

1 **Empirical streamflow simulation for water resource**
2 **management in data-scarce seasonal watersheds**

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4 **J. E. Shortridge,¹ S. D. Guikema,² and B. F. Zaitchik³**

5 [1]{Department of Geography and Environmental Engineering, Johns Hopkins University,
6 Baltimore, USA}

7 [2]{Department of Industrial and Operations Engineering, University of Michigan, Ann
8 Arbor, USA}

9 [3]{Department of Earth and Planetary Sciences, Johns Hopkins University, Baltimore, USA}

10 Correspondence to: J. E. Shortridge (jshortridge@jhu.edu)

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1 **Abstract**

2 In the past decade, machine-learning methods for empirical rainfall-runoff modeling have
3 seen extensive development and been proposed as a useful complement to physical
4 hydrologic models, particularly in basins where data to support process-based models are
5 limited. However, the majority of research has focused on a small number of methods, such as
6 artificial neural networks, despite the development of multiple other approaches for non-
7 parametric regression in recent years. Furthermore, this work has often evaluated model
8 performance based on predictive accuracy alone, while not considering broader objectives
9 such as model interpretability and uncertainty that are important if such methods are to be
10 used for planning and management decisions. In this paper, we use multiple regression and
11 machine-learning approaches (including generalized additive models, multivariate adaptive
12 regression splines, artificial neural networks, random forests, and M5 cubist models) to
13 simulate monthly streamflow in five highly-seasonal rivers in the highlands of Ethiopia and
14 compare their performance in terms of predictive accuracy, error structure and bias, model
15 interpretability, and uncertainty when faced with extreme climate conditions. While the
16 relative predictive performance of models differed across basins, data-driven approaches were
17 able to achieve reduced errors when compared to physical models developed for the region.
18 Methods such as random forests and generalized additive models may have advantages in
19 terms of visualization and interpretation of model structure, which can be useful in providing
20 insights into physical watershed function. However, the uncertainty associated with model
21 predictions under extreme climate conditions should be carefully evaluated, since certain
22 models (especially generalized additive models and multivariate adaptive regression splines)
23 become highly variable when faced with high temperatures.

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1 1 Introduction

2 Hydrologists and water managers have made use of observed relationships between
3 rainfall and runoff to predict streamflow ever since the creation of the rational method in the
4 19th century (Beven, 2011). However, the development of increasingly sophisticated machine
5 learning techniques, combined with rapid increases in computational ability, has prompted
6 extensive research into advanced methods for data-driven streamflow prediction in the past
7 decade. Artificial neural networks (ANNs), regression trees, and support vector machines
8 have been shown to be powerful tools for predictive modeling and exploratory data analysis,
9 particularly in systems that exhibit complex, non-linear behavior (Solomatine and Ostfield,
10 2008; Abrahard and See, 2007).

11 While distributed physical models that accurately represent hydrologic processes can still
12 be considered the gold standard for rainfall runoff modeling, empirical models can be a useful
13 tool in contexts where there is limited data on physical watershed processes but long time-
14 series of precipitation and streamflow (Iorgulescu and Beven, 2004). The development of
15 historical data centers and more recent efforts to merge satellite data with *in situ* observations
16 to monitor climate and hydrology has made acceptable climate and streamflow data more
17 widely available in data poor regions. Because obtaining measurement-based estimates of soil
18 hydraulic parameters or details on hydrologically-relevant land management activities can be
19 more difficult, empirical models may be particularly useful in these locations. While many
20 criticize these approaches as “black boxes” with no relationship to underlying physical
21 processes (See et al., 2007), a number of studies have demonstrated how empirical approaches
22 can be used to gain insights about physical system function (e.g., Han et al., 2007; Galelli and
23 Castelletti, 2013a). Additionally, improvements in interpretation and visualization methods
24 can make complex models more easily interpretable (Sudheer and Jain, 2004; Jain et al.,
25 2004). Finally, data-driven models can be useful in identifying situations where observed data
26 disagree with what would be predicted based on conceptual models, and thus identify
27 assumptions regarding runoff generation processes that may be incorrect (Beven 2011).

28 While there have been some applications of alternative machine learning methods, such
29 as support vector machines (Asefa et al., 2006; Lin et al., 2006) and regression-tree based
30 approaches (Iorgulescu and Beven, 2004; Galelli and Castelletti, 2013a) for streamflow
31 simulation, the vast majority of research has focused on artificial neural networks (Solomatine
32 and Ostfield, 2008). While they have demonstrated impressive predictive accuracy in a

1 number of different contexts, excessive parameterization of ANNs can result in overfit
2 models that are not generalizable to unseen data (Iorgulescu and Beven, 2004; Gaume and
3 Gosset, 2003). While methods exist to avoid overfitting, such as cross validation and
4 bootstrapping, these methods are not always employed (Solomatine and Ostfield, 2008). A
5 review by Maier et al. (2010) found that relatively few studies evaluated model performance
6 based on parameters such as Akaike information criterion that would lead to parsimonious
7 models that are likely to be more generalizable and interpretable. This can lead to complex
8 models that only result in modest improvements (or no improvements at all) over much
9 simpler approaches (Gaume and Gosset, 2003; Han et al., 2007).

10 Even outside of a hydrology context, it has been argued that ANNs are better suited for
11 problems aimed at prediction without any need for model interpretation, rather than those
12 where understanding the process generating predictions and the role of input variables is
13 important (Hastie et al., 2009). Given the importance that this interpretation plays in
14 understanding the contexts in which a hydrologic model is appropriate and reliable, the strong
15 opinions surrounding the use of ANNs for water resources management are perhaps not
16 surprising. To address this issue, a number of studies have focused on highlighting the
17 structure and mechanism by which machine learning models make predictions to confirm
18 their physical realism and gain insight into physical watershed function. For example, some
19 studies have demonstrated how internal ANN structure corresponds to physical hydrologic
20 processes (Wilby et al. 2003; Jain et al., 2004; Sudheer and Jain, 2004), while others have
21 shown how variable selection and importance can be used to gain insights about model
22 structure and runoff generating processes (Galelli and Castelletti, 2013a and 2013b). While
23 these studies demonstrate that a number of methods exist for characterizing model structure,
24 they generally focus on a single model type and thus provide little insight into the
25 comparative ease with which different model types can be interpreted.

26 While a number of comparison studies exist that apply multiple empirical models to a
27 given problem, finding generalizable insights from these studies is hindered because of the
28 limited number of models and datasets evaluated. Perhaps the most comprehensive
29 comparison to date is that of Elshorbagy et al. (2010a and 2010b), who compared six methods
30 for data-driven modeling of daily discharge in the Ourthe River in Belgium. This work found
31 that linear models were able to perform comparably to much more complex methods when the
32 data content of the models were limited, or when system input-output behavior was close to

1 linear. However, other studies have demonstrated the value of using more complex
2 approaches when modeling more complex rainfall-runoff behavior (e.g., Abrahart and See,
3 2007; Asefa et al., 2006). The differing results obtained across these studies indicate that no
4 single method is likely to be suitable for all basins, timescales, or applications.

5 However, it is important to recognize that predictive accuracy alone is not necessarily
6 sufficient justification for applying a model to a given problem. Models should not only be
7 accurate, but also be fit-for-purpose (Beven, 2011; Van Griensven et al., 2012). For instance,
8 accurate representation of low return period flows is more important in a flood forecasting
9 model than one aimed at predicting average amounts of water available for withdrawal and
10 human consumption. Similarly, the ability to provide insights into physical watershed
11 function may be more important in basins where land-use change could alter the hydrologic
12 regime, compared to a basin that is heavily urbanized and expected to remain so. The use of
13 multiple objective functions in training data-driven models can address this to some degree by
14 identifying models that provide sufficient balance between different performance objectives,
15 such as accurate representation of different portions of the flow hydrograph (De Vos and
16 Rientjes, 2008). However, more refined model training procedures will not necessarily
17 address other aspects of model performance that make it suitable for planning purposes, such
18 as interpretability (Solomatine and Ostfield, 2008). More comprehensive consideration of
19 model strengths and limitations should be standard practice in model development and
20 selection, rather than simply evaluating global error metrics.

21 In this work, we compare six methods for empirical streamflow prediction (linear models,
22 generalized additive models, multivariate adaptive regression splines, random forests, M5
23 model trees and ANNs) in their ability to predict monthly streamflow in five rivers in the
24 Lake Tana basin in Ethiopia. This study region was selected as it provides insights into the
25 use of data-driven models for streamflow simulation in tropical regions of the world that are
26 underrepresented in existing studies; for instance, a review of 210 articles on water resource
27 applications of ANNs found that over three quarters of the studies evaluated were conducted
28 in North America, Europe, Australia, or temperate East Asia (Maier et al., 2010). Existing
29 studies conducted in tropical regions generally apply a single methodology to the basin of
30 interest and evaluate predictive accuracy alone (see for instance, Machado et al., 2011;
31 Chibanga et al., 2003; Antar et al., 2006; Aqil et al., 2007), making it difficult to find
32 generalizable insights into the relative advantages of different modeling approaches in these

1 regions. Better development of data-driven models for these regions has the potential to be
2 particularly valuable because data limitations and complex hydrodynamic processes often
3 hinder the use of physical watershed models, but relatively long time series of streamflow,
4 precipitation and temperature may be available at a monthly timescale. These data, combined
5 with information on relevant landscape change (in particular, the expansion of agricultural
6 land cover), can be leveraged to create reasonably accurate empirical models.

7 Models are compared not only in terms of their predictive accuracy, but also in terms of
8 model error structure and the implications that this structure may have for water resource
9 applications. Additionally, we evaluate the methods by which model structure and predictor
10 variable influence can be evaluated to gain insights into physical system function for each
11 model type. Finally, we assess the suitability of using different model types for climate
12 change impact assessment by comparing model uncertainty in projections made for
13 increasingly extreme climate conditions. The overall objective of this research is not to
14 identify a single “best” model, but rather to highlight some of the strengths and limitations of
15 different approaches, as well as demonstrate important issues that should be kept in mind for
16 model comparisons in the future

17 **2 Data and Methods**

18 **2.1 Study Area**

19 Lake Tana is located at an elevation of approximately 1800 meters in the highlands of
20 northwest Ethiopia (Fig. 1). The catchment draining to the lake encompasses approximately
21 12,000 square kilometers, and the four main tributaries providing water to the lake are the
22 Gilgel Abbay (including its tributary, the Koga River), Ribb, Gumara, and Megech Rivers.
23 Collectively, these rivers account for 93% of the inflow to the lake (Alemayehu et al., 2010).
24 Ninety percent of rainfall in the basin occurs during the wet season from May until October,
25 and there is significant interannual variability in precipitation with annual rainfall levels
26 ranging from below 1000 mm to over 1800 mm (Achenef et al., 2013). Population growth and
27 expansion of agricultural and pastoral land use in the region has resulted in substantial
28 deforestation and land degradation, with agricultural, pastoral and settled land cover
29 comprising over 70% of the basin’s surface area (Rientjes et al., 2011; Garede and Minale,
30 2014; Gebrehiwot et al., 2010). There is some evidence that this has impacted the hydrology

1 of the rivers draining into the lake (Gebrehiwot et al., 2010). A summary of basin
2 characteristics for the evaluation period of 1960-2004 is presented in Table 1.

3 Approximately 2.6 million people live in the basin, and are largely settled in rural
4 areas and reliant on rainfed subsistence agriculture. This makes the region quite vulnerable to
5 climate variability and change, and a number of water resources infrastructure projects are
6 planned to better manage this vulnerability and support economic development (Alemayehu
7 et al., 2010). This includes the recent construction of the Tana-Beles hydropower transfer
8 tunnel and the Koga River irrigation reservoir, as well as five other reservoirs planned for
9 construction in the next 10 to 20 years (Alemayehu et al., 2010). To better understand the
10 potential implications of this development, extensive effort has been put towards developing
11 rainfall-runoff models for the Lake Tana basin, as well as other areas of the Ethiopian
12 highlands with similar characteristics (van Griensven et al., 2012). Many of these studies rely
13 on Soil and Water Assessment Tool (SWAT) models, although there are some that use water
14 balance approaches (Van Griensven et al., 2012). While these models have in some cases
15 demonstrated reasonably high accuracy, previous evaluations were largely based on Nash-
16 Sutcliffe Efficiency (NSE; Nash and Sutcliffe, 1970) which can be a flawed performance
17 metric in highly seasonal watersheds (Schaefli and Gupta, 2007; Legates and McCabe, 1999).
18 More importantly, the limited data available for physical parameterization of these models
19 required a heavy reliance on model calibration, which sometimes resulted in parameterization
20 schemes that are inconsistent with physical understanding of the region's hydrology
21 (Steenhuis et al., 2009; van Griensven et al., 2012). Furthermore, a number of studies relied
22 on empirical relationships such as curve numbers and the Hargreave's equation that were
23 developed for temperate regions (e.g., Mekonnen et al., 2009; Setegne et al., 2009). While
24 these limitations are likely to introduce considerable uncertainty into model projections,
25 particularly in situations where climatic or environmental conditions differ from those
26 experienced in the calibration period, few studies from this region of Ethiopia include any sort
27 of uncertainty analysis in model predictions. Empirical models could provide a useful
28 complement to physical models developed for the region by providing insights into physical
29 system function and allowing for more comprehensive uncertainty analysis.

30 **2.2 Data and Model Development**

31 Models were developed using monthly streamflow, climate, and land cover data for
32 the period from 1961 to 2004, resulting in 528 monthly observations. In each of the five major

1 rivers in the basin, we developed empirical models that estimated monthly streamflow as a
 2 function of climate conditions and agricultural land cover in each basin. Monthly streamflow
 3 data were taken from historic stream gauge records for each basin, as reported in feasibility
 4 studies developed for proposed irrigation projects (Alemayehu, 2010). Historic data for
 5 monthly average temperature, monthly total precipitation, and monthly wet days in each river
 6 basin were derived from the University of East Anglia Climate Research Unit (CRU) TS3.10
 7 gridded meteorological fields (Harris et al., 2014), which are based on meteorological station
 8 observations. Historic estimates of rainfall intensity were also calculated by dividing monthly
 9 total precipitation by CRU TS3.10 records of the number of wet days in that month, but was
 10 found to be highly correlated with monthly precipitation and did not result in significant
 11 improvements to the predictive accuracy of tested models. Thus, it was not included in the
 12 final model formulations. Finally, to account for historic increases in agricultural and pastoral
 13 land cover that have occurred in the basin, the percentage of land cover used for any crop or
 14 grazing was estimated from historic land cover analyses described by Rientjes et al. (2011),
 15 Gebrehiwot et al. (2010), and Garede and Minale (2014). These studies used historic aerial
 16 photos and satellite images to estimate land cover changes in the Ribb, Gilgel Abbay, and
 17 Koga basins from the periods of 1957 to 2011. The percentage of agricultural land cover was
 18 interpolated for years when data weren't available, and the value of agricultural land cover in
 19 the two basins without data was assumed to be equal to average agricultural land cover in the
 20 basins with data. Land cover was assumed to change on an annual, rather than monthly basis.
 21 While this approach is prone to errors that could stem from differing rates of land use change
 22 through time and between basins, it does provide a mechanism for capturing the long-term
 23 trend of expanding agricultural land cover that has been observed throughout the Ethiopian
 24 highlands when detailed land-cover data are unavailable. Including this data improved out-of-
 25 sample predictive accuracy of the models, further suggesting that it was a valuable addition.

26 Two general formulations for the empirical models were evaluated. The first (referred
 27 to below as the standard model formulation) was

$$28 \quad \log(Q_{b,t}) = f(P_{b,t}, P_{b,t-1}, P_{b,t-2}, T_{b,t}, T_{b,t-1}, T_{b,t-2}, AgLC_{b,t}) + \varepsilon_{b,t} \quad (1)$$

29 where $Q_{b,t}$ is the monthly streamflow in river b at time period t , $P_{b,t}$ and $T_{b,t}$ are the monthly
 30 total precipitation and average temperature in river basin b at time period t , $AgLC_{b,t}$ is the total
 31 percentage of agricultural land cover in basin b at time t , and $\varepsilon_{b,t}$ is the model error. The

1 subscripts $t-1$ and $t-2$ indicate lagged measurements from one and two months prior, and were
2 included to roughly account for storage times longer than one month that could impact
3 streamflow in each river. While the exact time of concentration is not known in each basin,
4 the minor influence of of climate conditions at two months prior suggest that climate
5 conditions from beyond this time period do not contribute significantly to flow variability.
6 The function f represents a general function that differed between the specific models assessed
7 and is discussed in more detail below. The logarithm of monthly streamflow was used as a
8 response variable to keep model predictions positive.

9 In the second formulation, streamflow and climate anomalies were used as the
10 response and predictor variables to better account for the highly seasonal nature of streamflow
11 and precipitation in the region. Streamflow anomalies were calculated for each observation by
12 subtracting the long-term average streamflow for that month (m) from the observed value and
13 dividing this number by the long-term standard deviation of that month's streamflow as in Eq.
14 (2). This procedure was repeated for precipitation and temperature, and these values were then
15 used to fit models of the form described in Eq. (3). It should be noted that although this
16 formulation uses long-term averages and standard deviations to convert anomaly values to
17 flow volumes, the anomaly values themselves are calculated based on climatic and land cover
18 conditions that are nonstationary through time.

$$19 \quad Q_{b,t}^{AN} = \frac{Q_{b,t} - \bar{Q}_{b,m}}{sd(Q_{b,m})} \quad (2)$$

$$21 \quad Q_{b,t}^{AN} = f(P_{b,t}^{AN}, P_{b,t-1}^{AN}, P_{b,t-2}^{AN}, T_{b,t}^{AN}, T_{b,t-1}^{AN}, T_{b,t-2}^{AN}, AgLC_{b,t}) + \varepsilon_{b,t} \quad (3)$$

22 Six different types of models were compared using each formulation in each basin:

- 23 1. A Gaussian linear regression model (GLM) using the basic stats package in the R
24 statistical computing software (R Development Core Team, 2014)
- 25 2. Gaussian generalized additive model (GAM): GAMs are a semi-parametric
26 regression approach where the response variable is estimated as the sum of
27 smoothing functions applied over predictor variables. These functions allow the
28 model to capture non-linear relationships between the predictor and response
29 variables without *a priori* assumptions about the form (eg., quadratic, logarithmic)
30 of these functions, and are fit using penalized likelihood maximization to prevent

1 model overfitting (Hastie and Tibshirani, 1990). GAMs were fit using the mgcv
2 package in R (Wood, 2011).

- 3 3. Multivariate adaptive regression splines (MARS): MARS are a non-parametric
4 regression approach where the response variable is estimated as the sum of basis
5 functions fit to recursively partitioned segments of the data (Friedman, 1991).
6 MARS models were fit using the earth package in R (Milborrow, 2015).
- 7 4. Artificial neural network (ANN): ANNs are a non-parametric regression approach
8 represented by a network of nodes and links that connects predictor variables to
9 the response variable. Each link in the network represents a function that maps the
10 input nodes into the output node (Ripley, 1996). ANN models were fit using the
11 nnet package in R (Venables and Ripley, 2013).
- 12 5. Random forest (RF): Random forests are a rule-based, non-parametric regression
13 approach where the model prediction is created by averaging the predicted value
14 from multiple regression trees which are trained on separate bootstrapped
15 resamples of the data. Each tree is fit using a small, randomly selected subset of
16 predictor variables, resulting in reduced correlation between trees (Breiman,
17 2001). Random forest models were fit using the randomForest package in R (Liaw
18 and Wiener, 2002).
- 19 6. M5 model: M5 models are a rule-based, non-parametric regression approach that
20 fits a linear regression model to each terminal node of a regression tree (Quinlan,
21 1992). M5 models were fit using the Cubist package in R (Kuhn et al., 2014).
- 22 7. Climatology model: A climatology model that simply predicted each month's
23 streamflow as equivalent to the long-term average streamflow for that month was
24 included for comparison purposes.

25 **2.3 Model Evaluation**

26 When using non-parametric regression approaches, it is important to avoid overfitting a
27 model to a given dataset because this can result in large errors in out-of-sample predictions
28 (Hastie et al., 2009). To avoid model overfit, the caret package in R (Kuhn, 2015) was used to
29 determine model parameters for the MARS, ANN, RF and M5 models. This package uses
30 resampling to evaluate the effect that model parameters have on the model's predictive
31 performance and chooses the set of parameters that minimizes out-of-sample error (Kuhn
32 2015). In this evaluation, 25 bootstrap resamples of the training dataset were generated for

1 each parameter value to be assessed. A model was fit using each bootstrap sample and used to
2 predict the remaining observations, and the parameter values that minimized average RMSE
3 across all resamples. Details on the specific parameters evaluated for each model are
4 presented in Table 2. While the development of more complex structures are possible for
5 some models, this process can result in over-parameterization and poor model performance
6 (Gaume and Gosset, 2003; Han et al., 2007). Additionally, the use of a standardized
7 parameterization procedure allows for a more even comparison between different model
8 types.

9 The predictive ability of each model was assessed using 50 random holdout cross-
10 validation samples. In each sample, a random selection of years were chosen, and
11 observations from these years were removed (“held-out”) from the dataset. The size of the
12 held-out sample ranged from 1 to 9 years. Each model was then fit to the remaining portion of
13 the data, using the caret package described above to determine model parameters for the
14 MARS, ANN, RF and M5 models. These models were then used to predict streamflow for the
15 held-out portion of the data, and both the mean absolute error (MAE) and NSE were
16 calculated after transforming model predictions after back to the original streamflow units.
17 Mean MAE and NSE were calculated for each model across the 50 cross-validation samples
18 and used to choose the model with the highest predictive accuracy in each basin. This cross-
19 validation procedure provides a mechanism for evaluating how well a model will generalize
20 to an unseen set of data while avoiding some of the problems that can arise from the use of a
21 single calibration and validation dataset (Elshorbagy et al., 2010a; Han et al., 2007).

22 MAE was included as an error metric because it provides a simple and easily
23 interpretable measure of error on the same scale as observed flow volumes. While NSE values
24 are acknowledged to be a flawed performance metric in highly seasonal watersheds where
25 seasonal fluctuations contribute to a substantial portion of flow variability (Schaeffli and
26 Gupta, 2007; Legates and McCabe, 1999), this metric was included to provide a rough
27 comparison of how empirical model performance compared to the performance of physical
28 models developed for the region. The use of alternative error metrics has been discussed
29 extensively in the literature (for instance Pushpalatha et al., 2012; Mathevet et al., 2006; Criss
30 and Winston, 2008), and could provide additional insights into what contributes to predictive
31 capabilities of different model formulations. However, this work examined predictive

1 accuracy based on MAE and NSE alone to allow for greater focus on how models differ in
2 terms of error structure and uncertainty.

3 As a rough point of comparison for the statistical models developed in this research, we
4 also evaluated discharge estimates derived from a process-based hydrological model. The
5 model used in this application is the Noah Land Surface Model version 3.2 (Noah LSM; Ek
6 et. al, 2003; Chen et al., 1996). Noah LSM was implemented for offline simulations of the
7 Lake Tana basin at a gridded spatial resolution of 5km for the period 1979-2010 using a time
8 step of 30 minutes. Meteorological forcing was drawn from the Princeton 50-year reanalysis
9 dataset (Sheffield et al. 2006), downscaled to account for Ethiopia's steep terrain using
10 MicroMet elevation correction equations (Liston & Elder 2006). The Princeton reanalysis was
11 selected because it provides relatively high resolution meteorological fields, including all
12 variables required to run a water and energy balance LSM like Noah, for the period 1948-
13 present. While higher resolution and possibly higher quality datasets are available for recent
14 years, this longer dataset was utilized to compare the process-based model to statistical
15 models developed for a long historical period. Soil parameters for the Noah simulation were
16 drawn from the FAO global soil database, land use was defined according to the United States
17 Geological Survey (USGS) global 1km land cover product, and vegetation fraction was
18 derived from MODerate Imaging Spectroradiometer (MODIS) imagery. Land cover was
19 treated as a static parameter over the full length of the simulation, as spatially complete
20 estimates of historical land use were not available at the required resolution and specificity.

21 The highest performing model in each basin based on MAE was retained for more
22 detailed evaluation of model error structure, covariate influence, and uncertainty in climate
23 change sensitivity analysis. To generate a complete time-series of out-of-sample model
24 predictions for error analysis, the holdout cross validation procedure was repeated for the
25 highest performing standard-formulation and anomaly-formulation models for each basin, but
26 this time holding out a single year of observations in each iteration. The predictions from this
27 cross validation were used to evaluate the how model error structure might impact model
28 predictions used for water resource applications. The influence of different predictor variables
29 on model predictions was also assessed for the highest performing model in each basin after
30 being fit to the complete dataset. Each predictor variable was assessed using metrics for
31 covariate importance and influence that are unique to that model type, demonstrating how
32 models could be used to gain physical insights about data-scarce regions and the mechanisms

1 for generating these insights for each type of model. Partial dependence plots (Hastie et al.,
2 2009) were also generated for each covariate for the highest performing model in each basin
3 to provide insights about how covariate influence compared across different basins and model
4 types.

5 Finally, two evaluations were conducted to assess uncertainty in model projections of
6 streamflow under increasingly extreme climate conditions to better understand the
7 implications of using different model formulations for climate change impact studies. Model
8 projections of streamflow in different climate conditions are likely to be accompanied by
9 considerable uncertainty, particularly when climate conditions exceed those experienced
10 historically. To assess this uncertainty, the best performing model in each basin was used to
11 generate streamflow predictions for 1) changes in temperature from 0 to 5° C, 2) changes in
12 precipitation from -30 to +30%, 3) an increase in temperature to 5° C combined with a
13 decrease in precipitation to -30%, and 4) an increase in temperature to 5° C combined with an
14 increase in precipitation to +30%. For each of the four assessments, the models generated
15 predictions for the 45-year historic climate record adjusted for a given degree of climate
16 change using the delta-change method (Gleick, 1986), while holding agricultural land cover
17 constant at 60%. In this method, monthly temperature values are simply added to the
18 temperature change value, and monthly precipitation values are multiplied by the precipitation
19 change percentage. Model predictions for the altered climate record were then used to
20 calculate the average annual streamflow in each river. This process was repeated 100 times
21 for models fit on random bootstrap resamples of the historic dataset to generate uncertainty
22 bounds surrounding model predictions and evaluated how the uncertainty in these predictions
23 increased as climate conditions became more extreme. It is important to recognize that these
24 should not be interpreted as a prediction or assessment of actual climate change impacts, but
25 rather a measurement of the sensitivity of modeled streamflow in the basin to different
26 climate conditions. Since one of the key motivations for using rainfall-runoff models is to
27 understand how climate change may impact water resources, it is important to understand
28 how model formulation contributes to this sensitivity and uncertainty.

1 **3 Results**

2 **3.1 Model Accuracy and Error Structure**

3 Table 3 shows the out-of-sample cross validation errors for each model assessed in each
4 basin. The random forest model had the lowest mean absolute error for the standard-
5 formulation model in four of the five basins, with the M5 model performing best for the Koga
6 basin. These models outperformed the Noah LSM simulations in all basins assessed. The
7 Noah LSM errors are for a single period of analysis and thus don't present an exact corollary
8 to the cross validation performed for the empirical models. Nevertheless, the significant
9 increases in errors associated with the Noah LSM model demonstrates the difficulty
10 associated with the use of process-based models in the region, particularly when relying on
11 global datasets that may be unreliable at the spatial and temporal resolutions required for
12 physical modeling. Physical models developed for monthly streamflow prediction in other
13 basins within the Ethiopian highlands have reported NSE values ranging from 0.53 to 0.92
14 (van Griensven et al., 2012), compared to values ranging from 0.71 to 0.87 for the random
15 forest models developed here. If this measure alone was used for model evaluation, these
16 empirical models would generally be classified as having good performance based on the
17 guidelines suggested by Moraisi et al. (2007). However, the climatology model outperforms
18 the best standard formulation models in all basins except Megech, indicating that in the
19 majority of basins the errors from the fitted empirical models are higher than those that result
20 from simply using the long-term monthly average for each month's prediction. This is due to
21 the fact that seasonality accounts for such a large portion of the variability in monthly flow
22 values, and demonstrates how high NSE values can be quite easy to obtain in seasonal basins.

23 Evaluation of anomaly model errors indicates that the models using this formulation
24 achieve better predictive accuracy than those using the standard formulation, and are able to
25 outperform the climatology model based on both NSE and MAE in all basins. However, the
26 highest performing models in each basin varies more when the anomaly formulation is used,
27 with the GLM, GAM, random forest, and M5 models all minimizing MAE in different basins.
28 In all basins except Koga, the highest performing model significantly outperformed the
29 climatology model based on paired Wilcoxon rank-sum tests (Bonferroni-corrected p-value <
30 0.01).

1 Further exploration of model residuals indicates another important advantage of using
2 the anomaly model formulation. In the standard model formulation, model residuals appear to
3 be non-random. Example autocorrelation plots are shown for the Gilgel Abbay and Ribb
4 Rivers in Fig. 2, and demonstrate that a positive autocorrelation exists at the 12 month time
5 lag. For brevity, only plots for two rivers are shown, although this autocorrelation existed in
6 the standard-formulation models for all basins except Megech (Table 4). This autocorrelation
7 occurs because the standard-formulation models consistently underestimate wet-season
8 streamflow while overestimating dry-season flows, as is apparent in hydrographs of observed
9 and predicted streamflow (Fig. 3). Because wet-season flows contribute such a large portion
10 of the total annual flow volume, this results in regular underestimation of aggregate values
11 such as mean annual flow (Table 4). This autocorrelation is reduced in the anomaly-
12 formulation models, meaning that they are better able to capture the peak flow volumes
13 experienced in the wet season and do not underestimate mean annual flow to the same degree
14 that the standard formulation models do.

15 **3.2 Model Structure and Covariate Influence**

16 Evaluating the relationship between predictor covariates and streamflow response can
17 lend insight into the physical processes underlying runoff generation in each basin. There are
18 two components of this relationship that can be evaluated: how much each covariate
19 contributes to model accuracy (covariate importance), and the direction and nature of the
20 relationship between covariate values and model response (covariate influence). In many
21 machine-learning models, complete description of the all of the mathematical relationships
22 within the model (for instance, through description of each tree comprising a random forest
23 model) is infeasible, requiring the use of other mechanisms for understanding covariate
24 importance and influence. However, because each model type is structured in a different way,
25 these mechanisms differ. This section first describes the mechanisms available for obtaining
26 insights about covariate influence in each of the highest performing models. To provide a
27 mechanism for comparing results across different basins, each basin model is then assessed
28 using the general approach of partial dependence plots.

29 In the Gilgel Abbay and Koga basins, the highest performing model was a simple
30 linear regression model. These models can be evaluated by reviewing model coefficients and
31 associated p-values, as shown in Table 5. In a standard linear regression, model coefficients
32 can be interpreted as the mean change in the response variable that results from a unit change

1 in that covariate when all others are held constant. These coefficients are for streamflow
2 anomalies rather than raw values, making their immediate interpretation less intuitive. For
3 instance, in the Gilgel Abbay model an increase of one standard deviation in precipitation
4 results in an increase of 0.22 standard deviations in flow. The associated p-value for each
5 coefficient evaluates a null hypothesis that the true coefficient value is equal to zero given the
6 other covariates in the model, and thus has no influence on the response variable.

7 Evaluating model structure based on regression coefficients is appealing due to their
8 simplicity and familiarity. However, it is important to keep in mind that the above
9 interpretations rely on specific assumptions regarding model error distributions. Examination
10 of fitted model residuals from both basins indicate that errors are autocorrelated in the Koga
11 basin and not normally distributed due to the presence of outliers in both basins. Non-
12 normality and autocorrelation both impact the t statistics and f statistics used to test for the
13 significance of model coefficients, and thus the p-values for these models are likely biased
14 (Montgomery et al., 2012).

15 Interpretation of variable influence in GAMs is based on the estimated degrees of
16 freedom (EDF) a covariate's smoothing function $s(X_i)$ uses within a model (Hastie and
17 Tibushini, 1986). An EDF value of one or below indicates a linear function relating the
18 response variable to that covariate, while values greater than one represent a non-linear
19 smoothing function. An EDF value of zero indicates that the covariate smoothing function is
20 penalized to zero (meaning it has no influence on model predictions). In the model for the
21 Megech River, the terms for lagged temperature at one and two months, as well as
22 precipitation lagged at two months were all smoothed to zero. Of the remaining covariates,
23 lagged precipitation has a linear impact on model response, while precipitation, temperature
24 and land cover have non-linear impacts. Smoothing functions can be plotted to gain more
25 insight about these relationships (Fig. 4). The functions for precipitation anomaly, lagged (one
26 month) precipitation anomaly, and agricultural land cover show a positive relationships with
27 streamflow, while the function for temperature anomaly predicts low streamflow at both high
28 and low anomalies.

29 P-values test the null hypothesis that a covariate's smoothing function is equal to zero,
30 but rest on the assumption that model residuals are homoscedastic and independent (Wood,
31 2012). Similar to the linear models, residuals in the Megech GAM model appear to be both
32 autocorrelated and heteroscedastic, meaning that a formal statistical interpretation of this

1 value may be inappropriate and that confidence bounds around smoothing functions might be
2 misleading.

3 The M5 cubist model fit for the Gumara basin is an ensemble of 100 small M5
4 regression trees. In each tree, the model splits observations based on logical rules related to
5 one or more covariates and fits a linear regression model to each set of observations. The final
6 model prediction is the average across all of the individual trees. Using this sort of ensemble
7 approach can reduce model variance and improve accuracy if the individual trees are
8 unbiased, uncorrelated predictors (Breiman 1996). This can be useful in avoiding models that
9 are overfit to the data, but can reduce model interpretability since direct visualization of
10 model structure becomes impractical as the number of trees increases. However, the
11 frequency with which individual covariates are used as splitting points within trees and as
12 regression coefficients can provide some insights about covariate importance (Table 5; note
13 that because multiple covariates can be used for rules and linear models, these don't
14 necessarily add to 100%). Model rules were largely based on land cover, with some rules
15 based on precipitation. These two covariates were also used most frequently in linear
16 regressions at model nodes, followed by temperature (current and 1-month lag) and 1-month
17 lagged precipitation. Notably, climate data from 2 months lagged were not used at all. While
18 this can be useful in identifying which covariates have the largest impact on model
19 predictions, it doesn't provide any information regarding the nature or direction of that
20 influence.

21 Similarly, the random forest model developed for the Ribb basin is an ensemble of
22 regression trees in which the final model prediction is the average of the predictions from
23 each individual tree. However, random forests use standard regression trees that do not
24 incorporate linear regression models at terminal nodes. Variable importance within the final
25 model is measured by recording the increase in out-of-sample MSE that results when a
26 covariate is randomly permuted for each tree in the ensemble. This increase in error is then
27 averaged across all trees in the ensemble. In our model, the largest increases in error resulted
28 from permutation of land cover and temperature, followed by 2-month lagged temperature
29 and precipitation. Covariate influence can be evaluated through the use of partial dependence
30 plots, which measure the change in model predictions that result from changing the value of
31 one parameter while leaving all other covariates constant (Hastie et al., 2009). Partial
32 dependence plots indicate that model predictions of streamflow are higher when the percent of

1 agricultural land cover is greater than approximately 75%, when temperatures anomalies are
2 low, and when precipitation anomalies are high. However, it appears that the plot for lagged
3 temperature might be sensitive to outliers at high temperature anomalies as evidenced by the
4 large increase that occurs above an anomaly of +2, in a region where very few data points are
5 present.

6 Many of the measures used to evaluate covariate importance and influence are model
7 specific, making inter-basin and inter-model comparisons difficult. However, the partial
8 dependence plots used in the randomForest R package can be developed for any model and
9 provide a mechanism for comparing the influence that covariates have in the different models
10 and basins (Shortridge et al., 2015). Partial dependence plots were generated for each basin's
11 best performing model and results are shown for climatic variables in Fig. 6. As expected,
12 models generally respond positively to increases in precipitation and negatively to increases
13 in temperature, with the greatest influence in the current month and decreasing influence at
14 one and two months prior. The influence of the current month's precipitation is linear in three
15 of the five basins; while this is constrained to be the case in the Gilgel Abbay and Koga
16 basins due to the use of a linear model, the linear response in Gumara is not required from the
17 M5 model structure. Interestingly, both Megech and Ribb demonstrate a linear response to
18 negative precipitation anomalies, but little response to positive anomalies. Streamflow
19 response to temperature is strongest in the Gumara basin; interestingly, this is the basin with
20 the smallest response to precipitation.

21 The partial dependence plots for the percentage of the basin classified as agricultural
22 land cover indicates a positive relationship between agricultural land cover and streamflow in
23 all basins except for the Gilgel Abbay (Fig. 7). This would be expected if deforestation had
24 contributed to a decrease in evapotranspiration in the contributing watersheds. The exact
25 nature of this response differs across the different rivers, with the relatively minor responses
26 in Koga and Ribb, and much stronger responses in the Gumara and Megech basins. However,
27 this plot also demonstrates some of the limitations associated with different model structures.
28 The plot for Gumara is highly erratic, indicating that the M5 model might be overfit to the
29 training dataset, despite the use of model averaging to reduce model variance. Additionally,
30 the GAM used in the Megech basin was only trained on agricultural land cover values up to
31 77%; while this model may be accurately representing the impact of land cover changes

1 within this range, extrapolating this relationship to higher values leads to predictions that may
2 not be physically realistic.

3 **3.3 Climate Change Sensitivity and Uncertainty Assessment**

4 Fig. 8 shows the results of the climate change sensitivity analysis for total flow from all
5 five tributaries, with dashed lines representing 95% confidence intervals obtained through 100
6 bootstrapped resamples of the data set. As would be expected, increasing temperature
7 independently of precipitation results in decreasing total flows while increasing precipitation
8 results in higher flows. However, the uncertainty surrounding temperature sensitivity
9 increases at higher changes in temperature, while the uncertainty surrounding precipitation
10 sensitivity remains relatively constant, even at extreme changes in annual precipitation. The
11 bottom panels of the figure show the sensitivity of total inflows to concurrent changes in
12 temperature and precipitation. Unsurprisingly, decreasing precipitation combined with higher
13 temperatures results in greater decreases in total flow than when temperature and precipitation
14 are varied independently. However, even if temperature increases are combined with higher
15 precipitation, total flows decline in the majority of bootstrap resamples.

16 The uncertainty surrounding temperature sensitivity is a key limitation to using data-
17 driven approaches for climate impact assessment. To better understand which models and
18 basins are contributing to this uncertainty, Fig. 9 shows how the coefficient of variation (the
19 standard deviation of predictions from all bootstrap samples divided by the mean of these
20 predictions) varies as a function of temperature change in each basin. From this figure, it is
21 apparent that the Megech model is by far the largest contributor to model uncertainty;
22 however, it is not clear whether this contribution is due to model structure (the GAM model
23 used for the Megech River) or characteristics associated with the basin itself. To investigate
24 how different model structures contributed to this uncertainty, the bootstrap resampling
25 procedure was used to assess uncertainty in streamflow predictions in the Gumara River from
26 all model types. This basin was chosen because all six models were able to outperform the
27 climatology model, and thus could be considered good choices for model selection based on
28 predictive accuracy alone. The results indicate that the increase in uncertainty is highest, and
29 increases non-linearly, in the GLM, GAM, and MARS models. Uncertainty increases more
30 slowly in the ANN and M5 models, and no noticeable increase in uncertainty is apparent in
31 the random forest model.

1 **4 Discussion**

2 The objective of this study was not to identify the “best” approach for empirical
3 rainfall-runoff modeling, as this is likely to be highly specific to the basin and problem to
4 which a model is applied. However, we hope that the comparison conducted here can
5 highlight some of the strengths and limitations of different approaches, as well as demonstrate
6 some important issues that should be kept in mind for model comparisons in the future. One
7 important finding was the limitation with using NSE as an error metric. Our results confirm
8 previous studies that found that even uninformative models able to capture basic seasonality
9 are able to achieve high NSE values (Legates and McCabe, 1999; Schaefli and Gupta, 2007),
10 and provide further evidence indicating that high NSE values should be considered a
11 necessary but not sufficient requirement for model usage in planning situations. For instance,
12 the simple climatology model used for comparison purposes here is able to achieve high NSE
13 values, but would be unsuitable for planning since it does not account for any interannual
14 variability nor the possibility for non-stationary conditions caused by changing climate and
15 land cover. In particular, understanding error structure can be valuable in evaluating whether
16 model biases might undermine the model’s suitability for management activities. In our
17 example, the autocorrelation present in the standard-formulation models meant that these
18 models were consistently underestimating wet-season flows, resulting in low estimates of the
19 total annual flow in the rivers. Since multiple reservoirs are planned for construction on these
20 rivers to support irrigation activities, this bias could lead to poor estimates of how much water
21 is available for agricultural use in the short term (ie., seasonal forecasting) and long-term (due
22 to climate change). Interestingly, difficulties in accurately capturing high flows has been
23 observed in physical hydrologic models for Ethiopia (e.g., Setegne et al., 2011; Mekonnen et
24 al., 2009) and more generally (e.g., Wilby, 2005). The implications of this limitation should
25 be carefully evaluated before using models for water resource planning or (more importantly)
26 flood risk evaluation.

27 Depending on the model type used, different mechanisms are available to evaluate
28 covariate importance and influence within the model. This evaluation can be useful in
29 confirming that the model is replicating physically realistic relationships between input and
30 output variables. While the relationships identified in this evaluation are fairly straightforward
31 (for example, increasing runoff with higher precipitation and lower temperatures), these
32 simple relationships are still important in highlighting the mechanisms by which the models

1 make predictions so that they are not “black boxes.” For instance, Han et al. (2007) explore
2 how ANN flood forecasting models responds to a double-unit input of rain, finding that some
3 formulations respond in a hydrologically meaningful way to increased rainfall intensity, while
4 others do not. Similarly, Galelli and Castelletti (2013a) describe how input variable
5 importance can be used to highlight differences in hydrologic processes between an urbanized
6 and forested watershed. The easy manner in which covariate relationships within the GAM
7 and random forest models can be visualized using a single command within their respective R
8 packages is a strong advantage to these approaches compared to methods such as M5 model
9 trees and artificial neural networks. Of course, partial dependence plots can be developed for
10 any model type (as was done in this research), but code must be written by the user and thus
11 requires a higher degree of effort than is necessary for in-package functions. A downside to
12 most machine-learning models is that they do not support the statistical formalism in
13 assessing variable importance that is possible when linear models and GAMs are used.
14 However, this formalism often rests on assumptions regarding model residuals that are
15 unlikely to be met in many hydrologic models (Sorooshian and Dracup, 1980).

16 Within the Lake Tana basin, evaluation of covariate influence indicates that each
17 basin’s model is performing in a physically realistic manner, with a runoff increasing with
18 higher precipitation levels and decreasing with higher temperatures. The influence of
19 precipitation and temperature is greatest in the current month, and progressively declines to a
20 very small influence after two months. This suggests that long-term (multi-month) storage
21 does not significantly contribute to variability in flow volumes. One interesting finding is the
22 non-linear relationship between concurrent month precipitation and runoff that exists in the
23 Megech and Ribb basins, which suggests that above a certain point increasing rainfall does
24 not result in a commiserate increase in streamflow. Other studies have noted the dampening
25 effect that wetlands and floodplains have had on river flows in the region (Dessie et al., 2014;
26 Gebrehiwot et al., 2010); this phenomenon could explain the non-linear relationship identified
27 in this work. The clearly negative relationship between temperature and runoff demonstrates
28 the degree to which upstream evapotranspiration impacts streamflow and suggests that
29 evapotranspiration is largely energy-limited, rather than water-limited. Increasing agricultural
30 land-use appears to be associated with higher runoff in all rivers except for Gilgel Abbay
31 (where no clear relationship between land cover and runoff was observed), and suggests that
32 agricultural expansion at the expense of forest cover has reduced the evaporative component
33 of the water balance in these basins. Finally, the relative performance of different model

1 formulations themselves can also be informative. For instance, the improved performance of
2 the anomaly-formulation models indicates that the relationship between precipitation and
3 runoff varies throughout the year and could point towards differences in runoff-generating
4 mechanisms in the wet and dry seasons that have been observed in other case studies (Wilby,
5 2005).

6 One limitation with data-driven approaches for streamflow prediction is that the
7 relationships they model can only generate reliable predictions for conditions that are
8 comparable to those experienced historically. Using these models to generate predictions for
9 conditions that exceed historic variability is likely to introduce considerable uncertainty into
10 their projections. Our results indicate that uncertainty in projections of streamflow under
11 changing precipitation is relatively constant, whereas uncertainty increases markedly in
12 projections of streamflow under increasing temperature. This result is not surprising when one
13 considers the basin's climate, which is characterized by highly variable rainfall but fairly
14 consistent temperatures (Table 6). A temperature increase of 3° C equates to almost two
15 standard deviations beyond the historic mean, whereas a change in precipitation of 30% is
16 well within the range of conditions experienced historically. One would expect that in other
17 climates (for example, temperate watersheds with only minor changes in rainfall throughout
18 the year), this relationship could be reversed. Despite the uncertainty that exists in projections
19 of streamflow under changing temperature, total annual flow appears to be quite sensitive to
20 increasing temperatures. In fact, the decreases in streamflow due to increasing temperature
21 appears likely to be more than enough to counteract any increases in streamflow resulting
22 from higher precipitation that is projected for the region in some global circulation models
23 (GCMs). This is consistent with the work of Setegne et al. (2011), who used projections from
24 multiple GCMs as input for a SWAT model developed for the region and found that
25 streamflow decreased in the majority of emissions scenarios and models, even when
26 precipitation increased. Unfortunately, this suggests that any hopes for a “windfall” of
27 additional water to support agriculture and hydropower in the region under climate change
28 may be unfounded.

29 Repeating the climate change sensitivity experiment with multiple models fit to the
30 Gumara watershed indicated that the MARS, GAM, and linear models all result in the largest
31 increase in uncertainty at high temperatures. This indicates that when models are fit to slightly
32 different bootstrap resamples of the historic dataset, the projected changes in streamflow at

1 high temperature changes can be highly erratic. This is likely due to the fact that extrapolating
2 the relationships that are observed between historic temperature and streamflow to higher
3 temperatures can lead to very large changes in streamflow. Fitting the models to bootstrap
4 resamples of the data results in minor changes to these relationships that can result in widely
5 varying projections when the models are used to predict streamflow at higher temperatures,
6 particularly when these relationships are nonlinear (as in the GAM). At the other end of the
7 spectrum, the random forest model exhibits almost no increase in uncertainty at high
8 temperatures, meaning that projections of streamflow at high temperatures are consistent
9 across the bootstrap resamples. This is likely the result of the random forest model structure.
10 The predicted value for each of a regression tree's terminal nodes is the average of all
11 observations that meet the conditions described for that node. Thus, the model will not predict
12 values beyond those experienced historically, even if covariate values exceed those contained
13 within the historic dataset. Thus, this model is likely to underestimate the change in
14 streamflow that results from increasing temperatures.

15 **5 Conclusions**

16 In this work, we compared multiple methods for data-driven rainfall-runoff modeling
17 in their ability to simulate streamflow in five highly-seasonal watersheds in the Ethiopian
18 highlands. Despite the popularity of ANNs in research on streamflow prediction to date,
19 ANNs were not found to be the most accurate model in any of the five basins evaluated. Other
20 methods, in particular GAMs and random forests, are able to capture non-linear relationships
21 effectively and lend themselves to simpler visualization of model structure and covariate
22 influence, making it easier to gain insights on physical watershed functions and confirm that
23 the model is operating in a physically realistic manner. However, it is important to carefully
24 evaluate model structure and residuals, as these can contribute to biased estimates of water
25 availability and uncertainty in estimating sensitivity to potential future changes in climate. In
26 particular, autocorrelation in model residuals can result in underestimation of aggregate
27 metrics such as annual flow volumes, even in models with high NSE performance.
28 Uncertainty in GAM projections was found to rapidly increase at high temperatures, whereas
29 random forest projections may be underestimating the impact of high temperatures on river
30 flows. Thorough consideration of this uncertainty and bias is important any time that models
31 are used for water planning and management, but especially crucial when using such models
32 to generate insights about future streamflow levels. By considering multiple model

1 formulations and carefully assessing their predictive accuracy, error structure and
2 uncertainties, these methods can provide an empirical assessment of watershed behavior and
3 generate useful insights for water management and planning. This makes them a valuable
4 complement to physical models, particularly in data-scarce regions with little data available
5 for model parameterization, and warrants additional research into their development and
6 application.

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21

1 Table 1. Study basin characteristics over the evaluation period of 1961 to 2004.

| Basin | Drainage area above gauge (km ²) | Average annual streamflow at gauge (MCM) | Standard deviation of annual streamflow (MCM) | Coefficient of variation of annual streamflow | Average temp (°C) | Average monthly rainfall [mm] | |
|--------------|--|--|---|---|-------------------|-------------------------------|---------|
| | | | | | | May-Oct | Nov-Apr |
| Gilgel Abbay | 2664 | 1883 | 217 | 0.12 | 15.7 | 206 | 39.3 |
| Gumara | 385 | 236 | 71 | 0.30 | 17.7 | 186 | 29 |
| Koga | 200 | 114 | 31 | 0.27 | 15.7 | 206 | 39.3 |
| Megech | 424 | 172 | 66 | 0.31 | 20.6 | 234 | 41.4 |
| Ribb | 677 | 210 | 83 | 0.36 | 18.2 | 263 | 45.8 |

2

1 Table 2. Model parameters evaluated through cross validation.

| Model type | R package | Parameters defined in model formulation | Parameters selected through cross validation |
|------------|--------------|--|--|
| GLM | stats | family = Gaussian | NA |
| GAM | mgcv | family = Gaussian method = generalized cross validation variable selection = true basis dimension $k = 3$ epsilon = 10^{-7} maxit = 200 | |
| MARS | earth | nk = 21 thresh = 0.001 fast.k = 20 pmethod = backward | degree = {1, 2, 3} nprune = {5, 10, 15, 20, 25} |
| ANN | nnet | weights = 1 rang = 0.7 maxit = 100 maxNWts = 1000 abstol = 10^{-4} reltol = 10^{-8} | size = {1, 2, 4, 8, 20} decay = {0.0, 0.1, 0.5, 1.0, 2.0} |
| RF | randomForest | ntree = 500 sampsize = 528 nodesize = 5 nPerm = 1 | mtry = {2, 3, 4, 5, 6, 7} |
| M5 | Cubist | rules = 100 extrapolation = 100 sample = 0 | committees = {10, 50, 100} neighbors = {0, 5, 9} |

2

1

2 Table 3. Cross validation errors for each assessed model.

| Standard Formulation | | GLM | GAM | MARS | RF | M5 | ANN | Climatology | Noah LSM |
|----------------------|--------------|-------|-------|-------|-------|-------|-------|-------------|----------|
| MAE | Gilgel Abbay | 30.78 | 18.54 | 16.75 | 14.89 | 15.11 | 17.22 | 10.42 | 28.11 |
| | Gumara | 4.29 | 3.41 | 3.28 | 2.67 | 2.96 | 3.15 | 2.57 | 3.95 |
| | Koga | 1.50 | 1.30 | 1.38 | 1.20 | 1.17 | 1.23 | 1.06 | 1.97 |
| | Megech | 4.45 | 2.64 | 2.83 | 2.37 | 2.53 | 3.04 | 2.54 | 4.09 |
| | Ribb | 4.69 | 2.98 | 3.50 | 2.97 | 3.27 | 3.17 | 2.81 | 7.01 |
| NSE | Gilgel Abbay | -0.02 | 0.81 | 0.83 | 0.87 | 0.86 | 0.84 | 0.95 | 0.59 |
| | Gumara | 0.04 | 0.51 | 0.61 | 0.80 | 0.66 | 0.70 | 0.81 | 0.48 |
| | Koga | 0.45 | 0.71 | 0.65 | 0.76 | 0.77 | 0.76 | 0.83 | 0.25 |
| | Megech | -1.85 | 0.63 | 0.46 | 0.73 | 0.65 | 0.52 | 0.71 | 0.41 |
| | Ribb | -1.14 | 0.71 | 0.39 | 0.71 | 0.31 | 0.67 | 0.73 | -0.75 |
| Anomaly Formulation | | GLM | GAM | MARS | RF | M5 | ANN | Climatology | Noah LSM |
| MAE | Gilgel Abbay | 9.73 | 9.82 | 10.10 | 10.12 | 9.94 | 9.79 | 10.42 | 28.11 |
| | Gumara | 2.22 | 2.25 | 2.43 | 2.23 | 2.16 | 2.22 | 2.57 | 3.95 |
| | Koga | 1.03 | 1.06 | 1.08 | 1.09 | 1.05 | 1.05 | 1.06 | 1.97 |
| | Megech | 2.49 | 2.48 | 2.63 | 2.66 | 2.69 | 2.50 | 2.54 | 4.09 |
| | Ribb | 2.79 | 2.76 | 2.84 | 2.70 | 2.78 | 2.77 | 2.81 | 7.01 |
| NSE | Gilgel Abbay | 0.95 | 0.95 | 0.95 | 0.95 | 0.95 | 0.95 | 0.95 | 0.59 |
| | Gumara | 0.85 | 0.85 | 0.82 | 0.85 | 0.86 | 0.86 | 0.81 | 0.48 |
| | Koga | 0.83 | 0.82 | 0.81 | 0.81 | 0.82 | 0.82 | 0.83 | 0.25 |
| | Megech | 0.73 | 0.72 | 0.65 | 0.66 | 0.61 | 0.72 | 0.71 | 0.41 |
| | Ribb | 0.73 | 0.75 | 0.72 | 0.75 | 0.73 | 0.74 | 0.73 | -0.75 |

3

1

2 Table 4. Residual autocorrelation factors at a 12-month lag for the standard formulation and
3 anomaly formulation models, and resulting mean annual observed and predicted flow.

4

| | Autocorrelation Factors | | Mean Annual Flow (MCM) | | |
|--------|-------------------------|---------|------------------------|----------|---------|
| | Standard | Anomaly | Observed | Standard | Anomaly |
| Gilgel | 0.33 | 0.11 | 22,925 | 20,703 | 22,958 |
| Gumara | 0.29 | 0.07 | 2,870 | 2,392 | 2,734 |
| Koga | 0.04 | 0.10 | 1,383 | 1,333 | 1,386 |
| Megech | 0.05 | 0.04 | 2,035 | 1,637 | 2,028 |
| Ribb | 0.21 | -0.01 | 2,575 | 1,969 | 2,615 |

5

1 Table 5. Covariate importance measurements from each basin's model

| Model type | Linear model | | | | Generalized additive model | | M5 model tree | Random forest | |
|----------------------|--|---------|----------------------|---------|--|---------|--|---|-------------------------|
| Measure of influence | Linear regression coefficients and associated p-values | | | | Estimated degrees of freedom (EDF) and associated p-values | | Covariate usage in tree rules and model coefficients | Increase in MSE when covariate is randomly permuted | |
| Basin | Gilgel Abbay | | Koga | | Megech | | Gumara | | Ribb |
| Covariate | Coefficient estimate | P-value | Coefficient estimate | P-value | EDF | P-value | Tree rules | Model coefficients | Percent increase in MSE |
| Prec | 0.22 | < 0.01 | 0.24 | < 0.01 | 1.346 | < 0.01 | 5% | 58% | 7.71% |
| Prec (lag 1) | 0.10 | 0.03 | 0.16 | < 0.01 | 0.624 | 0.08 | 0% | 19% | 2.79% |
| Prec (lag 2) | 0.01 | 0.74 | 0.05 | 0.26 | 0 | 0.29 | 0% | 0% | 1.10% |
| Temp | -0.09 | 0.08 | -0.07 | 0.17 | 1.023 | 0.07 | 0% | 47% | 12.74% |
| Temp (lag 1) | -0.04 | 0.49 | -0.06 | 0.22 | 0 | 0.32 | 0% | 46% | 4.97% |
| Temp (lag 2) | -0.01 | 0.81 | -0.09 | 0.08 | 0 | 0.56 | 0% | 0% | 8.16% |
| Agr. LC | 0.00 | 0.33 | 0.02 | 0.01 | 1.986 | < 0.01 | 86% | 73% | 15.21% |

2

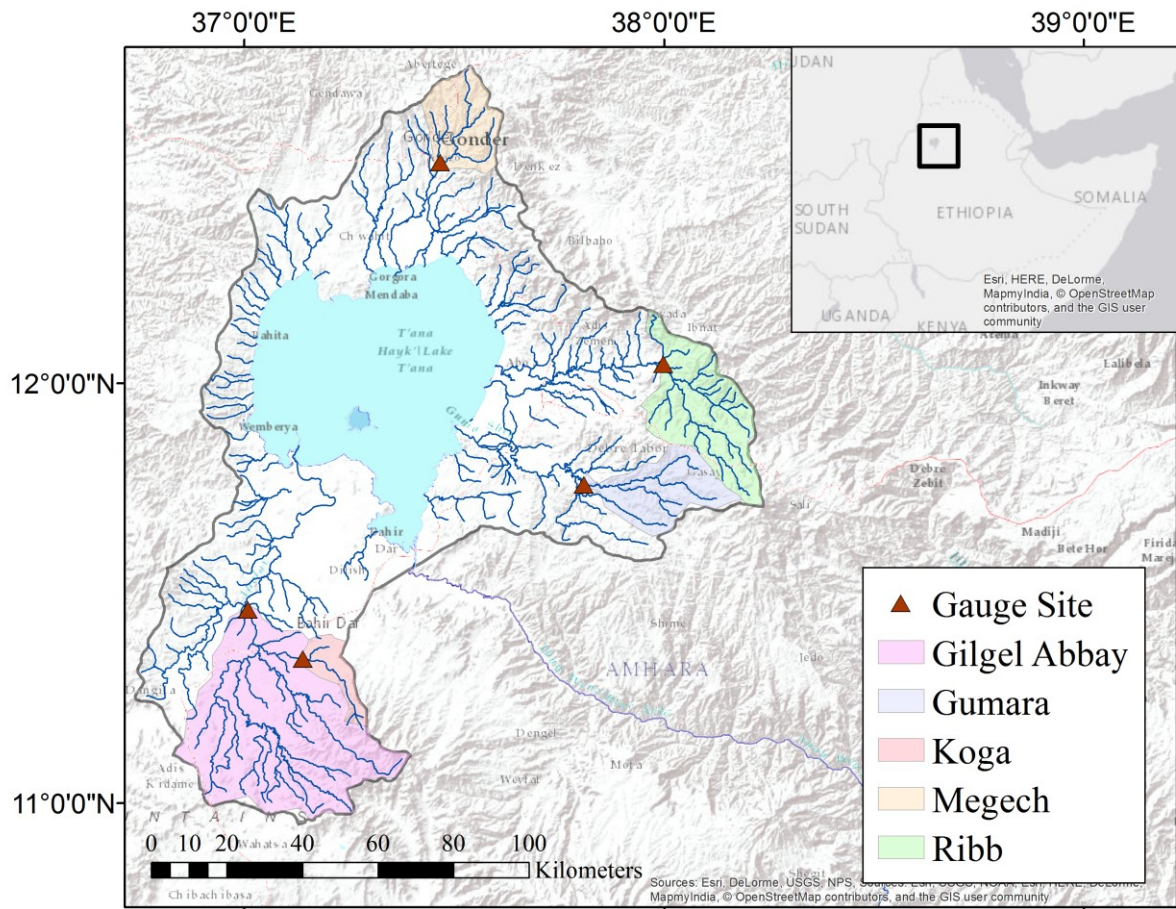
1 Table 6. Mean and standard deviation values for temperature, wet-season rainfall, and dry-
 2 season rainfall in each basin.

3

| | Temperature (°C) | | Wet season rainfall (mm/month) | | Dry season rainfall (mm/month) | |
|--------------|---------------------|------|--------------------------------------|-----|--------------------------------------|------|
| | Mean | SD | Mean | SD | Mean | SD |
| Gilgel Abbay | 15.7 | 1.54 | 206 | 145 | 39.3 | 56.5 |
| Gumara | 17.7 | 1.55 | 186 | 137 | 29.0 | 43.6 |
| Koga | 15.7 | 1.54 | 206 | 145 | 39.3 | 56.5 |
| Megech | 20.6 | 1.75 | 234 | 118 | 41.4 | 60.9 |
| Ribb | 18.2 | 1.61 | 263 | 115 | 45.8 | 57.0 |

4

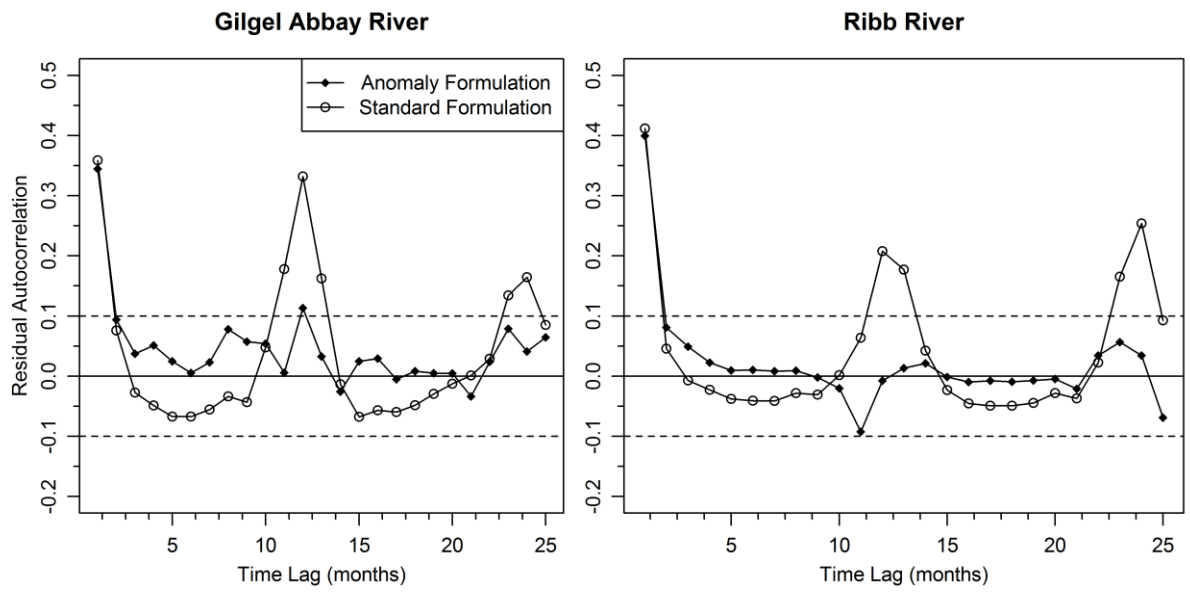
1 Figure 1. Map of Lake Tana and surrounding rivers



2

3

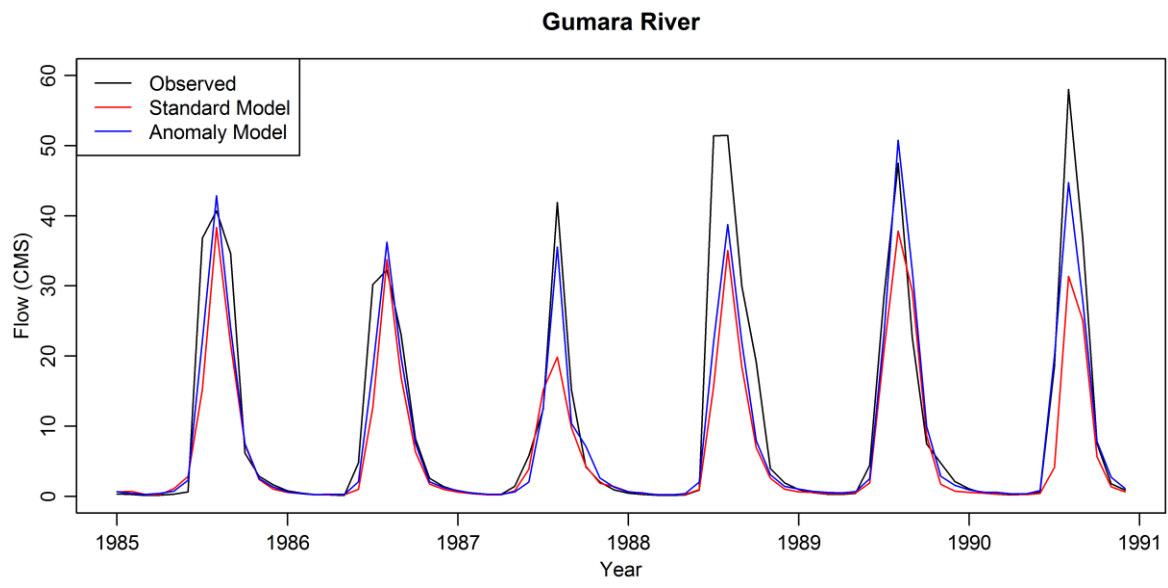
1 Figure 2. Autocorrelation in model residuals for the Gilgel Abbay and Ribb Rivers



2

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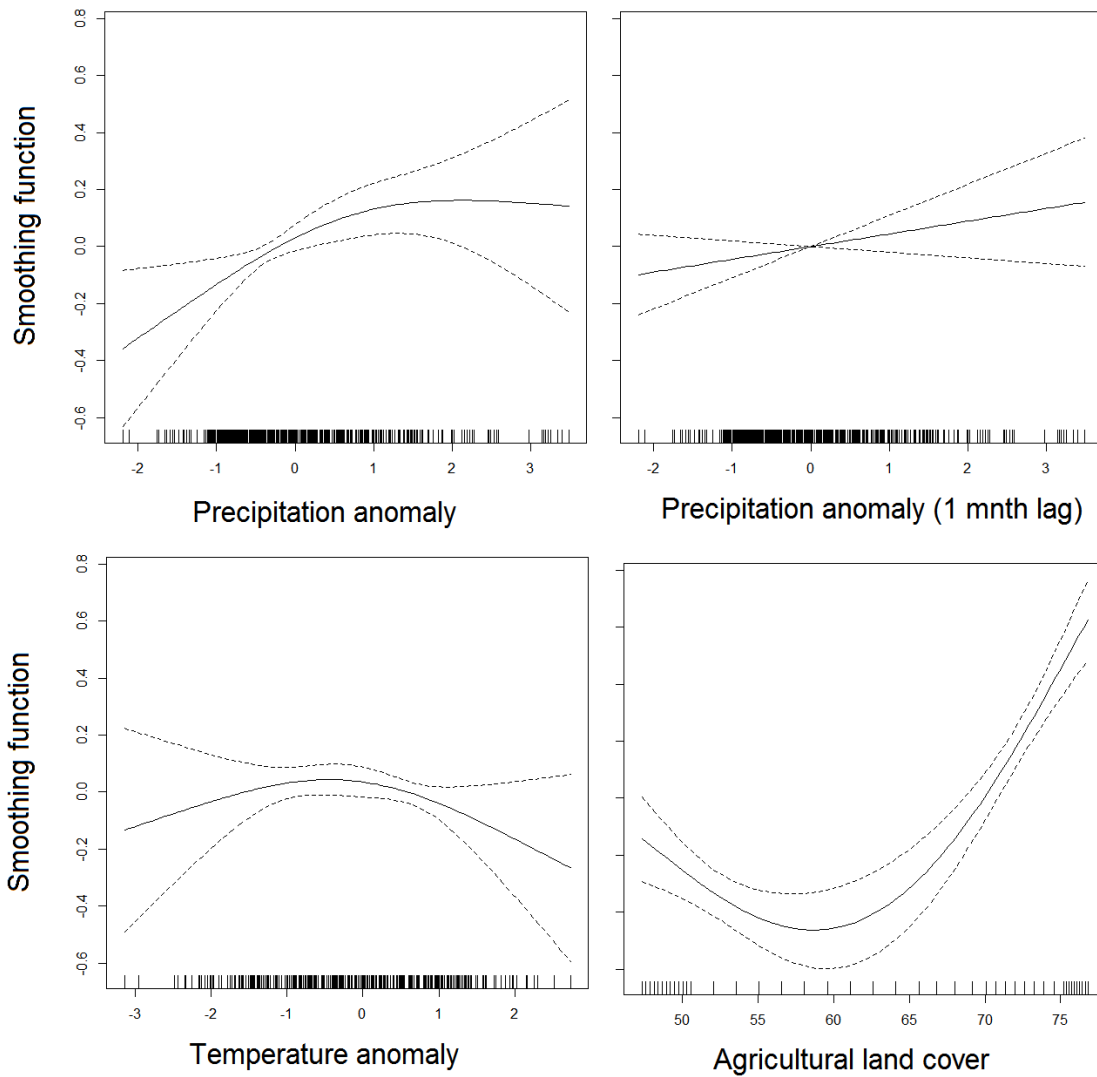
1 Figure 3. Example observed and predicted flows for Gilgel Abbay River from 1995 to 2000.



2

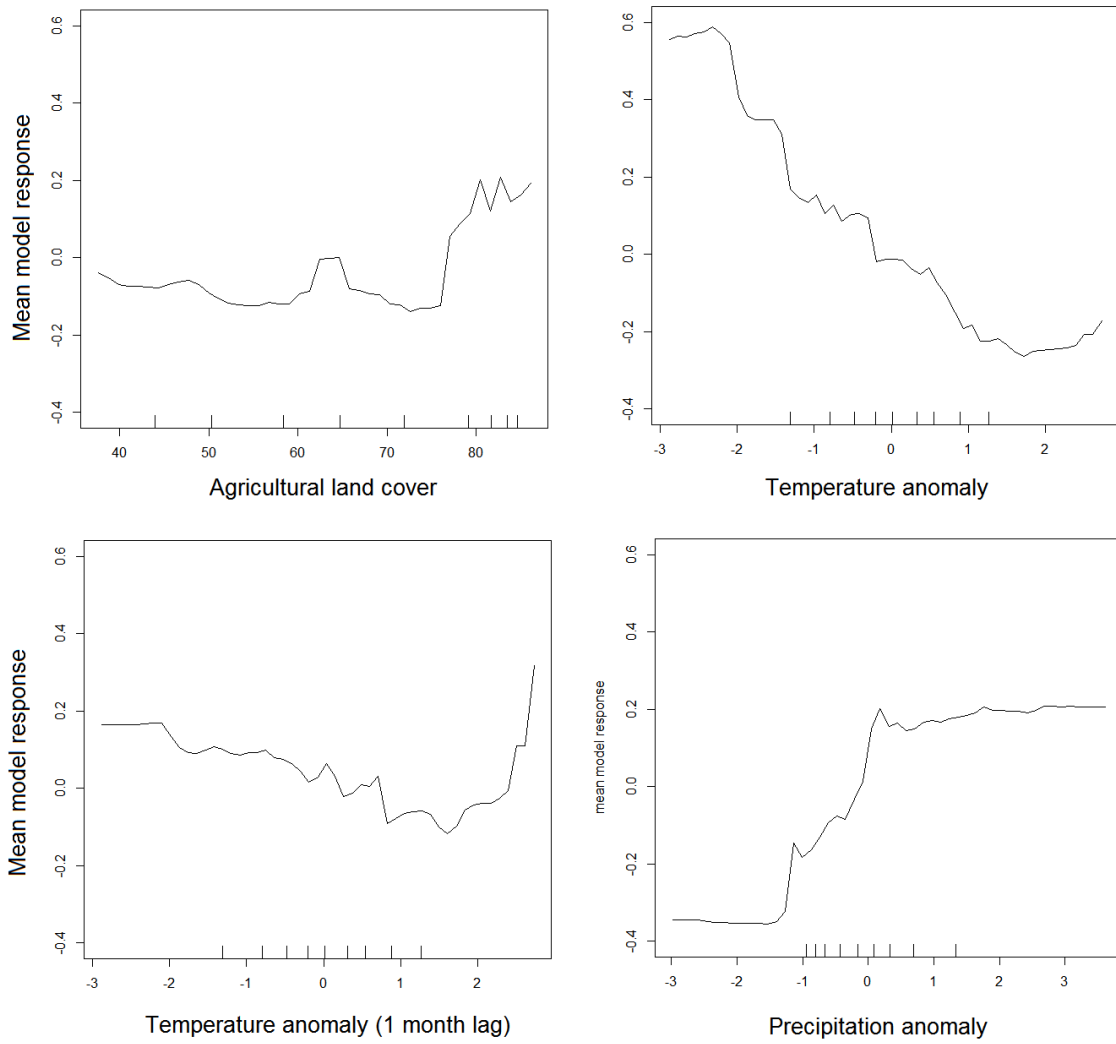
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1 Figure 4. Plots of the smoothing functions used in the Megech River GAM. Hash marks along
2 the x-axis indicate observation values of each covariate.



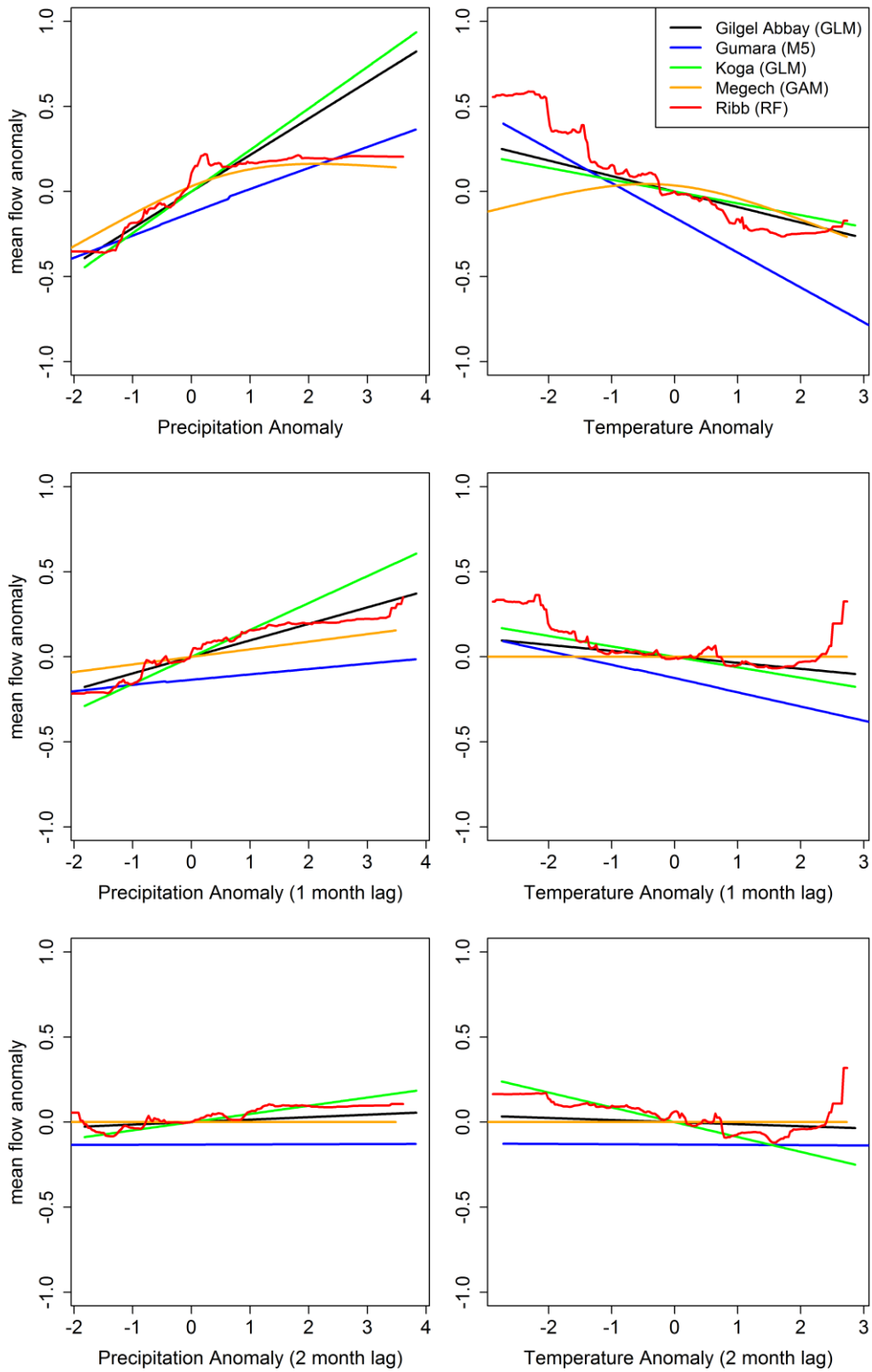
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1 Figure 5. Partial dependence plots for the Ribb River random forest model. Hash marks along
2 the x-axis show covariate sample decile values.



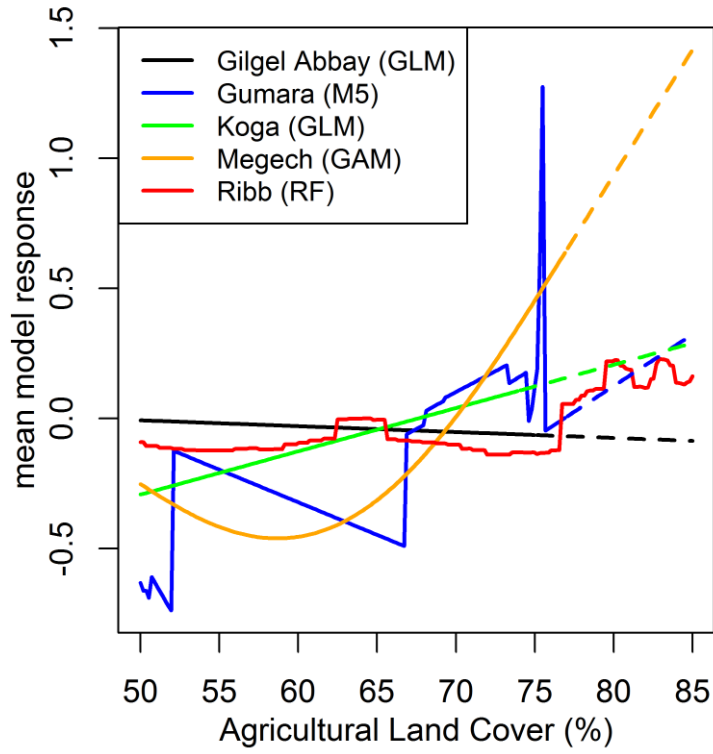
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1 Figure 6. Partial dependence plots for climate covariates in the highest performing model in
2 each basin. Model type is indicated in parentheses.



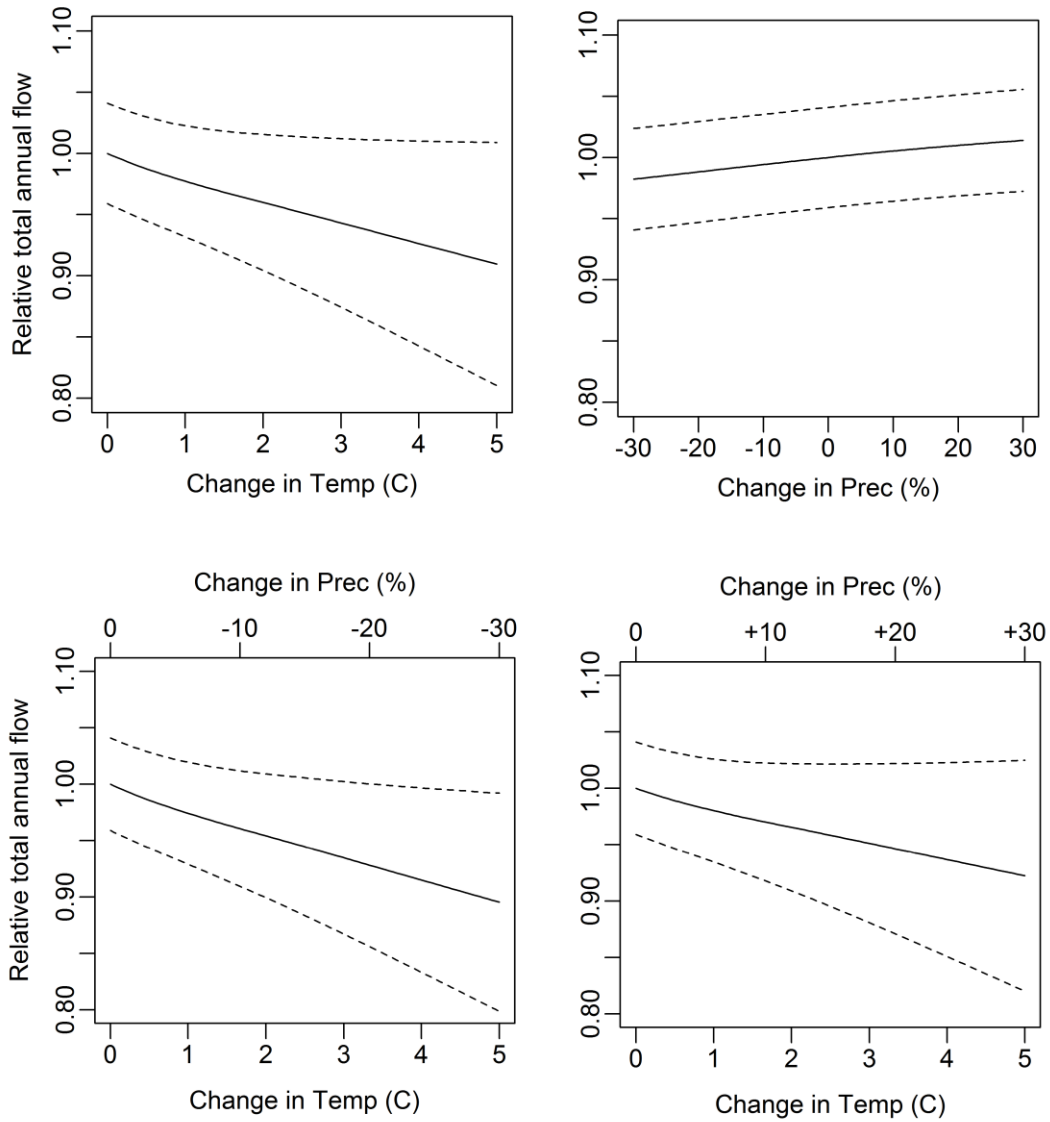
3

1 Figure 7. Partial dependence plot for agricultural land cover in the highest performing model
2 in each basin. Model type is listed in parentheses for each basin. Dashed lines
3 indicate values that exceed historic levels of agricultural land cover experienced in
4 that basin.



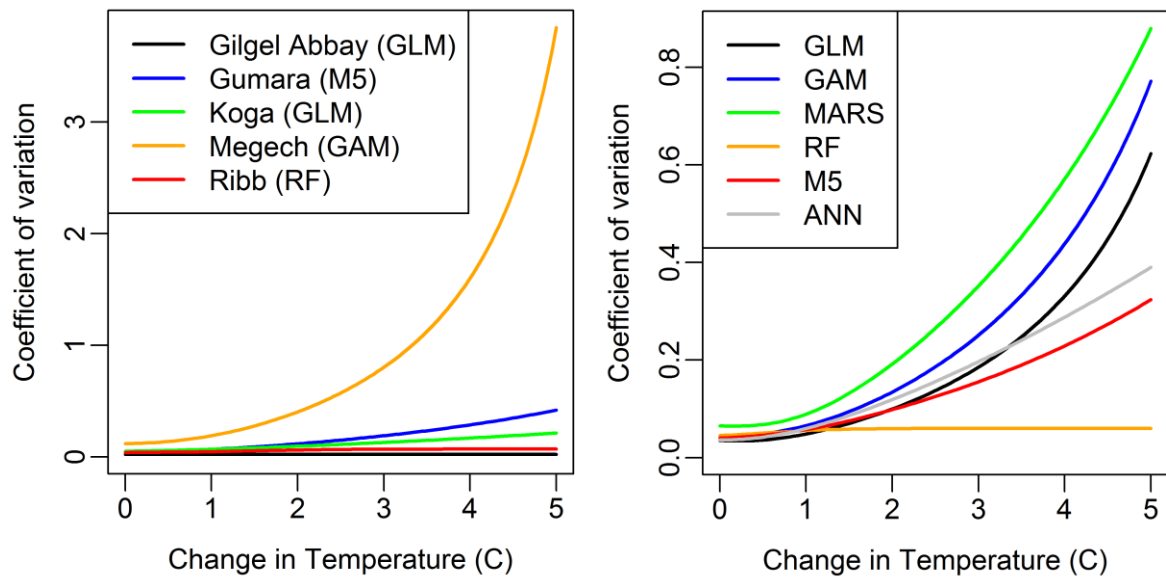
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6

1 Figure 8. Projected changes in total streamflow (relative to current long-term average) under
 2 changing climate conditions. The top two panels show the sensitivity to changes in
 3 temperature and precipitation when they are varied independently. The bottom panel
 4 shows sensitivity to changing temperature in conjunction with decreasing (left
 5 panel) and increasing (right panel) precipitation. Dashed lines represent 95%
 6 confidence bounds from bootstrap resampling.



7
8

1 Figure 9. Changes in the coefficient of variation across bootstrap resamples from the highest
2 performing model in each basin (left panel) and multiple models all applied to the
3 Gumara basin (right panel).



4