



Flood frequency  
analysis of historical  
flood data

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This discussion paper is/has been under review for the journal Hydrology and Earth System Sciences (HESS). Please refer to the corresponding final paper in HESS if available.

# Flood frequency analysis of historical flood data under stationary and non-stationary modelling

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Received: 14 December 2014 – Accepted: 16 December 2014 – Published: 14 January 2015

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Published by Copernicus Publications on behalf of the European Geosciences Union.

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## Abstract

Historical records are an important source of information about extreme and rare floods with a great value to establish a reliable flood return frequency. The use of long historic records for flood frequency analysis brings in the question of flood stationarity, since climatic and land-use conditions can affect the relevance of past flooding as a predictor of future flooding. In this paper, a detailed 400 year flood record from the Tagus River in Aranjuez (Central Spain) was analysed under stationary and non-stationary flood frequency approaches, to assess their implications on hazard studies. Historical flood records in Aranjuez were obtained from documents (Proceedings of the City Council, diaries, chronicles, memoirs, etc.), epigraphic marks, and indirect historical sources and reports. The water levels associated with different floods (derived from descriptions or epigraphic marks) were computed into discharge values using a one-dimensional hydraulic model. Secular variations on flood magnitude and frequency, found to respond to climate and environmental drivers, showed a good correlation between high values of historical flood discharges and a negative mode of the North Atlantic Oscillation index (NAO index). Over the systematic gauge record (1913–2008), an abrupt change on flood magnitude was produced in 1957 due to constructions of three major reservoirs in the Tagus headwaters (Bolarque, Entrepeñas and Buendia) controlling 80 % of the watershed surface draining to Aranjuez. Two different models were used for the flood frequency analysis: (a) a stationary model estimating statistical distributions incorporating imprecise and categorical data based on maximum likelihood estimators; (b) a time-varying model based on “generalized additive models for location, scale and shape” (GAMLSS) modelling, that incorporates external covariates related to climate variability (NAO index) and catchment hydrology factors (in this paper a reservoir index; RI). Flood frequency analysis using documentary data (plus gauged record) improved the estimates of the probabilities of rare floods (return intervals of 100 year and higher). Under non-stationary modelling flood occurrence associated with an exceedance probability of 0.01 (i.e. return period of 100 year) has changed over the last 500 year due

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to decadal and multi-decadal variability of the NAO. Yet, frequency analysis under stationary models was successful on providing an average discharge around which value flood quantiles estimated by non-stationary models fluctuate through time.

## 1 Introduction

5 Throughout Europe, national legislations on flood hazard assessment are based on flood-frequency analysis which estimate discharges associated with different return periods (usually 10, 25, 50, 100 and 500 year). This assessment is also gathered in the European Flood Directive (2007/60/CE) together with the use parameters (see Article 6.3 on flood hazard maps and flood risk maps), which are common practice in 10 several continents and countries such as USA, Australia and several South America countries. The common procedure involves the extrapolation from gauged hydrological registers, documenting 30–40 year records, of observed floods (usually not comprising an extraordinary event) and estimate quantiles for floods. The straightforward handling of these simple statistical methods that makes them so widely used hides important 15 uncertainties as well as scientific and technical problems which have been extensively discussed in the literature (Merz and Blöschl, 2008). This conventional method of addressing flood hazard assessment can be improved by including information of past floods (historical floods and palaeofloods), which should be accomplished using rigorous procedures of data collection and statistical modelling. Documentary records 20 provide a catalogue of the largest flood events that occurred during periods of human settlement, and provide evidence of all other events below or above specific flow stages or thresholds (Brázdil et al., 2006). Long records of historical extreme floods have been applied successfully in hazard analysis together with the more traditional empirical, statistical and deterministic methods to estimate the largest floods (Stedinger and Cohn, 1986; Frances et al., 1994; England et al., 2003). Information on these extreme floods 25 is highly demanded by planners and engineers and yet seldom registered in the observational record due to its short time length (Enzel et al., 1993; Benito et al., 2004).

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A major issue of using long-term flood data in flood frequency analysis is the assumption of stationarity i.e. the idea that “natural systems fluctuate within an unchanging envelope of variability” (Milly et al., 2008). The analysis of long-term historical flood records frequently reveals the occurrence of flood clusters and trends lasting from few decades to centuries and, subsequently, the fact that the frequency of flood magnitude changes over time (Benito et al., 2003). Most of these trends observed in historical flood series are related to decadal climate variability (e.g. influence of the North Atlantic Oscillation or ENSO-El Niño) which may affect the assumption of stationary in conventional flood frequency analysis (Merz et al., 2014). Milly et al. (2008) in their study on the recent impacts of climate change in the hydrological cycle, conclude that the assumption of stationary in hydrology is dead and should no longer serve as a central, default assumption in water-resource risk assessment and planning. Yet, the application of non-stationary concepts by the engineering community has been rather challenging due to the limited number of non-stationary models and scarce long-term flood records to test the performance of such models. Centennial historical flood data provides long-term real data to test and compare stationary and non-stationary models under climate variability. The results of this analysis together with the selection of appropriate covariates explaining the changes in flood series are relevant in climate-related impacts on flood risk assessment. Moreover, the optimal merging of instrumental, historical and palaeo-information is of great importance for understanding flood hazard and the appraisal of flood incidence through multi-decadal climate variability (Kundzewicz et al., 2014).

This paper aims to address stationary and non-stationary flood frequency analysis of historical floods (data censored above threshold), and to test their performance for flood hazard analysis. The specific objectives are: (1) analyse the long-term historical flood variability and identify the major drivers and covariates (climatic or human-induced environmental factors) describing the temporal changes, (2) test the stationary of flood series and, if so, estimate flood distribution of systematic (gauged) and non-systematic (historical) data under stationary assumption, (3) apply non-stationary models to histor-

ical flood data incorporating covariate indexes describing flood producing atmospheric circulation patterns (e.g. North Atlantic Oscillation, NAO) and other major human disturbance parameters (e.g. reservoir construction); and (4) compare results of using stationary and non-stationary models in flood hazards assessment.

## 2 Study area

The Tagus River drains the central Spanish Plateau (Meseta) and flows east–west into the Atlantic Ocean at Lisbon (Portugal). It is the longest river of the Iberian Peninsula (1200 km) and the third largest in catchment area (81 947 km<sup>2</sup>). The studied flood records are located in Aranjuez (Fig. 1a), in the upper part of the catchment (9340 km<sup>2</sup> in drainage surface). In Aranjuez, the Tagus River channel has a meandering pattern on a floodplain 800–1000 m in width. Present day mean annual discharge is 35 m<sup>3</sup> s<sup>-1</sup> with ordinary streamflow completely regulated by the upstream dams of Bolarque and the Entrepeñas-Buendia reservoir system. The natural streamflow is characterised by extreme seasonal and annual variability including severe flash floods with peak discharges more than 30 times the mean. In the Tagus basin, eastern and northeastern tributaries have a mixed hydrological regime from snowmelt and rainwaters from the Iberian and eastern Central Range areas, whereas the southern and northwestern tributaries are dominated by rainwaters. General discharge characteristics are: (1) maximum discharge from February to March, (2) minimum discharge in August, (3) a peak in December, and (4) a discharge reduction in January.

The Tagus River flood regime is influenced by the rainfall associated to the Atlantic fronts that cross the western Iberian Peninsula during low zonal circulation over the Atlantic (at 35–40° N) mostly during winter months (Capel, 1981; Trigo and Palutikof, 2001). The eastern tributaries and the Tagus headwaters show a second flood maximum during autumn related to cold pool cells developing mainly along the Mediterranean coast producing intense precipitation. Flood-producing atmospheric circulation patterns are closely related with the North Atlantic Oscillation (NAO), with a low (neg-

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ative) NAO index being associated to suppressed westerlies and storm track moving southerly toward the Mediterranean Sea (Walker and Bliss, 1932; van Loon and Rogers, 1978). Recent studies on the influence of the North Atlantic Oscillation (NAO) on the Tagus River flooding show evidence that the largest floods (average recurrence intervals > 25 year) are associated with negative mode of NAO during the 20–25 days (of a total 40 days period length) before the flood peak (Salgueiro et al., 2013). Moreover, the analysis of flood response under natural and dam-regulated regimes (before and after the construction of reservoirs ca. 1957–1960) revealed no changes in the behaviour of major floods (responding to a period of 25 days with a dominant negative NAO index mode). However, moderate flooding (return intervals of 10–25 year) was blurred during the post-dam period due to flood peak discharge attenuation by reservoirs.

### 3 Materials and methods

#### 3.1 Flood records database

A historical database was compiled from direct and indirect sources (Proceedings of the City Council, diaries, chronicles, memoirs, etc.). Direct written documents and maps were obtained from the General Archive of the Royal Palace with data about the administration of the Spanish Royal Heritage (S. XVI–XX), Simancas General Archive, Spanish National Library, and Geographic Service of the Spanish Army. Further bibliographical sources consulted include scientific and technical reports, local history works and non-systematic compilations by Rico Sinobas (1850), Bentabol (1900), Masachs (1948, 1950), Fontana-Tarrats (1977), González (1977), López-Bustos (1981), Comisión Técnica de Inundaciones (1985), Font (1988), and Canales (1989). All indirect information was checked by cross-referencing between different sites (Benito et al., 2003), following the standardised methodology proposed by Barriendos and Coeur (2004).

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Water stage data related to documentary descriptions and epigraphic marks was interpreted in terms of stage indicators as follow: (1) sites or landmarks reached by the flood (e.g. churches, bridges, etc.) were assumed to provide a exact discharge level (equal to the flood stage), (2) flooded areas (e.g. orchards, floodplain areas, gardens) provided a minimum flood stage, (3) non-flooded areas or landmarks (e.g. Royal Palace surrounded by water) were interpreted as a non reached maximum flood stage, (4) relative importance of the event with respect to previous floods (e.g. the 1840 flood was 2 m higher than the flood occurring in 1820) was quoted as a range of discharge in the case of two recorded levels.

The water levels (flood stage) required to reach the described flooded areas or landmarks (buildings, streets, gardens, etc) were converted into discharge values (O'Connor and Webb, 1988) using the HEC-RAS one dimensional hydraulic model (Hydrological Engineering Center, 2010). The HEC-RAS software, allows for rapid calculation of water-surface profiles for specified discharges, and energy loss coefficients. This conversion is an inverse problem, where the minimum discharge is obtained by matching the modelled flood water levels to those obtained from the elevation of sites inundated during the historical floods, as described in the documentary sources. Hydraulic modelling requires the estimation of key hydraulic characteristics of the river reaches (energy slope, roughness and cross sectional topography) as well as the boundary conditions upstream or downstream depending on the flow type selected in the model. Detail cross-section topography of the river channel was obtained from echo-sounder and GPS field survey performed by the LINDE Project (Spanish Ministry of Environment) for flood risk mapping. On the floodplain, cross-sections were completed when needed from detailed topographical maps (1 : 500 or 1 : 1000 in scale). In the study reach, river channel modification resulting of constructions of dykes, river lining, gardens setting and other engineering-architectural works were reconstructed based on historical maps and technical reports described among others by García Tapia, (1980), González Perez (1987) and Teran (1949).

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Systematic streamflow data of annual peak discharge were obtained from gauge records, namely from the Aranjuez station (9340 km<sup>2</sup>) covering the period between 1913 and 1985 although with several data gaps (1930–1954). In some years, annual daily maximum flows from Aranjuez station were transformed into peak discharges using a regression equation (Jiménez Alvarez et al., 2013). The missing annual flow data of the Tagus annual maximum flows were completed using data from Bolarque gauge station upstream of Aranjuez (7400 km<sup>2</sup>) and since 1985 to present using Villarubia de Santiago gauge station (9048 km<sup>2</sup>) located 20 km upstream of Aranjuez.

### 3.2 Statistical analysis under assumption of stationary flood series

Historical records potential is based not only in their value as documentary registers of the occurrence of a rare flood, but also as repositories of the magnitude of these floods by documenting, and sometimes mapping or graphically signal, the limits of every individual rare flood, over a centennial length period. In statistical analysis, streams flow monitoring data are considered systematic information whereas flood observations reported as having occurred above some threshold are known as non-systematic (censored) data sets (Leese, 1973). Fitting a distribution to these systematic and non-systematic data sets provides a compact representation of the frequency distribution, a task that requires a method to estimate the distribution parameters so that quantiles can be calculated. The parameters set for three statistical models, namely the Log-normal (LN), Generalised Extreme Value (GEV) and Two Component Extreme Value (TCEV) distribution functions have been estimated by the Maximum Likelihood Estimation (MLE) method. This method was selected based on its statistical features performance for large samples, and also because of its capacity to easily incorporate in the estimation process any additional non-systematic quantified data (Leese, 1973; Stedinger and Cohn, 1986).

Most historical information is non-systematic information of censored type since only floods of certain magnitudes (typically producing damages) are registered in documentary records. The minimum flood stage (threshold levels) required to be reported in



documentary is most commonly reported in areas with urban settlements or economically important human activities. In the case of the Iberian Peninsula sites studied till now, mostly located in floodplain areas (Benito et al., 2004). The flooding of sensitive areas, or perception threshold, may change through time, according to the progressive human occupation of the riverine areas and the socio-economic context. In the case of Aranjuez the perception threshold did not register significant changes through historical time. The minimum discharge ( $250 \text{ m}^3 \text{ s}^{-1}$ ) to inundate the floodplain in Aranjuez is related to the channel bankfull discharge because gardens, roads and buildings (including Royal Palace) are placed on the Tagus' floodplain (Fig. 2).

In our study, flood discharge in years 1919, 1954 and 1941 were estimated based on epigraphic flood marks. All the historical flood data in Aranjuez provide minimum discharge estimates based on elevation points affected by flooding, although the exact water depth in those sites is unknown (lower bound type after Naullet et al., 2005). A review of these various methods using historical data was presented by Ouarda et al. (1998) and by Francés (2004), and case study applications in Europe can be found e.g. in Frances et al. (1994), Naullet et al. (2005), Calenda et al. (2009) and Botero and Francés (2010), among others.

Statistical analyses rely on the general characteristics and stationarity of the flood. However, the temporal changes in the trajectory and statistics of a state variable may correspond to natural, low-frequency variations of the climate hydrological system or to non-stationary dynamics related to anthropogenic changes in key parameters such as land use and climate. Flood record stationarity from censored samples (systematic and/or non-systematic) was checked using Lang's test (Lang et al., 1999, 2004). This test assumes that the flood series can be described by a homogenous Poisson process. The 95 % tolerance interval of the accumulative number of floods above a threshold, or censored level is computed. Stationary flood series are those remaining within the 95 % tolerance interval (Naullet et al., 2005).

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### 3.3 Models with time-varying dependence parameters

Modelling of flood time series under a non-stationary assumption was carried out incorporating external covariates that explains the time changes of the statistical parameters (mean and variance) for the selected frequency distribution. In this study, the North Atlantic Oscillation (NAO) index and the reservoir index (RI) are introduced as explanatory variables of the marginal distribution parameters. The modelling framework was based on the “generalized additive models for location, scale and shape” (GAMLSS), as proposed by Rigby and Stasinopoulos (2005). GAMLSS has been successfully used in hydrological studies by Villarini et al. (2009a, b, 2010a, b, 2011) and López and Francés (2013).

In GAMLSS the response random variable  $Y$  (annual maximum peak discharge or historic flood above threshold discharge) has a parametric cumulative distribution function and its parameters can be modelled as a function of selected covariates, which in our study are represented by the winter NAO index (NAOw) and reservoir index (RI). A GAMLSS model assumes that independent observations  $y^t$  at a given time  $t = 1, 2, 3, \dots, n$  follow a cumulative probability distribution function  $F_Y = (y^t | \Theta^t)$ , where the evolution of its vector parameter  $\Theta^t$  can be expressed as a function of the explanatory variables  $x_j^t (j = 1, 2, 3, \dots, m)$  via a monotonic link function  $g_k(\cdot)$  as follows:

$$g_k = \Phi_k \beta_k + \sum_{j=1}^m h_{jk} (x_{jk}^t) \quad (1)$$

where  $\Phi_k$  is a matrix of explanatory variables (i.e. covariates) of order  $n \times m$ ,  $\beta_k$  is a parameter vector of length  $m$ , and  $h_{jk}(x_{jk}^t)$  represents the functional dependence of the distribution parameters (linear or smooth dependence) on explanatory variables  $x_j^t$ . In this study the smooth dependence is based on cubic spline functions. The addition of smoothing terms in the GAMLSS model has many advantages, such as identifying nonlinear dependence between the parameters of the parametric distributions and the explanatory variables (López and Francés, 2013).

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Due the characteristics of the selected probability distributions only the non-stationarity in the first two moments are considered. If the relation between distribution parameters and selected covariates follows a smooth dependence, this tends to increase the complexity of the model. According to Eq. (1), four different models can be considered for the time varying marginal distributions: (model a) two distribution parameters are constant; (model b) only the location parameter is time varying; (model c) only the scale parameter is time varying; and (model d) location and scale parameters are time varying. The final model is selected by comparing the Corrected Akaike Information Criterion (AICc; Hurvich and Tsai, 1989). This AICc criterion is more strict than the classical Akaike Information Criterion (AIC). In the absence of a statistic to evaluate the goodness of fit of the selected models as a whole, verification was made analyzing the normality and independence of the residuals of each model, following the recommendations of Rigby and Stasinopoulos (2005). With these criteria, a final model is provided with a balance between accuracy and complexity (López and Francés, 2013).

After defining the functional dependence between distribution parameters and each selected covariates and the effective degrees of freedom for the cubic spline, the distribution function  $F_Y(y_j|\vartheta_j)$  were selected, according to the largest value of the maximum likelihood. In the present study the best model performance was obtained for a Lognormal (LN) distribution function of two parameters. Detailed discussion on model fitting and selection can be found in Rigby and Stasinopoulos (2005), Stasinopoulos and Rigby (2007) and Villarini et al. (2009b).

The modelling of flood frequency over the systematic record should consider the high incidence of dam regulation on flood peak discharges. The Bolarque reservoir, together with the Entrepeñas and Buendia reservoirs complex, has substantially altered the stream flow regime of the upper Tagus watershed. Thus, under altered regime (1957–2008) the reservoir regulation is an important variable to explain the abrupt change on flood magnitude and frequency observed in the systematic record. A reservoir index (RI) was used to describe the dependence of flood variability on reservoir capacity and regulation strategies, as proposed by López and Francés (2013) for floods and recently

used for low flows by Jiang et al. (2014):

$$RI = \sum_{i=1}^N \left( \frac{A_i}{A_T} \right) \times \left( \frac{C_i}{C_T} \right) \quad (2)$$

Where:  $N$  is the total number of reservoirs upstream of the gauge station;  $A_i$  is the drainage surface controlled by the reservoir  $i$ ;  $A_T$  is the drainage surface upstream of the gauge station;  $C_i$  is the total reservoir capacity  $i$ ; and  $C_T$  is the mean annual runoff at the gauge station.

#### 4 Historical flood occurrence and discharge estimates

The historical flood record of Aranjuez constitutes, together with the Toledo one, the best registers to be found along the Tagus River, once they both gather a continuous data record on floods over a time period of up to 600 year. The Aranjuez Royal Palace and nearby extensive gardens were commissioned by Phillip II as a winter residence of the Spanish Kings. The oldest garden named as Isla Garden (AD 1387–1409) is bounded by the River Tagus and a diversion canal from a mill dam next to the Royal Palace. The other gardens (Prince's Gardens of ca. 150 ha) have been progressively extended along the Tagus River floodplain since AD 1545, together with a series of bridges, fountains, museums, as well as facilities for boat trips and deer hunting. In the middle of the Garden and next to the Tagus River is located The Farmer's Lodge, a neo-classic Palace (built for Charles IV in AD 1790–1803) highly vulnerable to flooding and frequently referred to in the documentary record when the floodplain is inundated. Currently, only two epigraphic flood marks are displayed at the entrance *patio*: the 1916 and 1924 flood levels.

The hydraulic model comprises a 13 km reach, from upstream of the Tagus's river confluence with the Jarama River to the upstream of the Prince's Garden at Embocador mill dam, along which 57 cross-sections were surveyed (Fig. 2). Hydraulically,

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the Aranjuez reach is relatively complex due to its wide floodplain and the near junction to the Jarama River. Discharge estimations using one-dimensional hydraulic modelling for a floodplain river reach may contain some uncertainties in the final result due to potential channel migrations, avulsion and meandering cut-off. However, historical maps show that meanders of the Tagus River along the studied reach were stabilised since AD 1795 when a 2 km reach was straightened and numerous dykes were built at either side of the Tagus River banks. Another source of uncertainty is related with the numerous hydraulic constructions (e.g. dams, bridges, dikes, dam mills) that have been set up along the Tagus River over time. The excellent documentation of the work descriptions and their emplacements (e.g. Díaz-Marta, 1992; García Tapia, 1980) were enough to include the corresponding temporal geometric changes of the cross-sections used in the hydraulic modelling. The model was calibrated using flood marks at the Labrador House Palace of the 20 December 1916 and 31 March 1924 floods. It was assumed for the calculations that flow was subcritical along the modelled reach. Manning's  $n$  values of 0.035 for the channel and 0.045–0.05 for the floodplain (crops and gardens with trees) were assigned. A sensitivity test performed on the model shows that for a 25% variation in roughness values, an error of 15% was introduced into the discharge results.

The historical flood reports quote over 100 locations or sites (buildings, streets, orchards, gardens) to have been flooded during the historical period. A database was implemented comprising: flood dates, flooded sites, GPS locations, maximum and minimum elevation and nearby cross-section used in the hydraulic modelling. As a result, for each historical flood, information of flooded and non-flooded affected sites and their elevation (maximum and/or minimum) is given, thus making possible to calculate the discharge associated with the flooded sites. Discharge values associated with each historical flood were estimated upon calculated water surface profiles bracketing elevations of flooded/non-flooded historical sites along the longitudinal profile (Fig. 2). Most of the documentary descriptions provide information on the minimum flood stage as written reports usually do not provide information of water depth at a site. Non-flooded

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5 sites explicitly indicated provide a maximum bound of discharge non-exceeded by the flood. Figure 2 shows the calculated water-surface profiles along the study reach with the best match to seven documentary flood indicators of flooded areas during the 1866 flood that provided a minimum discharge of  $650 \text{ m}^3 \text{ s}^{-1}$ . This individual analysis was carried out to each historical flood to estimate the associated discharge (Fig. 1c). Only 10 out of the 59 historical floods record did not have an associated stage reference. In this case, the stage reference for these ten floods was assumed to exceed the Tagus River channel capacity i.e. a minimum bankfull discharge ca.  $250 \text{ m}^3 \text{ s}^{-1}$ .

10 The flood record for Aranjuez comprises 59 historical floods (1557–1912) and 95 gauged floods (1913–2008). In terms of flood frequency distribution, four periods can be distinguished (Fig. 1) in Aranjuez: 1560–1611, 1739–1750, 1850–1899 and 1917–1928. The most important period in terms of flood number and magnitude corresponds to 1850–1899 and includes a total of 18 flood events, six of which showed minimum peak discharges of  $500 \text{ m}^3 \text{ s}^{-1}$  (Fig. 1c). The largest flood within this period occurred in December 1878 showing a minimum peak discharge of  $1200 \text{ m}^3 \text{ s}^{-1}$ , probably the largest at basin scale within the last 750 year. The next period in terms of flood magnitude corresponds to 1560–1611 with 7 floods; four reached a minimum discharge of  $400 \text{ m}^3 \text{ s}^{-1}$  and the largest (May 1611) over  $950 \text{ m}^3 \text{ s}^{-1}$ . During the 16th Century, intense precipitation with large floods alternating with severe drought occurred from 1585 to 1599 (Bullón, 2011). A third historical flood period corresponds to 1739–1750 with seven flood events, three of which yielded values of minimum discharge over  $350 \text{ m}^3 \text{ s}^{-1}$ , and the largest (December 1747) with an estimated minimum discharge of  $750 \text{ m}^3 \text{ s}^{-1}$  (Fig. 1c). During the 20th century, nine floods during the period 1913–1927 exceeded a discharge of  $350 \text{ m}^3 \text{ s}^{-1}$ , the largest one on 20 December 1616 reached a discharge of  $841 \text{ m}^3 \text{ s}^{-1}$ . An epigraphic mark of this flood is placed at a column at the entrance *patio* of The Farmer's Lodge, together with a lower level mark corresponding to the flood of 27 March 1924 (discharge estimate of  $700 \text{ m}^3 \text{ s}^{-1}$ ), the latter receding on 2 April 1924. However, according to our documentary data and modelling, the largest

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floods during the 20th Century occurred on 25 January 1941 with an estimated discharge of  $1100 \text{ m}^3 \text{ s}^{-1}$ , and on 8 March 1947, with a recorded discharge of  $930 \text{ m}^3 \text{ s}^{-1}$ .

## 5 Climatic, environmental and human drivers of flood variability

The temporal and spatial variability of flood patterns are closely related to global change drivers, climate and human activity (e.g. Benito et al., 2008; Hall et al., 2013). A major challenge in non-stationary flood frequency analysis is the identification of major external and internal parameters describing the non-stationarity pattern of the flood series.

At annual time-scale, moisture influx variability and therefore river flow is largely modulated by North Atlantic circulation, indicated by the NAO (Trigo et al., 2004). In the Tagus River flooding is highly related to persistent rainfall (several weeks) due to the successive passage of Atlantic cold fronts over the Iberian Peninsula in winter months (Cortesi et al., 2013). It is, therefore, natural to expect a strong impact of the NAO on flooding in large Atlantic river catchments. Most studies linking NAO index with floods (Benito et al., 2008; Machado et al., 2011; Silva et al., 2012) have used seasonal and monthly correlations leading to lack of significant connection and/or a delayed response between oscillation mode and hydrological event. Floods are produced by rainfall excess with an immediate hydrological response, requiring a daily to monthly resolution (Salgueiro et al., 2013). Figure 1b shows the relationship between the average winter (DJF) NAO index reconstructed by Luterbacher et al. (1999, 2002) since AD 1500 and the maximum discharges recorded during those months from instrumental (gauge stations) and documentary sources for the Tagus River at Aranjuez. A strong correlation is generally observed between negative monthly NAO index and flood discharges above  $400 \text{ m}^3 \text{ s}^{-1}$ . The correlation NAO-flood discharge in Aranjuez is not as robust as in other sites downstream (e.g. Alcantara and Talavera; Benito et al., 2008; Salgueiro et al., 2013) since in Aranjuez some autumn floods are connected with Mediterranean cyclogenesis.

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The temporal analysis of flood-NAO relationships (Fig. 1c) shows that periods of decreasing flood frequency and magnitude are associated with positive mode of NAO (e.g. 1650–1700, 1750–1800). The largest flood discharges occurred at periods with negative NAO phase or at periods with high NAO variability characteristic of strong meridional flow over the Atlantic. However, note that negative winter NAO index values are not always related to the existence of extraordinary floods.

Other factors affecting the stationarity of flood series are land-use changes and reservoir construction in the upper Tagus basin. Detailed analysis of land-use changes during the historical period is highly complex and out of scope of this paper. However, the upper Tagus catchment is a mountain region with a relatively low population density, covering a relatively high drainage surface of 9340 km<sup>2</sup>. Therefore, we can assume that local land-use changes can be neglected as major drivers of change in flood patterns. Analysis of the series of annual maximum discharges recorded at gauging stations indicates a decrease in the peaks of ordinary floods over the last 50 year (Fig. 1). This decrease in peak discharge is mainly due to the construction of dams between the 1950s and 1960s.

## 6 Flood frequency analyses

### 6.1 Flood frequency analysis under stationarity assumption

Stationarity tests (Lang et al., 1999 and 2004) of the combined documentary and instrumental (under natural regime) flood series above 250 and 350 m<sup>3</sup> s<sup>-1</sup> show significant non-stationary conditions over the period 1559–1956. In the case of discharge above 250 m<sup>3</sup> s<sup>-1</sup> the flood record was stationary over the period 1850–1956, whereas for discharges above 350 m<sup>3</sup> s<sup>-1</sup> the stationary period comprises 1779–1956. The test performed on the instrumental record (1913–2008) resulted in non-stationarity, due to a sharp decrease in flood discharges after 1957 related with stream flow regulation by reservoirs. In order to include in the analysis the maximum length of historical floods,

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the FFA analysis was carried out with 32 historical floods exceeding  $350 \text{ m}^3 \text{ s}^{-1}$  over the period 1779–1912, and 45 year of annual maximum floods (Fig. 3a). The LN (two parameters), TCEV and GEV distribution functions were applied to both, systematic data (dashed line) and systematic plus non-systematic (documentary flood) data (solid line) (Fig. 3b). Discharge values associated to different return periods ( $T$ ) that results of the distribution fitting are shown in Table 1. The visual matching as well as the statistical parameters indicates that distribution fitting is good in all cases, with the Log-normal distribution providing the more realistic quantile values. In fact, only one flood over the last 230 year reached a discharge of  $1200 \text{ m}^3 \text{ s}^{-1}$  (1878 flood event).

The incorporation of the historical data into the FFA results only in slightly higher values in the magnitude of the flood quantiles than those obtained with the systematic record before reservoir constructions (Fig. 4; Table 1). Major differences are found when the combined FFA quantiles are compared with annual maximum discharge series under the regulated systematic period. For example, the 0.01 annual probability flood of the combined historical and natural-systematic records is  $1450 \text{ m}^3 \text{ s}^{-1}$  whereas during the regulated systematic period is  $600 \text{ m}^3 \text{ s}^{-1}$ .

### 6.2 Modelling flood data under non-stationary approach

The implementation of the non-stationary model consisted on establish the type of dependence of the distribution parameters with the associated external covariates (winter NAO index and Reservoir index) as a function of time. For the selected Lognormal distribution (LN with 2 parameters:  $\mu^*$  and  $\sigma^*$ , the mean and the SD of the logarithms respectively) the modelled changes (linear or smooth dependence) on the distribution parameters are applied to obtain the best fitted distribution parameters for the discharge above the historical discharges threshold. It should be noticed that, for a LN, the variance of the original random variable is a function of its two parameters.

The reservoir index for the Aranjuez gauge station was calculated by Eq. (2). As previously indicated, the streamflow from Tagus basin headwaters is regulated by three

major reservoirs (Entrepeñas, Buendia and Bolarque). Construction of the Bolarque reservoir (capacity of  $31 \times 10^6 \text{ m}^3$ ) was completed in 1955 and the Entrepeñas-Buendia ( $2394 \times 10^6 \text{ m}^3$ ) were built between 1956–1958 (Dams and reservoir database of the Spanish Ministry of Environment; <http://sig.magrama.es/snczi/>). As shown in Fig. 5, in 1956, the reservoir index (RI) has an abrupt change point, because the construction of these dams.

Modelling of maximum annual discharge ( $Q$ ) followed a two step process. Firstly, the non-stationary models incorporated only the winter (December to March) NAOw index (period with highest correlation with flooding) with two models implemented for the periods 1913–1956 (Fig. 6a) and 1957–2008 (Fig. 6b). Secondly, the reservoir index (RI) was added as a second covariate for the whole systematic period (1913–2008).

The non-stationary model with NAOw as external covariate over the period 1913–1956 shows a linear dependence on both parameters the mean and the variance of the logarithms of annual flood peak discharge (Table 2) and therefore on the mean and variance of the floods (Fig. 6a). However, for the period 1957–2008 the linear dependence was only found in relation with the  $\log-Q$  mean (Table 2 and Fig. 6b), indicating the strong incidence of the reservoir regulation on the peaks of annual maximum floods. Also is interesting to underline that, for the whole period 1913–2008, the heteroscedasticity of the  $\log-Q$  data set is only described by the climate, and not by the reservoir construction. However, in terms of the original variable, the heteroscedasticity of the floods is described by both the climate and the reservoirs construction.

The models adequately describe the changes in the percentiles of the annual maximum flood peaks although, for some periods, high negative NAOw index do not implies high flood discharges. Moreover, there is a higher variability for the highest percentiles (75 and 95 %) whereas a low variability is observed for the 5 and 25 % percentiles. Figure 7 shows the observed discharges and the estimated quantiles plotted in relation with the NAOw index. The highest annual discharges and their associated 95 % percentile show a high correlation with negative NAO index for the period with natural

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or low altered stream regime (1913–1956), although some negative NAO index values resulted on moderate peak discharges.

The modelling of flood frequency including the NAOw and RI as external covariates explains adequately the observed flood variability for the Aranjuez station (Fig. 8). The flood modelling incorporating additionally the reservoir index (RI) allowed a good characterisation of the sharp change on the flood regime associated with the construction of the upstream reservoirs system, which controls about 80% of catchment surface draining to Aranjuez. These results show the potential of the RI as a covariate that explains the sharp change in the median and the variance, and how to incorporate this effect into the model.

### 6.3 Comparison between stationary and non-stationary models

The modelling of non-stationary series was carried out considering continuous annual maximum discharge records as a variable of response ( $Y$ ) to be explained in the model. In the following analysis the focus is on modelling flood frequency and magnitude of censored historical floods for which a minimum discharge threshold can be determined. In the study of historical (documentary) hydrology the most critical point is to establish a quantitative threshold of discharge, which in most cases depend on the human occupation of the floodplain (perception level) that may vary through time. In the case of Aranjuez, documented floods are in relation to the gardens and singular buildings (Royal Palace, Lodge etc.) settled on the floodplain, and their situation has remained almost invariant over the last 400 year. The threshold of discharge is therefore related to the bankfull discharge and the minimum discharge required inundating the floodplain, which has been established in  $250 \text{ m}^3 \text{ s}^{-1}$ . As indicated previously, the relationship between the average winter NAO index, reconstructed by Luterbacher et al. (1999, 2002) since AD 1500, and the historical flood discharges estimated in our study support the use of NAOw as a covariate to explain non-stationary through the historical period.

Figure 9 shows the results of the estimates of quantiles since AD 1700 assuming the modelling relationships obtained from the systematic record with natural regime

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(1913–1956). The comparison of the estimated quantiles with the non-systematic historical data shows that the reconstructed NAO index is a highly significant covariate to describe the temporal variability of historical floods recorded in Aranjuez. Similar results were obtained for the estimated quantiles with the non-stationary model based on the systematic record for the period 1913–2008 (Fig. 9), where both the NAOw index and the IR were used as covariates of the statistical parameters of the parametric distribution.

The results obtained from modelling flood frequency under non-stationary conditions while incorporating external forcing captures more adequately the dispersion of flood values, and shows the effect of climate indices modulating the frequency and magnitude of floods (Fig. 9). The historical non-stationary model also captures the presence of multi-decadal temporal trends, characteristic of the North Atlantic atmospheric circulation, which are incorporated in the flood frequency model. Figure 10 shows the results of FFA in stationary conditions and non-stationary conditions for an exceedance probability of 0.01 (i.e. return period of 100 year). It can be seen that non-stationarity models indicate the existence of periods in which flood frequency experienced significant variability (decreases and increases). On the long-term, stationary FFA based on historical data reflects the average discharge for the 100 year flood although in terms of flood hazards during some periods the stationary 100 year flood discharge ( $1450 \text{ m}^3 \text{ s}^{-1}$ ) has associated a higher probability of occurrence. A clear decrease in flood frequency can be seen during the period 1957–2008, which has been caused by the construction of the upper Tagus reservoirs, with a subsequent decrease in the frequency of occurrence of the floods in Aranjuez.

## 7 Discussion

Flood hazard analysis is critical for the elaboration of risk maps, and it is the basis of the design of hydraulic structures (e.g. dam spillways, diversion canals, dikes), urban drainage systems, land and urban planning and cross-drainage structures (e.g. bridges

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and culverts and dips). This probabilistic assessment of hazards and risks are routinely done assuming the stochastic nature of stream flow and their extremes. The reliability of the quantile estimator can be improved either by using the best statistical model or by increasing the amount of information (e.g. Botero and Frances, 2010). The use of historical censored data can increase the information length at the gauge station, providing, as it has been mentioned previously, that for each study case information error and the underlying stationary hypothesis are explored.

Flood frequency analysis with systematic and additional non-systematic information have been widely demonstrated to reduce the quantile estimate uncertainty (e.g. Steidinger and Cohn, 1986; Frances et al., 1994; O'Connell, 2005). In recent decades, land use change and anthropogenic climate change impacts on hydrology has largely questioned the stationarity hypothesis (Milly et al., 2008), and several non-stationary models have been applied in the study of annual maximum flood series (e.g. Cunderlik and Burns, 2003; Ouarda and El-Aldouni, 2011; López and Francés, 2013). So far, there is a lack of research assessing the reliability of non-stationary flood frequency models to represent hydrological disturbances due to low-frequency climate variability over centennial historical periods.

In the Aranjuez case study, the Lang's stationary test for a discharge threshold of  $350 \text{ m}^3 \text{ s}^{-1}$  was successfully passed covering the historical period (1779–1912) and the instrumental record under natural streamflow regime (1913–1957). The oldest historical period (1557–1778) showed a non-stationary behaviour, mainly related with a reduction on the flood frequency over the period 1630–1710, that climatically corresponds to the minimum Maunder, known by the existence of dry periods and characterised by a high hydrological variance and a decrease in flood activity over the western Iberian Peninsula (Rodrigo et al., 2000).

A detailed analysis of the stationary test shows several periods of accumulated flood events at the limit of the confidence interval due to a reduction of flood frequency (Fig. 3a). A reduction on flood frequency of large floods occurred over the first decades of the 19th Century (1810–1850). On the contrary, a high flood frequency of large floods

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was recorded over the second half of the 19th Century and beginning of the 20th Century. A similar flood reactivation during this period is recorded in other Iberian rivers such as Guadalentín (Benito et al., 2010), Júcar (Francés, 1998), Llobregat (Thorndy-craft et al., 2005) and Guadalquivir Rivers (Benito et al., 2005). Over the whole record, three of the largest five floods of the Tagus River occurred during the historical period record (1787, 1611 and 1830). The use of this non-systematic data improved the precision of the frequency analysis for the exceedance probability of 0.01 (i.e. return period of 100 year) for describing long-term average flood risk. The frequency analysis based solely in the systematic gauge record provided lower discharge values, both under natural flow regime (1911–1956) and to a greater degree during the substantially altered stream flow post-reservoir regime (1957–2008). In fact, the flood quantile estimates over the systematic period under non-stationary model (Fig. 8) shows a sharp change in the flood regime after the construction of the reservoirs system.

In the western Iberian Peninsula, moisture influx variability, and therefore river flow, is largely modulated by North Atlantic circulation, characterized by the NAO (Trigo et al., 2004). In the Tagus River a strong correlation is generally observed between flood discharges above  $400 \text{ m}^3 \text{ s}^{-1}$  (floods of  $\sim 10$  year return interval) and negative monthly NAO index, including both the historic and instrumental periods. However, negative winter NAO index values are not always related to the existence of extraordinary floods. These relationship between flood occurrence and negative NAO phase provides robust basis for using the winter NAO index as covariate explaining expected changes on flood stationarity towards the historical period, and validation of the changing discharge for an exceedance probability quantile. The modeling of the dependence on the occurrence rate and the dispersion coefficient in relation with the climatic variability during the instrumental period identified the NAO index as an explanatory covariate, allowing the modeling of the most likely changes on the statistical parameters towards the past. The implementation of a non-stationary model that includes the effects of the North Atlantic circulation mode in relation to climate variability provides an excellent description of the interannual flood variability associated to the 0.01 annual exceedance probabil-

ity flood (100 year return period) over the last 300 year. This variability in the case of the 100 year flood discharge obtained with the non-stationary model reproduces many of the observed changes on historical floods (with higher flood magnitudes recorded for the periods 1920–1950, 1860–1890 and 1700–1730) as well as other with below average flood magnitudes (e.g.1750–1800). The projection of NAO index in relation to climate change is still unclear with nearly half of the models predict a positive intensification of the index associated with global change (Osborn, 2004). Moreover, some studies have shown that the association between the NAO circulation mode (NAO index) and local or regional climate variables (rainfall and temperature) has changed over time (Rodó et al., 1997; Goodess and Jones, 2002).

The application of these models as a predictive tool requires the introduction of scenarios as well as an improved modelling of low frequency atmospheric circulation under anthropogenic climate change (Hall et al., 2013; Merz et al., 2014). Our results indicate that stationary frequency analysis using historical data provides a robust reference value of the “average” probabilistic discharge during the considered period. For instance, the peak flood for an annual exceedance probability of 0.01 during the historical (1779–1912) and the systematic period (1913–1956) under stationary model (log-normal distribution) is  $\sim 1450 \text{ m}^3 \text{ s}^{-1}$  whereas under the non-stationary model for the period 1700–1956 the 100 year flood discharge ranges from a maximum value of 4180 to a low of  $560 \text{ m}^3 \text{ s}^{-1}$  (Fig. 10) This interannual variability in probabilistic discharge reflects the high variability on streamflow characteristic of regions with contrasting inter-annual and inter-decadal rainfall amounts, such as the Iberian Peninsula. These changes on annual flood peaks have been also observed in the historical periods with discharges that ranged between  $1200 \text{ m}^3 \text{ s}^{-1}$  in 1878 to a low gauged discharge of  $59 \text{ m}^3 \text{ s}^{-1}$  in 1929. Our analyses indicate that a stationary model based on long-term historical flood records provides reliable average flood discharge quantiles, although the probabilities of exceedance changed from year to year. In the case of hazard-sensitive infrastructures, the hypothesis of a non-stationary model should be

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integrated in their design and implementation, once it is able to address a more accurate estimate regarding the worst case scenario for the Safety Check Flood.

## 8 Conclusions

Documentary records of the Tagus River in Aranjuez (Central Spain) provide evidence of continuous and homogenous data of extreme floods over the last 300 year. Documented reference to flooded and non-flooded sites (gardens, streets) and buildings (Royal Lodge) were successfully assessed and compute into flood discharges. The universality of this methodology depends however on the quality and continuity of the documented registers, as well as of the characteristics of the human settlement (spatial variation along time, burying processes, destruction or changes of landmarks) and the geomorphological dynamics of the area.

The discharge estimates show evidence that during the historical flood record (1557–1912) flood events of greater magnitude than the ones recorded in the gauging station (1913–2008) were documented. The documentary record contains descriptions of 59 historical floods with the largest event produced in 1878 with an estimated discharge of  $1200 \text{ m}^3 \text{ s}^{-1}$ . Documentary evidence illustrates the high sensitivity of flood magnitude and frequency to the climatic variability during the last millennium. Unusually high flood frequencies were registered in the periods: 1580–1610, 1730–1760, 1800–1810, 1870–1920, and 1940–1950. The construction of large reservoirs (Entrepeñas-Buendia and Bolarque dams), in full operation since 1957, changed dramatically the flood regime of the upper Tagus River. During the post-reservoir period (1957–nowadays) the maximum peak flow discharge reached  $280 \text{ m}^3 \text{ s}^{-1}$  (in 1964), with most of the events recording values below  $160 \text{ m}^3 \text{ s}^{-1}$ .

The identification of flood clusters during the historical record uncover the potential problems of assuming a stationarity in the flood frequency analysis. In this study, two different models were used for the flood frequency analysis of the Aranjuez flood data sets: (1) the stationary model, in which the distribution parameters do not depend on

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covariates, i.e. the parameters are constant in time; and (2) the time-varying model (non-stationary model) that incorporates external covariates, where the distribution parameters can vary as a function of climate variability (North Atlantic Oscillation index during winter months, NAOw) and the stream flow regulation by dams (Reservoir Index; RI). Regarding the stationary flood frequency analysis, a stationarity test associated to different discharge thresholds (perception levels) was efficient to detect anomalous flood periods/records. The Aranjuez flood record (historical and systematic) for discharges exceeding  $350 \text{ m}^3 \text{ s}^{-1}$  showed problems of stationarity with two cutting dates in AD 1778 and AD 1957, the former associated with an increase in flood frequency and the second as a result of a decreasing flooding after the upstream reservoirs system construction. The stationary test (Lang's test; Lang et al., 1999) was successfully passed over the period 1779–1956, covering part of the historical non-systematic record (1779–1912) and gauged record under natural stream flow conditions (1913–1956).

Flood frequency analysis using documentary data (plus gauged record) improved the estimates of the probabilities or rare floods (return intervals of 100 year and higher). Moreover, historical information increases and tests the representation of outliers in the systematic data. A flood frequency analysis with 32 historical floods ( $> 350 \text{ m}^3 \text{ s}^{-1}$ ) and 45 annual maximum flood from the gauged record provided the best fitting results to a Log-normal distribution, with a discharge of  $1450 \text{ m}^3 \text{ s}^{-1}$  associated to a 0.01 % annual exceedence probability (average recurrence interval of 100 year). Our analysis shows a major impact of river regulation by large reservoirs in the occurrence of large peak flows in the upper Tagus River basin, with major impact on the stationarity of the flood series.

Although on the long-run FFA under stationary models provides good average values, historical flood record indicate that secular flood occurrence related to climate variability has changed over the last 300 year, during which flood quantiles have fluctuated through time at decadal time spans. The modelling of flood time series under a non-stationary model incorporating the NAO index and Reservoir Index as exter-

nal covariates, demonstrate the existence of a high decadal variability on flood percentiles reflecting the internal climatic variability in the flood occurrence. According to non-stationary models, the peak flood associated with a 0.01 annual exceedance probability may range between 4180 and 560 m<sup>3</sup> s<sup>-1</sup>.

In terms of the hydrological effects of climate change, future global circulation model projections incorporate too much uncertainty to accurately specify expected patterns of precipitation change, and even less to estimate expected changes in the frequency and magnitude of extreme storm and flood events. Predictions can be improved by incorporating long-term flood records (several millennia) in climatic modelling and statistical analysis. The study of temporal variability of past climate–flood links can establish short- and long-term relationships at regional levels and in areas within different climatic zones. Regional studies of long-term climate–flood links involve calibrating the relationships, detecting trends (where they exist) and revising estimates of return periods. This integration will greatly improve our understanding of flood frequency and magnitude in the context of changing climates where the assumption of stationarity (implicit in most current flood risk models) is being questioned. The analysis of flood quantiles under stationary and non-stationary models over long-term historical periods is critical for the analysis of flood hazards of sensible infrastructures (dams, nuclear power plants) in the context of climate change.

*Acknowledgements.* This research was funded by the Spanish Ministry of Economy and Competitiveness through the research projects FLOODMED (CGL2008-06474-C02/BTE), SCARCE-CONSOLIDER (CSD2009-00065) and CLARIES (ref. CGL2011-29176), and by the CSIC PIE Intramural Project (ref. 201430E003).

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**Table 2.** Results of the marginal distributions with NAO and RI as the explanatory variables of distribution parameters.

Series	Probability distribution	Distribution parameters		AICc
		$\mu^*$	$\sigma^*$	
1913–1956	Lognormal	5.373–0.028NAO	exp (–0.2059–0.258NAO)	586.51
1957–2008	Lognormal	4.219–0.193NAO	exp (–0.6538)	521.45
1913–2008	Lognormal	5.46–0.153NAO-0.924RI	exp (–0.4209–0.1356NAO)	1109.59

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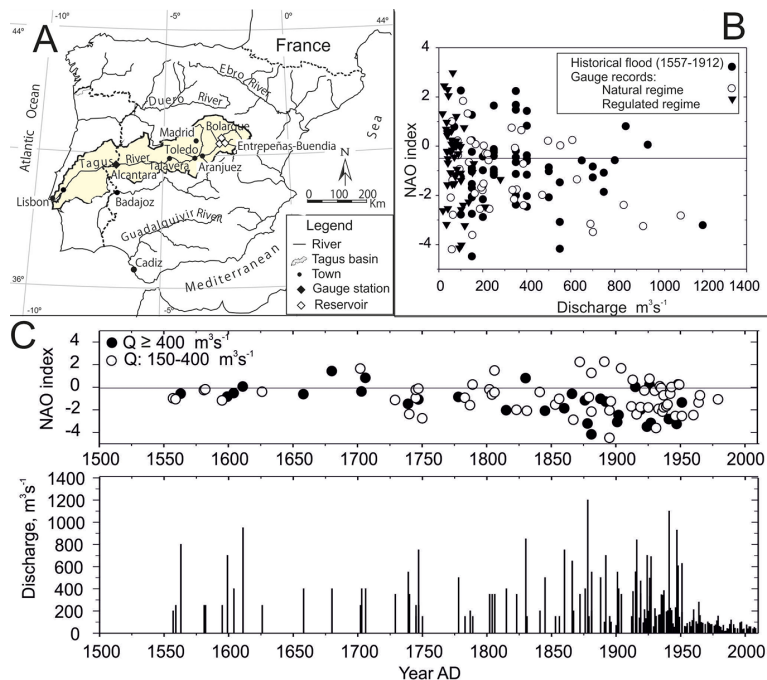
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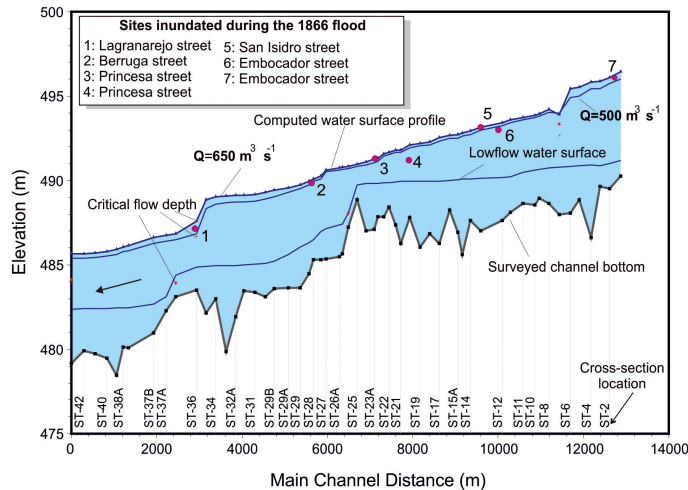


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**Figure 1.** The Tagus River Basin in the Iberian Peninsula. **(a)** Location of the historical and instrumental records used; **(b)** Relationship between the annual flood peak discharge at Aranjuez and the corresponding month value of NAO index. Discharges over  $400 \text{ m}^3 \text{ s}^{-1}$  are in most cases related with negative NAO phase; **(c)** Reconstructed flood record in Aranjuez during the historical period (1557–1912) and annual maximum flood series during the instrumental periods (1913–2008), with vertical bars showing the peak discharge. NAO index associated with flood dates are indicated (historical NAO index after Luterbacher et al., 1999, 2002).



**Figure 2.** Above: “Royal Palace of Aranjuez” with a view of the Tagus River and gardens painted by Antonio Joli, ca. 1753. Bottom: longitudinal profile of the channel bottom along the study reach and calculated water surface profiles fitting historical flood water indicators described as inundated during the 1866-flood. An estimated discharge of  $650 \text{ m}^3 \text{ s}^{-1}$  was assigned to this flood. Note the water surface profile for base flow conditions.

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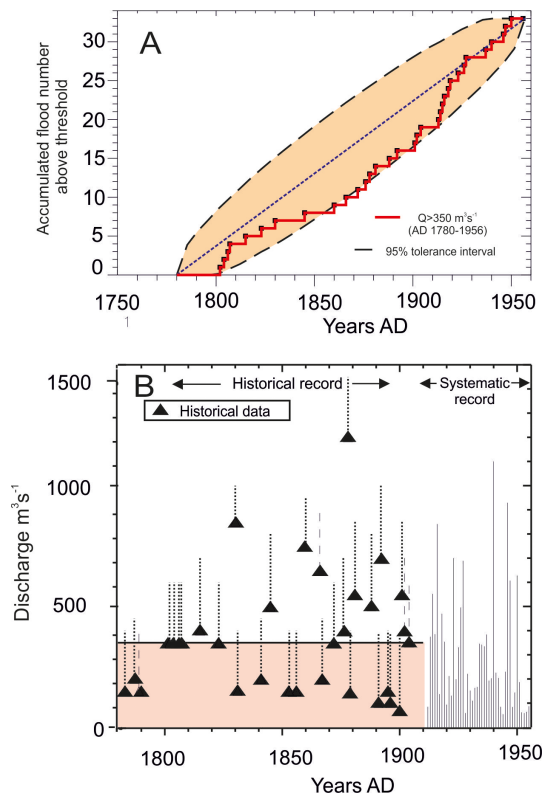
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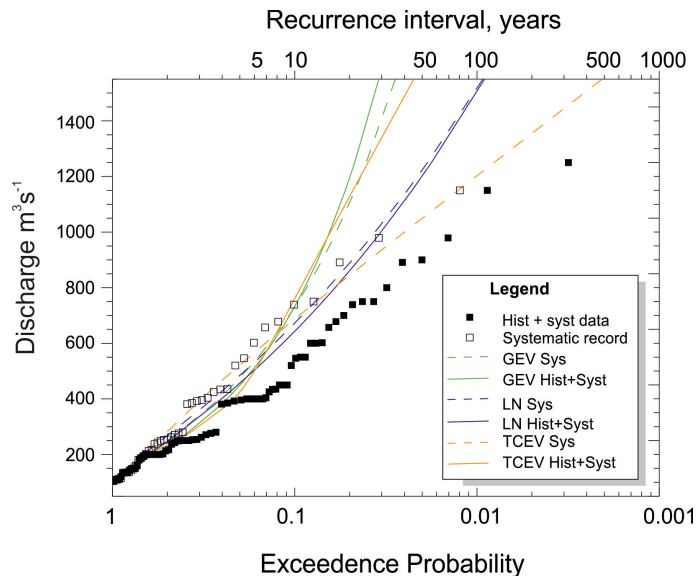
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**Figure 3.** (a) Poisson test on the time flood process in the Aranjuez record for the period 1780–1956, for floods exceeding a discharge threshold of  $350 \text{ m}^3 \text{ s}^{-1}$ . Note that the flood series (central line connecting points) remain within the 95 % tolerance interval (outside enveloped curves). (b) Organisation of historical and systematic flood data, for flood frequency analysis.

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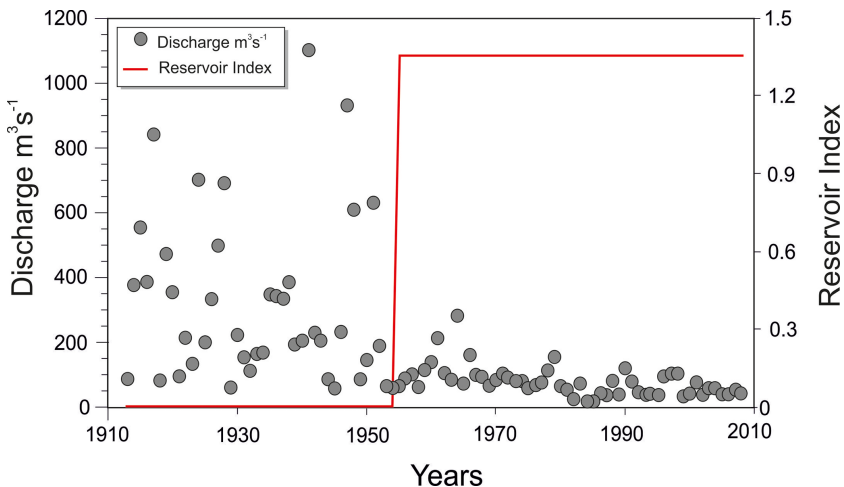
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**Figure 4.** Log-normal (LN), Generalised Extreme Value (GEV) and two component extreme-value (TCEV) distributions fitted with (1) systematic data and (2) systematic and historical flood data. Dots represent the plotting positions for the systematic record (white dots) and historical plus systematic records (black dots).

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**Figure 5.** Variations of the Reservoir Index (RI) in the Aranjuez gauge station. Note the sharp change in RI value produced in 1957.

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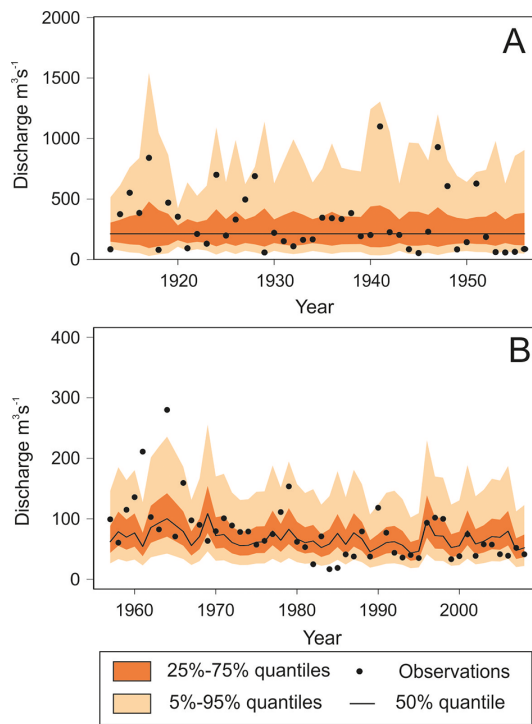
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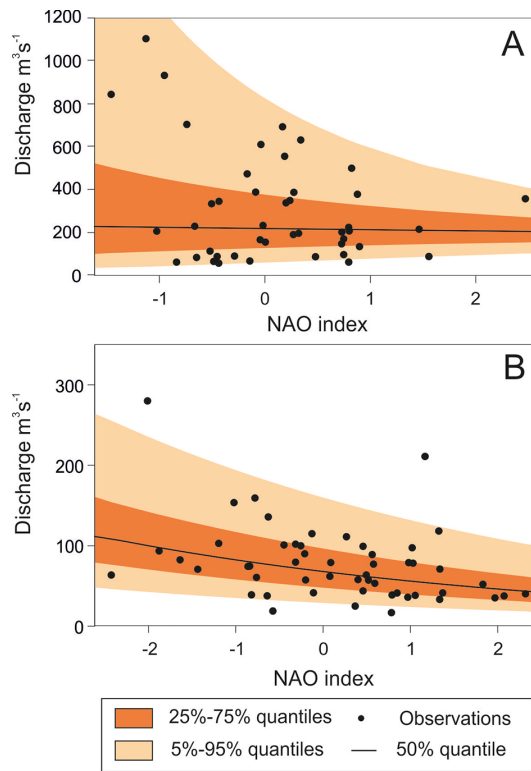
**Figure 6.** Modelling results of the variation in the marginal distributions of the annual maximum streamflow with the NAO index as the explanatory variable of the distribution parameters. **(a)** Quantile plot for the Aranjuez gauge in the period 1913–1956; **(b)** Quantile plot in the period 1957–2008.

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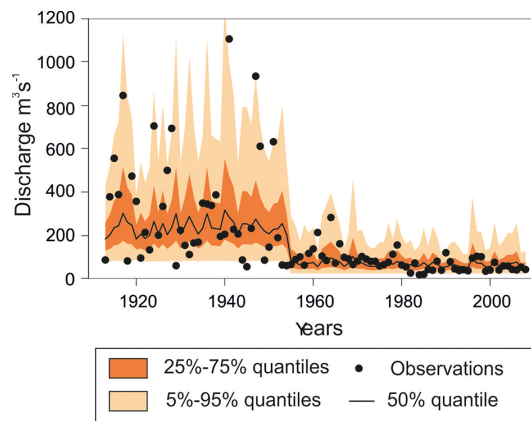


**Figure 7.** Estimated quantiles for the model with the NAO index as the explanatory variable plotted against NAO index. **(a)** Plot for the Aranjuez gauge in the period 1913–1956; **(b)** Plot during the period 1957–2011.

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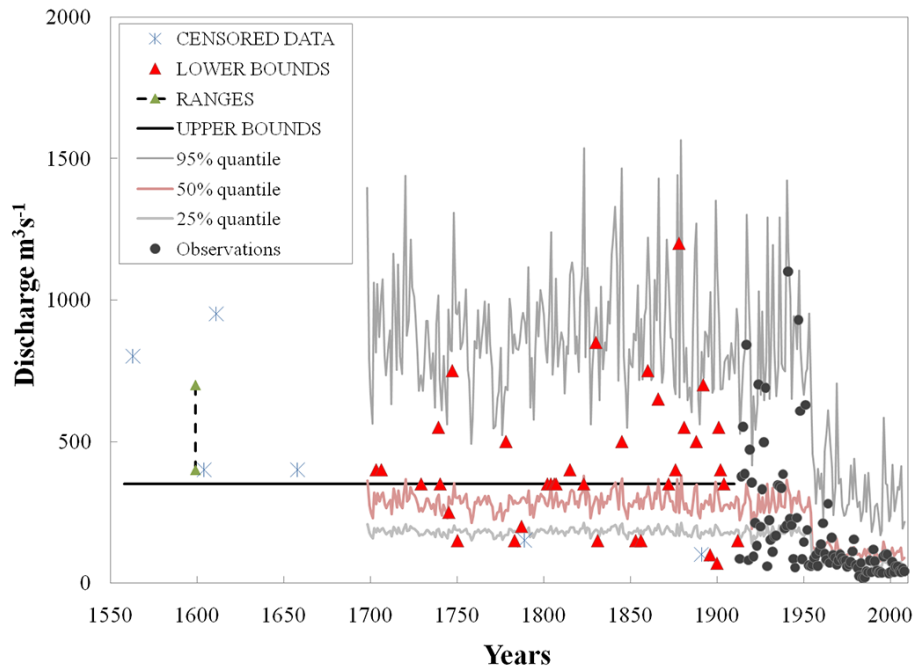
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**Figure 8.** Estimated quantiles for the non-stationary model with the NAO and RI (reservoir index) as explanatory variables of the distribution parameters. Note a sharp change on the maximum annual floods dividing the record in two sets (1913–1956 and 1957–2011) due to construction of a complex of three reservoirs in the Tagus headwaters.

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**Figure 9.** Estimated quantiles for the non-stationary model using systematic information under natural regimen (1913–1956) and the systematic information (1913–2008), with the NAO index and reservoir index as the explanatory variables of the distribution parameters.

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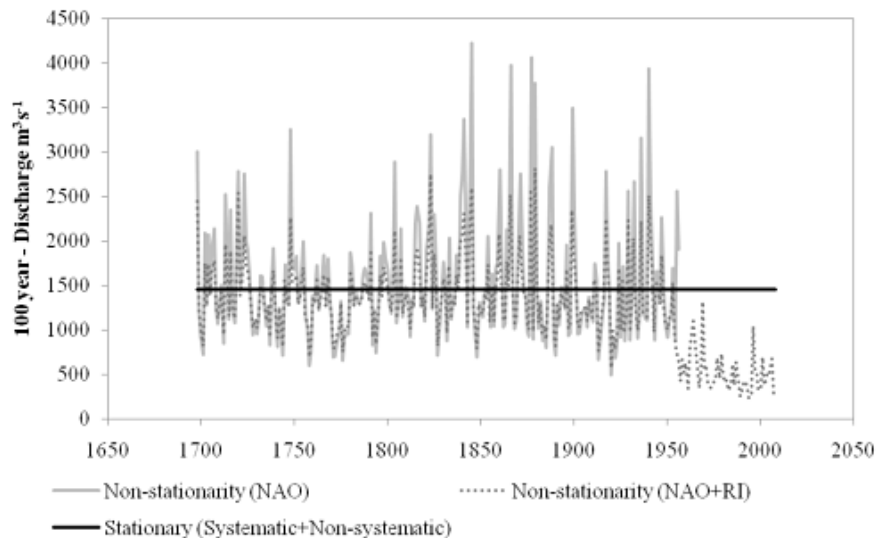
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**Figure 10.** Quantile estimates of the annual maximum floods with 0.01 annual exceedance probability based on stationary and non-stationary models.

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