



Hydrologic
calibration of paired
watersheds using
a MOSUM approach

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Hydrologic calibration of paired watersheds using a MOSUM approach

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Abstract

Paired watershed studies have historically been used to quantify hydrologic effects of land use and management practices by concurrently monitoring two neighboring watersheds (a control and a treatment) during the calibration (pre-treatment) and post-treatment periods. This study characterizes seasonal water table and flow response to rainfall during the calibration period and tests a change detection technique of moving sums of recursive residuals (MOSUM) to select calibration periods for each control-treatment watershed pair when the regression coefficients for daily water table elevation (WTE) were most stable to reduce regression model uncertainty. The control and treatment watersheds included 1–3 year intensively managed loblolly pine (*Pinus taeda* L.) with natural understory, same age loblolly pine intercropped with switchgrass (*Panicum virgatum*), 14–15 year thinned loblolly pine with natural understory (control), and switchgrass only. Although monitoring during the calibration period spanned 2009 to 2012, silvicultural operational practices that occurred during this period such as harvesting of existing stand and site preparation for pine and switchgrass establishment may have acted as external factors, potentially shifting hydrologic calibration relationships between control and treatment watersheds. Results indicated that MOSUM was able to detect significant changes in regression parameters for WTE due to silvicultural operations. This approach also minimized uncertainty of calibration relationships which could otherwise mask marginal treatment effects. All calibration relationships developed using this MOSUM method were quantifiable, strong, and consistent with Nash–Sutcliffe Efficiency (NSE) greater than 0.97 for WTE and NSE greater than 0.92 for daily flow, indicating its applicability for choosing calibration periods of paired watershed studies.

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1 Introduction

Forty percent of the projected growth of global energy demand in 2035 will come from low carbon sources such as biofuels, geothermal, solar, wind, and nuclear (Biol, 2014). By 2035, the use of biofuels is projected to triple compared to the 2011 baseline of 1.3 million barrels of oil equivalent per day (mboed^{-1}) with advanced biofuels supplying 20% of this demand (Biol, 2014). However, achievement of such sustainable bioenergy industry requires accurate assessment of effects of biofuel-driven land use changes as constrained by competing land demands for food and fiber production and urban sprawl. Therefore, the need to effectively optimize current land use practices to meet future biofuels production demands with limited environmental impact requires studies to quantify effects of such transitions of land use change on local and regional hydrology and water quality (Georgescu and Lobell, 2010).

Vast areas of southeastern US Coastal plain and Gulf Coast regions are covered by pine forests (*Pinus* spp.) planted primarily for timber production. A potential option for biofuel feedstock production within planted pine systems is intercropping of a perennial biofuel crop, such as switchgrass (*Panicum virgatum*) between rows of planted pine. However, documentation of sustainability of this system, including impact on water resources, must be quantified and compared to current forest management practices. In a traditional setting, pine is planted in rows, often in raised bed, and the space between beds are occupied by natural vegetation. Replacing the natural understory with switchgrass introduces a relatively uniform vegetation structure between pine beds and, thus, reduces intensity of runoff (Blanco-Canqui et al., 2004; Schmer et al., 2011). Most recently, Albaugh et al. (2014) studied the effects of intercropping switchgrass in pine stands on water use and gross primary productivity using 3 year (2009–2011) data from plot-scale experiments in an upper coastal plain site of North Carolina, USA. They reported an increase in water uptake (transpiration) by switchgrass during the peak growing season, with the total annual evapotranspiration (ET)

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higher from the traditional pine stand than switchgrass. However, there have been no such studies completed on a watershed-scale basis.

A number of studies (McCarthy et al., 1991, 1992; Skaggs et al., 2011; Amatya et al., 1996, 1998a, 2000, 2003, 2006; Beltran et al., 2010) have been conducted on pine plantations with improved drainage in low gradient coastal North Carolina to evaluate effects of both silvicultural and water management practices on downstream water quantity and quality. Most of these studies were based on a classical paired watershed approach, where two neighboring watersheds (one control and one treatment) were concurrently monitored during calibration (pre-treatment) and post-treatment periods (Clausen and Spooner, 1993; Loftis et al., 2001; Andreassen, 2004). A statistically significant relationship between control and treatment watersheds is established during calibration such that any significant shift detected during treatment is attributable to treatment effects (Zegre et al., 2010). The paired watershed approach also offers the ability to identify roles of forest cover, internal watershed behavior, and climate variability to establish a “baseline” for reference (Zegre, 2008).

The paired watershed approach has been extensively used to assess effectiveness of conservation practices (King et al., 2008; Tomer and Schilling, 2009; Jokela and Casler, 2011; Lemke et al., 2011), changes in water yield due to afforestation and harvesting or deforestation (Bosch and Hewlett, 1982; Fahley and Jackson, 1997; Brown et al., 2005; Amatya et al., 2006; Edwards and Troendle, 2008; Chescheir et al., 2009; Bren and Lane, 2014), and best management practices for controlling sediment transport, nitrogen and phosphorus runoff and leachate (Jaynes et al., 2004). This approach continues to be used on low-order watersheds as the primary method for impact assessments (Bren and Lane, 2014) although its validity for predicting effects on large flooding events has been challenged (Alila et al., 2009). In a recent study Bren and Lane (2014) analyzed high-quality data sequences from the pre-treatment phase of two Australian paired catchment projects to answer key questions on (a) the gain in information over time as the calibration period extends, in other words, an “optimal length” of calibration, (b) the relative gain or loss of information when using daily or monthly or

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annual data, and the effect of autocorrelation of residuals in calibration models, and (c) how the calibration development should be monitored in real time, to obtain the most efficient calibration.

Structural changes of regression parameters in the calibration regression models of paired watersheds due to pest infestations or anthropogenic activities other than the treatment may yield erroneous inferences of treatment effects (e.g., Vogl and Lopes, 2009). This issue is similar to question (c) addressed by Bren and Lane (2014) to obtain the most efficient calibration. Several studies have demonstrated significant changes in hydrology and nutrient concentrations and loads on forest sites due to silvicultural and water management operations (Amatya et al., 1996, 1998a, 2000; Beltran et al., 2010; Arthur et al., 1998; Lebo and Herrmann, 1998). Bliss and Comerford (2002) identified a decrease in water table right after harvesting and later an increase in water for about four months following harvesting of flat woods in Florida. Arthur et al. (1998), using a paired watershed approach to quantify impacts associated with tree harvesting and best management practices on water yield and water quality, reported increased water yield in the year following clear cutting of woody species in a Kentucky forest. Xu et al. (2002) attributed the rise of water table following vegetation removal to reduced transpiration because impacts of harvesting on forested watershed hydrology were more pronounced during the growing season. However, Laurén et al. (2009) demonstrated how uncertainty in pre-treatment data and thus the calibration relationship of paired watershed studies may influence estimates of the magnitude and duration of the treatment effects. Their monitoring of phosphorous loads on two independent paired catchments in Finland before and after clear-cutting showed that small treatment effects may be masked by uncertainty of the pre-treatment data.

In this study, a field experiment was set up in early 2009 to quantify watershed-scale effects of managing pine forests intercropped with switchgrass on poorly drained soils in Carteret County, NC on hydrology (i.e. water table elevation and flow) of the treatment watersheds using a paired watershed approach. Accordingly, this study required initial silvicultural operations to establish switchgrass growth during the period when

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monitoring was going on for calibration. These silvicultural operations included harvest-
ing, site preparation (bedding, shearing and raking) as well as herbaceous broadleaf
control as described by Albaugh et al. (2012) in their plot-scale study. All or some of
these activities may have potential to impact the surface soil hydraulic properties and,
thus, impact the water table elevation, drainage outflow, and nutrient and sediment dy-
namics (Albaugh et al., 2012). For example, Skaggs et al. (2006) found 20–30 times
higher effective hydraulic conductivity for the top 90 cm of the Deloss fine sandy loam
soil for matured plantation forest compared to the data published by NRCS Soil Sur-
vey prior to harvest at the Carteret County, NC study site. They observed that harvest
did not appear to affect those values but site preparation for regeneration, including
bedding, reduced the effective hydraulic conductivity to values typically assumed for
this series, Developing a calibration regression relationship with data collected from
the control and treatment watersheds undergoing such disturbances violates the as-
sumption of the paired watershed approach when both watersheds should be free of
external disturbances other than the treatment to be implemented.

Therefore, the specific objective of this study was to develop regression models that
minimize effects of external factors by minimizing uncertainty in the regression param-
eters based on their structural stability. Data monitoring for the pre-treatment calibra-
tion period were collected during 2009 to 2012 during which time silvicultural prac-
tices were conducted; these may have acted as external factors, affecting pre-treatment hy-
drologic relationships between control and treatment watersheds. A change detection
technique of moving sums of recursive residuals (MOSUMS: Bauer and Hackl, 1978;
Chu et al., 1995; Zeileis et al., 2013) is used to minimize the effects of these external
factors on the uncertainty of parameters of the calibration models. Utilizing MOSUMS
approach, pre-treatment regression models were developed using a sub-set of the
2009–2012 data when the effect of external factors was minimal and the regression
coefficients were most stable and also significant.

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by manual gauge measurements in an open area. Air and soil temperatures ($^{\circ}\text{C}$), relative humidity (%), wind speed (km h^{-1}), and solar radiation (W m^{-2}) were measured and recorded on a 30 min interval using an onsite weather station. The wind speed anemometer and the relative humidity sensor were stationed at 3.4 and 2.3 m above ground, respectively. The 30 min weather data was integrated to obtain daily average which was then used to calculate the Penman–Monteith-based potential evapotranspiration for a standard grass reference (REF-ET) for the site.

An adjustable height 120° V-notch weir, located at the outlet of each watershed, measured drainage outflow by continuously recording water levels (stage) upstream and downstream of the weir where the bottom of the V-notch was approximately 100 cm below average soil surface elevation for each watershed. Automatic stage recorders were installed to make stage measurements every 12 min upstream and downstream of a V-notch weir set about 0.3 m above the bottom of outlet ditch. A pump was installed in 1990 at the roadside collector ditch downstream of all watershed outlets to minimize weir submergence during large events (Amatya et al., 1996). Flow rates ($\text{m}^3 \text{s}^{-1}$) were computed using discharge–stage relationships for non-submerged and submerged weir conditions (Brater and King, 1996). 12 min flow rates were integrated to obtain daily totals normalized by watershed area (mm day^{-1}). Calculated daily flow values greater than the capacity of the downstream culvert during large storm events were capped to 45 mm day^{-1} , the approximate culvert drainage capacity (Amatya and Skaggs, 2011). Such data were excluded from analysis of treatment effects because of uncertainty due to highly submerged conditions. Computed flow data during weir submergence are susceptible to high uncertainty (USGS, 1997; Amatya et al., 1998b). Water table elevation (WTE) is continuously recorded on an hourly basis at the front and back experimental plots of each watershed (Fig. 1). The average of the back and front WTE was assumed as the representative WTE for each watershed.

2.3 Silvicultural operational and management practices

The control watershed D2 (loblolly pine of 14–15 year old stands) was thinned (50 % reduction in basal area) in early 2009 (Fig. 2). Watersheds D0 and D1 were clear-cut by April 2009, while 85 % of D3 was harvested by November 2009 and 100 % by May 2010. Pine planting on D0 and D1 was completed by January 2010. Watershed D0 was left with a natural understory while on D1 the natural understory between the pine rows was cleared in December 2010 by shearing between the rows for switchgrass planting. Therefore, the watershed land cover conditions for D0 and D1 were similar up to December 2010. Initial broadcast of switchgrass seeds on D1 and D3 in August 2011 did not germinate due to excess wet conditions. However, the second phase of switchgrass seed broadcast between March and April 2012 yielded a much better germination resulting in a stand establishment between pine rows of D1, but coverage at D3 was still only about 15–20 %. Other management practices included broadleaf control on D1 and D3 in August 2011.

2.4 Structural stability of calibration relationships

To minimize effects of silvicultural operations as external factors on uncertainty of calibration regression models and provide a more reliable calibration with adequate length, as noted by Bren and Lane (2014), a change detection technique of moving sums (MOSUM) of recursive residuals (Bauer and Hackl, 1978; Chu et al., 1995) was used to select the longest calibration periods for each control-treatment watershed pair (D1–D2 and D3–D2). There were no management operations on D0 after pine planting. However, MOSUM test was carried out for the D0–D2 calibration model as a reference for testing false identification of structural breaks. The MOSUM is a variant of the cumulative sums (CUSUM) method (Brown et al., 1975). Both methods test the null hypothesis that the regression coefficients of a linear regression are constant over time against an alternative hypothesis that the coefficients change over time due to extraneous factors. The CUSUM and the MOSUM tests have been applied to detect temporal changes in

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areas of landuse and land cover, hydrology, and national economic indicators (Ghosh, 2009; Vogl and Lopes, 2009; Verbesselt et al., 2010, 2012; Olmo et al., 2011; Tiwari et al., 2012; Webb et al., 2012; de Jong et al., 2013). This study used the MOSUM of recursive residuals (Eq. 1) because it is more sensitive to parameters that are temporarily unstable than the CUSUM because cumulated sums become less sensitive as the number of residuals becomes larger (Chu et al., 1995).

$$M_r = \sum_{i=r-w+1}^r \frac{W_i}{\sigma}; \quad r = \rho + w, \dots, N \quad (1)$$

$$W_r = \frac{y_r - \mathbf{x}'_r b_{r-1}}{\sqrt{\left(1 + \mathbf{x}'_r (\mathbf{X}'_{r-1} \mathbf{X}_{r-1})^{-1} \mathbf{x}_r\right)}}$$

$$\sigma^2 = \frac{1}{N - \rho} \sum_{i=\rho+1}^N (W_i - \bar{W})^2 \text{ and } \bar{W} = \frac{1}{N - \rho} \sum_{i=\rho+1}^N W_i,$$

where M_r is the r th moving sum of recursive residuals with a data window size of w , ρ is the total number of regression coefficients, N is the total number of data samples, W_r is the r th recursive residual, y_r is the r th observation of the response variable, \mathbf{x}'_r is the r th row vector of the explanatory variables, b_{r-1} is the ordinary least squares estimate of parameter b using data before the r th time step, σ^2 and \bar{W} are the variance and mean estimates of the recursive residuals W_i . The \mathbf{X}_{r-1} and \mathbf{X}'_{r-1} are the $[(n-1) \times \rho]$ regressor matrix and its transpose using data before the r th time step.

The MOSUM test for change detection was implemented in R-software environment using the strucchange package (Zeileis et al., 2013), which follows a three step procedure. The first step checks for existence of structural change based on the assumption that variability of the moving sums of recursive residuals under structural stability follows a Brownian motion (a random walk) with an expected mean of zero. Therefore,

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if the MOSUM crosses the 95 % confidence boundary, then structural change is detected. For details on the technical basis and the asymptotic function of the 95 % confidence boundary, refer to Zeileis et al. (2013). When structural change is detected in the first step, then steps two and three determine the number and location of change points also known as break points or break dates. Break points and corresponding 95 % confidence intervals are estimated based on methods developed by Bai (1994, 1997) and Bai and Perron (1998, 2003a, b) and implemented by Zeileis et al. (2013). The second step determines the number of break points by minimizing the Bayesian Information Criterion. However, one can predefine the maximum number of breakpoints for a given time series. For this analysis, four breakpoints were arbitrarily assumed to correspond to potential effects of clear-cutting, shearing and bedding, pine planting, and planting of switchgrass. The third step iteratively determines the location of break points by minimizing the regression sum of squares.

Based on the above procedural implementation of MOSUM and the fact that analyses are made on moving sums, the location where the MOSUM cross the 95 % confidence boundary is not always the location of the breakpoints. Also, when the MOSUM return inside the boundary, it does not mean the relationship has regained the previous structural stability. Finally, strength of the linear relationship, size of the moving window, and number of predetermined breakpoints influence the location of the breakpoints. A data window size of 30 days was used to calculate moving sums of recursive residuals and a significance level of 5 % ($\alpha = 0.05$) was used to detect structural changes in regression coefficients. Only the water table elevation data was used to determine pairwise control-treatment calibration periods for both daily water table elevation and daily flow data because for these low-gradient artificially drained coastal plain soils, the depth of the water table is the main driver of the hydrology (Amatya et al., 1996; Skaggs et al., 2011).

2.5 Regression analysis

Based on data spanning the stable calibration periods as estimated by the MOSUM test, a non-parametric bootstrap geometric mean regression (Plotnick, 1989; Efron and Tibshirani, 1994; Elkinton et al., 1996) was used to determine the calibration regression coefficients and their respective 95 % confidence intervals. This regression approach was used because it minimizes effects of temporal autocorrelation and consistent with a recent recommendation by Bren and Lane (2014) that any paired watershed approach should address non-normality and autocorrelation of residuals. Furthermore, bootstrapped standard errors and confidence intervals are smaller than ordinary least squared errors. The geometric mean regression, also known as the reduced major axis regression is suited for paired watershed analysis because it assumes errors are associated with both dependent (treatment watershed) and the independent (control watershed) variables (Friedman et al., 2013). One thousand bootstrap samples were used to estimate regression coefficients and corresponding confidence intervals.

3 Results

3.1 Water table elevation response to rainfall

Rainfall during 2009 and 2012 were relatively similar and well distributed throughout each year with some large events in the fall. Rainfall distributions of 2010 and 2011 were also similar with most events in the fall and the winter. The driest year was 2011 with annual rainfall of 1181.7 mm and a net precipitation (difference between rainfall and REF-ET) deficit of 53 mm while 2010 was the wettest with annual rainfall of 1420.6 mm and a net precipitation surplus of 275.7 mm. Analysis of rainfall measured at each watershed shows similar rainfall distributions across all watersheds each year with slight differences in actual rainfall amount. These annual differences were similar to the rainfall observed at the weather station. However, on average, there was

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a general negative linear trend of average annual rainfall from D0 with the highest rainfall (1348.6 mm) and D3 with the lowest rainfall (1234.2 mm). A similar trend was also observed for the same site by Amatya et al. (1996).

Seasonal climatic and vegetation dynamics affected WTE (Figs. 3–5), such that the WTE dropped below 1.5 m (above sea level) in the summer with occasional large rainfall events that temporarily raised the water table to the soil surface with ponding. For example, a large area of the watershed D3 had ponded conditions as a result of large events with daily rainfall exceeding 150 mm in late September 2010 and 100 mm in mid-October 2011 on wet antecedent conditions. Vegetation effects were reflected by higher WTE on all three treatment watersheds (young pine or emergent vegetation with a shallow root system and less evaporative demand) than the control watershed (D2: 14–15 year pine with deep root system and high evaporative demand) during the growing season. However, water table response to large storm events, characterized by a rise to the soil surface, was similar among all watersheds; an observation consistent with previous studies at this site (Amatya and Skaggs, 2011; Skaggs et al., 2011). The average difference of WTE between D0 and D2 was 17.3 ± 0.9 cm ($\pm 95\%$ CI), 22.5 ± 0.8 cm between D1 and D2, and 9.1 ± 0.6 cm between D3 and D2. The above differences between control and treatment watersheds for the 2009 to 2012 pre-treatment calibration period were significant and greater than the historical differences (Ssegane et al., 2013), except between (D0 and D2) as D0 was established only in 2009.

The greater average WTE difference between D1 and D2 was due to clear-cutting of D1 in 2009 and shearing of the understory between pine-rows in December 2010 and switchgrass seed broadcast in August 2010, and April 2012 (Fig. 2). The average WTE difference between D3 and D2 was the least given the fact that harvesting at D3 was later (between November 2009 to May 2010) than at D1 (Table 1). Another reason for this difference was the high variability between the WTE of the front and back plots for D3 (soil heterogeneity) compared to similar plots for D0, D1, and D2. A similar trend was evident in the maximum single day WTE difference between the control and

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tions, the role of vegetation became less dominant compared to the tendency of one watershed to produce less flow than the other in response to these events.

Similar seasonal shifts observed on D1 and D2 were also observed on WTE and flow differences between D3 and D2 (Table 2). This seasonal shift in hydrologic behavior was further moderated by seasonal variation of climatic variables, site preparation for switchgrass seed broadcasting on D1 and D3, and weir submergence and backflow mostly on D3. For example, the weir outlets at all four watersheds were submerged (downstream stage greater than the v-notch of the weir) on 27 September to 4 October 2010 and 26 January 2011 with backflow on D3 (26 January 2011) due to failure of the downstream pump which was installed to minimize such incidences. The total incidences of submergence and backflow on all watersheds occurred on 0.80 and 0.04 % of the time, respectively, based on 12 min stage data during this time. The D3 flow did not momentarily increase compared to D2 after 85 % of harvesting on D3 in November 2009 because it was during the dormant season with minimal ET demands at the end of the harvesting, such that any observed differences in D2 and D3 flows were similar to pre-harvest conditions (D2 flow greater than D3 flow). Note that prior to its harvest D3 was a 35 year old stand compared to D2 which was a 12 year old stand thinned in early 2009. Effects of vigorous emerging vegetation with increased leaf area index (LAI), particularly in 2010, a year after harvest of D1 and D3 may have also resulted in some negative differences in some months of the growing season. Sampson et al. (2011) noted that in coastal loblolly pine stands, herbaceous and arborescent species can dominate the site LAI for many years after a harvest (followed by planting). The observed seasonal shifts require minimum uncertainty in the regression parameters of calibration equations such that any small changes in hydrologic conditions are not masked by large standard errors.

3.3 MOSUM change detection in pairwise calibration periods

Movement of MOSUM outside the 95 % confidence bounds (Fig. 6; horizontal dotted lines) is indicative of structural break in stability of regression coefficients. However,

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return of the MOSUM within the confidence bounds does not imply return to previous state of stability. The points when the MOSUM crosses the confidence bounds may not be the actual point when structural changes occurred because of the moving sums. Such points are represented by the vertical dotted lines known as break points or break dates (with corresponding 95 % confidence interval). There were no operational management practices on D0 after pine planting (Fig. 2) and the MOSUM (Fig. 6a) shows no structural break because it does not cross the 95 % confidence interval. Therefore, the calibration period for D0 and D2 WTE and flow relationships included the period after pine planting on D0 to prior to the second phase of switchgrass planting on D1 and D3. For the relationship between WTE of D1 and D2 (Fig. 6b), the first break point (between 26 and 28 March 2009) coincides with harvesting of D1 (Fig. 2 and Fig. 6), the second break point (between 1 and 10 October 2009) is one month after shearing and bedding on D1, the third break point (between 17 February and 4 March 2010) is one month after pine planting on D1, and the fourth break point (between 24 and 30 August 2012) is about 3 months after the second re-broadcast of switchgrass seed on D1. Therefore, the actual calibration period with a minimal operational disturbance for D1 and D2 WTE and flow relationships is assumed as one month after pine planting on D1 (1 March 2010) to before preparations for the second phase of switchgrass seed broadcasting (31 March 2012) with 762 days altogether. Recently, Bren and Lane (2014) concluded that more complex models other than simple linear regression, achieved good calibration after 200 days of data. So this calibration period defined by the MOSUM method should be adequate.

The first break point for D3 and D2 WTE linear relationship (between 28 and 30 April 2009) was two months after 50 % thinning of D2 (Figs. 2 and 6c) and coincided with the start of the growing season after thinning. The second break point (between 6 and 8 September 2009) was one month before 85 % harvesting of D3. The authors believe this break point was most likely a false positive because no known operational activity occurred prior to this period. The third break point (between 9 and 14 November 2009) coincided with the 85 % harvesting of D3, and the fourth break

point (between 26 October and 10 November 2011) is about 2 months after the first seeding of switchgrass on D3 (15 August 2011). Harvesting of the final 15 % of trees on D3 in May 2010 did not significantly alter the D3 and D2 WTE relationship due to emergent vegetation on previously harvested 85 % portion of the watershed. Therefore, the calibration period for D3 and D2 WTE and flow relationships was between the end of 85 % harvesting of D3 (1 December 2009) and prior to preparations for the first phase of switchgrass planting (31 July 2011), with 608 days altogether. Again this is much longer than a threshold of 200 days recently suggested by Bren and Lane (2014).

3.4 Pairwise calibration with stable regression relationships

The coefficient of determination (R^2), Nash–Sutcliffe efficiency (NSE), and the root mean squared error (RMSE) were indicative of strong, quantifiable, and predictable relationships between hydrologic responses of control and treatment watersheds (Fig. 7). For R^2 ($0 \leq R^2 \leq 1.0$), and NSE ($-\infty \leq NSE \leq 1.0$), a value of 1.0 is indicative of an optimal model. Therefore, the calibration equations show that the two hydrologic responses of the control watershed (WTE and flow) are strong predictors of similar responses at treatment watersheds and, thus can serve as significant predictors of treatment effects using a paired watershed approach.

Use of all 2009 to 2012 data (the period after harvesting to prior to switchgrass planting) gave significantly different calibration equations with relatively weaker but not significantly different regression statistics (lower R^2 , and NSE, and higher RMSE). For the WTE relationships between D1 and D2, and between D3 and D2 using all data, the slopes are not significantly different from the slopes of the MOSUM data but the intercepts were significantly different. The daily flow relationships using all data, however, gave significantly different intercepts and slopes compared to MOSUM data based calibration relationships (Fig. 7e and f). Consistently, the RMSE of the MOSUM calibration equations were smaller than those based on all data. Therefore, the statistically significant differences in WTE and flow calibration relationships based on all data compared

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wider than the upper plot (MOSUM data), indicating the reduced uncertainty in regression parameters for the stable period. The model uncertainty is also reflected in the model fit performance statistics of NSE (0.97 vs. 0.87) and RMSE (0.084 vs. 0.162 m). Similar uncertainty trends were observed in daily WTE and flow calibration equations between D3 and D2. For example, calibration relationships with a similar slope, but different intercepts, will consistently over-predict or under-predict the expected values during the post-treatment period.

5 Conclusions

Seasonal water table and flow response to rainfall were influenced by both the silvicultural management operations and vegetation on four treatment watersheds in the coastal plain of North Carolina. This study laid out protocols to develop significant, efficient, stable, and predictable calibration relationships required to quantify hydrologic effects of intercropping switchgrass and pine, and conversion of pine forest to entirely switchgrass on a watershed scale. Although the calibration data starts from 2009 (after clear-cutting of three pine forest stands, D0, D1 and D3) to 2012 (after broadcast of switchgrass seeds for its establishment in April to May 2012), a moving sums of recursive residuals (MOSUM) test was used to detect and omit the periods with instability in regression coefficients potentially due to external factors like silvicultural operations to establish switchgrass. The analysis procedure using MOSUM demonstrated three important findings. (1) Use of all 2009 to 2012 data gave significantly different calibration relationships between the control and treatment watersheds compared to use of only the data spanning the period when the linear regression coefficients of the WTE were stable (2) Use of the MOSUM test to determine calibration relationships with a data period potentially uninfluenced by external factors minimized uncertainty of calibration relationships. Length of the calibration periods determined by our analysis was also consistent with a recent finding by Bren and Lane (2014). (3) All calibration relationships were quantifiable, strong, and consistent (R^2 and NSE greater than

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0.97 for WTE and R^2 and NSE greater than 0.92 for flow), a requirement for use of paired watershed approach. Therefore, the developed calibration relationships using a bootstrap geometric mean regression are adequate (with minimum uncertainty) to test significance of hydrologic effects of intercropping pines with switchgrass or 100 % conversion of a pine stand to switchgrass on the daily WTE and flow hydrologic responses at a watershed scale. This study will also have broader implications on similar eco-hydrologic studies where significant calibration relationships that exclude periods with effects of non-treatment factors are warranted to quantify actual treatment effects using the paired watershed approach. The change detection approach using the MOSUM method should be particularly helpful to develop optimum and significant calibration relationship in studies where a long-term monitoring for the calibration period may be cost and time prohibitive.

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Table 1. Monthly flow and differences in water table elevation (WTE) and flow of D1 (1–3 year young pine to be intercropped with switchgrass) and D2 (14–15 year pine; control watershed). The left and right columns represent months during the dormant (November to April) and growing (May to October) season, respectively. *Italic rows* are months with high occurrences of weir submergence (downstream weir stage above the V-notch of the weir).

Year	Month	FLOW, mm		Difference, [D1–D2]		D2 Rain mm	Month	FLOW, mm		Difference, [D1–D2]		D2 Rain mm
		D1	D2	WTE cm	FLOW mm			D1	D2	WTE cm	FLOW mm	
Pre-harvest/Post-harvest of D1												
2009	Nov	89	90	27	–1	209	May	11	0	47	11	81
	Dec	77	53	27	24	180	Jun	0	0	66	0	100
	Jan	8	16	15	–8	53	Jul	0	0	77	0	117
	Feb	3	12	16	–10	69	Aug	39	14	58	25	244
	Mar	5	24	23	–19	48	Sep	138	135	32	3	287
	Apr	18	22	42	–4	88	Oct	12	8	27	4	75
Post-harvest of D1												
2010	Nov	0	0	9	0	45	May	0	0	17	0	33
	Dec	7	11	11	–3	82	Jun	0	0	22	0	47
	Jan	96	139	17	–43	177	Jul	0	0	26	0	111
	Feb	80	66	24	15	111	Aug	0	0	31	0	175
	Mar	42	41	15	1	138	<i>Sep</i>	<i>42</i>	<i>49</i>	<i>21</i>	<i>–7</i>	<i>421</i>
	Apr	13	17	14	–4	47	<i>Oct</i>	<i>98</i>	<i>112</i>	<i>13</i>	<i>–13</i>	<i>25</i>
2011	Nov	7	6	15	1	68	May	0	0	19	0	17
	Dec	2	3	13	–1	37	Jun	0	0	22	0	67
	<i>Jan</i>	<i>68</i>	<i>78</i>	<i>13</i>	<i>–9</i>	<i>147</i>	Jul	0	0	28	0	50
	Feb	54	58	18	–4	99	Aug	59	18	24	41	256
	Mar	8	11	9	–4	83	Sep	53	39	30	14	185
	Apr	3	5	12	–2	38	Oct	60	46	23	13	131
2012	Nov	0	1	12	–1	20	May	0	0	16	0	130
	Dec	52	45	14	7	159	Jun	1	1	19	1	52
	Jan	20	19	13	1	62	Jul	0	0	25	0	172
	Feb	11	12	10	–1	65	Aug	62	68	25	–6	318
	Mar	23	24	10	–2	84	Sep	40	44	15	–4	132
	Apr	11	11	10	1	96	Oct	2	1	10	1	110

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Table 2. Monthly differences in water table elevation (WTE) and flow of D3 (clear-cut for planting switchgrass only) and D2 (14–15 year pine; control watershed). The left and right columns represent months during the dormant (November–April) and growing (May–October) season, respectively. *Italic rows* are months with high occurrences of weir submergence (downstream weir stage above the V-notch of the weir) and **bold rows** are months when backflow (downstream weir stage greater than upstream weir stage) occurred on D3.

Year	Month	FLOW, mm		Difference, [D3–D2]		D2 Rain mm	Month	FLOW, mm		Difference, [D3–D2]		D2 Rain mm
		D2	D3	WTE cm	FLOW mm			D2	D3	WTE cm	FLOW mm	
Pre-harvest/Post-harvest of D3												
2009	Nov	90	86	15	–4	209	May	0	0	–14	0	81
	Dec	53	51	23	–2	180	Jun	0	0	–19	0	100
	Jan	16	15	13	–1	53	Jul	0	0	–23	0	117
	Feb	12	12	15	–1	69	Aug	14	13	–24	–1	244
	Mar	24	23	14	–1	48	Sep	135	130	–3	–5	287
	Apr	22	21	12	–1	88	Oct	8	8	4	0	75
Post-harvest of D3												
2010	Nov	0	0	14	0	45	May	0	0	24	0	33
	Dec	11	8	14	–3	82	Jun	0	0	19	0	47
	Jan	139	30	16	–109	177	Jul	0	0	18	0	111
	Feb	66	26	18	–40	111	Aug	0	0	16	0	175
	Mar	41	25	11	–17	138	Sep	49	40	19	–9	421
	Apr	17	16	11	–1	47	Oct	112	94	16	–17	25
2011	Nov	6	11	9	6	68	May	0	0	18	0	17
	Dec	3	7	–8	3	37	Jun	0	0	12	0	67
	Jan	78	56	10	–22	147	Jul	0	0	6	0	50
	Feb	58	52	17	–5	99	Aug	18	26	9	8	256
	Mar	11	8	13	–4	83	Sep	39	57	22	18	185
	Apr	5	5	17	0	38	Oct	46	55	17	9	131
2012	Nov	1	7	10	6	20	May	0	3	17	3	130
	Dec	45	57	9	13	159	Jun	1	5	12	4	52
	Jan	19	17	8	–2	62	Jul	0	0	5	0	172
	Feb	12	13	1	2	65	Aug	68	88	0	20	318
	Mar	24	27	4	3	84	Sep	44	47	8	3	132
	Apr	11	5	5	–6	96	Oct	1	8	10	7	110

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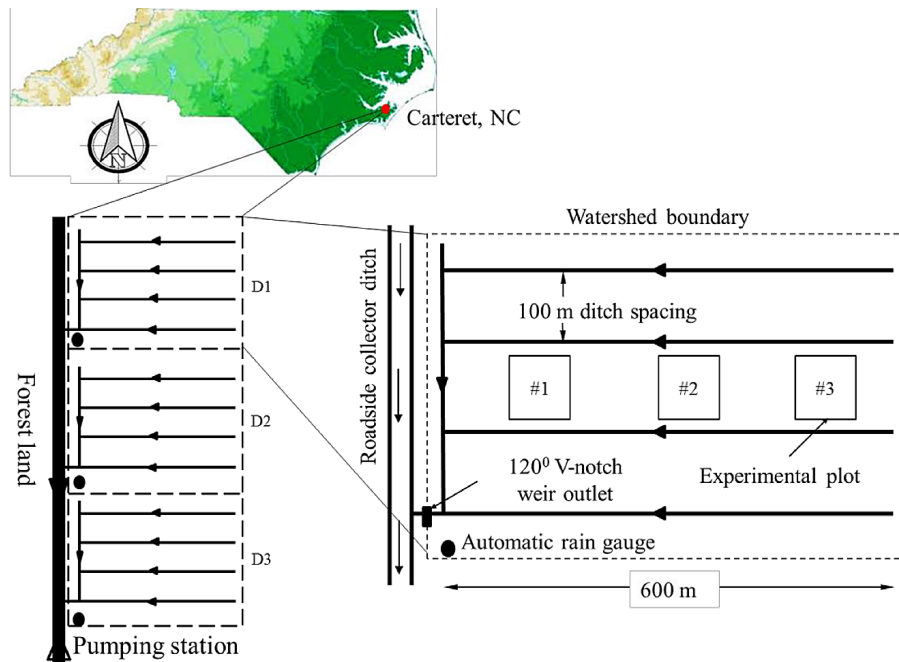


Figure 1. Location and layout of artificially drained pine-forest experimental watersheds with monitoring stations, Carteret County, North Carolina. The water table elevation (WTE) was monitored in plots #1 and #3.

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WS	Operational practice	comments	2009					2010					2011					2012																						
			J	F	M	J	J	A	S	O	N	D	J	F	M	J	J	A	S	O	N	D	J	F	M	J	J	A	S	O	N	D	J	F	M	J	J	A	S	O
D0	Harvesting	Jan-05 to April-01	█																																					
	Shearing & Bedding	Jun-13 to Jul-30							█																															
	Pine planting	Jan-18 to Jan-20												█																										
D1	Harvesting	Jan-01 to April-29	█																																					
	Shearing & Bedding	Jun-13 to Aug-07																																						
	Pine planting	Jan-18 to Jan-20																																						
	Shearing between rows	Dec-16 to Dec-17																																						
	Broad leaf control	Aug-11, 2011																																						
	Switchgrass seeding	Aug-15, 2011																																						
Switchgrass re-seeding	Apr-09, 2012																																							
D2	Thinning	End: Jan-22, 2009	█																																					
D3	85 % - Harvesting	Oct-19 to Nov-30																																						
	100 % - Harvesting	May-01 to May-10																																						
	Shearing and Raking	Apr-12 to Apr-19																																						
	Broad leaf control	Aug-11, 2011																																						
	Switchgrass seeding	Aug-15, 2011																																						
	Switchgrass re-seeding	Apr-09, 2012																																						

Figure 2. Timeline of operational management practices.

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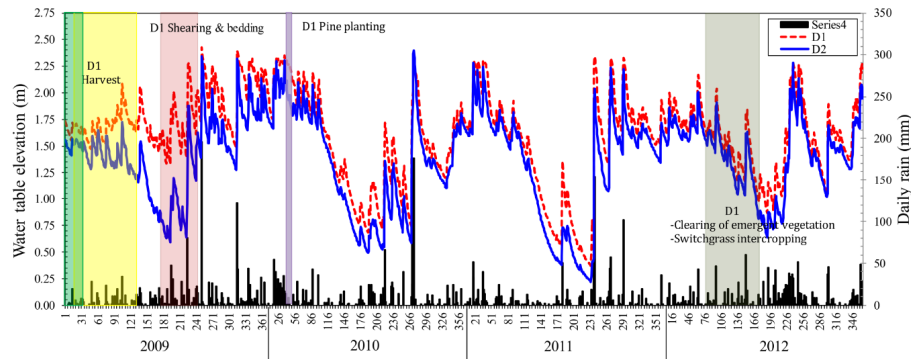


Figure 3. Daily average water table elevation (WTE) for D0 (young pine of 1–3 year old stands with natural understory) and D2 (control: pine of 14–15 year old stands). Daily rain is average of rain on watersheds D0, D1, D2, and D3.

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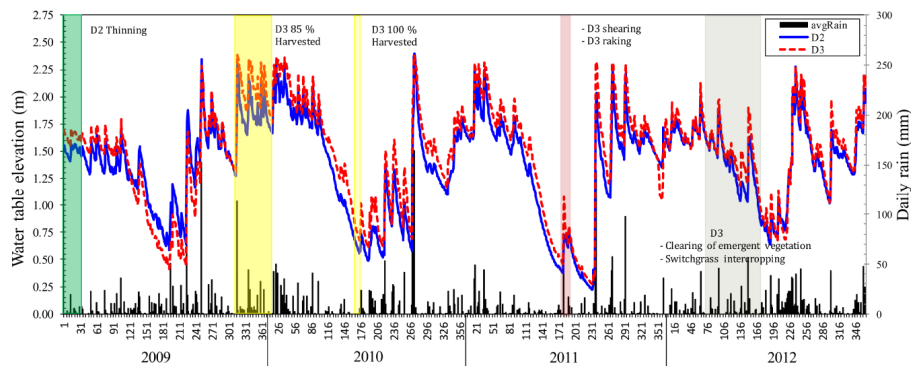


Figure 4. Daily average water table elevation (WTE) for D1 (young pine of 1–3 year old stands where spacing between pine rows were intercropped with switchgrass) and D2 (control: pine of 14–15 year old stands). Daily rain is average of rain on watersheds D0, D1, D2, and D3.

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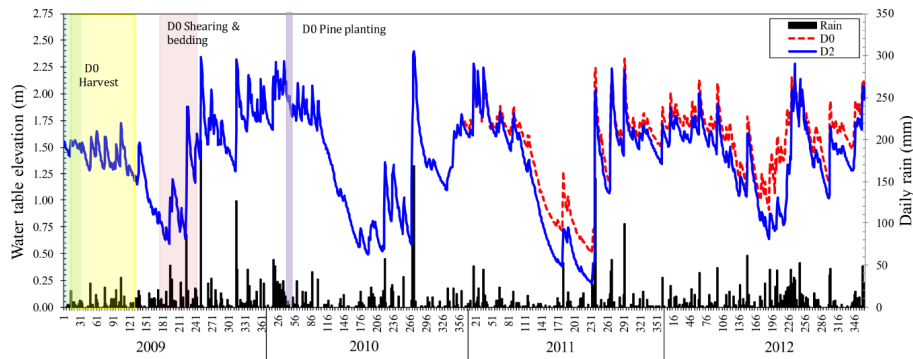


Figure 5. Daily average water table elevation (WTE) for D3 (clear-cut with emergent vegetation and planted with switchgrass only) and D2 (control: pine of 14–15 year old stands). Daily rain is average of rain on watersheds D0, D1, D2, and D3.

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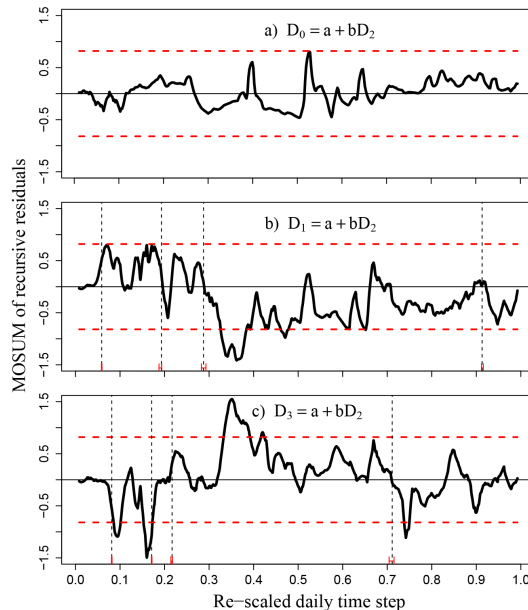


Figure 6. Graphs of the moving sums (MOSUM) of recursive residuals for the linear relationships between the water table elevation (WTE) of **(a)** D0 and D2, **(b)** D1 and D2, and **(c)** D3 and D2. A shift of the MOSUM outside the 95 % confidence intervals (long horizontal dotted lines) is indicative of a structural break in the linear relationship. The vertical dotted lines are estimated breakpoints (break dates). The corresponding small horizontal lines that cross each break date are the respective 95 % confidence intervals for each break date. Because the analysis is on moving sums, the location where the MOSUM cross the 95 % confidence boundary is not always the location of the breakpoints. Also, when the MOSUM return inside the boundary, it does not mean the relationship has regained the previous structural stability. No structural break on D0–D2 but there are structural breaks on D1–D2 and D3–D2 WTE linear regression models.

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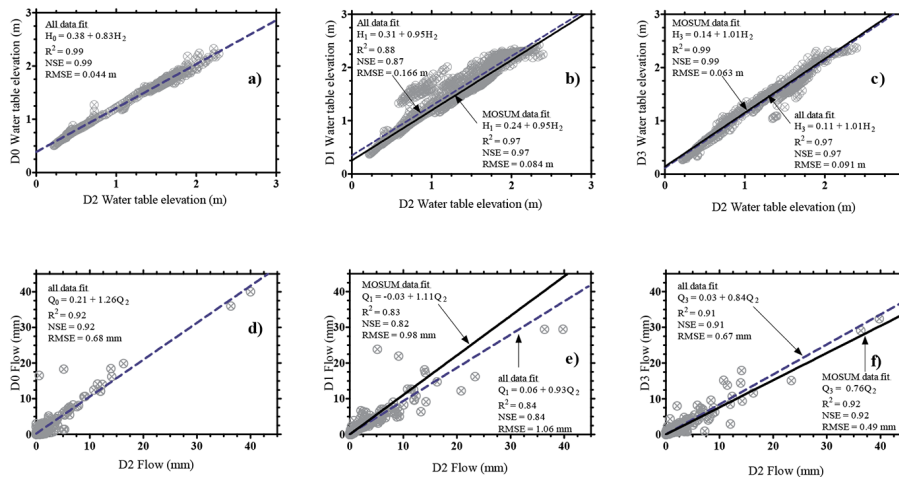


Figure 7. Calibration relationships between the daily water table elevation (WTE, **a–c**) and daily flow (**d–f**) of control (D2) and treatment (D0, D1, and D3) watersheds. All relationships are based on data period with no structural changes in the linear regression coefficients of the respective water table elevations (MOSUM data) and data from harvesting on D1 and D3 to second phase of switchgrass planting (all data).

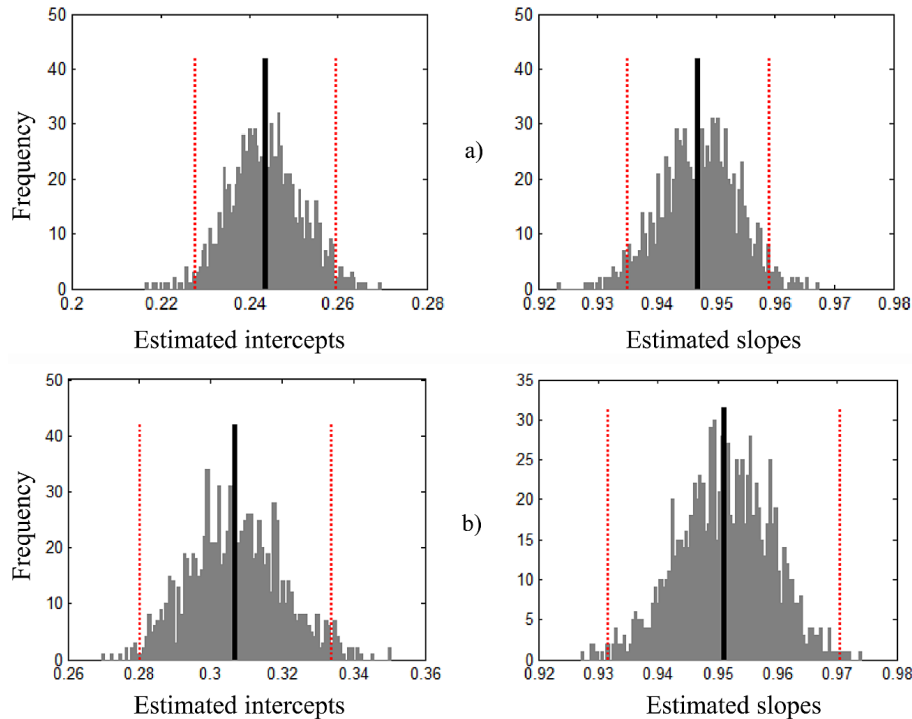


Figure 8. Uncertainty of the regression coefficients for the water table elevation (WTE) calibration equations between D1 and D2. The top graphs **(a)** are the intercept and slope (and corresponding 95% confidence interval; dotted vertical lines) using data only from the stable period of 1 March 2010 to 31 March 2012, as determined by moving sums of recursive residuals (MOSUM). The bottom graphs **(b)** are based on all data after harvesting (1 May 2009 to 31 March 2012). Distribution of coefficients is based on 1000 bootstrap resamples and a bin size of 100. The uncertainty bands of the lower plot (all data) are wider than the upper plot (MOSUM data) indicating the reduced uncertainty in regression parameters for the stable period.

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