Dear Prof. Schaefli,

Thank you very much for your comments again.

We have made the following corrections for this corrected copy:

(1) We shortened the figure captions for not repeating information.

However, we still repeat the explanations of some variables to make each figure relatively independent.

(2) Thank you for suggesting us a good paper. We cited it at Line 396 when discuss the multiple testing problem.

(3) We have made all variable names in italic. These changes have covered the whole manuscript including text, equations, figures, figure captions, tables, and supplement.

(4) We have figured out the latex problem at Line 392 (previous Line 393).

(5) In the S4 of supplement, we provide the names of all Matlab codes included in the package provided by A. Grinsted which can be available from http://www.glaciology.net/wavelet-coherence. Please see at Lines 588-590 which reads as " In the package provided by A. Grinsted, Matlab codes included are *anglemean.m*, *ar1.m*, *ar1nv.m*, *boxpdf.m*, *formatts.m*, *normalizepdf.m*, *phaseplot.m*, *smoothwavelet.m*, *wt.m*, *wtc.m*, *wtcdemo.m*, *wtcsignif.m*, *xwt.m*. ". By this way, readers can find what they need if the website will be updated.

(6) In the acknowledgement part, we added two funding supporters. So "The project was partially funded by the Natural Sciences and Engineering Research Council of Canada (NSERC) and Agriculture Development Fund of Saskatchewan." was changed to " The project was funded by the National Natural Science Foundation of China (41371233), the Natural Sciences and Engineering Research Council of Canada (NSERC), Agriculture Development Fund of Saskatchewan, and the New Zealand Institute for Plant & Food Research under the Land Use Change and Intensification programme"

(7) In the S2-S4 of the supplement, we suggest readers cite this work if they used our codes for publication. We leave the full citation of this manuscript and website of the supplement in red to be determined at Lines 101, 104, 394, 397, 596, and 602. Would you please be able to figure this out for me before publication? Or would you please guide me how to figure it out by ourselves?

The marked-up manuscript is attached in the following.

We hope this manuscript have reached the level for publication.

Thank you so much again.

Sincerely,

Wei Hu and Bing Si

- 1 Technical Note: Multiple wavelet coherence for untangling scale-specific
- 2 and localized multivariate relationships in geosciences
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## 10 Abstract

The scale-specific and localized bivariate relationships in geosciences can be 11 12 revealed using bivariate wavelet coherence. The objective of this study was to develop 13 a multiple wavelet coherence method for examining scale-specific and localized multivariate relationships. Stationary and non-stationary artificial datasets, generated 14 with the response variable as the summation of five predictor variables (cosine waves) 15 with different scales, were used to test the new method. Comparisons were also 16 17 conducted using existing multivariate methods, including multiple spectral coherence and multivariate empirical mode decomposition (MEMD). Results show that multiple 18 spectral coherence is unable to identify localized multivariate relationships, and 19 underestimates the scale-specific multivariate relationships for non-stationary 20 processes. The MEMD method was able to separate all variables into components at 21

the same set of scales, revealing scale-specific relationships when combined with 22 multiple correlation coefficients, but has the same weakness as multiple spectral 23 24 coherence. However, multiple wavelet coherences are able to identify scale-specific and localized multivariate relationships, as they are close to 1 at multiple scales and 25 26 locations corresponding to those of predictor variables. Therefore, multiple wavelet coherence outperforms other common multivariate methods. Multiple wavelet 27 coherence was applied to a real dataset and revealed the optimal combination of 28 29 factors for explaining temporal variation of free water evaporation at Changwu site in China at multiple scale-location domains. Matlab codes for multiple wavelet 30 coherence were developed and are provided in the supplement. 31

#### 32 **1. Introduction**

Geoscience data such as topography, climate, and ocean waves usually present 33 34 cyclic patterns, with high-frequency (small-scale) processes being superimposed on low-frequency (large-scale) processes (Si, 2008). More often than not, geoscience 35 data are transient, consisting of a variety of frequency regimes that may be localized 36 in space or time (Torrence and Compo, 1998; Si and Zeleke, 2005; Graf et al., 2014). 37 The transient characteristics exist widely in non-stationary processes, but also 38 sometimes occur in stationary processes (Feldstein, 2000). The wavelet method is a 39 common tool for detecting multi-scale and localized features of transient processes in 40 41 geosciences. Bivariate wavelet coherency has been widely used for untangling scale-specific and localized relationships for transient processes in areas including 42

geophysics (Lakshmi et al., 2004; Müller et al., 2008), hydrology (Labat et al., 2005;
Das and Mohanty, 2008; Tang and Piechota, 2009; Carey et al., 2013; Graf et al.,
2014), soil science (Si and Zeleke, 2005; Biswas and Si, 2011), meteorology
(Torrence and Compo, 1998), and ecology (Polansky et al., 2010). This method,
however, is limited to two variables. Processes in geosciences are usually complex
and may be affected by more than two environmental factors. A method is needed for
analyzing multivariate (>2 variables) and localized relationships at multiple scales.

Several methods have been used for characterizing multivariate relationships. For 50 example, multiple spectral coherence (MSC) has been used to explore the 51 52 scale-specific relationships between soil saturated hydraulic conductivity  $(K_s)$  and multiple soil physical properties (Koopmans, 1974; Si, 2008), but requires a stationary 53 data series, which is rare in geosciences. Multivariate empirical mode decomposition 54 (MEMD), a data-driven method, decomposes each variable into different components 55 56 (intrinsic mode functions (IMFs)) with each IMF corresponding to a "common scale" inherent in multiple variables (Rehman and Mandic, 2010). The MEMD method is 57 meritorious due to its ability to deal with both transient and nonlinear systems. The 58 combination of the squared multiple correlation coefficient and MEMD (MCC<sub>memd</sub>) 59 has been used to explore the multivariate control of soil water content and  $K_s$  at 60 61 multiple scales (Hu and Si, 2013; She et al., 2013, 2015; Hu et al., 2014). However, the sum of variances from different components did not typically equal the total 62 variance of the original series, which may produce misleading MCC<sub>memd</sub> results. In 63 addition, multivariate relationships in geosciences are most likely to change with time 64

or space due to the transient nature of the processes involved. However, localized multivariate relationships are not available using any of the existing multivariate methods. Therefore, extending the wavelet coherence from two variables to multiple variables is required .

69 An attempt to extend wavelet coherence from two to three variables has been made by Mihanović et al. (2009). Their method was also applied later in the marine sciences 70 (Ng and Chan, 2012a, b). Limitations arise when using the trivariate wavelet 71 72 coherence: first, only two predictor variables are considered; second, the two predictor variables must be orthogonal. Otherwise, extremely high or low (spurious) 73 coherence (>>1 or <0) may be produced. This spuriousness is inconsistent with the 74 definition of coherence, which may limit the application of this method in geosciences 75 where environmental variables are usually cross-correlated. Therefore, a robust 76 method for calculating MWC, which produces coherence in the closed interval of [0, 77 78 1], is needed.

The objective of this paper is to develop an MWC that applies to cases where there 79 80 are multiple environmental variables, of which may be cross-correlated. This method is first tested with artificial datasets to demonstrate its advantages over existing 81 multivariate methods. The superiority of the new method over others can be assessed 82 by determining whether the known major features of the artificial data are 83 demonstrated by these methods. The new method is then applied to a temporal series 84 of evaporation (E) from free water surface and meteorological factors at Changwu site 85 in Shaanxi, China. 86

# 87 **2. Theory**

Bivariate wavelet coherence can be understood as the traditional correlation coefficient localized in the scale-location domain (Grinsted et al., 2004). Just as correlation coefficients can be extended from two variables to multiple (>2) variables, wavelet coherence between two variables may also be extended to multiple variables. Similar to bivariate wavelet coherence, MWC is based on a series of auto- and cross-wavelet power spectra, at different scales and spatial (or temporal) locations, for the response variable and all predictor variables.

Following Koopman (1974), a matrix representation of the smoothed auto- and cross-wavelet power spectra for multiple predictor variables X ( $X = \{X_1, X_2, ..., X_q\}$ ) can be written as

98 
$$\overline{W}^{X,X}(s,\tau) = \begin{bmatrix} \overline{W}^{X_1,X_1}(s,\tau) & \overline{W}^{X_1,X_2}(s,\tau) & \cdots & \overline{W}^{X_1,X_q}(s,\tau) \\ \overline{W}^{X_2,X_1}(s,\tau) & \overline{W}^{X_2,X_2}(s,\tau) & \cdots & \overline{W}^{X_2,X_q}(s,\tau) \\ \vdots & \vdots & & \vdots \\ \overline{W}^{X_q,X_1}(s,\tau) & \overline{W}^{X_q,X_2}(s,\tau) & \cdots & \overline{W}^{X_q,X_q}(s,\tau) \end{bmatrix},$$
 (1)

99 where  $\overrightarrow{W}^{x_i, x_j}(s, \tau)$  is the smoothed auto-wavelet power spectra (when i=j) or 100 cross-wavelet power spectra (when  $i\neq j$ ) at scale *s* and spatial (or temporal) location 101  $\tau$ , respectively. For the detailed calculation of smoothed auto- and cross-wavelet 102 power spectra, see Supplement, Sect. S1.

103 The matrix of smoothed cross wavelet power spectra between response variable Y

- 104
- and predictor variables  $X_i$  can be defined as

105 
$$\overrightarrow{W}^{Y,X}(s,\tau) = \left[\overrightarrow{W}^{Y,X_1}(s,\tau) \quad \overrightarrow{W}^{Y,X_2}(s,\tau) \quad \cdots \quad \overrightarrow{W}^{Y,X_q}(s,\tau) \right],$$
(2)

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106 where  $\overline{W}^{Y,X_i}(s,\tau)$  is the smoothed cross-wavelet power spectra between Y and  $X_i$  at 107 scale s and spatial (or temporal) location  $\tau$ . 108 The smoothed wavelet power spectrum of response variable Y is  $\overline{W}^{Y,Y}(s,\tau)$ . 109 Following Koopmans (1974), the MWC at scale s and location  $\tau$ ,  $\rho_m^{-2}(s,\tau)$ , can

110 be written as

111 
$$\rho_{m}^{2}(s,\tau) = \frac{\overline{W}^{Y,X}(s,\tau)\overline{W}^{X,X}(s,\tau)^{-1}\overline{W}^{Y,X}(s,\tau)}{\overline{W}^{Y,Y}(s,\tau)}.$$
(3)

112 When only one predictor variable (e.g.,  $X_{a}$ ) is included in *X*, Eq. (3) is the equation 113 for bivariate wavelet coherence,  $\rho_b^2(s,\tau)$ , which can be expressed as (Torrence and 114 Webster, 1999; Grinsted et al., 2004):

115 
$$\rho_b^2(s,\tau) = \frac{\overline{W}^{Y,X_1}(s,\tau)}{\overline{W}^{X_1,X_1}(s,\tau)\overline{W}^{Y,Y}(s,\tau)}.$$
(4)

116 Therefore, bivariate wavelet coherence is consistent with multiple wavelet 117 coherence if only one predictor variable is included. In addition, the wavelet phase 118 between a response variable (Y) and a predictor variable ( $X_{1}$ ) is

119 
$$\phi(s,\tau) = \tan^{-1}\left(\operatorname{Im}\left(W^{Y,X_{1}}(s,\tau)\right) / \operatorname{Re}\left(W^{Y,X_{1}}(s,\tau)\right)\right), \qquad (5)$$

where Im and Re denote the imaginary and real part of  $W^{Y,X_1}(s,\tau)$ , respectively. Note that the phase information between a response variable *Y* and multiple predictor variables *X* cannot be obtained.

Multiple wavelet coherence at the 95% confidence level is calculated using the Monte Carlo method (Grinsted et al., 2004). Surrogate spatial series (i.e., red noise) of all variables are generated with a Monte Carlo simulation based on their first-order autocorrelation coefficient (AR1). The MWC at each scale and location is calculated Formatted: Subscript

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using the simulated spatial series. This is repeated an adequate number of times (e.g.,
1000) (Grinsted et al., 2004). At each scale, MWCs at all locations outside the cones
of influence, from all simulations are ranked in ascending order. The value at the 95th
percentile represents the 95% confidence level for the MWC at that scale. The Matlab
codes and user manual document for calculating MWC and significance level are
provided in the Supplement (Sect. S2–S4).

# 133 **3. Data and analysis**

# 134 **3.1 Artificial data for method test**

135	The method is tested using a stationary and non-stationary artificial dataset,
136	generated following Yan and Gao (2007). The response variable (y for the stationary
137	case and z for the non-stationary case) encompasses five cosine waves ( $y_1$ to $y_5$ for the
138	stationary case and $z_{1}$ to $z_{5}$ for the non-stationary case), with different dimensionless
139	scales (Fig. 1). For the stationary case, $y_1 = \cos(2\pi x/4)$ , $y_2 = \cos(2\pi x/8)$ , $y_3 = \cos(2\pi x/16)$ ,
140	$y_4 = \cos(2\pi x/32)$ , and $y_5 = \cos(2\pi x/64)$ , where $x=0, 1, 2,, 255$ . There is one regular
141	cycle every 4, 8, 16, 32, and 64 locations, representing dimensionless scales of 4, 8,
142	16, 32, and 64 for $y_1$ , $y_2$ , $y_3$ , $y_4$ , and $y_5$ , respectively (Fig. 1a). The regular cycles
143	make each predictor and response series stationary. For the non-stationary case,
144	$z_1 = \cos(500\pi(x/1000)^{0.5}), \qquad z_2 = \cos(250\pi(x/1000)^{0.5}), \qquad z_3 = \cos(125\pi(x/1000)^{0.5}),$
145	$z_4 = \cos(62.5\pi(x/1000)^{0.5})$ , and $z_5 = \cos(31.25\pi(x/1000)^{0.5})$ , where x=0, 1, 2,, 255.
146	The equation containing the square root of the location term results in the gradual
147	change in frequency (scale), with the greatest dimensionless scales of 4, 8, 16, 32, and

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148	64 at the right hand side for $z_1$ , $z_2$ , $z_3$ , $z_4$ , and $z_5$ , respectively (Fig. 1b). The average
149	scales for these predictor variables are 3, 5, 9, 17, and 32, respectively. The
150	location-varying scales make each predictor and response variable non-stationary.
151	For both the stationary and non-stationary series, the variance of the response
152	variable is 2.5. The predictor variables, each with a variance of 0.5, are orthogonal to
153	each other, and contribute equally to the total variance of the response variable. The
154	cosine-like artificial datasets mimic many time series such as seismic signals,
155	turbulence, air temperature, precipitation, hydrologic fluxes, and the El
156	Niño-Southern Oscillation. They also mimic geoscientific spatial series such as ocean
157	waves, seafloor bathymetry, land surface topography, and soil water content along a
158	hummocky landscape. Therefore, they are representative of a geoscience data series
159	and are suitable for testing the new method.
160	Multiple wavelet coherence between the response variable $y$ (or $z$ ) and two ( $y_2$ and

 $y_4$ , or  $z_2$  and  $z_4$ ) or three  $(y_2, y_3, y_4, y_4, y_4, y_5, z_3, y_4, y_4)$  predictor variables were 161 calculated. The advantage of the artificial data is that the known scale- and localized 162 163 features for all variables, and the known relationships between the response and each 164 predictor variable, are exact. By definition, the coherence is 1 at scales corresponding to those of the included predictor variables, and 0 at other scales. 165

To demonstrate the advantages of MWC in dealing with abrupt changes (a type of 166 transient and localized feature), the second half of the original series of  $y_2$  (or  $z_2$ ) or  $y_4$ 167 (or  $z_4$ ) are replaced by 0, and MWC between the response variable and new set of 168 predictor variables is calculated. We anticipate that the coherence changes from 1 to 0 169

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at the location where the new predictor variable becomes 0.

171 Predictor variables may not be as regular as that shown in Fig. 1, and may also be 172 cross-correlated to one another. For these reasons, zero-mean white noises with standard deviations of 0.3, 1, and 4 are added to the predictor variables of  $y_2$  (or  $z_2$ ) 173 and  $y_4$  (or  $z_4$ ). The resulting noised series have correlation coefficients of 0.9, 0.5, and 174 0.1, respectively, with their original predictor variable. Therefore, we will refer to 175 them as weakly, moderately, and highly noised series, respectively. Multiple wavelet 176 177 coherences between the response variable and different predictor variables (original and noised series) are calculated to demonstrate the performance of MWC when 178 noised or correlated predictor variables are involved. Only the non-stationary case 179 will be demonstrated, because the performances of MWC for stationary and 180 non-stationary cases are similar. 181

The MWC is compared to the MSC (Koopmans, 1974; Si, 2008) and MCC<sub>memd</sub> 182 183 (Hu and Si, 2013), which are widely used for spatial or temporal series analysis in 184 different disciplines. The advantages of the new method over these two methods will 185 be demonstrated mainly in terms of relationships between response and predictor variables at various scales of the response variable. The MSC is calculated based on 186 the calculated auto- and cross- power spectra, using an equation similar to Eq. (3). 187 188 The detailed introduction of this method can be found in Si (2008). For the calculation of MCC<sub>mend</sub>, a set of response and predictor variables form a multivariate data series 189 for MEMD. The MEMD is a data driven method and has the ability to align "common 190 scales" present within multivariate data. Please refer to Rehman and Mandic (2010) 191

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Hu and Si (2013) for the MEMD analysis, 192 and and the website (http://www.commsp.ee.ic.ac.uk/~mandic/research/emd.htm) for the related Matlab 193 194 codes. The original series of response and predictor variables can be decomposed by the MEMD, into different components (IMFs) with varying scales. For IMFs at the 195 196 same scale, multiple stepwise regressions are conducted between response and predictor variables, and the multiple correlation coefficients for each scale-specific 197 IMF are calculated. 198

#### 199 **3.2 Real data for application**

Daily evaporation (*E*) from free water surfaces in an E601 evaporation pan (pan diameter of 61.8 cm), and other meteorological factors (i.e., relative humidity, mean temperature, sun hours, and wind speed) were collected from January 1, 1979 to December 31, 2013, at Changwu site in Shaanxi, China. The Changwu site is a transition area between semi-arid and subhumid climates, where agricultural productivity is mainly limited by water. Monthly averages of all variables were used in this study, because we are mainly interested in seasonal and inter-annual variability.

#### 207 4. Results and discussion

## 208 4.1 MWC with orthogonal predictor variables

For the stationary data, there are two narrow, horizontal bands (red color) representing an MWC value of around 1, at the respective scales of 8 and 32 for all locations (Fig. 2a). These two bands also correspond to the scales of 8 and 32, respectively, for the two predictor variables. When an additional predictor variable with the scale of 16 is introduced, a wide band appears from 6 to 40, signifying that the MWC equals approximately 1 at all locations, at the scales of 8, 16, and 32. As anticipated, when all five predictor variables with scales ranging from 4 to 64 are included, coherence values of close to 1 are found in the whole scale-location domain (data not shown).

The application of MWC to the non-stationary datasets shows that the scales with 218 significant MWC values gradually increase as distance increases. This increase in the 219 scales is due to the non-stationarity of the variables (Fig. 2b). For example, when 220 221 predictor variables of  $z_2$  and  $z_4$  are included, scales of the two bands corresponding to MWC around 1 increase from 4 to 8 and from 8 to 32, respectively. Furthermore, as 222 expected, for only one predictor variable (stationary and non-stationary), MWC 223 reduces to bivariate wavelet coherence; there is only one band of coherence around 1, 224 225 which corresponds to the scale of that predictor variable (data not shown). Note that the significant MWC values for both stationary and non-stationary cases are not 226 227 exactly 1 at all scales or locations, due to the smoothing effect along both scales and locations. However, the mean MWC values of the significant bands are very high (i.e., 228 0.94–1.00), and the MWC values at the centre of the significant band are 1, which 229 230 corresponds to the exact scale of a predictor variable.

When the point values in the second half of the data series of a predictor variable are replaced by 0, the MWC values in that half of the data series are almost 0 at scales corresponding to that predictor variable (Fig. 3). For the stationary case, when the Formatted: Font: Italic Formatted: Subscript Formatted: Font: Italic Formatted: Subscript point values in the second half of the data series of predictor variable  $\underline{y_2}$  (or  $\underline{y_4}$ ) are replaced by 0, the MWC values are around 1 at the scale of 8 (or 32) in the first half of the transect, and 0 in the second half (Fig. 3a). Similar results are also found for the non-stationary case (Fig. 3b). This is expected because the constant series of 0 is not correlated to the response variables at any scale. Much like bivariate wavelet coherence, the MWC method is able to detect abrupt changes in the data series, and has the advantages of dealing with localized multivariate relationships.

241 4.2 MWC with noised and correlated predictor variables

242	When $z_2$ and a noised series derived from $z_2$ are included as predictor variables,
243	there is only one band of coherence close to 1 at scales corresponding to $z_2$ ,
244	irrespective of the correlation between $z_2$ and a noised series of $z_2$ (Fig. 4a). When $z_2$
245	and a noised series of $z_4$ are included as predictor variables, the coherence depends on
246	the degree of the noise (Fig. 4b). For weakly noised series, there are two bands of
247	coherence of around 1, corresponding to the scales of $z_2$ and $z_4$ , respectively. The
248	percentage area of significant coherence (PASC) is 23%, which equals that of when $z_2$
249	and $z_{4}$ are included. With the increasing magnitude of noise, the coherence and
250	corresponding PASC at the scales corresponding to $z_4$ decrease. When $z_2$ and a
251	strongly noised series of $z_4$ are considered, the band of coherence around 1, at scales
252	corresponding to $Z_4$ , disappears.
253	The inclusion of a third noised $z_4$ variable substantially increases the area with high
254	coherence (in red) as compared to the case when only $z_2$ and $z_4$ are included (Fig. 4c).

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This indicates that MWC will increase as the number of predictor variables increases, 255 with the highest coherence less or equal to 1, irrespective of the number of predictor 256 257 variables. However, the area of significant coherence may not necessarily increase because of the simultaneously increased statistical significance threshold (Ng and 258 259 Chan, 2012a). In fact, the PASC values for three predictor variables (19–20%) are lower than those of the two predictor variables (23%). This indicates that, in this case, 260 261 two predictor variables are better than three in terms of explaining the variations of 262 the response variable. This occurs because the variance of the response variable that is explained by the noised variable is already accounted for by other variables. Therefore, 263 264 only an additional variable that can independently explain a fair amount of variance could contribute significantly to explaining variations of a response variable (Fig. 4b). 265 This may also explain why there is only one band of coherence around 1 at scales 266 corresponding to  $z_2$ , when  $z_2$  and a noised series of  $z_2$  are included (Fig. 4a). This 267 268 information is helpful in choosing predictor variables for developing scale-specific predictions, especially when predictor variables are correlated. 269

# 270 **4.3 Comparison with other multivariate methods**

271 4.3.1 MSC

The MSC as a function of scale is shown in Fig. 5a. For the stationary case, when  $y_2$  and  $y_4$  are included as predictor variables, there are two plateaus centered at the scales of 8 and 28, representing a coherence of 1. As expected, when an additional predictor variable  $y_3$  is added, the corresponding scale of 16 also shows coherence of

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1. The MSC produces similar scale-specific relationships, as MWC does for a 276 stationary dataset, with exception given to the centered scale (i.e., 28) with a 277 278 coherence of 1. Here, the scale with a unity MSC deviates from the expected value 279 (i.e., 32) for predictor variable  $y_4$ . For the non-stationary case, however, the MSC is 280 much lower than 1 for the predictor variables of  $z_2$  and  $z_4$ ; an MSC of 1 is present only at the scale of 8 when an additional predictor variable z3 is added. Obviously, the 281 MSC underestimates the multivariate relationships, and is not suitable for 282 non-stationary processes (Si, 2008) due to its inability to deal with localized features. 283 The MSC at a specific scale provides the average of multivariate relationships, across 284 285 all locations. Due to the change in scale of a predictor variable with location for the non-stationary case, the MSC deviates greatly from 1. 286

The MSC decreases at scales when the second half of the included predictor 287 variable series are replaced by 0 for both the stationary and non-stationary series (Fig. 288 289 5b). For example, when the second half of the  $y_4$  series in the stationary case are replaced by 0, the MSC at scales of around 32 decreases from 1 to 0.52. Although the 290 291 MSC, throughout the second half of the series, can detect the decrease of coherence at the scales corresponding to the 0 values, the exact locations for the decrease cannot be 292 293 identified. In fact, the coherence decreases only in the second half of the series, and 294 does not change in the first half of the series. The location for the decrease can be easily identified by the MWC, but not by MSC. This further demonstrates the 295 inability of the MSC to deal with localized features. 296

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#### 297 4.3.2 MCC<sub>memd</sub>

298 Five intrinsic mode functions (IMFs) with non-negligible variance, are obtained for multivariate data series. While the obtained scales for the response variable y are in 299 300 agreement with the true scales for the stationary case, the obtained scales (i.e., 3, 6, 11, 21, and 43) for the response variable z deviate slightly from the average scales for the 301 non-stationary case. For the response variable, the contribution of IMFs to the total 302 variance generally decreases (20% to 13% for stationary, and 27% to 11% for 303 304 non-stationary) from IMF1 to IMF5. This disagrees with the fact that each scale contributes equally (i.e., 20%) to the total variance. In addition, the sum of variances 305 over all IMFs for each variable is less than 100% (ranging from 84% to 93%), 306 307 indicating that MEMD cannot capture all the variances. For the detailed results of MEMD, see Supplement, Sect. S5. 308

309 The MCC<sub>memd</sub> as a function of scale, is shown in Fig. 6a. For the stationary case, 310 when predictor variables of  $y_2$  and  $y_4$  are included, the MCC<sub>memd</sub> values are 0.98 and 311 0.93, respectively, at scales corresponding to those of  $y_2$  and  $y_4$ . When a predictor 312 variable of y<sub>3</sub> is included, the MCC<sub>memd</sub> values are 1.00, 1.00, and 0.96, respectively, at scales corresponding to those of  $y_2, y_3$ , and  $y_4$ . For the non-stationary, two predictor 313 variable case, the corresponding MCC<sub>mend</sub> values are 0.80 and 0.85. For the 314 315 non-stationary, three predictor variable case, the corresponding MCC<sub>mend</sub> values are 0.95, 0.99, and 0.91, respectively. Therefore, the MCC<sub>memd</sub> can be used to determine 316 317 the scale-specific multivariate relationships. Similar to MSC, however, the MCC<sub>mend</sub> underestimates the multivariate relationships, especially for the non-stationary case 318

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with less predictor variables. On the contrary, the  $MCC_{memd}$  also overestimates the multivariate relationships. For example, when considering only predictor variables corresponding to scales of 8, 16, and 32, the  $MCC_{memd}$  value for the stationary case is 0.47 at the scale of 64. This deviates much from the expected  $MCC_{memd}$  value of 0 (Fig. 6a). The possible underestimation and overestimation by the  $MCC_{memd}$  may come from the decomposition errors inherent in the MEMD algorithm (Rehman and Mandic, 2010).

Similar to MSC, the localized multivariate relationships cannot be obtained from MCC<sub>memd</sub>. This can be better explained by the decrease of MCC<sub>memd</sub> when half of the series of the predictor variables are replaced by 0 (Fig. 6b). Take the stationary case for example, the MCC<sub>memd</sub> values at the scales corresponding to  $y_2$  and  $y_4$  decrease from 0.98 to 0.49, and from 0.93 to 0.62, respectively, when the second half of the  $y_2$ and  $y_4$  series are replaced by 0.

332 As explained above, the MWC has advantages in untangling localized multivariate 333 relationships as compared to the common multivariate methods. It is important to 334 reveal the multivariate relationships which vary with time or space, that are associated 335 with different processes. For example, discharge usually occurs on knolls, while 336 recharge usually occurs in neighboring depressions (Gates et al., 2011). Therefore, the 337 controlling factors of soil water storage may vary with the land element characteristics of a location. Local controls may be more important on knolls, while non-local 338 controls may be more important in depressions (Grayson et al., 1997). In a temporal 339 domain, vegetation transpiration contributes more to the evapotranspiration in the 340

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growing seasons, which may result in the changes of environmental factors explainingtemporal variations of evapotranspiration in different seasons.

343 4.4 Application of the MWC

Each meteorological factor was significantly correlated to *E*, but the dominant factors explaining variations in *E* differed with scale. For example, the relative humidity was the dominating factor at small (2–8 months) and large (>32 months) scales, while temperature was the dominating factor at the medium (8–32 months) scales. Overall, the relative humidity corresponded to the greatest mean MWC (0.62) and PASC value (40%) at multiple scale-location domains. For the detailed relationships between *E* and each factor, see Supplement, Sect. S6.

351 The MWC analysis shows that the combination of relative humidity and mean temperature produced the greatest mean MWC (0.82) and PASC (49%) among all 352 353 two-factor cases. This suggested that relative humidity and mean temperature were the most appropriate factors for explaining variations in E at multiple scale-location 354 355 domains (Fig. 7a). However, adding an additional factor such as sun hours, which was 356 the best among all three-factor cases, increased the average coherence (0.91), but 357 slightly decreased the PASC to 48% (Fig. 7b). This indicated that sun hours was not 358 significantly different from red noise in explaining additional variation in E. Similar 359 results were found when the wind speed was added. This occurs because most areas 360 with significant coherence between E and sun hours or wind speed were a subset of areas with significant coherence between E and relative humidity or mean 361

temperature (see Supplement, Sect. S3). Therefore, relative humidity and mean temperature were adequate for explaining the temporal variation of E at various scales at this site. This was consistent with Li et al. (2012), who indicated that relative humidity and mean temperature were the two main contributors to the temporal change of potential evapotranspiration on the Chinese Loess Plateau.

#### 367 5. Conclusions

Multiple wavelet coherence was developed to determine scale-specific and 368 369 localized multivariate relationships in geosciences. The new method was tested and 370 compared with existing multivariate methods, using an artificial dataset. The new method can be used to determine the proportion of the variance of a response variable 371 372 that is explained by predictor variables, at a specific scale and location (spatially or temporally). As compared with bivariate wavelet coherence, more variation may be 373 explained at multiple scale-location domains by the MWC. Including more variables 374 is only beneficial if the variables are not strongly cross-correlated, and can 375 independently explain a fair amount of variability in a response variable. Therefore, 376 377 the best combinations of variables that explain multivariate, spatial or temporal variability at multiple scales can be determined. This is important for optimizing 378 379 variables to develop scale-specific prediction.

The MSC and  $MCC_{memd}$  can determine multivariate relationships at multiple scales, but localized multivariate relationships are not available. Furthermore, both MSC and MCC<sub>memd</sub> are likely to underestimate the degree of multivariate relationships for non-stationary processes. In addition, the performance of  $MCC_{memd}$  relies on the performance of MEMD, which needs further development. Application of the MWC into the real dataset indicates that the combination of relative humidity and mean temperature are the optimal factors that can be used to explain temporal variations of *E* at the Changwu site in China.

Limitations of the new method also exist. Theoretically, any number of predictor 388 variables can be included in the multiple wavelet analysis. However, the statistical 389 390 significance threshold usually increases with the number of predictor variables (Grinsted et al., 2004; Ng and Chan, 2012a). In addition, the inclusion of too many 391 predictor variables may result in the statistical significance significance threshold at 392 particular wavelet scales (e.g., the lowest and largest scales) to approach unity. This 393 would restrict the availability of statistical information. Furthermore, similar to 394 bivariate wavelet analysis, the new method also suffers from the multiple-testing 395 396 problem (Maraun and Kurths, 2004; Maraun et al., 2007; Schaefli et al., 2007; Schulte et al., 2015; Schulte, 2016). Therefore, a more robust statistical significance testing 397 398 method may be beneficial to the new method.

In summary, multiple wavelet coherence has advantages over existing multivariate methods, and provides an effective vehicle for untangling complex spatial or temporal variability for multiple controlling factors at multiple scales and locations. It may also be used as a data-driven tool for modeling and predicting various processes in the area of geosciences, such as precipitation, drought, soil water dynamics, stream flow, and atmospheric circulation.

# 405 Acknowledgements

406	The Matlab codes for calculating multiple wavelet coherence are developed based on
407	the codes provided by A. Grinsted (http://www.glaciology.net/wavelet-coherence) and,
408	together with the user manual, are available in the Supplement (Sect. S2-S4). The
409	project was partially funded by the National Natural Science Foundation of China
410	(41371233), the Natural Sciences and Engineering Research Council of Canada
411	(NSERC),-and-Agriculture Development Fund of Saskatchewan, and the New
412	Zealand Institute for Plant & Food Research under the Land Use Change and
413	Intensification programme. We thank the two anonymous reviewers for their
414	constructive comments.

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517	Figure 1. (a) Stationary and (b) non-stationary series of response variables (y for	
518	stationary and z for non-stationary case) encompassing five cosine waves ( $y_1$ to $y_5$ for	
519	stationary and $z_1$ to $z_5$ for non-stationary case) with different dimensionless scales.(a)	
520	Stationary and (b) non stationary series of response variables (y for stationary and z	$\leq$
521	for non stationary case) encompassing five cosine waves (y1 to y5 for stationary and	$\leq$
522	<i>z1</i> to <i>z5</i> for non-stationary case) with different dimensionless scales.	$\leq$
523	Figure 2. Multiple wavelet coherence (a) between response variable y and predictor	C
524	variables $y_2$ and $y_4$ ; (b) between response y and predictors $y_2$ , $y_3$ , and $y_4$ ; (c) between	
525	response z and predictors $z_2$ and $z_4$ ; and (d) between response z and predictors $z_2$ , $z_3$ ,	
526	and $z_4$ . The artificial data series (y) encompasses five cosine waves ( $y_1$ , $y_2$ , $y_3$ , $y_4$ , and	
527	$y_5$ ) with different scales for the stationary case, and the artificial data series (z)	
528	encompasses five cosine waves $(z_1, z_2, z_3, z_4, and z_5)$ with different scales for the	
529	non-stationary case. The predictor variables, connected by a hyphen, are shown in the	
530	top right corner of each subplot. Thin solid lines demarcate the cones of influence,	
531	and thick solid lines show the 95% confidence levels. Multiple wavelet coherence (a)	
532	between response variable y and predictor variables y2 and y4; (b) between response y	
533	and predictors $y_2$ , $y_3$ , and $y_4$ ; (c) between response z and predictors $z_2$ and $z_4$ ; and (d)	
534	between response <u>z</u> and predictors <u>z</u> 2, <u>z</u> 3, and <u>z</u> 4. The artificial data series (y)	
535	encompasses five cosine waves $(y1, y2, y3, y4, and y5)$ with different scales for the	
536	stationary case, and the artificial data series (z) encompasses five cosine waves (z1, z2,	$V_{\lambda}$
537	$z^3$ , z4, and z5) with different scales for the non-stationary case. The predictor	4

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variables, connected by a hyphen, are shown in the top right corner of each subplot.	
Thin solid lines demarcate the cones of influence, and thick solid lines show the 95%	
confidence levels.	
<b>Figure 3.</b> <u>Multiple wavelet coherence (a) between y and <math>y_2h_0</math> and <math>y_4</math>; (b) between y</u>	
and $y_2$ and $y_4h_0$ ; (c) between z and $z_2h_0$ and $z_4$ ; and (d) between z and $z_2$ and $z_4h_0$ .	
The variables $y_2h_0$ (or $z_2h_0$ ) and $y_4h_0$ (or $z_4h_0$ ) refer to the new series of $y_2$ (or $z_2$ ) and	
<u><math>y_4</math> (or <math>z_4</math>), in which the second half are replaced by 0. Multiple wavelet coherence (a)</u>	
between y and y2h0 and y4; (b) between y and y2 and y4h0; (c) between z and z2h0	$\leq$
and <u>z</u> 4; and (d) between z and z2 and z4h0. The artificial data series (y)	
encompasses five cosine waves (y1, y2, y3, y4, and y5) with different scales for the	
stationary case and the artificial data series (z) encompasses five cosine waves (z1, z2,	
z3, z4, and z5) with different scales for the non stationary case. The variables <u>y2h0 (or</u>	
z2h0) and y4h0 (or z4h0) refer to the new series of y2 (or z2) and y4 (or z4), in which	
the second half are replaced by 0. The predictor variables, connected by a hyphen, are	
shown in the top right corner of each subplot. Thin solid lines demarcate the cones of	
influence and thick solid lines show the 95% confidence levels.	ľ
Figure 4. <u>Multiple wavelet coherence of an artificial data series (z) encompassing five</u>	
cosine waves $(z_1, z_2, z_3, z_4, and z_5)$ with different scales and (a) $z_2$ and noised $z_2$ , (b)	
$z_2$ and noised $z_4$ , and (c) $z_2$ , $z_4$ , and noised $z_4$ for the non-stationary case. $z_2wn$ ( $z_4wn$ ),	
$z_2mn$ ( $z_4mn$ ), and $z_2sn$ ( $z_4sn$ ) indicate weakly, moderately, and strongly noised $z_2$ ( $z_4$ )	
series, respectively. Weakly, moderately, and strongly noised series are correlated with	
original series, having with correlation coefficients of 0.9, 0.5, and 0.1,	
	<ul> <li>variables, connected by a hyphen, are shown in the top right corner of each subplot.</li> <li>Thin solid lines demarcate the cones of influence, and thick solid lines show the 95% confidence levels.</li> <li>Figure 3. Multiple wavelet coherence (a) between y and y<sub>2</sub>h<sub>0</sub> and y<sub>4</sub>; (b) between y and y<sub>2</sub> and y<sub>4</sub>h<sub>0</sub>; (c) between z and z<sub>2</sub>h<sub>0</sub> and z<sub>4</sub>; and (d) between z and z<sub>2</sub> and z<sub>4</sub>h<sub>0</sub>.</li> <li>The variables y<sub>2</sub>h<sub>0</sub> (or z<sub>2</sub>h<sub>0</sub>) and y<sub>4</sub>h<sub>0</sub> (or z<sub>4</sub>h<sub>0</sub>) refer to the new series of y<sub>2</sub> (or z<sub>4</sub>) in which the second half are replaced by 0. Multiple wavelet coherence (a)</li> <li>between y and y<sub>2</sub>h<sub>0</sub> and y<sub>4</sub>h; (b) between y and y<sub>2</sub> and y<sub>4</sub>h<sub>0</sub>; (c) between z and z<sub>2</sub><sup>2</sup> and z<sub>4</sub><sup>2</sup>h<sub>0</sub></li> <li>and z<sub>4</sub><sup>4</sup>; and (d) between z and z<sub>2</sub><sup>2</sup> and z<sub>4</sub>h<sub>0</sub>. The artificial data series (y)</li> <li>encompasses five cosine waves (y1, y2, y3, y4, and y5) with different scales for the stationary case and the artificial data series (z) encompasses five cosine waves (z1, z2, z<sub>3</sub>, z4, and z5) with different scales for the non stationary case. The variables <u>y2h0</u> (or z4h0) refer to the new series of y<sub>2</sub> (or z<sub>2</sub>) and y<sub>4</sub> (or z4h), in which the second half are replaced by 0. The predictor variables, connected by a hyphen, are shown in the top right corner of each subplot. Thin solid lines demarcate the cones of influence and thick solid lines show the 95% confidence levels.</li> <li>Figure 4. Multiple wavelet coherence of an artificial data series (z) encompassing five (cosine waves (z<sub>1</sub>, z<sub>2</sub>, z<sub>3</sub>, z<sub>4</sub>, and z<sub>5</sub>) with different scales and (a) z<sub>2</sub> and noised z<sub>4</sub>, and (c) z<sub>2</sub>, z<sub>4</sub>, and noised z<sub>4</sub> for the non-stationary case. z<sub>2</sub>wn (z<sub>4</sub>wn), z<sub>1</sub>un noised z<sub>4</sub>, and (c) z<sub>2</sub>, z<sub>4</sub>, and noised z<sub>4</sub> for the non-stationary case. z<sub>2</sub>wn (z<sub>4</sub>wn), (cosine waves (z<sub>1</sub>, z<sub>2</sub>, (z<sub>4</sub>, y<sub>1</sub>) indicate weakly, moderately, and strongly noised series are correlated with original series, having with correlation coefficients of 0.9, 0.5, and 0.1,</li> </ul>

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560	respectivery. wumple wavelet concretice of an artificial data series ( <u>z) encompassing</u>
561	five cosine waves ( <u>z1, z2, z3, z4, and z5) with different scales and (a) z2 and noised z2,</u>
562	(b) <u>z2 and noised z4, and (c) z2, z4, and noised z4 for the non stationary case. The</u>
563	predictor variables are connected by a hyphen and shown in the top right corner of
564	each subplot. <u>z2wn (z4wn), z2mn (z4mn), and z2sn (z4sn) indicate weakly, moderately,</u>
565	and strongly noised <u>z2 (z4)</u> series, respectively. Weakly, moderately, and strongly
566	noised series are correlated with original series, having correlation coefficients of 0.9,
567	0.5, and 0.1, respectively. Thin solid lines demarcate the cones of influence and thick
568	solid lines show the 95% confidence levels.
569	Figure 5. Multiple spectral coherence (MSC) of an artificial data series (y or z)
570	encompassing five cosine waves ( $y_1$ to $y_5$ ; or $z_1$ to $z_5$ ) with different scales and (a)
571	two ( $y_2$ and $y_4$ ; or $z_2$ and $z_4$ ) or three ( $y_2$ , $y_3$ , and $y_4$ ; or $z_2$ , $z_3$ , and $z_4$ ) data series, and
572	(b) two ( $y_2$ and $y_4$ ; or $z_2$ and $z_4$ ) data series when the second half of one data series are
573	replaced by 0. The variables $y_2h_0$ (or $z_2h_0$ ) and $y_4h_0$ (or $z_4h_0$ ) refer to the new series
574	of $y_2$ (or $z_2$ ) and $y_4$ (or $z_4$ ) in which the second half are replaced by 0. Multiple spectral
575	coherence (MSC) of an artificial data series (v or z) encompassing five cosine waves
576	(y1 to y5; or z1 to z5) with different scales and (a) two (y2 and y4; or z2 and z4) or
577	three (y2, y3, and y4; or z2, z3, and z4) data series, and (b) two (y2 and y4; or z2 and ///
578	$z^{4}$ ) data series when the second half of one data series are replaced by 0. The variables
579	$y^{2h0}$ (or $z^{2h0}$ ) and $y^{4h0}$ (or $z^{4h0}$ ) refer to the new series of $y^2$ (or $z^2$ ) and $y^4$ (or $z^4$ )
580	in which the second half are replaced by 0.
581	Figure 6. Multiple correlation coefficient between multivariate empirical mode

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582	decomposition (MCC <sub>memd</sub> ) of an artificial series (y or z) and (a) two ( $y_2$ and $y_4$ ; or $z_2$
583	and $z_4$ ) or three $(y_2, y_3, and y_4; or z_2, z_3, and z_4)$ data series, and (b) two $(y_2 and y_4; or z_4)$
584	$z_2$ and $z_4$ ) data series when the second half of one data series are replaced by 0. The
585	variables $y_2h_0$ (or $z_2h_0$ ) and $y_4h_0$ (or $z_4h_0$ ) refer to the new series of $y_2$ (or $z_2$ ) and $y_4$
586	(or $z_4$ ) in which the second half are replaced by 0 Multiple correlation coefficient
587	between multivariate empirical mode decomposition (MCC <sub>memd</sub> ) of an artificial series
588	$(y \text{ or } z)$ and (a) two $(y^2 \text{ and } y^4; \text{ or } z^2 \text{ and } z^4)$ or three $(y^2, y^3, \text{ and } y^4; \text{ or } z^2, z^3, \text{ and } z^4)$
589	data series, and (b) two (y2 and y4; or z2 and z4) data series when the second half of
590	one data series are replaced by 0. The variables <u>y2h0 (or z2h0) and y4h0 (or z4h0)</u>
591	refer to the new series of $y^2$ (or $z^2$ ) and $y^4$ (or $z^4$ ) in which the second half are
592	replaced by 0.
593	Figure 7. Multiple wavelet coherence between evaporation (E) from water surfaces
594	and meteorological factors ((a) relative humidity and mean temperature and (b)
595	relative humidity, mean temperature, and sun hours) at Changwu site in Shaanxi,
596	China. Thin solid lines demarcate the cones of influence, and thick solid lines show
597	the 95% confidence level.

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Figure 2. Multiple wavelet coherence (a) between response variable y and predictor variables  $y_2$  and  $y_4$ ; (b) between response y and predictors  $y_2$ ,  $y_3$ , and  $y_4$ ; (c) between response z and predictors  $z_2$  and  $z_4$ ; and (d) between response z and predictors  $z_2$ ,  $z_3$ and  $z_4$ . The artificial data series (y) encompasses five cosine waves ( $y_4$ ,  $y_2$ ,  $y_3$ ,  $y_4$ , and  $y_5$ ) with different scales for the stationary case, and the artificial data series (z) encompasses five cosine waves  $(z_1, z_2, z_3, z_4, and z_5)$  with different scales for the non-stationary case. The predictor variables, connected by a hyphen, are shown in the top right corner of each subplot. Thin solid lines demarcate the cones of influence, and thick solid lines show the 95% confidence levels.

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**Figure 3.** Multiple wavelet coherence (a) between  $\underline{y}$  and  $\underline{y_2h_0}$  and  $\underline{y_4}$ ; (b) between  $\underline{y}$  and  $\underline{y_2}$  and  $\underline{y_4h_0}$ ; (c) between  $\underline{z}$  and  $\underline{z_2h_0}$  and  $\underline{z_4}$ ; and (d) between  $\underline{z}$  and  $\underline{z_2}$  and  $\underline{z_4h_0}$ . The artificial data series (y) encompasses five cosine waves (y1, y2, y3, y4, and y5) with different scales for the stationary case and the artificial data series (z) encompasses five cosine waves (z1, z2, z3, z4, and z5) with different scales for the non-stationary case. The variables  $\underline{y_2h_0}$  (or  $\underline{z_2h_0}$ ) and  $\underline{y_4h_0}$  (or  $\underline{z_4h_0}$ ) refer to the new series of  $\underline{y_2}$  (or  $\underline{z_2}$ ) and  $\underline{y_4}$  (or  $\underline{z_4}$ ), in which the second half are replaced by 0. The predictor variables, connected by a hyphen, are shown in the top right corner of each subplot. Thin solid lines demarcate the cones of influence and thick solid lines show the 95% confidence levels.



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**Figure 4.** Multiple wavelet coherence of an artificial data series (*z*) encompassing five cosine waves (*z*<sub>1</sub>, *z*<sub>2</sub>, *z*<sub>3</sub>, *z*<sub>4</sub>, and *z*<sub>5</sub>) with different scales and (a) *z*<sub>2</sub> and noised *z*<sub>2</sub>, (b) *z*<sub>2</sub> and noised *z*<sub>4</sub>, and (c) *z*<sub>2</sub>, *z*<sub>4</sub>, and noised *z*<sub>4</sub> for the non-stationary case. The predictor variables are connected by a hyphen and shown in the top right corner of each subplot. *z*<sub>2</sub>*wn* (*z*<sub>4</sub>*wn*), *z*<sub>2</sub>*mn* (*z*<sub>4</sub>*mn*), and *z*<sub>2</sub>*sn* (*z*<sub>4</sub>*sn*) indicate weakly, moderately, and strongly noised *z*<sub>2</sub> (*z*<sub>4</sub>) series, respectively. Weakly, moderately, and strongly noised series are correlated with original series, having with correlation coefficients of 0.9, 0.5, and 0.1, respectively. Thin solid lines demarcate the cones of influence and thick solid lines show the 95% confidence levels.

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**Figure 5.** Multiple spectral coherence (MSC) of an artificial data series (y or z) encompassing five cosine waves ( $y_4$  to  $y_5$ ; or  $z_4$  to  $z_5$ ) with different scales and (a) two ( $y_2$  and  $y_4$ ; or  $z_2$  and  $z_4$ ) or three ( $y_2$ ,  $y_3$ , and  $y_4$ ; or  $z_2$ ,  $z_3$ , and  $z_4$ ) data series, and (b) two ( $y_2$  and  $y_4$ ; or  $z_2$  and  $z_4$ ) data series when the second half of one data series are replaced by 0. The variables  $y_2h_0$  (or  $z_2h_0$ ) and  $y_4h_0$  (or  $z_4h_0$ ) refer to the new series of  $y_{24}$  (or  $z_2$ ) and  $y_4$  (or  $z_4$ ) in which the second half are replaced by 0.

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**Figure 6.** Multiple correlation coefficient between multivariate empirical mode decomposition (MCC<sub>memd</sub>) of an artificial series (y or z) and (a) two ( $y_2$  and  $y_4$ ; or  $z_2$  and  $z_4$ ) or three ( $y_2$ ,  $y_3$ , and  $y_4$ ; or  $z_2$ ,  $z_3$ , and  $z_4$ ) data series, and (b) two ( $y_2$  and  $y_4$ ; or  $z_2$  and  $z_4$ ) data series when the second half of one data series are replaced by 0.-The variables  $y_2h_0$  (or  $z_2h_0$ ) and  $y_4h_0$  (or  $z_4h_0$ ) refer to the new series of  $y_2$  (or  $z_2$ ) and  $y_4$  (or  $z_4$ ) in which the second half are replaced by 0.

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**Figure 7.** Multiple wavelet coherence between evaporation (*E*) from water surfaces and meteorological factors ((a) relative humidity and mean temperature and, (b) relative humidity, mean temperature, and sun hours) at Changwu site in Shaanxi, China. This solid lines demarcate the cones of influence, and thick solid lines show the 95% confidence level.