less than 30% for parameters CI, SM, and Kc, while the ratio exceeds 30% for parameters B, WM, C, EX, CG, and Im. It implies that reducing the initial ranges can improve the calibration for parameters whose values obey normal distributions.

25 4.2 Effect of range adjustment pattern on calibration results

Since the probability distribution of a single parameter has a direct relation with the parameter range selection, the range adjustment pattern of a single parameter was discussed on the basis of the parameter probability distribution in the investigation.

For a normal distribution, reducing the range was used to find the optimal range. Figure 8 shows the calibration results of parameter CI when the different ranges are selected. The MINR (0.679–0.713) and the MAXR (0.623–0.694) were picked out based on the cumulative frequency curve derived from calibrations with the initial range (0–0.900). From the cumulative curves and the histograms in Fig. 8a, 8b and 8c, it is found that the probability distribution of parameter CI values is converted from a normal distribution to a uniform distribution when the initial range is reduced to the MINR, whereas the probability

distribution approximates an exponential one when the MAXR is used. Figure 8d reveals the contribution of the parameter range selection to $E_{\rm NS}$. It is found that the minimum $E_{\rm NS}$ except extreme outliers rises from 0.881 to 0.884 and the $E_{\rm NS}$ concentrates at a higher value range, when the MINR is used instead of the initial range. Using the reduced range of high probability density is, therefore, helpful to make calibration more stable and more efficient.

To an exponential distribution, both reduced ranges and extended ranges of reasonable meaning were used to select the optimal range for parameter calibration. Figure 9 shows the calibration results under three different input ranges of parameter KI. Since the initial range of parameter KI cannot be extended, the two reduced ranges (i.e. the MINR (0.660–0.700) and the MAXR (0.522–0.660)) were picked out by using the cumulative frequency curve. From the cumulative curves and the histograms in Fig. 9a, 9b and 9c, it is found that the probability distribution of parameter KI values is similar to a uniform distribution in the case of the MINR, whereas that is still exponential in the case of the MAXR. The contributions of the three parameter ranges to E_{NS} are shown in Fig. 9d. Thus, the MINR is best for calibration of parameter KI when compared with the MAXR or the initial range, which is similar to the calibration result of parameter CI. In general, the MINR is better than the MAXR for parameter calibration, because the parameter values that may achieve a higher E_{NS} can be easily picked out from the MINR of higher probability density.

Figure 10 shows the calibration results of parameter B whose initial range can be extended. Parameter B generally ranges from 0.1 to 0.4 for most areas, but it is quite different for karst areas where the soil moisture storage varies remarkably with space. As a result, the value of parameter B can be greater than 0.4. From Fig. 10a and 10b, it is shown that the probability distribution of parameter B is converted from an exponential distribution to a normal distribution when the initial range is extended to new one (B=0.1–0.6). After the MINR selection was performed on the initial range and the extended range respectively, the two ranges, i.e. the MINR (B=0.36–0.40) and the extension-MINR (B=0.379–0.488) were obtained and then used to calibrate parameter B. From Fig. 10c and 10d, it is found that the probability distribution of parameter B approximates a uniform distribution when the MINR or the extension-MINR is used. The box-plots of E_{NS} for different ranges are shown in Fig. 10e. It is shown that there is little improvement in maximum E_{NS} when MINR is used for calibration instead of the initial range. There is an increase of 0.0003 in maximum E_{NS} if the initial range is replaced with the extension range or the extension-MINR. As for minimum E_{NS} (except outliers), an increase of 0.001 in the case of the MINR, a decrease of 0.003 in the case of the extension-MINR are found when the initial range is substituted with the three ranges respectively. It suggests that an appropriate range extension followed by a MINR selection is helpful to improve calibration for the parameter whose probability distribution is exponential and initial range can be extended.

4.3 Effect of multiple parameters ranges combination on calibration results

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The S-PRS method was employed to determine the optimal range for each parameter. According to the optimal ranges and the corresponding initial ranges, indexed R_C and S_E were quantified to understand parameter correlation and sensitivity. It is obvious from Table 4 that R_C values in the columns of parameters CI and WM are positive, but most R_C values in the column of parameter Im are negative. The negative R_C related to two parameters means that using the optimal range of one parameter

is adverse to calibrate the other. Both $R_{\text{C EX,Im}}$ and $R_{\text{C Im,EX}}$ are negative in spite of small values, which implies that using the optimal ranges of parameters EX and Im simultaneously is not conductive to calibrate these two parameters. The mean of R_{C} ($R_{\text{C mean}}$) varies with parameters. Parameter CI has the maximum $R_{\text{C mean}}$ of 0.465 and parameter Im the minimum $R_{\text{C mean}}$ of – 0.026. Furthermore, all parameters have positive $R_{\text{C mean}}$ values except for parameter Im, owing to the accumulative negative correlation between parameter Im and the others.

To coordinate with negatively related parameters, the index S_E was used to pick out parameters of higher sensitivity to E_{NS} . From Table 4, it is found that parameter CI has the maximum S_E of 54.7% and parameter Im the minimum S_E of 0.3%. Most S_E values are more than 20% except those of parameters C, EX and Im. It suggests that parameters C, EX and Im are of low sensitivity to $E_{\rm NS}$ and the others highly sensitive to $E_{\rm NS}$. Parameter CI is the most sensitive while Im the most insensitive, which agrees with the work of Lü et al. (2013) and Song et al. (2013). For the well-developed karst areas, the thin layer of soil and strong permeability of limestone make rainfall easy to penetrate into the ground. Moreover, the existence of karst caves and subsurface streams contribute to great interflow storage which accounts for a large proportion of streamflow. As a result, the calibration of parameters KI (representing the penetrability of free water to interflow) and CI (representing recession capacity of interflow storage) has a significant influence on rainfall-runoff simulation results. Hence, parameters KI and CI are highly sensitive in the investigation. Thus, the optimal ranges of parameters of higher sensitivity should be used to improve calibration. In order to determine the optimal range combination of multiple parameters, seven cases were investigated with different range combinations of parameters (Table 5). Case 1 was defined as the initial case using all initial ranges. Cases 2-4 were defined as the single parameter range selection (S-SPR) cases. Cases 5–7 were set as the multiple parameters ranges selections (M-SPR) cases. The box-plots of E_{NS} for different cases are given in Fig. 11. There is a little decrease in E_{NS} when Case 4 is separately compared with Case 1, Case 2 and Case 3. It can be explained that both R_{C EX,Im} and R_{C Im,EX} are negative and the combination of the optimal ranges corresponding to the two parameters leads to a worse calibration result. As the S_E value of parameter Im is less than that of parameter EX, parameter EX is given priority to use the optimal range. It is the reason why the calibration result of Case 3 is better than that of Case 2. As for the cases with the multi-parameter range selection (i.e. Cases 5–7), the $E_{\rm NS}$ values are much greater than those of cases 1-4. There is approximately an increase of 0.001 in maximum $E_{\rm NS}$ and an increase of 0.01 in minimum $E_{\rm NS}$ when the multi-parameter range selection is performed. There are some differences in $E_{\rm NS}$ with the comparison between Cases 5–7 in a magnified box-plot chart. Case 6 has the most concentrated values of $E_{\rm NS}$ and the largest mean value of E_{NS} among the three cases. It means that the combination of optimal ranges of all parameters (see Case 7) is not the optimum to calibrate a multi-parameter model inasmuch as some parameters like Im have negative correlation on other parameters. Hence, the initial ranges of parameters having negative mean values of R_C and low values of

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Through a calibration run, a set of calibrated values of all parameters and the corresponding E_{NS} are obtained. Figure 12 shows the variation curves of maximum and minimum values of E_{NS} with number of runs by using a GA method and a proposed PRS method, respectively. It is indicated from Figure 12 that no mater it is maximum or minimum E_{NS} , the value calculated by using a proposed method is almost the same as that by using a GA method when the number of runs does not exceed 100.

 $S_{\rm E}$ are supposed to be used to calibrate parameters instead of the corresponding optimal ranges.

If a proposed method is used for calibration instead of a GA method, there are approximately an increase of 0.001 in maximum $E_{\rm NS}$ and an increase of 0.01 in minimum $E_{\rm NS}$ when the number of runs is greater than 100. Thus, for any particular run number, the value of $E_{\rm NS}$ calculated by using a PRS method is not less than that by using a GA method. The application of a proposed method, therefore, contributes to a relatively efficient calibration.

5 5. Conclusions

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Considering that there is a relation between the selection of multi-parameter ranges and the calibration effect of a hydrological model, an approach to determine an optimal combination of ranges for the multi-parameter calibration was put forward by analysing parameter probability distribution, parameter sensitivity and correlation between parameters. The newly proposed method was applied for the calibration of a Xinanjiang model for karst areas, and some findings are presented as follows.

In the Xinanjiang model for karst areas, parameters CI, Kc, SM and B approximately obey normal probability distributions and parameters WM, C, EX, KI, CG and Im exponential probability distributions after 100 independent calibration runs. For the parameters of a normal distribution, the MINR defined by using a cumulative frequency curve of calibrated values is preferred to be selected as the optimal parameter range for calibration. For the parameters of an exponential distribution, the extension-MINR is recommended to be used for calibration if the initial range can be extended towards the high-probability side; otherwise the MINR is selected as the optimal range for calibration.

The proposed parameter range selection (PRS) method improves the minimum and mean values of $E_{\rm NS}$. The application of the proposed methodology results in an increase of 0.01 in minimum $E_{\rm NS}$, compared with that of the pure GA method. The rising of minimum $E_{\rm NS}$ with little change of the maximum may shrink the range of the possible solutions. As a result, the uncertainty of the model performance can be effectively controlled.

The M-SPR method is superior to the S-SPR one for calibrating hydrologic models with multiple parameters. The $R_{\rm C}$ and $S_{\rm E}$ are two important indexes that can help to analyse the sensitivity and correlation between parameters and consequently to coordinate with the negatively related parameters. The initial ranges of parameters of relatively low $S_{\rm E}$ and negative $R_{\rm C}$ mean and the optimal ranges of parameters of positive $R_{\rm C}$ mean should be preferred to be chosen for the multi-parameter model calibration.

Acknowledgements. The investigation is supported by special funds for public welfare industry research projects of the Ministry of Water Resources of the People's Republic of China (No. 201401057) and the Scientific Research Foundation for the Returned Overseas Chinese Scholars, State Education Ministry (No. 2013–1792).

Reference

Abbott, M. B., Bathurst, J. C., Cunge, J. A., O'Connell, P. E., and Rasmussen, J.: An introduction to the European Hydrological System — Systeme Hydrologique Europeen, "SHE", 1: History and philosophy of a physically-based, distributed modelling system, J. Hydrol., 87, 45-59, doi:10.1016/0022-1694(86)90114-9, 1986.

- Abebe, N. A., Ogden, F. L., and Pradhan, N. R.: Sensitivity and uncertainty analysis of the conceptual HBV rainfall–runoff model: Implications for parameter estimation, J. Hydrol., 389, 301-310, doi:10.1016/j.jhydrol.2010.06.007, 2010.
- Bao, H., Wang, L., Li, Z., Zhao, L., and Zhang, G.: Hydrological daily rainfall-runoff simulation with BTOPMC model and comparison with Xin'anjiang model, Water Sci. Eng., 3, 121-131, doi:10.3882/j.issn.1674-2370.2010.02.001, 2010.
- 5 Beck, M. B.: Water quality modeling: A review of the analysis of uncertainty, Water Resour. Res., 23, 1393-1442, doi:10.1029/WR023i008p01393, 1987.
 - Ben Touhami, H., Lardy, R., Barra, V., and Bellocchi, G.: Screening parameters in the Pasture Simulation model using the Morris method, Ecol. Modell., 266, 42-57, doi:10.1016/j.ecolmodel.2013.07.005, 2013.
 - Boyle, D. P., Gupta, H. V., and Sorooshian, S.: Toward improved calibration of hydrologic models: Combining the strengths of manual and automatic methods, Water Resour. Res., 36, 3663-3674, doi:10.1029/2000wr900207, 2000.

- Cai, Y., Esaki, T., Liu, S., and Mitani, Y.: Effect of Substitute Water Projects on Tempo-Spatial Distribution of Groundwater Withdrawals in Chikugo-Saga Plain, Japan, Water Resour. Manage., 28, 4645-4663, doi:10.1007/s11269-014-0764-2, 2014.
- Caviedes-Voullième, D., García-Navarro, P., and Murillo, J.: Influence of mesh structure on 2D full shallow water equations and SCS Curve Number simulation of rainfall/runoff events, J. Hydrol., 448-449, 39-59, doi:10.1016/j.jhydrol.2012.04.006, 2012.
- Chandwani, V., Vyas, S. K., Agrawal, V., and Sharma, G.: Soft Computing Approach for Rainfall-runoff Modelling: A Review, Aquat. P., 4, 1054-1061, doi:10.1016/j.aqpro.2015.02.133, 2015.
- Cheng, C.-T., Zhao, M.-Y., Chau, K. W., and Wu, X.-Y.: Using genetic algorithm and TOPSIS for Xinanjiang model calibration with a single procedure, J. Hydrol., 316, 129-140, doi:10.1016/j.jhydrol.2005.04.022, 2006.
- 20 Chu, W., Gao, X., and Sorooshian, S.: Improving the shuffled complex evolution scheme for optimization of complex nonlinear hydrological systems: Application to the calibration of the Sacramento soil-moisture accounting model, Water Resour. Res., 46, W09530, doi:10.1029/2010wr009224, 2010.
 - Cooper, V. A., Nguyen, V. T. V., and Nicell, J. A.: Evaluation of global optimization methods for conceptual rainfall-runoff model calibration, Water Sci. Technol., 36, 53-60, doi:10.1016/S0273-1223(97)00461-7, 1997.
- 25 Cullmann, J., Krausse, T., and Saile, P.: Parameterising hydrological models Comparing optimisation and robust parameter estimation, J. Hydrol., 404, 323-331, doi:10.1016/j.jhydrol.2011.05.003, 2011.
 - Duan, Q., Sorooshian, S., and Gupta, V. K.: Optimal use of the SCE-UA global optimization method for calibrating watershed models, J. Hydrol., 158, 265-284, doi:10.1016/0022-1694(94)90057-4, 1994.
- Esmaeili, S., Thomson, N. R., Tolson, B. A., Zebarth, B. J., Kuchta, S. H., and Neilsen, D.: Quantitative global sensitivity analysis of the RZWQM to warrant a robust and effective calibration, J. Hydrol., 511, 567-579, doi:10.1016/j.jhydrol.2014.01.051, 2014.
 - Freeze, R. A., and Harlan, R. L.: Blueprint for a physically-based, digitally-simulated hydrologic response model, J. Hydrol., 9, 237-258, doi:10.1016/0022-1694(69)90020-1, 1969.

- Getirana, A. C. V.: Integrating spatial altimetry data into the automatic calibration of hydrological models, J. Hydrol., 387, 244-255, doi:10.1016/j.jhydrol.2010.04.013, 2010.
- Graziani, F., Tong, C., and Graziani, F.: A Practical Global Sensitivity Analysis Methodology for Multi-Physics Applications, in, edited by: Graziani, F., Lect. Notes Comput. Sci. Eng., 277-299, doi:10.1029/WR021i004p00473, 2008.
- 5 Gupta, V. K., and Sorooshian, S.: The Automatic Calibration of Conceptual Catchment Models Using Derivative-Based Optimization Algorithms, Water Resour. Res., 21, 473-485, doi:10.1029/WR021i004p00473, 1985.
 - Hao, F., Sun, M., Geng, X., Huang, W., and Ouyang, W.: Coupling the Xinanjiang model with geomorphologic instantaneous unit hydrograph for flood forecasting in northeast China, Int. Soil Water Conserv. Res., 3, 66-76, doi:10.1016/j.iswcr.2015.03.004, 2015.
- Huang, X., Liao, W., Lei, X., Jia, Y., Wang, Y., Wang, X., Jiang, Y., and Wang, H.: Parameter optimization of distributed hydrological model with a modified dynamically dimensioned search algorithm, Environ. Modell. Softw., 52, 98-110, doi:10.1016/j.envsoft.2013.09.028, 2014.
 - Jha, M. K., Kumar, A., Nanda, G., and Bhatt, G.: Evaluation of traditional and nontraditional optimization techniques for determining well parameters from step-drawdown test data, J. Hydrol. Eng., 11, 617-630, doi:10.1061/(asce)1084-0699(2006)11:6(617), 2006.
 - Jiang, Y., Liu, C., Huang, C., and Wu, X.: Improved particle swarm algorithm for hydrological parameter optimization, Appl. Math. Comput., 217, 3207-3215, doi:10.1016/j.amc.2010.08.053, 2010.
 - Jiang, Y., Liu, C., Li, X., Liu, L., and Wang, H.: Rainfall-runoff modeling, parameter estimation and sensitivity analysis in a semiarid catchment, Environ. Modell. Softw., 67, 72-88, doi:10.1016/j.envsoft.2015.01.008, 2015.
- Jin, X., Xu, C.-Y., Zhang, Q., and Singh, V. P.: Parameter and modeling uncertainty simulated by GLUE and a formal Bayesian method for a conceptual hydrological model, J. Hydrol., 383, 147-155, doi:10.1016/j.jhydrol.2009.12.028, 2010.
 - Ju, Q., Yu, Z., Hao, Z., Ou, G., Zhao, J., and Liu, D.: Division-based rainfall-runoff simulations with BP neural networks and Xinanjiang model, Neurocomputing, 72, 2873-2883, doi:10.1016/j.neucom.2008.12.032, 2009.
 - Kennedy, J.: The particle swarm: social adaptation of knowledge, IEEE Int. Conf., 1997.

- 25 Kucherenko, S., Tarantola, S., and Annoni, P.: Estimation of global sensitivity indices for models with dependent variables, Comput. Phys. Commun., 183, 937-946, doi:10.1016/j.cpc.2011.12.020, 2012.
 - Lü, H., Hou, T., Horton, R., Zhu, Y., Chen, X., Jia, Y., Wang, W., and Fu, X.: The streamflow estimation using the Xinanjiang rainfall runoff model and dual state-parameter estimation method, J. Hydrol., 480, 102-114, doi:10.1016/j.jhydrol.2012.12.011, 2013.
- Li, L., Xia, J., Xu, C., and Singh, V. P.: Evaluation of the subjective factors of the GLUE method and comparison with the formal Bayesian method in uncertainty assessment of hydrological models, J. Hydrol., 390, 210-221, doi:10.1016/j.jhydrol.2010.06.044, 2010.
 - Lin, K., Lv, F., Chen, L., Singh, V. P., Zhang, Q., and Chen, X.: Xinanjiang model combined with Curve Number to simulate the effect of land use change on environmental flow, J. Hydrol., 519, 3142-3152, doi:10.1016/j.jhydrol.2014.10.049, 2014.

- Madsen, H.: Automatic calibration of a conceptual rainfall–runoff model using multiple objectives, J. Hydrol., 235, doi:276-288, 10.1016/s0022-1694(00)00279-1, 2000.
- Nelder, J. A., and Mead, R.: A simplex method for function minimization, Comput. J., 4, 308-313, 1965.

20

- Papathanasiou, C., Makropoulos, C., and Mimikou, M.: Hydrological modelling for flood forecasting: Calibrating the post-fire initial conditions, J. Hydrol., 529, 1838-1850, doi:10.1016/j.jhydrol.2015.07.038, 2015.
- Rosenbrock, H.: An automatic method for finding the greatest or least value of a function, Comput. J., 3, 175-184, 1960.
- Sedki, A., Ouazar, D., and El Mazoudi, E.: Evolving neural network using real coded genetic algorithm for daily rainfall–runoff forecasting, Expert Syst. Appl., 36, 4523-4527, doi:10.1016/j.eswa.2008.05.024, 2009.
- Sherman, L. K.: Stream Flow from Rainfall by the Unit-Graph Method, Eng. News Rec., 108, 501-505, 1932.
- 10 Shin, M., Guillaume, J. H. A., Croke, B. F. W., and Jakeman, A. J.: Addressing ten questions about conceptual rainfall—runoff models with global sensitivity analyses in R, J. Hydrol., 503, 135-152, doi:10.1016/j.jhydrol.2013.08.047, 2013.
 - Song, X., Kong, F., Zhan, C., Han, J., and Zhang, X.: Parameter identification and global sensitivity analysis of Xin'anjiang model using meta-modeling approach, Water Sci. Eng., 6, 1-17, doi:10.3882/j.issn.1674-2370.2013.01.001, 2013.
- Song, X., Zhang, J., Zhan, C., Xuan, Y., Ye, M., and Xu, C.: Global sensitivity analysis in hydrological modeling: Review of concepts, methods, theoretical framework, and applications, J. Hydrol., 523, 739-757, doi:10.1016/j.jhydrol.2015.02.013, 2015.
 - Tefaruk, H., Practical computation of gamma frequency factors. Hydrol. Sci. J., 36, 599-610, doi:10.1080/02626669-109492546
 - Van Griensven, A., Meixner, T., Grunwald, S., Bishop, T., Diluzio, M., and Srinivasan, R.: A global sensitivity analysis tool for the parameters of multi-variable catchment models, J. Hydrol., 324, 10-23, doi:10.1016/j.jhydrol.2005.09.008, 2006.
 - Vincendon, B., Ducrocq, V., Saulnier, G.-M., Bouilloud, L., Chancibault, K., Habets, F., and Noilhan, J.: Benefit of coupling the ISBA land surface model with a TOPMODEL hydrological model version dedicated to Mediterranean flash-floods, J. Hydrol., 394, 256-266, doi:10.1016/j.jhydrol.2010.04.012, 2010.
 - Viola, F., Noto, L. V., Cannarozzo, M., and La Loggia, G.: Daily streamflow prediction with uncertainty in ephemeral catchments using the GLUE methodology, Phys. Chem. Earth, 34, 701-706, doi:10.1016/j.pce.2009.06.006, 2009.
 - Wang, J., Li, X., Lu, L., and Fang, F.: Parameter sensitivity analysis of crop growth models based on the extended Fourier Amplitude Sensitivity Test method, Environ. Modell. Softw., 48, 171-182, doi:10.1016/j.envsoft.2013.06.007, 2013.
 - Wang, Q. J.: The Genetic Algorithm and Its Application to Calibrating Conceptual Rainfall-Runoff Models, Water Resour. Res., 27, 2467-2471, doi:10.1029/91wr01305, 1991.
- Wang, Q. J.: Using genetic algorithms to optimise model parameters, Environ. Modell. Softw., 12, 27-34, doi:10.1016/s1364-8152(96)00030-8, 1997.
 - Xie, Y., Xia, L. V., Liu, R., Mao, L., and Liu, X., Research on port ecological suitability evaluation index system and evaluation model. Front. Struct. Civ. Eng., 9, 65-70, doi:10.1007/s11709-014-0258-6, 2015.

- Zambrano-Bigiarini, M., and Rojas, R.: A model-independent Particle Swarm Optimisation software for model calibration, Environ. Modell. Softw., 43, 5-25, doi:10.1016/j.envsoft.2013.01.004, 2013.
- Zanon, F., Borga, M., Zoccatelli, D., Marchi, L., Gaume, E., Bonnifait, L., and Delrieu, G.: Hydrological analysis of a flash flood across a climatic and geologic gradient: The September 18, 2007 event in Western Slovenia, J. Hydrol., 394, 182-197, doi:10.1016/j.jhydrol.2010.08.020, 2010.

- Zhang, S., and Shi, J.: A Microwave Wetland Surface Emissivity Calibration Scheme Using SCE-UA Algorithm and AMSR-E Brightness Temperature Data, P. Environ. Sci., 10, 2731-2739, doi:10.1016/j.proenv.2011.09.424, 2011.
- Zhang, X., Srinivasan, R., Zhao, K., and Liew, M. V.: Evaluation of global optimization algorithms for parameter calibration of a computationally intensive hydrologic model, Hydrol. Process., 23, 430-441, doi:10.1002/hyp.7152 %/ John Wiley & Sons, Ltd., 2009.
- Zhao, R. J., Zhang, Y. L., Fang, L. R., Liu, X. R., and Zhang, Q. S.: The Xinanjiang model, in Hydrological forecasting, P. Oxford Sym. IAHS, 129, 351–356, doi:10.1007/s11269-012-0244-5, 1980.

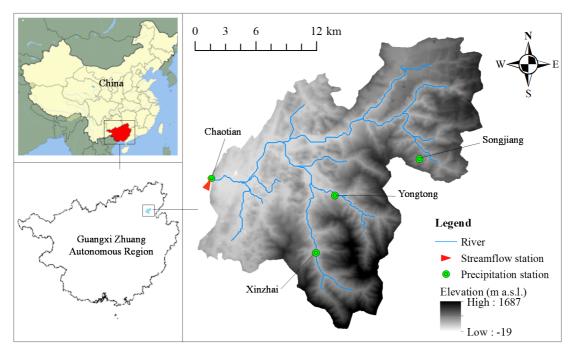


Fig. 1. Location of the study area

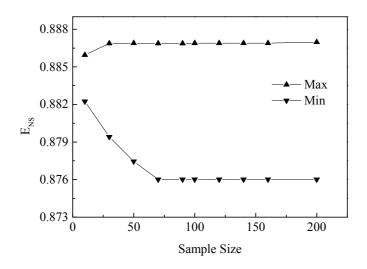


Fig. 2. Variation curves of maximum and minimum $E_{\rm NS}$ with sample sizes

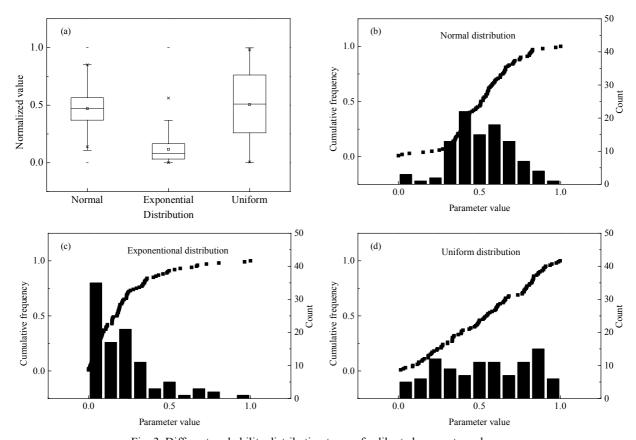


Fig. 3. Different probability distribution types of calibrated parameter values

(a) Box-plot charts of normal, exponential and uniform distribution;

Cumulative frequency curve and histogram for normal (b), exponential (c) and uniform (d) distributions

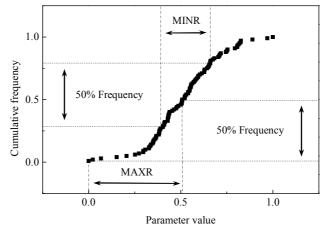


Fig. 4. Selection of minimum and maximum range (MINR and MAXR) with a cumulative frequency of 50%

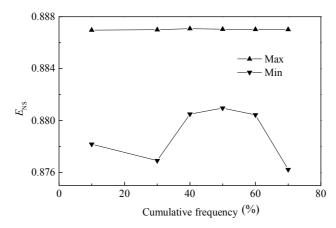


Fig. 5. Variation curves of maximum and minimum E_{NS} of a single parameter with cumulative frequency values

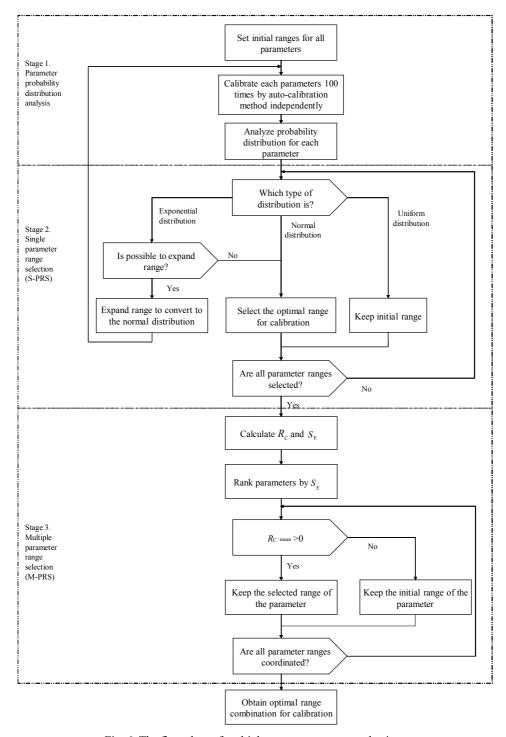


Fig. 6. The flow chart of multiple parameters ranges selections

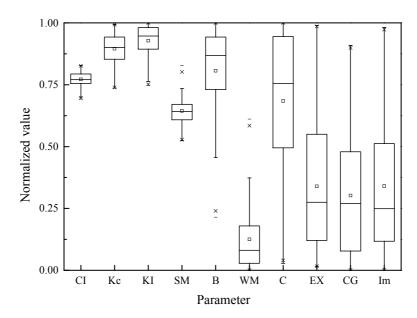


Fig. 7. The box-plot chart of normalized calibrated values for parameters of Xinanjiang model

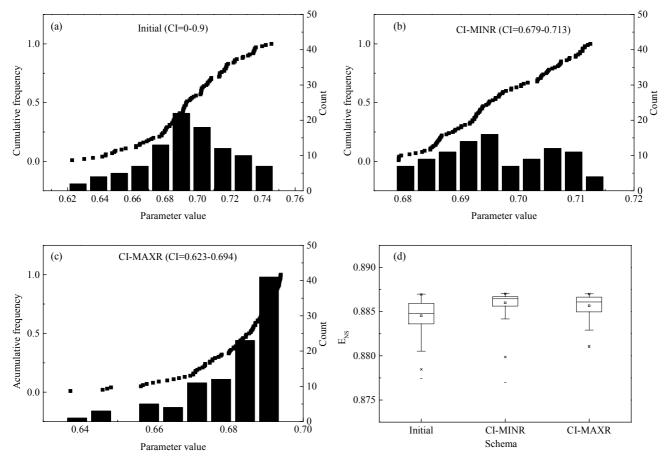


Fig. 8. Results of range selection of parameter CI Probability distribution of parameter values for schema initial range (a), CI-MINR (b) and CI-MAXR (c); (d) Box-plot chart of $E_{\rm NS}$ for three schemas

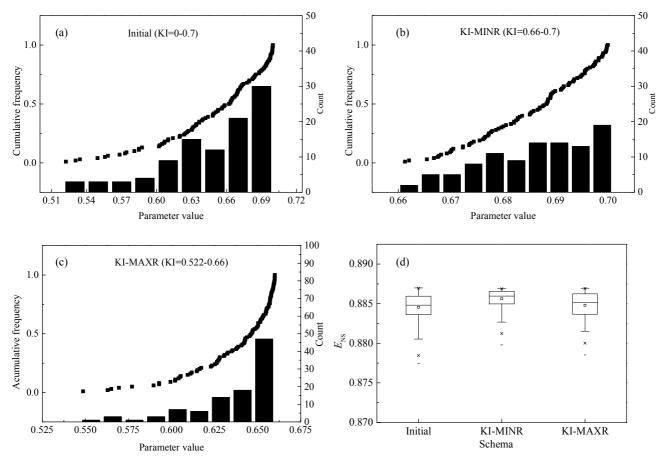


Fig. 9. Results of range selection of parameter KI Probability distribution of parameter values for schema initial range (a), KI-MINR (b) and KI-MAXR (c); (d) Box-plot chart of $E_{\rm NS}$ for three schemas

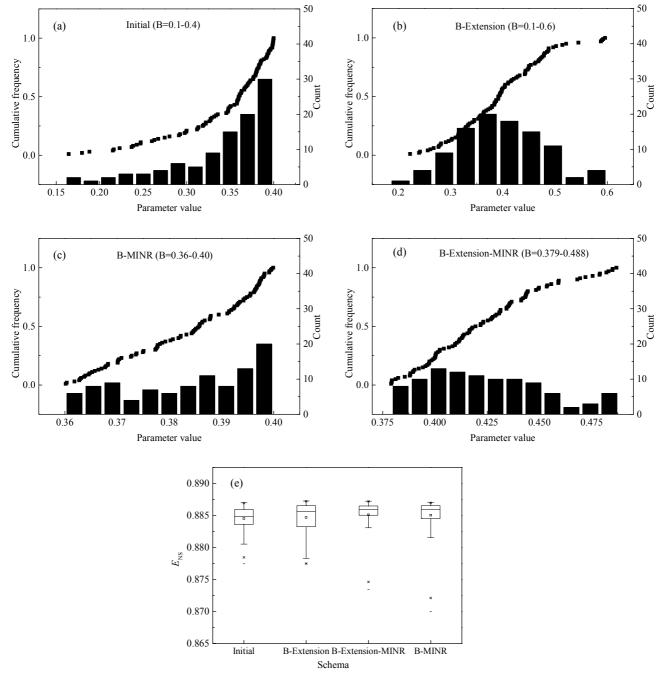


Fig. 10. Results of range selection of parameter B Probability distribution for schema initial range (a), B–Extension (b), B–MINR (c) and B–Extension–MINR (d); (e) Box–plot chart of $E_{\rm NS}$ for four schemas

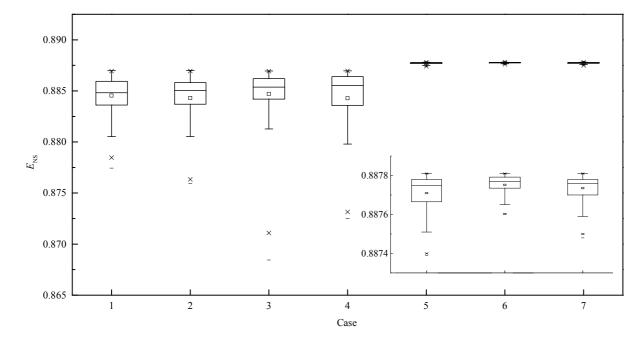


Fig. 11. The box–plot chart of $E_{\rm NS}$ for different cases

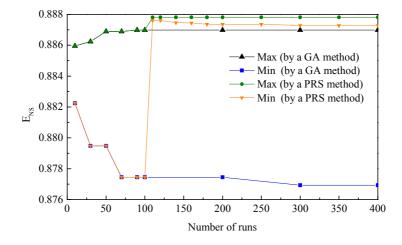


Figure 12. The variation curves of maximum and minimum $E_{\rm NS}$ with number of runs by using a GA method and a proposed PRS method

Table 1. Metro-hydrological statistical data of the study area

Year	Q_{Max} (m ³ /s)	$Q_{\rm Min}~({ m m}^3/{ m s})$	Q_{Avg} (m ³ /s)	P _{Max} (mm/d)
1996	719	0.76	14.38	235
1997	308	0.76	14.32	155
1998	369	0.66	13.67	157
1999	282	0.53	12.81	144
2000	339	1.14	11.37	107

 Q_{Max} , Q_{Min} and Q_{Avg} mean the maximum, minimum and average value of daily streamflow, respectively, and P_{Max} means the maximum value of daily precipitation.

Table 2. Parameters of Xinanjiang model

Parameter	Definition	Units
CI	Recession constants of the lower interflow storage	dimensionless
Kc	Ratio of potential evapotranspiration to pan evaporation	dimensionless
KI	Outflow coefficients of the free water storage to interflow	dimensionless
SM	Areal mean free water capacity of the surface soil layer, which represents the maximum possible deficit of free water storage	mm
В	Exponential parameter with a single parabolic curve, which represents the non-uniformity of the spatial	dimensionless
WM	Averaged soil moisture storage capacity of the whole layer	mm
C	Coefficient of the deep layer, that depends on the proportion of the basin area covered by vegetation with deep roots	dimensionless
EX	Exponent of the free water capacity curve influencing the development of the saturated area	dimensionless
CG	Recession constants of the groundwater storage relationships	dimensionless
KG*	Outflow coefficients of the free water storage to groundwater relationships	dimensionless
Im	Percentage of impervious and saturated areas in the catchment	dimensionless

^{*} The value of KG is calculated by the function 0.7-KI

Table 3. Range changes and K-S tests (α =0.2) of parameters in schema Initial

				Δ_{Max}^{***}				
Parameter	Initial range	Calibrated range*	Ratio** (%)	Normal	Expoential	Uniform distribution		
				distribution	distribution			
CI	0-0.9	0.630-0.745	12.78	0.062 (pass)	0.328 (fail)	0.115 (fail)		
Kc	0-1.1	0.81-1.09	25.45	0.076 (pass)	0.305 (fail)	0.089 (pass)		
KI	0-0.7	0.534-0.7	23.71	0.128 (fail)	0.076 (pass)	0.173 (fail)		
SM	10-50	31–39.4	21.00	0.060 (pass)	0.304 (fail)	0.110 (fail)		
В	0.1-0.4	0.238-0.4	54.00	0.180 (fail)	0.062 (pass)	0.203 (fail)		
WM	120-200	120-150	37.50	0.181 (fail)	0.072 (pass)	0.231 (fail)		
C	0.1-0.2	0.1-0.2	100.00	0.163 (fail)	0.082 (pass)	0.217 (fail)		
EX	1.0-1.5	1.0-1.5	100.00	0.118 (fail)	0.079 (pass)	0.135 (fail)		
CG	0.950-0.998	0.950-0.994	91.67	0.123 (fail)	0.102 (pass)	0.139 (fail)		
Im	0.01-0.04	0.01-0.04	100.00	0.134 (fail)	0.076 (pass)	0.148 (fail)		

^{*} the calibrated parameter range except the extreme outlier

^{**} the ratio is calculated by dividing the length of the range derived from 100 GA calibration runs by the initial range length

^{5 ***} the Δ_{Max} is calculated by using the normalnized parameter values

Table 4. The indexed R_C and S_E of parameters when the optimal range of each parameter is used for calibration

Para	meter*	CI	Kc	KI	SM	В	WM	С	EX	CG	Im
Optimal range of		0.679-	0.95-	0.66-	35–39	0.379-	105-	0.175	1-1.118	0.95-	0.01-
a single parameter		0.713	1.05	0.7		0.488	110	-0.2		0.966	0.0245
	CI	1.000	0.334	0.371	0.462	0.322	0.113	0.105	0.115	-0.128	0.272
	Kc	0.689	1.000	0.467	0.429	0.504	0.503	0.389	0.102	0.284	0.150
	KI	0.778	0.315	1.000	0.445	0.574	0.268	0.456	0.328	0.060	0.258
	SM	0.508	-0.199	0.422	1.000	-0.089	0.009	-0.063	0.383	0.218	-0.032
R_{C}	В	0.914	0.560	0.698	-0.017	1.000	0.972	-0.175	0.007	-0.319	-0.722
KC	WM	0.575	0.311	0.439	0.553	0.325	1.000	0.229	0.360	-0.069	-0.235
	С	0.208	0.273	0.083	0.151	0.277	0.335	1.000	0.077	0.200	0.210
	EX	0.054	0.047	-0.011	0.018	0.371	0.045	0.009	1.000	-0.021	-0.025
	CG	0.221	0.246	-0.135	0.022	0.010	0.198	-0.034	-0.009	1.000	-0.112
	Im	0.238	0.073	-0.025	0.045	0.031	0.030	-0.026	-0.020	0.001	1.000
Mear	n of R _C	0.465	0.218	0.257	0.234	0.258	0.275	0.099	0.149	0.025	-0.026
$S_{\rm E}$	(%)	54.7	47.9	36.6	41.7	48.1	39.9	10.8	14.7	21.9	0.3

^{*} The parameter represents parameter X in Eq. 2.

Table 5. Parameter ranges setting for different cases

Case _		Range setting of parameter									
	CI	Kc	KI	SM	В	WM	С	EX	CG	Im	
1	I	I	I	I	I	I	I	I	I	I	
2	I	I	I	I	I	I	I	I	I	O	
3	I	I	I	I	I	I	I	O	I	I	
4	I	I	I	I	I	I	I	O	I	O	
5	O	O	O	O	O	O	O	O	I	I	
6	O	O	O	O	O	O	O	O	O	I	
7	O	O	O	O	O	O	O	O	O	O	

The symbol 'I' represents the initial range of the parameter in Table 3, and 'O' the optimal range of the parameter in Table 4.