April 15, 2016

Dr. Dominic Mazvimavi Department of Earth Sciences Institute for Water Studies University of the Western Cape Bellville 7535 Republic of South Africa

Dear Dr. Mazvimavi,

We are submitting a revised manuscript for HESS-2015-413, originally entitled "Empirical streamflow simulation for water resource management in data-scarce seasonal watersheds" but now with the revised title (in response to reviewer comments) of "Machine learning methods for empirical streamflow simulation: a comparison of model accuracy, interpretability and uncertainty in seasonal watersheds." This revision addresses the comments provided by the two referees in the most recent iteration of review. Additionally, we've summarized how we responded to the specific comments in the text below, and attached a marked up version of the manuscript in track changes. We thank you for your time and consideration of our manuscript, and look forward to your decision.

Sincerely,

Julie Shortridge Johns Hopkins University

Reviewer 1:

- 1. I do not see a change to Figure 3 that demonstrates greater inter-annual variability. We respectfully disagree with the reviewer that Figure 3 does not show inter-annual variability; for instance, the observed peak wet season flow in 1986 was approximately 30 CMS, while the peak wet season flow in 1990 was almost twice that at approximately 60 CMS. While this variability may be relatively minor compared to seasonal variability in the flow regime, we think that the figure in its current form provides an accurate representation of both seasonal and inter-annual variability, as well as the ability of different model formulations in capturing this variability.
- 2. The rainfall intensity calculation is simplistic and that could affect its value as an input. It would be better just to say that intensity data are not available. *The text referring rainfall intensity has been removed from Section 2.2.*
- *3.* I still do not think that models that respond to higher rainfall with higher runoff (or similar statements) can be considered physical realism and I would like to see such comments changed. If a model does not respond in a sensible way that is just related to really bad

model structure. The statements in the paper represent a very broad interpretation of physical realism and perhaps this should be made a bit clearer.

The references to "physical realism" have been removed from the second and third paragraphs of discussion section, as well as the conclusion section. Instead, we just refer to models as performing in a "reasonable" fashion.

Reviewer 2:

The authors have addressed most of the comments, so I believe that the manuscript should be considered for publication as long as the following minor aspects are first taken into account.

 Title. The manuscript describes a hydrological modelling problem (i.e., fitting models, studying interpretability, bias and uncertainty), not a water resources management one. Water management entails the use of decision-making techniques to influence the way water resources are operated or planned in a given basin. None of these techniques is used in the study. I understand that the authors have a different opinion, but I believe that the title is misleading. I would leave this to the editor.

We propose the revised title "Machine learning methods for empirical streamflow simulation: a comparison of model accuracy, interpretability and uncertainty in seasonal watersheds"

- Line 24-27, page 5. The study is not only about comparing six models in terms of their predictive ability, as clearly explained in the last paragraph of the introduction. The authors may want to anticipate this important point.
 This text has been revised, removing the reference to predictive ability so it focuses only on the case study location.
- 3. Line 12, page 9. Please correct the typo 'of of'. *This has been corrected*.
- 4. Line 15-16, page 9. Please provide the technical details of the log transformation (see reviewer #1, comment no. 8. How are the streamflow data distributed? The distributions of streamflow data, log-transformed streamflow, and streamflow anomalies in each basin are shown in supplementary Figure S1.
- 5. Line 25-28, page 9. More technical details are needed here. How are the 'anomaly themselves calculated based on climatic and land cover conditions that are non-stationary through time? Text has been added elaborating on the flow calculations in the anomaly model formulation. The anomaly value for a given month in the time series (for instance, June 1990) are calculated based on temperature, precipitation, and land cover values for that specific month. The long-term average and standard deviation values (estimated from all observed June flows from 1961 to 2004) are only used to convert this anomaly value back into a flow value.

6. Figure 3 illustrates the predicted streamflow. Which model (of the six proposed) issues these predictions? It does not emerge from the caption. The same comment applies to Table 4. *The model types have been added to Table 4 and to the caption of Figure 3.*

- 1 Machine learning methods for eEmpirical streamflow
- 2 simulation: a comparison of model accuracy,
- 3 interpretability and uncertainty in for water resource
- 4 management in data-scarce seasonal watersheds
- 5

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1 Abstract

2 In the past decade, machine-learning methods for empirical rainfall-runoff modeling have 3 seen extensive development and been proposed as a useful complement to physical hydrologic models, particularly in basins where data to support process-based models are 4 5 limited. However, the majority of research has focused on a small number of methods, such as 6 artificial neural networks, despite the development of multiple other approaches for non-7 parametric regression in recent years. Furthermore, this work has often evaluated model 8 performance based on predictive accuracy alone, while not considering broader objectives 9 such as model interpretability and uncertainty that are important if such methods are to be used for planning and management decisions. In this paper, we use multiple regression and 10 machine-learning approaches (including generalized additive models, multivariate adaptive 11 regression splines, artificial neural networks, random forests, and M5 cubist models) to 12 13 simulate monthly streamflow in five highly-seasonal rivers in the highlands of Ethiopia and 14 compare their performance in terms of predictive accuracy, error structure and bias, model interpretability, and uncertainty when faced with extreme climate conditions. While the 15 relative predictive performance of models differed across basins, data-driven approaches were 16 able to achieve reduced errors when compared to physical models developed for the region. 17 18 Methods such as random forests and generalized additive models may have advantages in 19 terms of visualization and interpretation of model structure, which can be useful in providing 20 insights into physical watershed function. However, the uncertainty associated with model 21 predictions under extreme climate conditions should be carefully evaluated, since certain 22 models (especially generalized additive models and multivariate adaptive regression splines) 23 become highly variable when faced with high temperatures.

24

1 **1 Introduction**

2 Hydrologists and water managers have made use of observed relationships between 3 rainfall and runoff to predict streamflow ever since the creation of the rational method in the 4 19th century (Beven, 2011). However, the development of increasingly sophisticated machine 5 learning techniques, combined with rapid increases in computational ability, has prompted 6 extensive research into advanced methods for data-driven streamflow prediction in the past decade. Artificial neural networks (ANNs), regression trees, and support vector machines 7 have been shown to be powerful tools for predictive modeling and exploratory data analysis, 8 9 particularly in systems that exhibit complex, non-linear behavior (Solomatine and Ostfield, 10 2008; Abrahard and See, 2007).

11 While distributed physical models that accurately represent hydrologic processes can still 12 be considered the gold standard for rainfall runoff modeling, empirical models can be a useful 13 tool in contexts where there is limited data on physical watershed processes but long time-14 series of precipitation and streamflow (Iorgulescu and Beven, 2004). The development of 15 historical data centers and more recent efforts to merge satellite data with in situ observations to monitor climate and hydrology has made acceptable climate and streamflow data more 16 17 widely available in data poor regions. Because obtaining measurement-based estimates of soil hydraulic parameters or details on hydrologically-relevant land management activities can be 18 19 more difficult, empirical models may be particularly useful in these locations. While many criticize these approaches as "black boxes" with no relationship to underlying physical 20 21 processes (See et al., 2007), a number of studies have demonstrated how empirical approaches can be used to gain insights about physical system function (e.g., Han et al., 2007; Galelli and 22 23 Castelletti, 2013a). Additionally, improvements in interpretation and visualization methods can make complex models more easily interpretable (Sudheer and Jain, 2004; Jain et al., 24 25 2004). Finally, data-driven models can be useful in identifying situations where observed data disagree with what would be predicted based on conceptual models, and thus identify 26 27 assumptions regarding runoff generation processes that may be incorrect (Beven 2011).

While there have been some applications of alternative machine learning methods, such as support vector machines (Asefa et al., 2006; Lin et al., 2006) and regression-tree based approaches (Iorgulescu and Beven, 2004; Galelli and Castelletti, 2013a) for streamflow simulation, the vast majority of research has focused on artificial neural networks (Solomatine and Ostfield, 2008). While they have demonstrated impressive predictive accuracy in a

number of different contexts, excessive parameterization of ANNs can result in overfit 1 2 models that are not generalizable to unseen data (Iorgulescu and Beven, 2004; Gaume and 3 Gosset, 2003). While methods exist to avoid overfitting, such as cross validation and 4 bootstrapping, these methods are not always employed (Solomatine and Ostfield, 2008). A 5 review by Maier et al. (2010) found that relatively few studies evaluated model performance based on parameters such as Akaike information criterion that would lead to parsimonious 6 7 models that are likely to be more generalizable and interpretable. This can lead to complex 8 models that only result in modest improvements (or no improvements at all) over much 9 simpler approaches (Gaume and Gosset, 2003; Han et al., 2007).

10 Even outside of a hydrology context, it has been argued that ANNs are better suited for problems aimed at prediction without any need for model interpretation, rather than those 11 where understanding the process generating predictions and the role of input variables is 12 13 important (Hastie et al., 2009). Given the importance that this interpretation plays in 14 understanding the contexts in which a hydrologic model is appropriate and reliable, the strong 15 opinions surrounding the use of ANNs for water resources management are perhaps not surprising. To address this issue, a number of studies have focused on highlighting the 16 17 structure and mechanism by which machine learning models make predictions to confirm their physical realism and gain insight into physical watershed function. For example, some 18 19 studies have demonstrated how internal ANN structure corresponds to physical hydrologic 20 processes (Wilby et al. 2003; Jain et al., 2004; Sudheer and Jain, 2004), while others have 21 shown how variable selection and importance can be used to gain insights about model 22 structure and runoff generating processes (Galelli and Castelletti, 2013a and 2013b). While 23 these studies demonstrate that a number of methods exist for characterizing model structure, they generally focus on a single model type and thus provide little insight into the 24 25 comparative ease with which different model types can be interpreted.

While a number of comparison studies exist that apply multiple empirical models to a given problem, finding generalizable insights from these studies is hindered because of the limited number of models and datasets evaluated. Perhaps the most comprehensive comparison to date is that of Elshorbagy et al. (2010a and 2010b), who compared six methods for data-driven modeling of daily discharge in the Ourthe River in Belgium. This work found that linear models were able to perform comparably to much more complex methods when the data content of the models were limited, or when system input-output behavior was close to linear. However, other studies have demonstrated the value of using more complex
 approaches when modeling more complex rainfall-runoff behavior (e.g., Abrahart and See,
 2007; Asefa et al., 2006). The differing results obtained across these studies indicate that no
 single method is likely to be suitable for all basins, timescales, or applications.

5 However, it is important to recognize that predictive accuracy alone is not necessarily 6 sufficient justification for applying a model to a given problem. Models should not only be 7 accurate, but also be fit-for-purpose (Beven, 2011; Van Griensven et al., 2012). For instance, 8 accurate representation of low return period flows is more important in a flood forecasting 9 model than one aimed at predicting average amounts of water available for withdrawal and human consumption. Similarly, the ability to provide insights into physical watershed 10 function may be more important in basins where land-use change could alter the hydrologic 11 regime, compared to a basin that is heavily urbanized and expected to remain so. The use of 12 multiple objective functions in training data-driven models can address this to some degree by 13 14 identifying models that provide sufficient balance between different performance objectives, 15 such as accurate representation of different portions of the flow hydrograph (De Vos and Rientjes, 2008). However, more refined model training procedures will not necessarily 16 17 address other aspects of model performance that make it suitable for planning purposes, such as interpretability (Solomatine and Ostfield, 2008). More comprehensive consideration of 18 19 model strengths and limitations should be standard practice in model development and 20 selection, rather than simply evaluating global error metrics.

21 In this work, we compare six methods for empirical streamflow prediction simulation 22 (linear models, generalized additive models, multivariate adaptive regression splines, random 23 forests, M5 model trees and ANNs) in their ability to predict monthly streamflow in five 24 rivers in the Lake Tana basin in Ethiopia. This study region was selected as it provides 25 insights into the use of data-driven models for streamflow simulation in tropical regions of the 26 world that are underrepresented in existing studies; for instance, a review of 210 articles on water resource applications of ANNs found that over three quarters of the studies evaluated 27 were conducted in North America, Europe, Australia, or temperate East Asia (Maier et al., 28 2010). Existing studies conducted in tropical regions generally apply a single methodology to 29 the basin of interest and evaluate predictive accuracy alone (see for instance, Machado et al., 30 31 2011; Chibanga et al., 2003; Antar et al., 2006; Aqil et al., 2007), making it difficult to find generalizable insights into the relative advantages of different modeling approaches in these 32

regions. Better development of data-driven models for these regions has the potential to be particularly valuable because data limitations and complex hydrodynamic processes often hinder the use of physical watershed models, but relatively long time series of streamflow, precipitation and temperature may be available at a monthly timescale. These data, combined with information on relevant landscape change (in particular, the expansion of agricultural land cover), can be leveraged to create reasonably accurate empirical models.

7 Models are compared not only in terms of their predictive accuracy, but also in terms of 8 model error structure and the implications that this structure may have for water resource 9 applications. Additionally, we evaluate the methods by which model structure and predictor variable influence can be evaluated to gain insights into physical system function for each 10 model type. Finally, we assess the suitability of using different model types for climate 11 change impact assessment by comparing model uncertainty in projections made for 12 increasingly extreme climate conditions. The overall objective of this research is not to 13 14 identify a single "best" model, but rather to highlight some of the strengths and limitations of 15 different approaches, as well as demonstrate important issues that should be kept in mind for 16 model comparisons in the future

17 2 Data and Methods

18 **2.1 Study Area**

19 Lake Tana is located at an elevation of approximately 1800 meters in the highlands of northwest Ethiopia (Fig. 1). The catchment draining to the lake encompasses approximately 20 21 12,000 square kilometers, and the four main tributaries providing water to the lake are the 22 Gilgel Abbay (including its tributary, the Koga River), Ribb, Gumara, and Megech Rivers. Collectively, these rivers account for 93% of the inflow to the lake (Alemayehu et al., 2010). 23 Ninety percent of rainfall in the basin occurs during the wet season from May until October, 24 and there is significant interannual variability in precipitation with annual rainfall levels 25 26 ranging from below 1000 mm to over 1800 mm (Achenef et al., 2013). Population growth and 27 expansion of agricultural and pastoral land use in the region has resulted in substantial deforestation and land degradation, with agricultural, pastoral and settled land cover 28 29 comprising over 70% of the basin's surface area (Rientjes et al., 2011; Garede and Minale, 2014; Gebrehiwot et al., 2010). There is some evidence that this has impacted the hydrology 30

of the rivers draining into the lake (Gebrehiwot et al., 2010). A summary of basin
 characteristics for the evaluation period of 1960-2004 is presented in Table 1.

3 Approximately 2.6 million people live in the basin, and are largely settled in rural 4 areas and reliant on rainfed subsistence agriculture. This makes the region quite vulnerable to climate variability and change, and a number of water resources infrastructure projects are 5 6 planned to better manage this vulnerability and support economic development (Alemayehu 7 et al., 2010). This includes the recent construction of the Tana-Beles hydropower transfer 8 tunnel and the Koga River irrigation reservoir, as well as five other reservoirs planned for 9 construction in the next 10 to 20 years (Alemayehu et al., 2010). To better understand the potential implications of this development, extensive effort has been put towards developing 10 11 rainfall-runoff models for the Lake Tana basin, as well as other areas of the Ethiopian highlands with similar characteristics (van Griensven et al., 2012). Many of these studies rely 12 13 on Soil and Water Assessment Tool (SWAT) models, although there are some that use water 14 balance approaches (Van Griensven et al., 2012). While these models have in some cases demonstrated reasonably high accuracy, previous evaluations were largely based on Nash-15 Sutcliffe Efficiency (NSE; Nash and Sutcliffe, 1970) which can be a flawed performance 16 17 metric in highly seasonal watersheds (Schaefli and Gupta, 2007; Legates and McCabe, 1999). More importantly, the limited data available for physical parameterization of these models 18 19 required a heavy reliance on model calibration, which sometimes resulted in parameterization schemes that are inconsistent with physical understanding of the region's hydrology 20 21 (Steenhuis et al., 2009; van Griensven et al., 2012). Furthermore, a number of studies relied 22 on empirical relationships such as curve numbers and the Hargreave's equation that were 23 developed for temperate regions (e.g., Mekonnen et al., 2009; Setegne et al., 2009). While these limitations are likely to introduce considerable uncertainty into model projections, 24 25 particularly in situations where climatic or environmental conditions differ from those 26 experienced in the calibration period, few studies from this region of Ethiopia include any sort 27 of uncertainty analysis in model predictions. Empirical models could provide a useful 28 complement to physical models developed for the region by providing insights into physical 29 system function and allowing for more comprehensive uncertainty analysis.

30 2.2 Data and Model Development

31 Models were developed using monthly streamflow, climate, and land cover data for 32 the period from 1961 to 2004, resulting in 528 monthly observations. In each of the five major

rivers in the basin, we developed empirical models that estimated monthly streamflow as a 1 2 function of climate conditions and agricultural land cover in each basin. Monthly streamflow 3 data were taken from historic stream gauge records for each basin, as reported in feasibility 4 studies developed for proposed irrigation projects (Alemayehu, 2010). Historic data for 5 monthly average temperature and, monthly total precipitation, and monthly wet days in each river basin were derived from the University of East Anglia Climate Research Unit (CRU) 6 7 TS3.10 gridded meterological fields (Harris et al., 2014), which are based on meteorological 8 station observations. Historic estimates of rainfall intensity were also calculated by dividing 9 monthly total precipitation by CRU TS3.10 records of the number of wet days in that month, but was found to be highly correlated with monthly precipitation and did not result in 10 11 significant improvements to the predictive accuracy of tested models. Thus, it was not included in the final model formulations. Finally, to account for historic increases in 12 13 agricultural and pastoral land cover that have occurred in the basin, the percentage of land cover used for any crop or grazing was estimated from historic land cover analyses described 14 15 by Rientjes et al. (2011), Gebrehiwot et al. (2010), and Garede and Minale (2014). These 16 studies used historic aerial photos and satellite images to estimate land cover changes in the Ribb, Gilgel Abbay, and Koga basins from the periods of 1957 to 2011. The percentage of 17 18 agricultural land cover was interpolated for years when data weren't available, and the value 19 of agricultural land cover in the two basins without data was assumed to be equal to average 20 agricultural land cover in the basins with data. Land cover was assumed to change on an 21 annual, rather than monthly basis. While this approach is prone to errors that could stem from 22 differing rates of land use change through time and between basins, it does provide a 23 mechanism for capturing the long-term trend of expanding agricultural land cover that has 24 been observed throughout the Ethiopian highlands when detailed land-cover data are unavailable. Including this data improved out-of-sample predictive accuracy of the models, 25 26 further suggesting that it was a valuable addition.

Two general formulations for the empirical models were evaluated. The first (referred
to below as the standard model formulation) was

29
$$\log(Q_{b,t}) = f(P_{b,t}, P_{b,t-1}, P_{b,t-2}, T_{b,t}, T_{b,t-1}, T_{b,t-2}, AgLC_{b,t}) + \varepsilon_{b,t}$$
(1)

30 where $Q_{b,t}$ is the monthly streamflow in river *b* at time period *t*, $P_{b,t}$ and $T_{b,t}$ are the monthly 31 total precipitation and average temperature in river basin *b* at time period *t*, $AgLC_{b,t}$ is the total

percentage of agricultural land cover in basin b at time t, and $\varepsilon_{b,t}$ is the model error. The 1 2 subscripts t-1 and t-2 indicate lagged measurements from one and two months prior, and were included to roughly account for storage times longer than one month that could impact 3 4 streamflow in each river. While the exact time of concentration is not known in each basin, 5 the minor influence of of climate conditions at two months prior suggest that climate 6 conditions from beyond this time period do not contribute significantly to flow variability. 7 The function f represents a general function that differed between the specific models assessed 8 and is discussed in more detail below. The logarithm of monthly streamflow was used as a 9 response variable to keep model predictions positive. The distribution of streamflow data and log-transformed streamflow values in each basin are shown in supplementary Fig. S1. 10

11 In the second formulation, streamflow and climate anomalies were used as the 12 response and predictor variables to better account for the highly seasonal nature of streamflow 13 and precipitation in the region. Streamflow anomalies were calculated for each observation by 14 subtracting the long-term average streamflow for that month (m) from the observed value and 15 dividing this number by the long-term standard deviation of that month's streamflow as in Eq. (2). Anomaly values thus represent how streamflow in a given month compares to the long-16 17 term average flow for that month; for instance, an anomaly value of 1.0 for June of 1990 would indicate that streamflow in that month was one standard deviation higher than the 18 19 average June flow from 1961 to 2004. This procedure was repeated for precipitation and temperature, and these values were then used to fit models of the form described in Eq. (3). In 20 21 each month of the time series, the model estimates the flow relative to the long-term average 22 flow for that month, based on whether temperature and precipitation values were greater or 23 less than their long-term averages, as well as the percentage of agricultural land cover in that month of the time series. In this sense, the anomaly values are calculated based on climatic 24 25 and land cover conditions that vary through time. These anomaly values are then converted 26 back to raw flow values based on the long-term average and standard deviation of flow for that month. The distribution of streamflow anomaly values in each basin are shown in 27 28 supplementary Fig. S1. It should be noted that although this formulation uses long-term 29 averages and standard deviations to convert anomaly values to flow volumes, the anomaly values themselves are calculated based on climatic and land cover conditions that are 30 nonstationary through time. 31

32
$$Q_{b,t}^{AN} = \frac{Q_{b,t} - \bar{Q}_{b,m}}{sd(Q_{b,m})}$$
 (2)

| 1 | | |
|----|-------------|--|
| 2 | | $Q_{b,t}^{AN} = f(P_{b,t}^{AN}, P_{b,t-1}^{AN}, P_{b,t-2}^{AN}, T_{b,t}^{AN}, T_{b,t-1}^{AN}, T_{b,t-2}^{AN} AgLC_{b,t}) + \varepsilon_{b,t} $ (3) |
| 3 | Six differe | ent types of models were compared using each formulation in each basin: |
| 4 | 1. | A Gaussian linear regression model (GLM) using the basic stats package in the R |
| 5 | | statistical computing software (R Development Core Team, 2014) |
| 6 | 2. | Gaussian generalized additive model (GAM): GAMs are a semi-parametric |
| 7 | | regression approach where the response variable is estimated as the sum of |
| 8 | | smoothing functions applied over predictor variables. These functions allow the |
| 9 | | model to capture non-linear relationships between the predictor and response |
| 10 | | variables without apriori assumptions about the form (eg., quadratic, logarithmic) |
| 11 | | of these functions, and are fit using penalized likelihood maximization to prevent |
| 12 | | model overfitting (Hastie and Tibshirani, 1990). GAMs were fit using the mgcv |
| 13 | | package in R (Wood, 2011). |
| 14 | 3. | Multivariate adaptive regression splines (MARS): MARS are a non-parametric |
| 15 | | regression approach where the response variable is estimated as the sum of basis |
| 16 | | functions fit to recursively partitioned segments of the data (Friedman, 1991). |
| 17 | | MARS models were fit using the earth package in R (Milborrow, 2015). |
| 18 | 4. | Artificial neural network (ANN): ANNs are a non-parametric regression approach |
| 19 | | represented by a network of nodes and links that connects predictor variables to |
| 20 | | the response variable. Each link in the network represents a function that maps the |
| 21 | | input nodes into the output node (Ripley, 1996). ANN models were fit using the |
| 22 | | nnet package in R (Venables and Ripley, 2013). |
| 23 | 5. | Random forest (RF): Random forests are a rule-based, non-parametric regression |
| 24 | | approach where the model prediction is created by averaging the predicted value |
| 25 | | from multiple regression trees which are trained on separate bootstrapped |
| 26 | | resamples of the data. Each tree is fit using a small, randomly selected subset of |
| 27 | | predictor variables, resulting in reduced correlation between trees (Breiman, |
| 28 | | 2001). Random forest models were fit using the randomForest package in R (Liaw |
| 29 | | and Wiener, 2002). |
| 30 | 6. | M5 model: M5 models are a rule-based, non-parametric regression approach that |
| 31 | | fits a linear regression model to each terminal node of a regression tree (Quinlan, |
| 32 | | 1992). M5 models were fit using the Cubist package in R (Kuhn et al., 2014). |

Climatology model: A climatology model that simply predicted each month's
 streamflow as equivalent to the long-term average streamflow for that month was
 included for comparison purposes.

4 2.3 Model Evaluation

5 When using non-parametric regression approaches, it is important to avoid overfitting a 6 model to a given dataset because this can result in large errors in out-of-sample predictions (Hastie et al., 2009). To avoid model overfit, the caret package in R (Kuhn, 2015) was used to 7 8 determine model parameters for the MARS, ANN, RF and M5 models. This package uses resampling to evaluate the effect that model parameters have on the model's predictive 9 10 performance and chooses the set of parameters that minimizes out-of-sample error (Kuhn 11 2015). In this evaluation, 25 bootstrap resamples of the training dataset were generated for 12 each parameter value to be assessed. A model was fit using each bootstrap sample and used to predict the remaining observations, and the parameter values that minimized average RMSE 13 14 across all resamples. Details on the specific parameters evaluated for each model are 15 presented in Table 2. While the development of more complex structures are possible for 16 some models, this process can result in over-parameterization and poor model performance 17 (Gaume and Gosset, 2003; Han et al., 2007). Additionally, the use of a standardized 18 parameterization procedure allows for a more even comparison between different model 19 types.

20 The predictive ability of each model was assessed using 50 random holdout cross-21 validation samples. In each sample, a random selection of years were chosen, and 22 observations from these years were removed ("held-out") from the dataset. The size of the 23 held-out sample ranged from 1 to 9 years. Each model was then fit to the remaining portion of 24 the data, using the caret package described above to determine model parameters for the 25 MARS, ANN, RF and M5 models. These models were then used to predict streamflow for the 26 held-out portion of the data, and both the mean absolute error (MAE) and NSE were 27 calculated after transforming model predictions after back to the original streamflow units. 28 Mean MAE and NSE were calculated for each model across the 50 cross-validation samples 29 and used to choose the model with the highest predictive accuracy in each basin. This cross-30 validation procedure provides a mechanism for evaluating how well a model will generalize 31 to an unseen set of data while avoiding some of the problems that can arise from the use of a single calibration and validation dataset (Elshorbagy et al., 2010a; Han et al., 2007). 32

1 MAE was included as an error metric because it provides a simple and easily 2 interpretable measure of error on the same scale as observed flow volumes. While NSE values 3 are acknowledged to be a flawed performance metric in highly seasonal watersheds where 4 seasonal fluctuations contribute to a substantial portion of flow variability (Schaefli and 5 Gupta, 2007; Legates and McCabe, 1999), this metric was included to provide a rough 6 comparison of how empirical model performance compared to the performance of physical 7 models developed for the region. The use of alternative error metrics has been discussed 8 extensively in the literature (for instance Pushpalatha et al., 2012; Mathevet et al., 2006; Criss 9 and Winston, 2008), and could provide additional insights into what contributes to predicitive 10 capabilities of different model formulations. However, this work examined predicitve 11 accuracy based on MAE and NSE alone to allow for greater focus on how models differ in 12 terms of error structure and uncertainty.

13 As a rough point of comparison for the statistical models developed in this research, we 14 also evaluated discharge estimates derived from a process-based hydrological model. The 15 model used in this application is the Noah Land Surface Model version 3.2 (Noah LSM; Ek et. al, 2003; Chen et al., 1996). Noah LSM was implemented for offline simulations of the 16 17 Lake Tana basin at a gridded spatial resolution of 5km for the period 1979-2010 using a time step of 30 minutes. Meteorological forcing was drawn from the Princeton 50-year reanalysis 18 19 dataset (Sheffield et al. 2006), downscaled to account for Ethiopia's steep terrain using 20 MicroMet elevation correction equations (Liston & Elder 2006). The Princeton reanalysis was 21 selected because it provides relatively high resolution meteorological fields, including all 22 variables required to run a water and energy balance LSM like Noah, for the period 1948-23 present. While higher resolution and possibly higher quality datasets are available for recent years, this longer dataset was utilized to compare the process-based model to statistical 24 25 models developed for a long historical period. Soil parameters for the Noah simulation were drawn from the FAO global soil database, land use was defined according to the United States 26 27 Geological Survey (USGS) global 1km land cover product, and vegetation fraction was 28 derived from MODerate Imaging Spectroradiometer (MODIS) imagery. Land cover was 29 treated as a static parameter over the full length of the simulation, as spatially complete 30 estimates of historical land use were not available at the required resolution and specificity.

The highest performing model in each basin based on MAE was retained for more detailed evaluation of model error structure, covariate influence, and uncertainty in climate

1 change sensitivity analysis. To generate a complete time-series of out-of-sample model 2 predictions for error analysis, the holdout cross validation procedure was repeated for the 3 highest performing standard-formulation and anomaly-formulation models for each basin, but 4 this time holding out a single year of observations in each iteration. The predictions from this 5 cross validation were used to evaluate the how model error structure might impact model predictions used for water resource applications. The influence of different predictor variables 6 7 on model predictions was also assessed for the highest performing model in each basin after 8 being fit to the complete dataset. Each predictor variable was assessed using metrics for 9 covariate importance and influence that are unique to that model type, demonstrating how 10 models could be used to gain physical insights about data-scarce regions and the mechanisms 11 for generating these insights for each type of model. Partial dependence plots (Hastie et al., 12 2009) were also generated for each covariate for the highest performing model in each basin 13 to provide insights about how covariate influence compared across different basins and model 14 types.

15 Finally, two evaluations were conducted to assess uncertainty in model projections of streamflow under increasingly extreme climate conditions to better understand the 16 17 implications of using different model formulations for climate change impact studies. Model projections of streamflow in different climate conditions are likely to be accompanied by 18 19 considerable uncertainty, particularly when climate conditions exceed those experienced 20 historically. To assess this uncertainty, the best performing model in each basin was used to generate streamflow predictions for 1) changes in temperature from 0 to 5° C, 2) changes in 21 precipitation from -30 to +30%, 3) an increase in temperature to 5° C combined with a 22 23 decrease in precipitation to -30%, and 4) an increase in temperature to 5° C combined with an increase in precipitation to +30%. For each of the four assessments, the models generated 24 25 predictions for the 45-year historic climate record adjusted for a given degree of climate 26 change using the delta-change method (Gleick, 1986), while holding agricultural land cover 27 constant at 60%. In this method, monthly temperature values are simply added to the 28 temperature change value, and monthly precipitation values are multiplied by the precipitation 29 change percentage. Model predictions for the altered climate record were then used to calculate the average annual streamflow in each river. This process was repeated 100 times 30 31 for models fit on random bootstrap resamples of the historic dataset to generate uncertainty bounds surrounding model predictions and evaluated how the uncertainty in these predictions 32 33 increased as climate conditions became more extreme. It is important to recognize that these

should not be interpreted as a prediction or assessment of actual climate change impacts, but rather a measurement of the sensitivity of modeled streamflow in the basin to different climate conditions. Since one of the key motivations for using rainfall-runoff models is to understand how climate change may impact water resources, it is important to understand how model formulation contributes to this sensitivity and uncertainty.

6 3 Results

7 3.1 Model Accuracy and Error Structure

8 Table 3 shows the out-of-sample cross validation errors for each model assessed in each 9 basin. The random forest model had the lowest mean absolute error for the standardformulation model in four of the five basins, with the M5 model performing best for the Koga 10 11 basin. These models outperformed the Noah LSM simulations in all basins assessed. The Noah LSM errors are for a single period of analysis and thus don't present an exact corollary 12 to the cross validation performed for the empirical models. Nevertheless, the significant 13 14 increases in errors associated with the Noah LSM model demonstrates the difficulty 15 associated with the use of process-based models in the region, particularly when relying on 16 global datasets that may be unreliable at the spatial and temporal resolutions required for 17 physical modeling. Physical models developed for monthly streamflow prediction in other basins within the Ethiopian highlands have reported NSE values ranging from 0.53 to 0.92 18 19 (van Griensven et al., 2012), compared to values ranging from 0.71 to 0.87 for the random forest models developed here. If this measure alone was used for model evaluation, these 20 21 empirical models would generally be classified as having good performance based on the guidelines suggested by Moraisi et al. (2007). However, the climatology model outperforms 22 23 the best standard formulation models in all basins except Megech, indicating that in the 24 majority of basins the errors from the fitted empirical models are higher than those that result 25 from simply using the long-term monthly average for each month's prediction. This is due to 26 the fact that seasonality accounts for such a large portion of the variability in monthly flow values, and demonstrates how high NSE values can be quite easy to obtain in seasonal basins. 27

Evaluation of anomaly model errors indicates that the models using this formulation achieve better predictive accuracy than those using the standard formulation, and are able to outperform the climatology model based on both NSE and MAE in all basins. However, the highest performing models in each basin varies more when the anomaly formulation is used, with the GLM, GAM, random forest, and M5 models all minimizing MAE in different basins.
 In all basins except Koga, the highest performing model significantly outperformed the
 climatology model based on paired Wilcoxon rank-sum tests (Bonferroni-corrected p-value < 0.01).

5 Further exploration of model residuals indicates another important advantage of using 6 the anomaly model formulation. In the standard model formulation, model residuals appear to 7 be non-random. Example autocorrelation plots are shown for the Gilgel Abbay and Ribb 8 Rivers in Fig. 2, and demonstrate that a positive autocorrelation exists at the 12 month time 9 lag. For brevity, only plots for two rivers are shown, although this autocorrelation existed in 10 the standard-formulation models for all basins except Megech (Table 4). This autocorrelation 11 occurs because the standard-formulation models consistently underestimate wet-season streamflow while overestimating dry-season flows, as is apparent in hydrographs of observed 12 13 and predicted streamflow (Fig. 3). Because wet-season flows contribute such a large portion 14 of the total annual flow volume, this results in regular underestimation of aggregate values 15 such as mean annual flow (Table 4). This autocorrelation is reduced in the anomalyformulation models, meaning that they are better able to capture the peak flow volumes 16 17 experienced in the wet season and do not underestimate mean annual flow to the same degree 18 that the standard formulation models do.

19 **3.2 Model Structure and Covariate Influence**

20 Evaluating the relationship between predictor covariates and streamflow response can 21 lend insight into the physical processes underlying runoff generation in each basin. There are two components of this relationship that can be evaluated: how much each covariate 22 23 contributes to model accuracy (covariate importance), and the direction and nature of the 24 relationship between covariate values and model response (covariate influence). In many 25 machine-learning models, complete description of the all of the mathematical relationships 26 within the model (for instance, through description of each tree comprising a random forest model) is infeasible, requiring the use of other mechanisms for understanding covariate 27 28 importance and influence. However, because each model type is structured in a different way, 29 these mechanisms differ. This section first describes the mechanisms available for obtaining 30 insights about covariate influence in each of the highest performing models. To provide a mechanism for comparing results across different basins, each basin model is then assessed 31 using the general approach of partial dependence plots. 32

1 In the Gilgel Abbay and Koga basins, the highest performing model was a simple 2 linear regression model. These models can be evaluated by reviewing model coefficients and 3 associated p-values, as shown in Table 5. In a standard linear regression, model coefficients 4 can be interpreted as the mean change in the response variable that results from a unit change 5 in that covariate when all others are held constant. These coefficients are for streamflow 6 anomalies rather than raw values, making their immediate interpretation less intuitive. For 7 instance, in the Gilgel Abbay model an increase of one standard deviation in precipitation 8 results in an increase of 0.22 standard deviations in flow. The associated p-value for each 9 coefficient evaluates a null hypothesis that the true coefficient value is equal to zero given the 10 other covariates in the model, and thus has no influence on the response variable.

11 Evaluating model structure based on regression coefficients is appealing due to their simplicity and familiarity. However, it is important to keep in mind that the above 12 13 interpretations rely on specific assumptions regarding model error distributions. Examination 14 of fitted model residuals from both basins indicate that errors are autocorrelated in the Koga 15 basin and not normally distributed due to the presence of outliers in both basins. Nonnormality and autocorrelation both impact the t statistics and f statistics used to test for the 16 17 significance of model coefficients, and thus the p-values for these models are likely biased 18 (Montgomery et al., 2012).

19 Interpretation of variable influence in GAMs is based on the estimated degrees of 20 freedom (EDF) a covariate's smoothing function $s(X_i)$ uses within a model (Hastie and 21 Tibushini, 1986). An EDF value of one or below indicates a linear function relating the 22 response variable to that covariate, while values greater than one represent a non-linear 23 smoothing function. An EDF value of zero indicates that the covariate smoothing function is 24 penalized to zero (meaning it has no influence on model predictions). In the model for the 25 Megech River, the terms for lagged temperature at one and two months, as well as 26 precipitation lagged at two months were all smoothed to zero. Of the remaining covariates, lagged precipitation has a linear impact on model response, while precipitation, temperature 27 and land cover have non-linear impacts. Smoothing functions can be plotted to gain more 28 insight about these relationships (Fig. 4). The functions for precipitation anomaly, lagged (one 29 30 month) precipitation anomaly, and agricultural land cover show a positive relationships with 31 streamflow, while the function for temperature anomaly predicts low streamflow at both high 32 and low anomalies.

P-values test the null hypothesis that a covariate's smoothing function is equal to zero, but rest on the assumption that model residuals are homoscedastic and independent (Wood, Similar to the linear models, residuals in the Megech GAM model appear to be both autocorrelated and heteroscedastic, meaning that a formal statistical interpretation of this value may be inappropriate and that confidence bounds around smoothing functions might be misleading.

7 The M5 cubist model fit for the Gumara basin is an ensemble of 100 small M5 8 regression trees. In each tree, the model splits observations based on logical rules related to 9 one or more covariates and fits a linear regression model to each set of observations. The final model prediction is the average across all of the individual trees. Using this sort of ensemble 10 11 approach can reduce model variance and improve accuracy if the individual trees are unbiased, uncorrelated predictors (Breiman 1996). This can be useful in avoiding models that 12 13 are overfit to the data, but can reduce model interpretability since direct visualization of 14 model structure becomes impractical as the number of trees increases. However, the 15 frequency with which individual covariates are used as splitting points within trees and as regression coefficients can provide some insights about covariate importance (Table 5; note 16 17 that because multiple covariates can be used for rules and linear models, these don't 18 necessarily add to 100%). Model rules were largely based on land cover, with some rules 19 based on precipitation. These two covariates were also used most frequently in linear 20 regressions at model nodes, followed by temperature (current and 1-month lag) and 1-month 21 lagged precipitation. Notably, climate data from 2 months lagged were not used at all. While 22 this can be useful in identifying which covariates have the largest impact on model 23 predictions, it doesn't provide any information regarding the nature or direction of that influence. 24

25 Similarly, the random forest model developed for the Ribb basin is an ensemble of regression trees in which the final model prediction is the average of the predictions from 26 27 each individual tree. However, random forests use standard regression trees that do not incorporate linear regression models at terminal nodes. Variable importance within the final 28 29 model is measured by recording the increase in out-of-sample MSE that results when a 30 covariate is randomly permuted for each tree in the ensemble. This increase in error is then 31 averaged across all trees in the ensemble. In our model, the largest increases in error resulted 32 from permutation of land cover and temperature, followed by 2-month lagged temperature

1 and precipitation. Covariate influence can be evaluated through the use of partial dependence 2 plots, which measure the change in model predictions that result from changing the value of 3 one parameter while leaving all other covariates constant (Hastie et al., 2009). Partial 4 dependence plots indicate that model predictions of streamflow are higher when the percent of 5 agricultural land cover is greater than approximately 75%, when temperatures anomalies are 6 low, and when precipitation anomalies are high. However, it appears that the plot for lagged 7 temperature might be sensitive to outliers at high temperature anomalies as evidenced by the 8 large increase that occurs above an anomaly of +2, in a region where very few data points are 9 present.

10 Many of the measures used to evaluate covariate importance and influence are model 11 specific, making inter-basin and inter-model comparisons difficult. However, the partial dependence plots used in the randomForest R package can be developed for any model and 12 provide a mechanism for comparing the influence that covariates have in the different models 13 14 and basins (Shortridge et al., 2015). Partial dependence plots were generated for each basin's best performing model and results are shown for climatic variables in Fig. 6. As expected, 15 16 models generally respond positively to increases in precipitation and negatively to increases 17 in temperature, with the greatest influence in the current month and decreasing influence at 18 one and two months prior. The influence of the current month's precipitation is linear in three 19 of the five basins; while this is constrained to the be the case in the Gilgel Abbay and Koga 20 basins due to the use of a linear model, the linear response in Gumara is not required from the 21 M5 model structure. Interestingly, both Megech and Ribb demonstrate a linear response to 22 negative precipitation anomalies, but little response to positive anomalies. Streamflow 23 response to temperature is strongest in the Gumara basin; interestingly, this is the basin with the smallest response to precipitation. 24

25 The partial dependence plots for the percentage of the basin classified as agricultural 26 land cover indicates a positive relationship between agricultural land cover and streamflow in all basins except for the Gilgel Abbay (Fig. 7). This would be expected if deforestation had 27 28 contributed to a decrease in evapotranspiration in the contributing watersheds. The exact nature of this response differs across the different rivers, with the relatively minor responses 29 30 in Koga and Ribb, and much stronger responses in the Gumara and Megech basins. However, 31 this plot also demonstrates some of the limitations associated with different model structures. 32 The plot for Gumara is highly erratic, indicating that the M5 model might be overfit to the training dataset, despite the use of model averaging to reduce model variance. Additionally,
the GAM used in the Megech basin was only trained on agricultural land cover values up to
77%; while this model may be accurately representing the impact of land cover changes
within this range, extrapolating this relationship to higher values leads to predictions that may
not be physically realistic.

6 **3.3** Climate Change Sensitivity and Uncertainty Assessment

7 Fig. 8 shows the results of the climate change sensitivity analysis for total flow from all 8 five tributaries, with dashed lines representing 95% confidence intervals obtained through 100 9 bootstrapped resamples of the data set. As would be expected, increasing temperature 10 independently of precipitation results in decreasing total flows while increasing precipitation results in higher flows. However, the uncertainty surrounding temperature sensitivity 11 12 increases at higher changes in temperature, while the uncertainty surrounding precipitation 13 sensitivity remains relatively constant, even at extreme changes in annual precipitation. The 14 bottom panels of the figure show the sensitivity of total inflows to concurrent changes in 15 temperature and precipitation. Unsurprisingly, decreasing precipitation combined with higher 16 temperatures results in greater decreases in total flow than when temperature and precipitation are varied independently. However, even if temperature increases are combined with higher 17 18 precipitation, total flows decline in the majority of bootstrap resamples.

19 The uncertainty surrounding temperature sensitivity is a key limitation to using data-20 driven approaches for climate impact assessment. To better understand which models and 21 basins are contributing to this uncertainty, Fig. 9 shows how the coefficient of variation (the 22 standard deviation of predictions from all bootstrap samples divided by the mean of these 23 predictions) varies as a function of temperature change in each basin. From this figure, it is 24 apparent that the Megech model is by far the largest contributor to model uncertainty; 25 however, it is not clear whether this contribution is due to model structure (the GAM model 26 used for the Megech River) or characteristics associated with the basin itself. To investigate 27 how different model structures contributed to this uncertainty, the bootstrap resampling 28 procedure was used to assess uncertainty in streamflow predictions in the Gumara River from 29 all model types. This basin was chosen because all six models were able to outperform the 30 climatology model, and thus could be considered good choices for model selection based on 31 predictive accuracy alone. The results indicate that the increase in uncertainty is highest, and 32 increases non-linearly, in the GLM, GAM, and MARS models. Uncertainty increases more slowly in the ANN and M5 models, and no noticeable increase in uncertainty is apparent in
 the random forest model.

3 4 Discussion

4 The objective of this study was not to identify the "best" approach for empirical 5 rainfall-runoff modeling, as this is likely to be highly specific to the basin and problem to which a model is applied. However, we hope that the comparison conducted here can 6 highlight some of the strengths and limitations of different approaches, as well as demonstrate 7 8 some important issues that should be kept in mind for model comparisons in the future. One 9 important finding was the limitation with using NSE as an error metric. Our results confirm 10 previous studies that found that even uninformative models able to capture basic seasonality are able to achieve high NSE values (Legates and McCabe, 1999; Schaefli and Gupta, 2007), 11 12 and provide further evidence indicating that high NSE values should be considered a necessary but not sufficient requirement for model usage in planning situations. For instance, 13 14 the simple climatology model used for comparison purposes here is able to achieve high NSE values, but would be unsuitable for planning since it does not account for any interannual 15 16 variability nor the possibility for non-stationary conditions caused by changing climate and land cover. In particular, understanding error structure can be valuable in evaluating whether 17 18 model biases might undermine the model's suitability for management activities. In our 19 example, the autocorrelation present in the standard-formulation models meant that these 20 models were consistently underestimating wet-season flows, resulting in low estimates of the 21 total annual flow in the rivers. Since multiple reservoirs are planned for construction on these 22 rivers to support irrigation activities, this bias could lead to poor estimates of how much water 23 is available for agricultural use in the short term (ie., seasonal forecasting) and long-term (due 24 to climate change). Interestingly, difficulties in accurately capturing high flows has been observed in physical hydrologic models for Ethiopia (e.g., Setegne et al., 2011; Mekonnen et 25 26 al., 2009) and more generally (e.g., Wilby, 2005). The implications of this limitation should be carefully evaluated before using models for water resource planning or (more importantly) 27 flood risk evaluation. 28

Depending on the model type used, different mechanisms are available to evaluate covariate importance and influence within the model. This evaluation can be useful in confirming that the model is replicating physically realistic relationships between input and output variables in a reasonable manner. While the relationships identified in this evaluation

1 are fairly straightforward (for example, increasing runoff with higher precipitation and lower 2 temperatures), these simple relationships are still important in highlighting the mechanisms by 3 which the models make predictions so that they are not "black boxes." For instance, Han et al. 4 (2007) explore how ANN flood forecasting models responds to a double-unit input of rain, 5 finding that some formulations respond in a hydrologically meaningful way to increased rainfall intensity, while others do not. Similarly, Galelli and Castelletti (2013a) describe how 6 7 input variable importance can be used to highlight differences in hydrologic processes 8 between an urbanized and forested watershed. The easy manner in which covariate 9 relationships within the GAM and random forest models can be visualized using a single 10 command within their respective R packages is a strong advantage to these approaches 11 compared to methods such as M5 model trees and artificial neural networks. Of course, partial 12 dependence plots can be developed for any model type (as was done in this research), but 13 code must be written by the user and thus requires a higher degree of effort than is necessary 14 for in-package functions. A downside to most machine-learning models is that they do not 15 support the statistical formalism in assessing variable importance that is possible when linear models and GAMs are used. However, this formalism often rests on assumptions regarding 16 17 model residuals that are unlikely to be met in many hydrologic models (Sorooshian and Dracup, 1980). 18

19 Within the Lake Tana basin, evaluation of covariate influence indicates that each basin's model is performing in a physically realistic reasonable manner, with a runoff 20 increasing with higher precipitation levels and decreasing with higher temperatures. The 21 22 influence of precipitation and temperature is greatest in the current month, and progressively 23 declines to a very small influence after two months. This suggests that long-term (multimonth) storage does not significantly contribute to variability in flow volumes. One 24 25 interesting finding is the non-linear relationship between concurrent month precipitation and 26 runoff that exists in the Megech and Ribb basins, which suggests that above a certain point 27 increasing rainfall does not result in a commiserate increase in streamflow. Other studies have 28 noted the dampening effect that wetlands and floodplains have had on river flows in the 29 region (Dessie et al., 2014; Gebrehiwot et al., 2010); this phenomenon could explain the nonlinear relationship identified in this work. The clearly negative relationship between 30 31 temperature and runoff demonstrates the degree to which upstream evapotranspiration impacts streamflow and suggests that evapotranspiration is largely energy-limited, rather than 32 33 water-limited. Increasing agricultural land-use appears to be associated with higher runoff in

1 all rivers except for Gilgel Abbay (where no clear relationship between land cover and runoff 2 was observed), and suggests that agricultural expansion at the expense of forest cover has 3 reduced the evaporative component of the water balance in these basins. Finally, the relative 4 performance of different model formulations themselves can also be informative. For 5 instance, the improved performance of the anomaly-formulation models indicates that the relationship between precipitation and runoff varies throughout the year and could point 6 7 towards differences in runoff-generating mechanisms in the wet and dry seasons that have 8 been observed in other case studies (Wilby, 2005).

9 One limitation with data-driven approaches for streamflow prediction is that the relationships they model can only generate reliable predictions for conditions that are 10 11 comparable to those experienced historically. Using these models to generate predictions for conditions that exceed historic variability is likely to introduce considerable uncertainty into 12 13 their projections. Our results indicate that uncertainty in projections of streamflow under 14 changing precipitation is relatively constant, whereas uncertainty increases markedly in 15 projections of streamflow under increasing temperature. This result is not surprising when one 16 considers the basin's climate, which is characterized by highly variable rainfall but fairly consistent temperatures (Table 6). A temperature increase of 3° C equates to almost two 17 standard deviations beyond the historic mean, whereas a change in precipitation of 30% is 18 19 well within the range of conditions experienced historically. One would expect that in other 20 climates (for example, temperate watersheds with only minor changes in rainfall throughout 21 the year), this relationship could be reversed. Despite the uncertainty that exists in projections 22 of streamflow under changing temperature, total annual flow appears to be quite sensitive to 23 increasing temperatures. In fact, the decreases in streamflow due to increasing temperature appears likely to be more than enough to counteract any increases in streamflow resulting 24 25 from higher precipitation that is projected for the region in some global circulation models (GCMs). This is consistent with the work of Setegne et al. (2011), who used projections from 26 27 multiple GCMs as input for a SWAT model developed for the region and found that streamflow decreased in the majority of emissions scenarios and models, even when 28 29 precipitation increased. Unfortunately, this suggests that any hopes for a "windfall" of 30 additional water to support agriculture and hydropower in the region under climate change 31 may be unfounded.

1 Repeating the climate change sensitivity experiment with multiple models fit to the 2 Gumara watershed indicated that the MARS, GAM, and linear models all result in the largest 3 increase in uncertainty at high temperatures. This indicates that when models are fit to slightly 4 different bootstrap resamples of the historic dataset, the projected changes in streamflow at 5 high temperature changes can be highly erratic. This is likely due to the fact that extrapolating 6 the relationships that are observed between historic temperature and streamflow to higher 7 temperatures can lead to very large changes in streamflow. Fitting the models to bootstrap 8 resamples of the data results in minor changes to these relationships that can result in widely 9 varying projections when the models are used to predict streamflow at higher temperatures, particularly when these relationships are nonlinear (as in the GAM). At the other end of the 10 11 spectrum, the random forest model exhibits almost no increase in uncertainty at high temperatures, meaning that projections of streamflow at high temperatures are consistent 12 13 across the bootstrap resamples. This is likely the result of the random forest model structure. The predicted value for each of a regression tree's terminal nodes is the average of all 14 15 observations that meet the conditions described for that node. Thus, the model will not predict 16 values beyond those experienced historically, even if covariate values exceed those contained 17 within the historic dataset. Thus, this model is likely to underestimate the change in 18 streamflow that results from increasing temperatures.

19 **5** Conclusions

20 In this work, we compared multiple methods for data-driven rainfall-runoff modeling in their ability to simulate streamflow in five highly-seasonal watersheds in the Ethiopian 21 22 highlands. Despite the popularity of ANNs in research on streamflow prediction to date, 23 ANNs were not found to be the most accurate model in any of the five basins evaluated. Other 24 methods, in particular GAMs and random forests, are able to capture non-linear relationships 25 effectively and lend themselves to simpler visualization of model structure and covariate 26 influence, making it easier to gain insights on physical watershed functions and confirm that 27 the model is operating in a physically realistic reasonable manner. However, it is important to carefully evaluate model structure and residuals, as these can contribute to biased estimates of 28 29 water availability and uncertainty in estimating sensitivity to potential future changes in 30 climate. In particular, autocorrelation in model residuals can result in underestimation of aggregate metrics such as annual flow volumes, even in models with high NSE performance. 31 32 Uncertainty in GAM projections was found to rapidly increase at high temperatures, whereas

1 random forest projections may be underestimating the impact of high temperatures on river 2 flows. Thorough consideration of this uncertainty and bias is important any time that models 3 are used for water planning and management, but especially crucial when using such models 4 to generate insights about future streamflow levels. By considering multiple model 5 formulations and carefully assessing their predictive accuracy, error structure and uncertainties, these methods can provide an empirical assessment of watershed behavior and 6 7 generate useful insights for water management and planning. This makes them a valuable 8 complement to physical models, particularly in data-scarce regions with little data available 9 for model parameterization, and warrants additional research into their development and 10 application.

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| | Drainage area above | Average annual streamflow | Standard deviation of annual | Coefficient of variation | Average temp | Average monthly rainfall [mm] | | |
|--------------|-----------------------------|---------------------------------|------------------------------------|--------------------------------|-----------------|----------------------------------|---------|--|
| Basin | gauge (km ²) | at gauge (MCM) | (MCM) | of annual streamflow | (°C) | May-Oct | Nov-Apr | |
| Gilgel Abbay | 2664 | 1883 | 217 | 0.12 | 15.7 | 206 | 39.3 | |
| Gumara | 385 | 236 | 71 | 0.30 | 17.7 | 186 | 29 | |
| Koga | 200 | 114 | 31 | 0.27 | 15.7 | 206 | 39.3 | |
| Megech | 424 | 172 | 66 | 0.31 | 20.6 | 234 | 41.4 | |
| Ribb | 677 | 210 | 83 | 0.36 | 18.2 | 263 | 45.8 | |

| 1 Table 1. Study basin characteristics over the evaluation period of 1961 to 2 | 004. |
|--|------|
|--|------|

| Model type | R package | Parameters defined in model formulation | Parameters selected through cross validation |
|---------------|--------------|---|--|
| GLM | stats | family = Gaussian | NA |
| GAM | mgcv | family = Gaussian | |
| | | method = generalized cross validation | |
| | | variable selection = true | |
| | | basis dimension $k = 3$ | |
| | | $epsilon = 10^{-7}$ | |
| | | maxit = 200 | |
| MARS | earth | nk = 21 | degree = $\{1, 2, 3\}$ |
| | | thresh = 0.001 | nprune = $\{5, 10, 15, 20, 25\}$ |
| | | fast.k = 20 | |
| | | pmethod = backward | |
| ANN | nnet | weights = 1 | size = $\{1, 2, 4, 8, 20\}$ |
| | | rang = 0.7 | decay = $\{0.0, 0.1, 0.5, 1.0, 2.0\}$ |
| | | maxit = 100 | |
| | | maxNWts = 1000 | |
| | | $abstol = 10^{-4}$ | |
| | | $reltol = 10^{-8}$ | |
| RF | randomForest | ntree $= 500$ | $mtry = \{2, 3, 4, 5, 6, 7\}$ |
| | | sampsize = 528 | |
| | | nodesize = 5 | |
| | | nPerm = 1 | |
| M5 | Cubist | rules = 100 | committees = $\{10, 50, 100\}$ |
| | | extrapolation = 100 | neighbors = $\{0, 5, 9\}$ |
| | | sample = 0 | |

1 Table 2. Model parameters evaluated through cross validation.

2 Table 3. Cross validation errors for each assessed model.

| Standa | rd Formulation | GLM | GAM | MARS | RF | M5 | ANN | Climatology | Noah LSM |
|--------|-----------------|-------|-------|-------|-------|-------|-------|-------------|----------|
| | Gilgel Abbay | 30.78 | 18.54 | 16.75 | 14.89 | 15.11 | 17.22 | 10.42 | 28.11 |
| | Gumara | 4.29 | 3.41 | 3.28 | 2.67 | 2.96 | 3.15 | 2.57 | 3.95 |
| MAE | Koga | 1.50 | 1.30 | 1.38 | 1.20 | 1.17 | 1.23 | 1.06 | 1.97 |
| | Megech | 4.45 | 2.64 | 2.83 | 2.37 | 2.53 | 3.04 | 2.54 | 4.09 |
| | Ribb | 4.69 | 2.98 | 3.50 | 2.97 | 3.27 | 3.17 | 2.81 | 7.01 |
| | Gilgel Abbay | -0.02 | 0.81 | 0.83 | 0.87 | 0.86 | 0.84 | 0.95 | 0.59 |
| | Gumara | 0.04 | 0.51 | 0.61 | 0.80 | 0.66 | 0.70 | 0.81 | 0.48 |
| NSE | Koga | 0.45 | 0.71 | 0.65 | 0.76 | 0.77 | 0.76 | 0.83 | 0.25 |
| | Megech | -1.85 | 0.63 | 0.46 | 0.73 | 0.65 | 0.52 | 0.71 | 0.41 |
| | Ribb | -1.14 | 0.71 | 0.39 | 0.71 | 0.31 | 0.67 | 0.73 | -0.75 |
| Anoma | aly Formulation | GLM | GAM | MARS | RF | M5 | ANN | Climatology | Noah LSM |
| | Gilgel Abbay | 9.73 | 9.82 | 10.10 | 10.12 | 9.94 | 9.79 | 10.42 | 28.11 |
| | Gumara | 2.22 | 2.25 | 2.43 | 2.23 | 2.16 | 2.22 | 2.57 | 3.95 |
| MAE | Koga | 1.03 | 1.06 | 1.08 | 1.09 | 1.05 | 1.05 | 1.06 | 1.97 |
| | Megech | 2.49 | 2.48 | 2.63 | 2.66 | 2.69 | 2.50 | 2.54 | 4.09 |
| | Ribb | 2.79 | 2.76 | 2.84 | 2.70 | 2.78 | 2.77 | 2.81 | 7.01 |
| | Gilgel Abbay | 0.95 | 0.95 | 0.95 | 0.95 | 0.95 | 0.95 | 0.95 | 0.59 |
| | Gumara | 0.85 | 0.85 | 0.82 | 0.85 | 0.86 | 0.86 | 0.81 | 0.48 |
| NSE | Koga | 0.83 | 0.82 | 0.81 | 0.81 | 0.82 | 0.82 | 0.83 | 0.25 |
| | Megech | 0.73 | 0.72 | 0.65 | 0.66 | 0.61 | 0.72 | 0.71 | 0.41 |
| | Dibb | 0.73 | 0.75 | 0.72 | 0.75 | 0.73 | 0.74 | 0.73 | -0.75 |

| 2 | Table 4. Residual autocorrelation factors at a 12-month lag for the highest performing |
|---|--|
| 3 | standard formulation and anomaly formulation models in each basin (with model type in |
| 4 | parenthesis), and resulting mean annual observed and predicted flow. |

| | Autocorrel | ation Factors | Mean A | Mean Annual Flow (MCM) | | | | | |
|--------|------------------|-------------------|----------|------------------------|---------|--|--|--|--|
| | Standard | Anomaly | Observed | Standard | Anomaly | | | | |
| Gilgel | 0.33 <u>(RF)</u> | 0.11 <u>(GLM)</u> | 22,925 | 20,703 | 22,958 | | | | |
| Gumara | 0.29 <u>(RF)</u> | 0.07 <u>(M5)</u> | 2,870 | 2,392 | 2,734 | | | | |
| Koga | 0.04 <u>(M5)</u> | 0.10 <u>(GLM)</u> | 1,383 | 1,333 | 1,386 | | | | |
| Megech | 0.05 <u>(RF)</u> | 0.04 <u>(GAM)</u> | 2,035 | 1,637 | 2,028 | | | | |
| Ribb | 0.21 (RF) | -0.01 (RF) | 2.575 | 1.969 | 2.615 | | | | |

| Model type | be Linear model | | | | | eralized ve model | M5 : | model tree | Random forest |
|----------------------|-------------------------|----------------------------|--------------------------------|---------|---|---|---------------|-----------------------|---|
| Measure of influence | Line a | ar regressi nd associat | on coefficient: ed p-values | 8 | Esti degr freedo and ass va | Estimated degrees of freedom (EDF) and associated p- values Covariate usage in tree rules and model coefficients | | | Increase in MSE when covariate is randomly permuted |
| Basin | Gilgel A | bbay | Kog | a | Me | egech | C | Gumara | Ribb |
| Covariate | Coefficient estimate | P-value | Coefficient estimate | P-value | EDF | P-value | Tree rules | Model coefficients | Percent increase in MSE |
| Prec | 0.22 | < 0.01 | 0.24 | < 0.01 | 1.346 | < 0.01 | 5% | 58% | 7.71% |
| Prec (lag 1) | 0.10 | 0.03 | 0.16 | < 0.01 | 0.624 | 0.08 | 0% | 19% | 2.79% |
| Prec (lag 2) | 0.01 | 0.74 | 0.05 | 0.26 | 0 | 0.29 | 0% | 0% | 1.10% |
| Temp | -0.09 | 0.08 | -0.07 | 0.17 | 1.023 | 0.07 | 0% | 47% | 12.74% |
| Temp (lag 1) | -0.04 | 0.49 | -0.06 | 0.22 | 0 | 0.32 | 0% | 46% | 4.97% |
| Temp (lag 2) | -0.01 | 0.81 | -0.09 | 0.08 | 0 | 0.56 | 0% | 0% | 8.16% |
| Agr. LC | 0.00 | 0.33 | 0.02 | 0.01 | 1.986 | < 0.01 | 86% | 73% | 15.21% |
| 2 | | | | | | | | | |

1 Table 5. Covariate importance measurements from each basin's model

Table 6. Mean and standard deviation values for temperature, wet-season rainfall, and dry-season rainfall in each basin. 2

| | Temperature (°C) | | Wet s rain (mm/z | season nfall month) | Dry season rainfall (mm/month) | |
|--------------|---------------------|------|------------------------|---------------------------|--------------------------------------|------|
| | Mean | SD | Mean | SD | Mean | SD |
| Gilgel Abbay | 15.7 | 1.54 | 206 | 145 | 39.3 | 56.5 |
| Gumara | 17.7 | 1.55 | 186 | 137 | 29.0 | 43.6 |
| Koga | 15.7 | 1.54 | 206 | 145 | 39.3 | 56.5 |
| Megech | 20.6 | 1.75 | 234 | 118 | 41.4 | 60.9 |
| Ribb | 18.2 | 1.61 | 263 | 115 | 45.8 | 57.0 |



1 Figure 1. Map of Lake Tana and surrounding rivers





Figure 3. Example observed and predicted flows <u>from the standard formulation RF model and</u> <u>anomaly formulation M5 model</u> for <u>the Gumara River from 19895</u> to <u>19912000</u>.



Figure 4. Plots of the smoothing functions used in the Megech River GAM. Hash marks along
 the x-axis indicate observation values of each covariate.



Figure 5. Partial dependence plots for the Ribb River random forest model. Hash marks along
 the x-axis show covariate sample decile values.





Figure 6. Partial dependence plots for climate covariates in the highest performing model in each basin. Model type is indicated in parentheses.



Figure 7. Partial dependence plot for agricultural land cover in the highest performing model
 in each basin. Model type is listed in parentheses for each basin. Dashed lines
 indicate values that exceed historic levels of agricultural land cover experienced in
 that basin.



Figure 8. Projected changes in total streamflow (relative to current long-term average) under changing climate conditions. The top two panels show the sensitivity to changes in temperature and precipitation when they are varied independently. The bottom panel shows sensitivity to changing temperature in conjunction with decreasing (left panel) and increasing (right panel) precipitation. Dashed lines represent 95% confidence bounds from bootstrap resampling.



Figure 9. Changes in the coefficient of variation across bootstrap resamples from the highest
 performing model in each basin (left panel) and multiple models all applied to the
 Gumara basin (right panel).



