

G. Blöschl (Referee)

bloeschl@hydro.tuwien.ac.at

Received and published: 19 December 2015

The derived distribution approach (DDA) that estimates the annual rainfall distribution assisted by subannual rainfall information (Eagleson, 1978) is tested for two rainfall stations. Resampling with shortened records indicates that the DDA gives a smaller variability between subsamples in terms of mean and standard deviation than the direct estimation of the moments.

The paper is very well written and the method is interesting. It should be noted, however, (and acknowledged in the paper) that the variability between samples the authors use as a proxy of uncertainty is actually a conditional probability - conditional on the applicability of the assumptions of the rainfall model. These assumptions provide struc-

C5717

ture and therefore reduce the degrees of freedom of the estimation (similar to fitting a distribution with fewer parameters), hence the finding of the reduced uncertainty. Similar to any parameterisation, this reduction in uncertainty is real if the assumptions apply, but spurious if they don't.

The difficulty with the resampling is that it does not test the applicability of these assumptions. Maybe this is beyond the scope of the paper, but a more comprehensive test would be to use, say, only 3 years of data from each of a large number of stations and test the DDA estimates against the respective long records. This would constitute a full blind testing. I am not sure whether the same reduction in uncertainty would result from such a comprehensive testing. For example, I cannot see how long range dependence is captured in the DDA. Long range dependence is surely important in many parts of the world and makes drought and wet periods more persistent than a random sequence. There is a substantial literature on this which is relevant to the paper.

So, overall, based on the evidence shown in the paper I would be hesitant to conclude "that the DDA, in combination with high resolution gauging, provides more accurate and less uncertain estimates of long-term precipitation statistics such as interannual variability and quantiles of annual precipitation with high return periods even for records as short as 5 years." (p. 12988). I do not think this has been demonstrated in the paper and I can't see why this would be possible, given that long term climate processes (involving, eg., ocean dynamics) are quite different from shorter term processes. A more subtle statement acknowledging the assumptions would therefore be more appropriate. Alternatively, the authors may choose to provide a full blind testing on a large number of stations.

Interactive comment on Hydrol. Earth Syst. Sci. Discuss., 12, 12987, 2015.



Anonymous Referee #2

Received and published: 29 February 2016

General comments:

The authors test a method proposed by Eagleson (1978) to estimate the probability distribution of annual rainfall from short observed records of individual storms. Especially they address the questions how record length and temporal resolution of rainfall influence the estimation accuracy. They apply the method for one rainfall station in Conception/ Chile and one station in Lugano/ Switzerland. They found, that using short records of storms provides better estimates of annual rainfall statistics than using short annual rainfall records directly.

The idea and methods are not new. The special focus on record length and temporal resolution is only partly novel. However, the revival of this idea may have important practical relevance for dealing with non-stationary time series, which cannot be used in C6899

full length for future water resources management. This message would justify publication of the article. It is a pity, that the authors only mention this in the introduction and in their conclusions, but did not investigate this in more detail e.g. using non-stationary time series. The manuscript is well written and quite compact and can be published after some revisions.

Detailed comments:

1. Introduction: There is no reference where the method of Eagleson (1978) has been applied. It would be useful for the reader to discuss a couple of applications and to compare it with the results of the authors in the conclusions. Is this really the first time Eagleson's method has been used, which would be quite surprising for me?

2. Page 12994: The data are resampled 200 times. Is this done with or without replacement and why? The former is the bootstrap the latter the permutation approach. Using the former also allows to resample the full 25-year record and to estimate it's sampling uncertainty.

3. Page 12995: Small storms < 1mm are neglected. How much do those storms contribute to annual rainfall. If this is significant, how can this be estimated using the DDA approach to obtain the real total annual rainfall sum.

4. Page 12995-12996: The fitting of distributions (exponential, gamma, normal and lognormal) should be accompanied with results from a goodness of fit test. Especially, the choice of the normal distribution should be justified by a quantitative test measure.

5. Figs. 5 to 8: In these figures the highest density for the samples of the DDA approach appears always on the left side of the peak of the full sample. This looks like a bias? Please discuss this outcome.

6. Conclusions: The limits of this approach should be mentioned, i.e. it can only be used for estimating annual rainfall sums. It should be made very clear, that for many engineering applications extreme values of short durations are required and that for

this long records are still necessary and this method is not applicable.

7. Figure 10: It is not clear what is the difference of the data for the x and y axes. Both are labelled "Annual rainfall"?

Technical corrections:

- Equations: There seems to be a problem with many equations in the pdf-file. The sum sign is not readable.

- References: There are strange numbers at the end of many references.

Interactive comment on Hydrol. Earth Syst. Sci. Discuss., 12, 12987, 2015.

C6901



C. Meier

cimeier@memphis.edu

Received and published: 29 February 2016

We agree with the referee's suggestions, and will moderate some of our conclusions, as well as acknowledge in more detail the assumptions behind the model.

We are also conducting the "full blind tests" that he recommends, using 52 high-quality Swiss stations, each with 32 years of continuously-recorded rain data (totalised every 10 min); our preliminary conclusions indicate that the Derived Distribution Approach still brings an improvement when estimating the interannual variability of annual precipitation, but on average only for records shorter than 6 years. These results will be incorporated in the revised manuscript.

Interactive comment on Hydrol. Earth Syst. Sci. Discuss., 12, 12987, 2015. C6902



C. I. Meier et al.

jorgemoraga@udec.cl

Received and published: 29 February 2016

We thank Anonymous Referee #2 for his comments, which are addressed below in the same order as they were made:

1. We are actually not aware of any previous attempt at applying Eagleson's (1978) derived distribution analysis (DDA) in the study of the inter-annual variability of annual rainfall. The paper has been cited 195 times (per Scopus) with most applications related to ecohydrological modelling of soil moisture dynamics (and vegetation), derived distributions of runoff frequency, stochastic rainfall modelling (intra-storm or else seasonal/intra-annual variability), morphological evolution of drainage basins, and applications related to floods and urban stormwater management. Surprising as it might

C6903

seem, we do think that this is the first time that the effects of record length and resolution on the DDA have been systematically addressed.

2. Please note here that it is whole years that are resampled from the available record, so as to not seasonally bias our results. For example, when resampling 7-yr long series, we take the 25 (Concepción, CL) or 32 (Lugano, CH) year-long records, and randomly sample 7 complete chronological (Jan 1st to Dec 31st) years of storms, without replacement. In this way, we construct a "new" possible 7-yr long record, for which we test the DDA and compare it with the traditional approach of fitting a pdf to annual totals. We repeat this approach 200 times at every location, for each record length (n = 3, 5, 7, 10, and 15 years). As mentioned in the manuscript, this type of re-sampling destroys the correlation structure of annual precipitation. We did test the null hypothesis that there is no temporal correlation and it was not rejected, at both sites.

3. For brevity, we did not indicate that we only neglected storms < 1 mm for the case of Concepción, but not in Lugano. This was because it was difficult to obtain the detailed storm information from paper pluviograms, for such type of events (Lugano has digital data). On average, in Concepción, the contribution of such small events to total annual rainfall is of only 4.1 mm/yr (std. deviation of 1.8 mm), so we feel this is justified. We should also mention that sadly, the original pluviograms for Concepción were lost, so that the storm characteristics for these small (<1 mm) events cannot be recovered.

4. We will indicate results from goodness of fit tests in the revised version of the manuscript, as requested by the referee.

5. Indeed, we concur that Figures 5 to 8 for the DDA (right side, plots in blue) seem to have a bias, but the reason for this is quite simple, actually: on average, the derived distributions have a very low coefficient of asymmetry, with little variability, so that distributions are almost symmetrical. Moreover, the range also has little variability. Thus, when the mode (=peak density) of each derived distribution (in blue) is smaller than the full-sample mode (either the normal in black or the derived in light green), the area

under the pdf curve at the left of the full-sample mode will have to be strictly larger than 0.5, so that the area to its right is strictly smaller than 0.5 (so that total probability is still 1). Conversely, when the peak density is larger than that for the full-sample pdfs, the reverse will happen.

6. Even though we thought this was clear enough from the title onwards, we will add a "disclaimer" to our conclusions, stressing that our results only have to do with the interannual variability of annual rainfall, and are thus not useful for the study of seasonal variation or the analyses of extreme, short-duration events.

C6905

Interactive comment on Hydrol. Earth Syst. Sci. Discuss., 12, 12987, 2015.

Manuscript prepared for Hydrol. Earth Syst. Sci. with version 5.0 of the LATEX class copernicus.cls. Date: 15 June 2016

Describing the inter-annual variability of precipitation with the derived distribution approach: Effects of record length and resolutionon the derived distribution of annual precipitation

Claudio I. Meier¹, J. Sebastián Moraga², Geri Pranzini², and Peter Molnar³

¹Department of Civil Engineering, University of Memphis, USA ²Departamento de Ingeniería Civil, Universidad de Concepción, Chile ³Institute of Environmental Engineering, ETH Zürich, Switzerland

Correspondence to: C. I. Meier (cimeier@memphis.edu), P. Molnar (molnar@ifu.baug.ethz.ch)

Abstract.

Traditional frequency analysis of annual precipitation requires the fitting of Inter-annual variability of precipitation ³⁵ is traditionally described by fitting a probability model to

- ⁵ yearly precipitation totals. There are three potential problems with this approach: a long record (at least 25~30 yrs) is required in order to fit the model, years with missing rainfall data cannot be used, and the data need to be homogeneous. To overcome, i.e., one has to assume stationarity.
- ¹⁰ To overcome some of these limitations, we test an alternative methodology proposed by Eagleson (1978), based on the derived distribution approach (DDA). This allows for better (DD) approach. It allows estimation of the pdf ⁴⁵ probability density function (pdf) of annual rainfall with-
- out requiring long records, provided that high-resolution continuously gauged precipitation data are available to derive external storm properties. The DDA-DD approach combines marginal pdfs for storm depth depths and inter-arrival 50 time to arrive at times to obtain an analytical formulation of
- the distribution of annual precipitationunder the assumption , under the simplifying assumptions of independence between events . We tested the DDA at and independence between storm depth and time to the next storm. Because 55 it is based on information about storms and not on annual
- totals, the DD can make use of information from years with incomplete data; more importantly, only a few years of rainfall measurements should suffice to estimate the parameters of the marginal pdfs, at least at locations where it 60 rains with some regularity.
- For two temperate locations in different climates (ConcepciónConcepción, Chile, and Lugano, Switzerland), quantifying we randomly resample shortened time series to

evaluate in detail the effects of record length -on the DD, comparing the results with the traditional approach of fitting a Normal (or Lognormal) distribution. Then, at the same two stations, we assess the biases introduced in the DD when using daily, totalized rainfall, instead of continuously gauged data. Finally, for randomly selected periods between 3 and 15 years in length, we conduct full blind tests at 52 high-quality gauging stations in Switzerland, analyzing the ability of the DD to estimate the long-term standard deviation of annual rainfall, as compared to direct computation from the sample of annual totals.

Our results show that, as compared to the fitting of a normal or log-normal distribution, the DDA significantly Normal or Lognormal distribution (or equivalently, direct estimation of the sample moments), the DD approach reduces the uncertainty in annual precipitation estimates (especially interannual inter-annual variability) when only short records (below $6 \sim 8$ years) are available. The DDA In such cases, it also reduces the bias in annual precipitation quantiles with high return periods. We also show demonstrate that using precipitation data aggregated every 24-hr24 h, as commonly available at most weather stations, introduces a noticeable bias in the DDA. Our DD. These results point to the tangible benefits of installing high-resolution (hourlyor less, at least) precipitation gauges at previously ungauged locations. We show that the DDA, in combination with high resolution gauging, provides more accurate and less uncertain estimates of long-term precipitation statistics such as interannual variability and quantiles of annual precipitation with high return periods even for records as short as 5 yearspropose that the DD approach is a suitable tool for the statistical description and study of annual rainfall, not just when only

2 C. I. Meier et al.: Describing the inter-annual variability of precipitation with the derived distribution approach...

120

65 short records are available, but also when dealing with non-stationary time series of precipitation.

1 Introduction

- Total annual precipitation and its variability between 70 years are important climatic variables for water bal-125 ance studies, developing regional climatologies, planning and management of water resources, and assessing water stress in general. Interannual variability in precipitation Inter-annual variability in rainfall re-
- ⁷⁵ sults from many factors of which such as long term ¹³⁰ ecohydrological modelling of soil moisture and volume at a such as ENSO, NAO, etc. (e.g., Higgins et al., 1999; Barlow et al., 2001), (ENSO, NAO, etc.; see, e.g., Higgins et al., 1999; Barlow et al., 2001), the strength and persistence of seasonality (e.g., Fatichi et
- al., 2012), and stochasticity in weather and precipitation ¹³⁵ formationin general are key. Internanual variability . Inter annual variation in precipitation is an important descriptor of the climatic environment and which directly impacts the occurrence of droughts (e.g., Dai et al., 2004; Dai, 2011),
- vegetation productivity in water-limited ecosystems (e.g., ¹⁴⁰ Knapp and Smith, 2001; Reyer et al., 2013; Fatichi and Ivanov, 2014), as well as the distribution of rainfall extremes (e.g., Groisman et al., 2005).
- A traditional statistical analysis of annual precipitation typically consists of estimating key statistics (mean, 145 variance, skewness, etc.) and fitting a probability distribution model to the annual (or seasonal) data. In temperate zones , this could be for example a normal or log-normal distribution(e.g., Markovic, 1965; Linsley et al., 1982)According
- to Markovic (1965) and Linsley et al. (1982), in temperate 150 zones this would be a Normal or a Lognormal distribution, fitted to a sample of at least 25 ~ 30 yrs 25~30 years of data. This-However, this approach is often impractical, because at many locations only a few years of precipitation data are available and many records are incomplete. With short 155
- records, the estimated statistics and parameters of the fitted probability model are highly uncertain. Moreover, natural fluctuations in climate over decadal or longer time scales, now accentuated by anthropogenic change, imply that most
- long climate records are not statistically homogeneous and stationary (Milly et al., 2008). This leads to the problem that while long records are required to accurately estimate the statistics and probability distribution of annual rainfall, precipitation itself in fact may might in fact be non-stationary 160
- ¹¹⁰ over such long periods. Thus, a methodology an approach is needed that would allow for a better estimation of the probability distribution of annual precipitation without requiring long records.

Eagleson (1978) developed such a methodology by 165 deriving the distribution of annual precipitation from the properties of the individual storms making up the yearly totals. Given independent storm arrivals and taking typical using prescribed models for the marginal probability distributions of storm interarrival inter-arrival times and storm depths, the probability density function (pdf) of annual precipitation can be derived (Eagleson, 1978) analytically. Under this derived distribution approach (DDA) (DD) approach, only a few years of high resolution continuously gauged precipitation data, from which storm arrivals , depths and durations and depths can be extracted and their distributions estimated, are necessary to arrive at an accurate estimate of estimate the probability distribution of annual precipitation for a site. Even though Eagleson's (1978) original paper has a large number of citations, most of these relate to ecohydrological modelling of soil moisture and vegetation (e.g., Dufrêne et al., 2005; Ivanov et al., 2008), dynamics derived distributions of runoff and flood frequency rainfall modelling (for example, Onof et al., 1998; Willems, 2001), or morphological evolution of drainage basins (Tucker and Bras, 2000). We are not aware of any previous attempt at applying Eagleson's DD approach to the study of the inter-annual variability of precipitation, even though the method seems particularly well-suited to deal with locations with short records, as well as to account for non-stationarities introduced by a changing climate.

The main aim of this paper work is to investigate the DDA at two temperate locations in different elimates (Concepción, Chile, and Lugano, Switzerland) and to compare it with the traditional procedure of fitting a normal or log-normal distribution to a performance of the DD approach for describing inter-annual variability of rainfall. We do this by comparing it with traditional procedures based on long series of annual precipitation depthstotals. This paper specifically addresses two questions in detail: (a) What what is the effect of record length on the estimates of annual precipitation (mean, variance deviation, and quantiles) obtained with both the different methods? (b) What is the effect of rainfall temporal resolution (sampling timetime-step) on the results? The latter question is important for sites where only daily rainfall data are available, so that the accuracy of storm statistics for the DDA-DD is reduced.

2 Methods

2.1 Study sites and data

The DDA and normalDD and Normal/log-normal Lognormal probability distributions were fitted to precipitation data for two temperate locations with dissimilar climate. For the first site in Concepción, Chile, the data eonsists come from the Bellavista Research Weather Station, which was operated by Universidad de Concepción. They consist of 19 years of daily and 6 years of weekly pluviograms (paper rain charts), continuously recorded 235

255

over the period 1975–1999, 1975–1999 with a Lambrecht float-recording and siphoning rain gauge. For the second

- site in Lugano, Switzerland, there is we used a 32 year-long
 precipitation record 1981–2012 year long precipitation
 record (1981–2012) available from MeteoSwiss (the 220
 Swiss Federal Office of Meteorology and Climatology).
 These are 10-min precipitation depthsmeasured by a
- ¹⁷⁵ Lambrecht tipping-bucket instrument The full blind tests were conducted on 52 weather stations of the MeteoSwiss network, including that in Lugano. These are all 10 min ²²⁵ precipitation depths, with 0.1 mm resolution. The Swiss , measured over the exact same period (1981-2012) with
- the same Lambrecht tipping-bucket instruments, with standardized calibration and maintenance. These rainfall data are of very good quality and have recently been used in 230 other studies of storm properties in Switzerland (e.g., Gaál et al., 2014; Molnar et al., 2015).

185 2.2 Event definition

Using derived distributions first requires defining independent storms in the record, in order to obtain the parameters needed for the marginal distributions of storm depth and interarrival-inter-arrival time. Although there are many differ-

- ent approaches for selecting a criterion of for event indepen-240 dence (e.g., Dunkerley, 2008), we chose to discriminate successive independent storms based on a Minimum Inter-event Time, MIT (Restrepo-Posada and Eagleson, 1982; Driscoll et al., 1989; Gaál et al., 2014). With this approachescheme, any
- ¹⁹⁵ dry spell (i.e., between recorded precipitation) longer than ²⁴⁵ the MIT defines two independent storms. Conversely, if a gap without precipitation is shorter than the MIT, then we assign both precipitation pulses to the same storm event (see Figure 1).

For each independent storm, we obtained the following external storm properties from the data: storm depth H, rainfall event duration T_r , time elapsed between the end of the storm and the beginning of the next storm T_b , and time between the beginning of successive storms T_a (interarrival ²⁵⁰ inter arrival time). These variables are shown in Figure 1, where independent storms are simplified into rectangular pulses, in the lower panel.

2.3 Derived distribution of annual precipitation

Eagleson (1978) defined annual precipitation (P_a) as the sum of precipitation depths over the finite number of events that occur throughout a year (or wet season). It P_a can thus be considered a compound variable which depends on the number ν of storm events in a given year (wet season), as well as on the storm depths H_i contributed by each storm:

215
$$P_a(\nu) = \sum_{j=1}^{\nu} h_j$$
 (1)

Both ν and H_j , are random variables with probability distributions that can be estimated on the basis of available, high resolution continuously gauged precipitation data. We are interested in obtaining the probabilistic behaviour behavior of the compound variable P_a , knowing the pdfs of the storm properties external properties of independent storm events.

In his work, Eagleson (1978) assumed that both the interarrival time T_a and the rainfall depth per storm H_i are identically and independently distributed (iid) variables, which are in turn mutually independent. The first assumption ("identically distributed") means that the probabilistic behaviour of these variables is time-invariant, i.e., storms behave similarly in terms of their frequency and rainfall depth, every year, and throughout the year. Although weather disturbances are much more frequent and intense in certain seasons, both in Concepción and Lugano(central south, temperate part of Chile) and in Switzerland, there are no clear limits between dry and wet seasons (rainfall events occur all year round)so. Thus, it should be fine to assume that P_a corresponds to an integration of the precipitation process at the yearly scale. Homogeneity Instead, homogeneity could also be assumed at the seasonal scale, if there were evidence for this in the data. The independence In turn, the "independently distributed" assumption implies that the characteristics of a given storm are not affected by previous rainfall events, i. e., subsequent events result from different weather systems. Finally, the assumption of mutual independence entails that storm depths are not affected by the time elapsed since the previous event, and vice-versa. Under these assumptions three assumptions, the distribution of annual rainfall is given by:

$$f_{P_a}(y) = \sum_{\nu=1}^{\infty} f_{P_a(\nu)}(y) \cdot P_{\theta}(\nu)$$
(2)

where $f_{P_a}(y)$ is the probability density corresponding to an annual rainfall of exactly y mm; $f_{P_a(\nu)}(y)$ is the probability density corresponding to a rain depth of y mm occurring in ν storms; and $P_{\theta}(\nu)$ is the discrete probability mass of having exactly ν storms in a given year.

Equation 2 represents the probability density that the sum of the rainfall depths contributed by ν annual storms adds up to exactly y mm, weighted by the discrete probability of having ν storms in that year. We followed Eagleson (1978) in modelling the occurrence of storm events as a Poisson process in order to determine the discrete probability (or probability mass) of having ν storm events over a period of length t:

$$P_{\theta}(\nu) = \frac{(\omega t)^{\nu} \cdot e^{-\omega t}}{\nu!} \qquad \nu = 0, 1, 2, \dots$$
(3)

The single parameter in this distribution ω represents the average rate of arrival or occurrence of events, whilst its

inverse ω^{-1} , corresponds to the average time elapsed between the beginning of two consecutive events. In , i.e., the

265

- mean inter-arrival time. As explained above, in our analysis t is a whole calendar year, but in semi-arid and arid climates it would could represent the duration of the wet season $_{310}$ within the year-, as described in Eagleson (1978). Note that
- mathematically, the above choice is equivalent to fitting an Exponential distribution with parameter ω to the sample of inter-arrival times.

To obtain $f_{P_a(\nu)}(y)$ in Equation 2 it is necessary to pre-315 scribe the probability distribution of precipitation depths of the iid events. For this, Eagleson (1978) chose the Gamma distribution with two parameters λ and κ , because of its versatility and its regenerative property. The latter means that the sum of n iid Gamma(λ, κ) variables also has a Gamma dis-320 tribution with parameters ($\lambda, n\kappa$) such that the mean storm

depth is then $m_H = \kappa/\lambda$ and its variance $\sigma_H^2 = \kappa/\lambda^2$. The density function of total precipitation y from ν storms, $f_{P_a(\nu)}(y)$, can then be expressed as:

-

325

$$f_{P_a(\nu)}(y) = \frac{\lambda \cdot (\lambda y)^{\nu \kappa - 1} \cdot e^{-\lambda y}}{\Gamma(\nu \kappa)} \qquad y > 0 \tag{4}$$

where $\Gamma(x)$ is the gamma function.

Replacing the expressions for $P_{\theta}(\nu)$ and $f_{P_a(\nu)}(y)$ in ³³⁰ Equation 2 yields the probability density function of annual precipitation as (Eagleson, 1978):

$$f_{P_a}(y) = \sum_{\nu=1}^{\infty} \frac{\lambda \cdot (\lambda y)^{\nu \kappa - 1} \cdot e^{-\lambda y}}{\Gamma(\nu \kappa)} \cdot \frac{(\omega t)^{\nu} \cdot e^{-\omega t}}{\nu!} \qquad y > 0$$
³³
(5)

Integrating Equation 5 results in the cumulative distribution function (cdf) for annual precipitation (Eagleson, 1978):³⁴⁰

$$F_{P_a}(y) = e^{-\omega t} + \sum_{\nu=1}^{\infty} \frac{(\omega t)^{\nu} \cdot e^{-\omega t}}{\nu!} P[\nu \kappa, \lambda y] \qquad y \ge 0 \quad (6)$$

where $P[\nu\kappa,\lambda y]$ is Pearson's incomplete Gamma function.

In concluding this section, we should briefly discuss the choice of the models for the marginal distributions. Having rainfall data at a given location for a specific period, one ³⁵⁰ could certainly use the pdfs of best fit, instead of prescribing the Exponential and Gamma distributions for the storm inter-arrival times and depths, respectively. In such case, though, it would be highly improbable that a closed-form

- solution for $f_{P_{\alpha}}(y)$ could be found, so that numerical ³⁵⁵ methods would be needed. As we are primarily interested here in testing through comparisons the general ability of the DD approach for describing the inter-annual variability of
- precipitation, we stick to Eagleson's (1978) formulation, as
 described above. 360

2.4 Performance of the DDA

If long records are available, e.g., at least $25 \sim 30$ yrs, then key statistics and as well as the distribution of annual precipitation totals can be estimated from such data. For example, in temperate, humid areas, annual precipitation typically follows a normal or a log-normal distribution Normal or a Lognormal distribution (e.g., Linsley et al., 1982; Markovic, 1965), which can be fitted to data (e.g., Linsley et al., 1982; Markovic, 1965)the yearly totals. On the other hand, the DDA summarised DD approach summarized above allows for the description of annual precipitation based only on storm statistics (with parameters representing the mean number of storms in a year, as well as the mean and variance of event depth). At locations with sufficient storms, such statistics can be adequately estimated from much shorter records, which can be estimated from short recordsprovided these have a sufficiently high temporal resolution temporal resolution that is detailed enough to accurately describe event properties. We assess the performance of the DDA versus DD method vs. traditional model fitting to annual precipitation data with two important effects in mind.

First is the effect of record length, i.e., the uncertainty and bias which <u>comes</u> come from using short precipitation records both for the traditional and derived distribution approaches. To this end We address this issue using two different methods:

(i) After testing for one-year lag independence, we randomly subsample shorter records from the original series in Concepción Concepción and Lugano, without replacement, to which we apply both methods the DD and the traditional fitting of a distribution. Instead of using independent continuous records (Pranzini, 2000; Pranzini and Meier, 2001)(as done by Pranzini, 2000; P which would yield only a handful of subsamples, the analyses were are carried out after resampling assembling 200 *n*-year datasets, where *n*-year-long resampled records, where n is the number of different, randomly picked years that constitute each subsample (n = 3, 5, 7, 10, 15) (n = 3, 5, 7, 10, 15 years). This For example, one of the 200 5-yr-long resampled datasets for Concepción was assembled with the rainfall data for the years 1988-1991-1977-1981-1994. At each of the two locations we thus consider 200 shorter, resampled records for each one of the five durations n, on top of the full (25 or 32 year-long) original time-series. Next, for each one of the 1001 different records at each site, we identify all independent storm events in the series, extract their inter-arrival times and total precipitation depths, and then fit the respective Exponential and Gamma distributions, thus obtaining the DD of annual precipitation. We also fit Normal and Lognormal pdfs to each record, and then compare various statistics (mean, standard deviation, and skewness) and quantiles to assess the performance of the proposed DD versus the "traditional" methodology. This

procedure allows us to generate enough subsamples to draw statistically significant conclusions from the results. It is important to note though, that the resampling destroys any long-range dependency that could be present in the original record. In this method, replication (200 x) is achieved 420

365 record. In this method, replication (200 x) is achieved 420 through resampling.

Second (ii) In order to provide a more realistic setting but still allow for statistical comparisons, we also conduct full blind tests of the DD approach, at 52 different locations in

- ³⁷⁰ Switzerland, each with the same 32 years of high-quality rainfall data. In these, we analyze the ability of the DD to estimate the long-term (32-yr) standard deviation of annual rainfall, as compared to direct computation from the sample of annual totals, when only a shorter, continuous record is
- available. The tests are done for all possible short record durations N between 3 and 15 years (i.e., N = 3, 4, 5, ..., 14, 15 years). For each value of N, at each one of the 52 stations, we randomly choose a N-yr long, continuous record; for example, for N = 5 years at Genève-Cointrin,
- we randomly selected the record between January 1st, 1996, and December 31st, 2000. We next fit the DD to this short record and compute the standard deviations as obtained both from the DD and from the sample of size 5 years, and then compare them with the long-term (1981-2012) deviation at
- the station by computing their relative errors. In this second method, replication (52 x) is achieved by considering a large number of stations.

The second focus of our work is the effect of data resolution on the DD. At many locations precipitation is observed

- only once a dayand higher-resolution, continuously-gauged
 so that higher resolution, continuously gauged records are not available. Thus, it is interesting to test how applicable the DDA-DD approach is when using such low-resolution, daily data. To this end we aggregated aggregate the continuously
- gauged data for Concepción and Lugano every 24 hours (between 8h (between 08:00 LT in a given day and 808:00 next day) LT next day, as is commonly done in meteorological practice). When decreasing the data resolution, the MIT was is accordingly changed to 1 day, for both Concepción and Luganoat both locations, to accommodate the minimum

identifiable dry spell under the new scenario.

In summary, we worked with 2 types of precipitation data: continuously-recorded and aggregated every 24 hours. We also considered 5 shorter record durations ranging from

- ⁴⁰⁵ 3 to 15 years, on top of the full (25 or 32 year long) time-series. For each combination of data type and series length, we identified all storm events in the series, extracted their external characteristics, and then fitted exponential and Gamma distributions to their interarrival times and total ⁴⁶⁰
- ⁴¹⁰ precipitation depths, respectively. In this way, we obtained the derived distribution of annual precipitation for all of the datasets, at both study sites. We also fit normal and log-normal pdfs to the same datasets (the "traditional" approach), and then conducted two comparisons: (i) between
- 415 the traditional methodology and the proposed DDA, and

(ii) within the DDAIn this case we are interested in the differences within the DD method, between using continuously-gauged rainfall data and daily data, considering the same 200 n-year-long (n = 3, 5, 7, 10, 15 years) shortened records that were previously assembled.

3 Results

3.1 Event properties

Storm events in Concepción Concepción are dominated by mid-latitude extra-tropical cyclones, which produce fronts resulting in low to mid-intensity rainfall events that occur throughout the year, but with higher frequency and magnitude during the winter months (Falvey and Garreaud, 2007). When using a 12 hour h MIT in order to discriminate independent storms, we obtained obtain a total of 1350 rainfall events over the 1975–1999 period, neglecting 1975–1999 period. This count neglects storms with a total depth below 1 mm. In this case, because of the difficulties involved in extracting the properties for such small events from paper pluviograms. These discarded trace events amount on average for 4.1 mm per year, whereas mean annual rainfall in Concepción is around 1200 mm, so we neglect this source of bias. In Concepción, the Gamma distribution and Exponential distributions fit the data adequately, as visually shown in Fig. 2. In the case of the storm depths, we fit a Gamma with parameters $\lambda = 0.02784 \text{ mm}^{-1}$, and $\kappa = 0.6157$, fits the storm depth data very well. Likewise, an exponential ($\chi^2 = 58.5$, df = 10, p = 7 * 10⁻⁹). For the inter-arrival times, we choose an Exponential distribution with parameter $\omega = 0.006261 \text{ h}^{-1}$ adequately fits the interarrival times ($\chi^2 = 239.7$, df = 10, p = 0). Both marginal distributions are shown in Figure 2.

In Lugano, storms are dominated by local-scale convective systems in the summer and fall, which produce mid to high intensity showers during a few hourswhich can be, accentuated by orographic effects (Schiesser et al., 1995; Panziera et al., 2014). Consequently, a 4 hour h MIT was chosen, resulting in 1794 independent storms over the 1981–2012 period, neglecting storms with a total depth below 1 mm1981–2012 period. Storm depths are fitted with a Gamma distribution with parameters $\lambda = 0.0255 \text{ mm}^{-1}$, and $\kappa = 0.3281$, while the interarrival ($\chi^2 = 57.2$, df = 11, p = 3 * 10⁻⁸), whils the inter-arrival times are fitted with an exponential distribution with parameter $\omega = 0.0136 \text{ h}^{-1}$ ($\chi^2 = 541.5$, df = 11, p = 0), as shown in Figure 3.

3.2 Effect of record length

3.2.1 Resampling tests

At Concepción and Lugano, for the <u>entire complete</u> records, we obtained the <u>pdfs cumulative distribution functions</u> of annual precipitation using the <u>DDADD</u>, based on the marginal

- distributions for storm counts and depths, and compared 520 them with fitting normal and log-normal fitted Normal and Lognormal distributions (Figure 4). At both locations, we find that there are no significant differences between the three different distributions, at least in the range where
- the bulk of the data lies. We found that generally normal 525 and log-normal distributions yielded. In general, the Normal and Lognormal distributions yield very similar results. For thisreason,.; because of this, and considering that our focus is on the improvement that can be had by using the DD
- ⁴⁷⁵ approach, we omit the log-normal Lognormal results from ⁵³⁰ the rest of this paper. Moreover, it is basically impossible to reject any specific model for the shorter samples (3 to 15 years) that we are interested in.

The sampling of shorter records was conducted on an annual basis in order to maintain seasonal coherence within

- ⁴⁸⁰ nual basis in order to maintain seasonal coherence within a given year, i.e., years were selected at random and all storms within those years were sampled. This As this sam- 535 pling destroys any correlation in precipitation between years, if it exists. We we verified the lack of temporal correla-
- ⁴⁸⁵ tion in annual precipitation using Kendall's τ statistic (Ferguson et al., 2000) for lag-1 autocorrelation. In Concepción, Kendall's $\tau = -0.1522$ (p-value = 0.3128) for n = 25 years. ⁵⁴⁰ For Lugano Kendall's $\tau = 0.0151$ (p-value = 0.9195) for n = 32 years. The Thus, the null hypothesis that there is
- ⁴⁹⁰ no lag-1 autocorrelation is not rejected in either case. This justifies the sampling of storms into gives some support to the decision of resampling shorter records based on whole 545 years.

The consequences of reducing record lengths to n = 5 and

- n = 10 years are shown in Figures 5 and 6 for Concepción, and Figures 7 and 8 for Lugano, respectively. At both locations the results show a clear increase in dispersion of the 550 pdfs of annual precipitation with shorter records, but the variability is significantly lower when using the **DDADD**. The re-
- sults for other subsample sizes (n = 3, 7, 15) show the same tendency. The reproduction of the standard deviation of annual precipitation in Figure 9 clearly shows that the DDA dramatically DD approach reduces the uncertainty in this estimate of interannual inter-annual variability for shorter 555 record lengths.

In hydrological practice, we commonly fit distributions in order to estimate quantileswith a low frequency of occurrence but high magnitude. Because of the versatility in the shape of the derived distribution (Equation 5), the DDA 560

- 510 results in a lower bias and reduced uncertainty as compared with fitting a Normal distribution. These results are presented in Tables A1, A2, A3, and A4, where we show the values of annual precipitation for different return periods, as computed with both methods for all record lengths, us-565
- ⁵¹⁵ ing continuously recorded continuously-recorded data, both in Concepcion Concepción and Lugano. The tables show the mean, standard deviation and skewness, as well as 10 selected quantiles, obtained with the DDA and fitting a normal PDF both with the DD and by fitting a Normal pdf. These 570

statistics are computed for both the complete records and the resampled, shorter records, in which case the mean and standard deviation of 200 samples are presented.

The uncertainty in estimating the variability of annual precipitation from short records is very large if only annual totals are used. The DDA Based on the resampling tests, it would seem that the DD for the same short records , based on using event properties, significantly reduces this uncertainty. significantly reduces this uncertainty. Still, one should be careful with the results obtained from these resampled records, as they hinge on the applicability of the model's assumptions.

3.2.2 Full blind tests

These were conducted in order to test the ability of the DD approach to estimate the long-term inter-annual variability of precipitation based on short records, without the need for any assumptions. For each record length (3 to 15 years), at each one of the 52 locations, we computed the percentual relative errors between the 'true' long-term (32 yr) standard deviation of annual rainfall, and two different estimations from the short record: (i) using the DD, and (ii) direct computation from the (small-size) sample of annual totals. Figure 10 shows the mean (\pm standard deviation) and median (N=52) decrease in percentual relative error that is achieved when using the DD (positive values indicate that the DD performs better). Smoothing our results, we conclude that for these 52 Swiss stations, both in the mean and the median, the DD improves estimation of the long-term variability of annual rainfall when record length is shorter than about $6 \sim 8$ years, with a marked gain when $N \leq 4$ yrs. When records longer than 8 yrs are available, there is little difference between both estimation methods.

3.3 Effect of data resolution

Most weather stations world-wide are not equipped with continuously recording rain gauges; in such cases precipitation is usually measured only once per day. In order to test the suitability of daily rainfall data in the DDA for the DD method, we totalized our high resolution data over 24-hr 24 h long periods and then applied the DDA-DD approach to these daily data. Figure 11 shows that when we use daily instead of continuously recorded precipitation data we obtain similar values for the central tendency indicators, but the extremes of the distribution show important biases.

This occurs as well, at both locations, when combining the use of aggregated data and a shorter record length as wellshorter record lengths, as shown in Figures 12 and 13at both stations. The spread of annual precipitation increases both with when reducing the record length as well as by aggregating data when data are aggregated, which may lead to an important bias, especially in wet years. This result points to the benefits of installing high-resolution gauges to derive accurate storm statistics as well as to the limited value of short daily records for deriving reliable precipitation statistics at the annual scale, using the DD approach.

4 Discussion and conclusions

- ⁵⁷⁵ The distributions of annual precipitation obtained with ⁶³⁰ Eagleson's (1978) derived distribution approach , DDA, are very similar to the fitting of a normal or log-normal Normal or Lognormal distribution when using complete, continuously-recorded time series the complete,
- ⁵⁸⁰ continuously gauged record (Figure 4). Thus, whenever ⁶³⁵ a long and homogeneous (stationary) rainfall record is availableand the homogeneity issue has been dealt with, the traditional approach of fitting a normal or log-normal distribution is Normal or Lognormal distribution should be
- adequate. On the other hand, the amount of information 640
 used in the DDA DD approach is much greater, because it explicitly includes statistics from the many storm events that occur within the yearseach year with data, instead of considering only annual sums. This allows for a much better
 estimation of the three parameters of the model and reduces 645

the uncertainty in the estimates.

More importantly, the DDA-<u>DD</u> method still yields good results when attempting to estimate the annual rainfall distribution with shorter records, as long as these consist of

- continuous data, as shown in Figures 5 and 6continuously 650 gauged data. Shorter records yield larger variability in annual precipitation, but using the DD, measures of both the central tendency and dispersion are still more consistent with those estimated using all of the information available. This is be-
- cause even if, say only 3 years of data are available, there 655 is still a sufficiently large number of rainfall events (e.g., an average of 54 yearly storms for Concepción Concepción and 119 for Lugano) to allow for a good an adequate probabilistic description of their external characteristics.
- On the other hand, when one attempts to fit a normal distribution to only a few annual rainfall data points distribution (Figures 5and 6) -8), or estimate sample moments (Figures 9 and 10) with only a few annual rainfall totals, there is a large uncertainty in the estimates. Indeed,
- 610 interannual variability is estimated better with only 3 or 5 years of continuously recorded data by DDA than it is with 665 much longer annual records only (Figure 9). Furthermore, years with incomplete records may still be used with DDA the DD method in order to extract storm properties and esti-
- ⁶¹⁵ mate model parameters, while for fitting a distribution to annual totals only complete years can be used. Results requires ⁶⁷⁰ years with complete data. However, our results also show that a bias is introduced when only daily rainfall records are available , however, a bias is introduced (Figures 11, 12 and 11, 12 and 12).
- ⁶²⁰ 13). Therefore, -13), so that the use of low-resolution data for DDA-rainfall data cannot be recommended for the DD method.

Overall, our results show that Eagleson'Eagleson's derived distribution approach is a better way of estimating the probability distribution of annual precipitation, especially when only a short, high-resolution record is available, because the uncertainty in estimates is reduced. The importance of these results lies not only in the possibility of estimating annual rainfall and its variability when short time series are available; it also allows to study the only short records are available. When there is suspicion of non-stationarity in a rainfall record, the DD method should be useful for describing the long-term behaviour of annual precipitation (even when long time series are if a long series is available) by using shorter records that guarantee homogeneity of the databreaking the longer record into shorter series over which it is more tenable to assume stationarity. In turn, one could also think of using the DD approach as part of a test for homogeneity of rainfall records, under the basic assumption that if annual rainfall is showing trends, these should be reflected in event frequency and in the distribution of storm depths.

An important conclusion of this work is that installing high-resolution (hourly or less) precipitation gauges in previously ungauged locations, even for short periods, has tangible benefits for the estimates in the estimation of long-term precipitation statistics, such as interannual inter-annual variability and quantiles of annual precipitation with high return periodsat those locations. This is important because accurate gauge-level precipitation estimates remain vital for the correction of remotely sensed data and in merging different precipitation data types, e.g., weather radar, satellite, etc. (e.g., Xie and Arkin, 1996), and as well as for the spatial interpolation of precipitation, especially in areas with complex topography (e.g., Masson and Frei, 2014).

Acknowledgements. The ETH Global R4D Seed Money Grants scheme is acknowledged for providing financial support for this research. The publication of this manuscript has been funded through ETH Zürich membership in the open-access Copernicus Publishers. Conicyt (Chile) funded C. Meier and J. Moraga through project Fondef CA13I10190. We thank the Swiss Federal Office of Meteorology and Climatology, MeteoSwiss, provided for providing the rainfall data for Switzerland. We thank Mr. Freddy Echeverría and Prof., and Freddy Echeverría and Alberto Foppiano, Departamento de Física de la Atmósfera y el OcéanoFísica de la Atmósfera y el Océano, Universidad de Concepción, who gave us the precipitation data for the Bellavista weather station. Prof. Glenn Hofmann, Departamento de Estadística, Universidad de Concepción, helped improve the Concepción, for those for Concepción. Comments from Günter Blöschl and an anonymous reviewer much improved our original manuscript.

References

Barlow, M., Nigam, S., and Berbery, E.H.: ENSO, Pacific decadal variability, and US summertime precipitation, drought, and

8 C. I. Meier et al.: Describing the inter-annual variability of precipitation with the derived distribution approach...

streamflow, J. Climate, 14(9), 2105–2128, 2001.

- Benjamin, J.R., and Cornell, C.A.: Probability, Statistics, and Deci-735 sion for Civil Engineers, McGraw-Hill, New York, 1970.
 Dai, A. G.: Drought under global warming: a review, Wiley Interdisc. Rev. Climate Change, 2(1), 45–65, 2011.
- Dai, A. G., Trenberth, K.E., and Qian, T.T.,: A global dataset of Palmer Drought Severity Index for 1870-2002: Rela-740 tionship with soil moisture and effects of surface warming, J. <u>HydrometeorologyHydrometeorol.</u>, 5(6), 1117–1130, doi:10.1175/JHM-386.1, 2004.
- 685 Díaz-Granados, M.A., Valdés, J.B., Bras, R.L.: A physically based flood frequency distribution. Water Resources Research, 20 (7), 745 pp. 995-1002, 1984.
 - Driscoll, E. D., Palhegyi, G. E., Strecker, E. W., and Schelley, P. E.: Analysis of storm event characteristics for selected rainfall
- gages throughout the United States, U.S. Environmental Protection Agency, Washington D.C., 1989.
 Dufrêne, E., Davi, H., François, C., Le Maire, G., Le Dantec, V.,
- Granier, A.: Modelling carbon and water cycles in a beech forest. Part I: Model description and uncertainty analysis on modelled NEE . Ecological Modelling, 185 (2-4), pp. 407-436, 2005.
- Dunkerley, D.: Identifying individual rain events from pluviograph 755 records: a review with analysis of data from an Australian dryland site, Hydrol. Process., 22(26), 5024–5036, 2008.
- Eagleson, P.S.: Climate, soil, and vegetation. 2. The distribution of annual precipitation derived from observed storm sequences, Water. Resour. Res., 14(5), 713–721, 1978.
 - Falvey, M., and Garreaud, R.: Wintertime precipitation episodes in central Chile: Associated meteorological conditions and orographic influences, J. Hydrometeorol., 8, 171–193, doi:10.1175/jhm562.1, 2007.
- Fatichi, S., and Ivanov, V.: Interannual variability of 765 evapotranspiration and vegetation productivity, Water Resour. Res., 50, 3275–3294, doi:10.1002/2013WR015044, 2014.

705

- Fatichi, S., Ivanov, V. Y., and Caporali, E.: Investigating interannual variability of precipitation at the global scale: Is there a connection with seasonality? J. Climate, 25, 5512–5523, 2012.
 Fatichi, S., and Ivanov, V.: Interannual variability of evapotranspiration and vegetation productivity, Water Resour. Res., 50, 3275–3294, doi:10.1002/2013WR015044, 2014.
- Ferguson, T. S., Genest, C. and Hallin, M.: Kendall's tau for serial dependence, Can. J. Stat., 28, 587–604, 2000.
 Freeze, R.A.: A stochastic-conceptual analysis of rainfall-runoff processes on a hillslope. Water Resources Research, 16 (2), pp. 391-408, 1980.
- Gaál, L., Molnar, P., and Szolgay, J.: Selection of intense rainfall events based on intensity thresholds and lightning data 780 in Switzerland, Hydrol. Earth Syst. Sci., 18, 1561—1573, doi:10.5194/hess-18-1561-2014, 2014.
- Groisman, P. Y., Knight, R. W., Easterling, D. R., Karl, T. R.,
 Hegerl, G. C., and Razuvavev, V. A. N.: Trends in intense precipitation in the climate record, J. Climate, 18(9), 1326–1350, 785 2005.
- Higgins, R. W., Chen, Y., and Douglas, A. V.: Interannual variability of the North American warm season precipitation regime, J. Climate, 12(3), 653–680, 1999.
- Ivanov, V.Y., Bras, R.L., Vivoni, E.R.: Vegetation-hydrology 790 dynamics in complex terrain of semiarid areas: 1. A mechanistic approach to modeling dynamic feedbacks. Water Resources

Research, 44 (3), art. no. W03429, 2008.

- Knapp, A. K., and Smith, M. D.: Variation among biomes in temporal dynamics of aboveground primary production, Science, 291(5503), 481–484, 2001.
- Linsley, R.K., Kohler, M.A and Paulhus, J.L.H.: Hydrology for Engineers, 3rd Edition, McGraw-Hill, New York, 1982.
- Markovic, R. D.: Probability functions of best fit to distributions of annual precipitation and runoff, Hydrology Paper No 8, Colorado State University, Fort Collins, Colorado, 1965.
- Masson, D and C. Frei: Spatial analysis of precipitation in a high-mountain region: exploring methods with multi-scale topographic predictors and circulation types, Hydrol. Earth Syst. Sci., 18(11), 4543–4563, doi:10.5194/hess-18-4543-2014, 2014.
- Milly, P. C. D., Betancourt, J., Falkenmark, M., Hirsch, R. M., Kundzewicz, Z. W., Lettenmaier and D. P., Stouffer, R. J.: Stationarity is dead: Whither water management?, Science, 319, 573—574, 2008.
- Molnar, P., Fatichi, S., Gaál, L., Szolgay, J., and Burlando, P.: Storm type effects on super Clausius–Clapeyron scaling of intense rainstorm properties with air temperature, Hydrol. Earth Syst. Sci., 19, 1753–1766, doi:10.5194/hess-19-1753-2015, 2015.
- Onof, C., Mackay, N.G., Oh, L., Wheater, H.S.: An improved rainfall disaggregation technique for GCMs. Journal of Geophysical Research Atmospheres, 103 (D16), art. no. 98JD01147, pp. 19577-19586, 1998.
- Panziera, L., James, C. N. and Germann U.: Mesoscale organization and structure of orographic precipitation producing flash floods in the Lago Maggiore region, Q. J. Roy. Meteor. Soc., 141, 224– -248, doi: 10.1002/qj.2351, 2014.
- Pranzini, G.: Annual rainfall distribution in Concepción as derived from the characteristics of the storm sequence, Thesis presented for the professional degree of Civil Engineer, Depto. de Ingeniería Civil, Universidad de Concepción, Chile, 2000 (in Spanish).
- Pranzini, G. and C. I. Meier: Annual rainfall distribution in Concepción as derived from the characteristics of the storm sequence, in Proceedings of the XV Chilean Conference of Hydraulic Engineering, Concepción, 7-9 November 2001, 27–38, 2001 (in Spanish).
- Restrepo-Posada, P. J. and Eagleson, P. S.: Identification of independent rainstorms, J. Hydrol., 55, 303–319, doi:10.1016/0022-1694(82)90136-6, 1982.
- Reyer, C.P.O., Leuzinger, S., Rammig, A., Wolf, A. Bartholomeus, R.P., Bonfante, A., de Lorenzi, F., Dury, M., Gloning, P., Abou Jaoude, R., Klein, T., Kuster, T.M., Martins, M., Niedrist, G., Riccardi, M., Wohlfahrt, G., de Angelis, P., de Dato, G., Francois, L., Menzel, A., and Pereira, M.: A plant's perspective of extremes: terrestrial plant responses to changing climatic variability, Global Change Biology, 19(1) 75–89, doi:10.1111/gcb.12023, 2013.
- Schiesser, H. H., Houze, R. A. and Huntrieser, H.: The mesoscale structure of severe precipitation systems in Switzerland, Mon. Weather Rev., 123(7), 2070–2097, 1995.
- Tucker, G.E., Bras, R.L.: A stochastic approach to modeling the role of rainfall variability in drainage basin evolution. Water Resources Research, 36 (7), pp. 1953-1964, 2000.
- Willems, P.: A spatial rainfall generator for small spatial scales. Journal of Hydrology, 252 (1-4), pp. 126-144, 2001.
- Xie, P. P., and Arkin, P. A.: Analyses of global monthly precipitation using gauge observations, satellite estimates, and numerical

C. I. Meier et al.: Describing the inter-annual variability of precipitation with the derived distribution approach... 9

model predictions, J. Climate, 9(4), 840-858, 1996.



Fig. 1. External independent-storm properties obtained from the rain gauge records.



Fig. 2. Fitted distributions to rainfall depths interarrival times (left) and interarrival times rainfall depths (right) for 25 years of data in Concepción.



Fig. 3. Fitted distributions to rainfall depths interarrival times (left) and interarrival times rainfall depths (right) for 32 years of data in Lugano.



Fig. 4. Cumulative distributions derived with the DDA and fitted as a normal Normal and log-normal Lognormal distribution to annual precipitation totals for Concepción, Chile (left) and Lugano, Switzerland (right).



Fig. 5. Estimated PDFs for 10-year subsamples using Normal Distributions (red, left) and Derived Distributions (blue, right) in Concepción.



Fig. 6. Estimated PDFs for 5-year subsamples using Normal Distributions (red, left) and Derived Distributions (blue, right) in Concepción.



Fig. 7. Estimated PDFs for 10-year subsamples using Normal Distributions (red, left) and Derived Distributions (blue, right) in Lugano.



Fig. 8. Estimated PDFs for 5-year subsamples using Normal Distributions (red, left) and Derived Distributions (blue, right) in Lugano.



Fig. 9. Sample standard deviations of annual precipitation computed from yearly totals (in blue), compared to the corresponding population standard deviations estimated with the DDA (in green). For record lengths ≤ 15 yr, the whiskers show the range ± 1 std from resampling (n=200). Concepción is on the left and Lugano on the right.

Table A1. Quantiles of the distribution of annual rainfall in Concepción, as obtained with different record lengths resampled 200 times and continuous data using derived distributions (in mm).

Sample Size		Mean	STD	C1	QUANTILES										
				Skew.	Q .01	Q .02	Q .05	Q .10	Q .20	Q .80	Q .90	Q .95	Q .98	Q .99	
25 Years		1213.6	265.4	0.35	670	727	817	900	1006	1451	1579	1688	1814	1900	
15 Vacana	m	1209.8	264.9	0.35	667	725	814	897	1002	1447	1575	1683	1809	1895	
15 rears	s	39.5	8.0	0.01	27.1	28.4	30.3	32.1	34.5	45.3	48.5	51.4	54.7	57.0	
10 Years	m	1217.3	266.2	0.35	672	730	820	903	1009	1456	1584	1693	1820	1906	
	s	60.0	13.2	0.01	38.8	40.9	44.2	47.3	51.5	70.2	75.9	80.8	86.6	90.5	
7 Vaara	m	1214.0	264.8	0.35	672	729	818	902	1007	1451	1579	1687	1813	1899	
/ iears	s	82.1	17.6	0.02	53.9	56.6	61.0	65.2	70.7	95.5	103.0	109.5	117.1	122.4	
5 Vaara	m	1205.0	263.8	0.35	665	722	811	894	998	1442	1569	1677	1802	1888	
5 fears	S	99.1	21.2	0.02	66.4	69.6	74.6	79.5	85.9	114.9	123.8	131.4	140.6	146.9	
3 Years	m	1242.2	271.1	0.35	687	746	837	922	1030	1485	1616	1727	1856	1944	
	s	140.7	28.1	0.04	98.4	102.5	109.2	115.4	123.6	160.6	171.8	181.6	193.2	201.2	

 Table A2. Quantiles of the distribution of annual rainfall in Concepción, as obtained with different record lengths resampled 200 times fitting normal Normal distributions (in mm).

<u> </u>			CTTD	QUANTILES												
Sample Size		Mean	SID	Q .01	Q .02	Q .05	Q .10	Q .20	Q .80	Q .90	Q .95	Q .98	Q .99			
25 Years (DDA)		1213.6	265.4	670	727	817	900	1006	1451	1579	1688	1814	1900			
25 Years (Normal)		1194.1	269.8	566	640	750	848	967	1421	1540	1638	1748	1822			
15 Years	m	1189.8	270.2	561	635	745	843	962	1417	1536	1634	1745	1818			
	s	42.3	28.2	80.9	74.5	65.3	57.8	50.1	46.8	53.4	60.3	69.0	75.2			
10 V	m	1197.9	271.7	566	640	751	850	969	1427	1546	1645	1756	1830			
10 rears	s	63.8	42.6	123.7	113.9	99.9	88.4	76.6	69.7	79.3	89.5	102.5	111.9			
7	m	1194.9	264.4	580	652	760	856	972	1417	1534	1630	1738	1810			
/ Years	s	87.6	60.1	172.8	158.8	138.9	122.7	105.8	96.4	110.3	124.9	143.6	156.9			
5 V	m	1186.7	266.4	567	640	749	845	963	1411	1528	1625	1734	1807			
5 Years	s	104.1	75.5	215.6	197.8	172.2	151.2	128.9	114.5	132.4	151.1	174.9	191.9			
3 Years	m	1224.2	242.2	661	727	826	914	1020	1428	1535	1623	1722	1788			
	s	152.0	115.9	325.2	297.6	257.8	225.0	190.3	170.4	199.3	229.1	266.6	293.2			

Table A3. Quantiles of the distribution of annual rainfall in Lugano, as obtained with different record lengths resampled 200 times and continuous data using derived distributions (in mm).

Sample Size		Mean	STD	C1	QUANTILES											
				Skew.	Q .01	Q .02	Q .05	Q .10	Q.20	Q .80	Q .90	Q .95	Q .98	Q .99		
32 Years		1529.8	280.0	0.27	942	1006	1105	1197	1311	1786	1919	2032	2163	2252		
15 Vaara	m	1531.2	270.8	0.27	960	1023	1120	1208	1319	1778	1907	2016	2142	2227		
15 rears	S	60.4	8.3	0.03	44.1	45.9	48.8	51.4	54.7	68.9	72.9	76.3	80.3	83.1		
10 Years	m	1520.5	268.6	0.27	955	1017	1112	1200	1310	1765	1893	2001	2126	2210		
	s	71.5	10.2	0.03	52.7	54.8	58.1	61.1	64.9	81.4	86.2	90.3	95.0	98.3		
7 Vaara	m	1531.6	269.8	0.26	962	1024	1121	1210	1321	1779	1908	2017	2143	2228		
/ Teals	s	101.8	14.7	0.05	73.8	76.9	81.9	86.5	92.5	117.9	125.3	131.6	139.0	144.0		
5 Vacana	m	1532.0	268.8	0.25	964	1026	1123	1211	1322	1780	1908	2017	2142	2228		
5 years	s	121.1	16.5	0.07	89.0	92.7	98.4	103.9	110.7	140.1	148.6	155.9	164.4	170.1		
3 Years	m	1529.5	267.4	0.24	963	1025	1121	1210	1320	1778	1906	2015	2140	2225		
	s	150.9	21.4	0.09	110.2	114.9	122.2	129.2	137.9	175.8	186.8	196.3	207.2	214.8		

 Table A4. Quantiles of the distribution of annual rainfall in Lugano, as obtained with different record lengths resampled 200 times fitting normal Normal distributions (in mm).

G 1 G.		м	OTD	QUANTILES											
Sample Size		Mean	SID	Q .01	Q .02	Q .05	Q .10	Q .20	Q .80	Q .90	Q .95	Q .98	Q .99		
32 Years (DDA)		1529.8	280.0	942	1006	1105	1197	1311	1786	1919	2032	2163	2252		
32 Years (Normal)		1530.5	290.1	856	935	1053	1159	1286	1775	1902	2008	2126	2205		
15 Years	m	1530.5	280.2	879	955	1070	1171	1295	1766	1890	1991	2106	2182		
	s	56.8	43.7	101.9	92.1	78.4	68.0	58.6	75.6	90.0	103.1	118.6	129.3		
10 Vaara	m	1521.7	281.6	867	943	1058	1161	1285	1759	1883	1985	2100	2177		
10 rears	s	69.0	62.0	137.7	123.1	102.4	86.2	71.3	99.5	121.4	141.0	163.9	179.5		
7	m	1530.0	276.0	888	963	1076	1176	1298	1762	1884	1984	2097	2172		
/ rears	s	98.8	81.6	177.4	158.8	132.9	113.2	96.4	140.1	169.0	194.6	224.6	245.1		
5 V	m	1532.6	264.0	918	990	1098	1194	1310	1755	1871	1967	2075	2147		
5 Years	s	116.4	103.6	239.7	215.2	180.5	152.9	126.6	162.0	197.3	229.2	266.9	292.7		
3 Years	m	1530.0	258.8	928	999	1104	1198	1312	1748	1862	1956	2062	2132		
	s	141.9	135.2	322.9	290.1	243.2	205.0	166.6	196.1	241.5	282.9	332.1	365.8		



Fig. 10. Mean $(\pm \text{ std. dev.})$ and median decrease in relative percentual error when using short records (3 to 15 yrs long) and the derived distribution instead of direct computation from the sample annual totals, to estimate the long-term (32 yr) standard deviation of annual rainfall at 52 locations in Switzerland.



Fig. 11. Effects of data resolution on the distributions obtained with Derived Distributions for Concepción (left) and Lugano (right). The diagonal black line represents a perfect agreement.



Fig. 12. Estimated PDFs for 10-year (left) and 5-year (right) subsamples using a 24-h totalized record in Concepción.



Fig. 13. Estimated PDFs for 10-year (left) and 5-year (right) subsamples using a 24-h totalized record in Lugano.