# The Impact of Uncertain Precipitation Data on Insurance Loss Estimates Using a Flood Catastrophe Model

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#### 12 Abstract

Catastrophe risk models used by the insurance industry are likely subject to significant 13 14 uncertainty, but due to their proprietary nature and strict licensing conditions they are not available for experimentation. In addition, even if such experiments were conducted, these 15 would not be repeatable by other researchers because commercial confidentiality issues 16 prevent the details of proprietary catastrophe model structures from being described in public 17 domain documents. However, such experimentation is urgently required to improve decision 18 making in both insurance and re-insurance markets. In this paper we therefore construct our 19 own catastrophe risk model for flooding in Dublin, Ireland in order to assess the impact of 20 typical precipitation data uncertainty on loss predictions. As we consider only a city region 21 rather than a whole territory and have access to detailed data and computing resources 22 typically unavailable to industry modellers, our model is significantly more detailed than 23 24 most commercial products. The model consists of four components, a stochastic rainfall 25 module, a hydrological and hydraulic flood hazard module, a vulnerability module and a financial loss module. Using these we undertake a series of simulations to test the impact of 26 27 driving the stochastic event generator with four different rainfall data sets: ground gauge data, gauge corrected rainfall radar, meteorological re-analysis data (ERA-Interim) and a satellite 28 rainfall product (CMORPH). Catastrophe models are unusual because they use the upper 29 three components of the modelling chain to generate a large synthetic database of unobserved 30 31 and severe loss-driving events for which estimated losses are calculated. We find the loss estimates to be more sensitive to uncertainties propagated from the driving precipitation 32 datasets than to other uncertainties in the hazard and vulnerability modules, suggesting that 33 the range of uncertainty within catastrophe model structures may be greater than commonly 34 35 believed.

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#### 37 **1.0 Introduction and Literature Review**

38 The repeated occurrence of high profile flood events across the British Isles, such as Carlisle in January 2005, Gloucestershire in July 2007 and Dublin in October 2011, has resulted in 39 sustained public, commercial, political and scientific interest in flood risk. 40 Recent catastrophic flood events in other countries, such as the Indus floods in Pakistan (2010), the 41 42 Australian and Thai floods (2011), and the Central European Floods (2013) have further 43 raised the profile of flood risk through extensive global news coverage. The economic cost 44 associated with flooding is often high. It is estimated that the October and November 2000 floods in the UK caused insured losses of £1.3 billion (Pall et al., 2011), whilst household 45 losses resulting from the summer 2007 floods reached £2.5 billion, with business losses 46 47 accounting for a further £1 billion (Chatterton et al., 2010; Pitt, 2008). The reinsurance firm Munich Re estimates that total economic losses from the Australian and Thailand events were 48 USD 2.8 billion and USD 40 billion respectively (Munich Re, 2012), whilst the reinsurance 49 firm Swiss Re estimates these figures at USD 6.1 billion and USD 30 billion (Swiss Re, 50 51 2012). Much of the total insured loss was from business interruption and contingent business interruption claims, demonstrating the global impact of such events. 52

Due to the scale of potential losses the insurance and reinsurance industries require accurate 53 flood risk estimates, and the current accepted approach is to use calculation chains 54 comprising linked stochastic and physically-based models. These calculation chains, known 55 as catastrophe or 'CAT' models, are at the core of a methodological framework employed by 56 57 the insurance industry to produce probabilistic estimates of natural catastrophe risk. First developed in the late 1980s to model earthquake risk, the methodology was widely adopted 58 throughout the 1990s to model a range of hazards such as tropical cyclone windstorms and 59 storm-surge floods (Wood et al., 2005). Today, such models are relied upon by the 60 61 insurance and risk management industries to guide a wide range of financial decisions (Grossi et al., 2005). Whilst being applicable to a wide range of hazards, commercial 62 63 'vendor' CAT models typically share a common structure that can be broken down into four component parts: 64

- i. Stochastic module. The stochastic module is used to generate a database of plausible 65 event driving conditions. In the case of flooding, this could be a database of extreme 66 precipitation events over the catchment(s) that drive fluvial or pluvial risk where the 67 insured assets are located. The stochastic module is typically trained on historically 68 observed data. As observational records of natural hazards are typically short  $(10^{1})$ 69 years) relative to return periods of interest to the insurance industry ( $10^2$  to  $10^4$  years), 70 the module must be capable of simulating events whose magnitude exceeds that of the 71 72 largest observed event.
- ii. Hazard module. The hazard module is used to simulate a selection of events from the
  database generated by the stochastic module. The hazard module needs to produce an
  estimate of damage-driving characteristics across the area where insured assets are
  located. In the case of flooding this is likely to take the form of a map of water
  depths.
- iii. Vulnerability module. The vulnerability module calculates the expected damage to assets as a result of the event modelled by the hazard module. These damages are

80 expressed as a damage ratio that varies between 0 (no damage) and 1 (total loss). Factors influencing the susceptibility of an asset to damage may include terms such as 81 building age, occupancy type, construction materials, or height. These parameters are 82 typically uncertain, and thus vulnerability may be represented by an uncertain 83 84 measure that maps the expected damage to a particular asset against a continuously 85 variable hazard module output such as water depth and/or velocities. This is often done using a beta distribution with non-zero probabilities for damage ratios of 0 and 86 87 1.

iv. Financial module. The financial module transforms the per event damage estimates
produced by the vulnerability module into an estimate of insured loss. Estimates of
insured losses are generated by aggregating the losses from all assets being considered
and applying policy conditions such as limits and deductibles to the total estimate of
loss. The financial module resamples the database of simulated events to produce a
large number of different time series realisations from which time-aggregated loss
curves are produced.

As with any study that involves the modelling of environmental processes, it is important to 95 address the presence of uncertainty within the system. Previous studies that consider flood 96 risk using a model cascade framework have found the 'driving' component at the top of the 97 cascade to be the most significant source of uncertainty (Kay et al., 2008; McMillan and 98 99 Brasington, 2008). Cloke et al. (2012) also highlight the problem of uncertainty propagating from global and regional climate models when attempting to assess flood hazard on the River 100 Severn in the UK. Due to their focus on low frequency, high magnitude events, the 101 stochastic component of a CAT model inevitably has to extrapolate to event scales beyond 102 those in the observational record. As a result, the loss estimates produced by CAT models 103 may be particularly sensitive to the propagation of uncertainty in the data used to drive the 104 105 stochastic component. If true, this will indicate that CAT model cascades are even more sensitive to driving uncertainties than other previously studied hydrological model cascades. 106 107 As the stochastic module forms the driving component of a CAT model, this study attempts to assess the uncertainties derived from the choice of data used to calibrate, and therefore 108 govern, the behaviour of the stochastic module. In order to provide context for this analysis, 109 further limited analysis of the effect of parametric uncertainty within the hazard module and 110 111 uncertainty within the vulnerability model were performed.

When developing a CAT model, it is important to bear in mind that the recent Solvency II 112 legislation in Europe (European Parliament and European Council, 2009) requires that model 113 114 users are able to understand and communicate how their models function. Many users will not be specialists in the field of environmental sciences and thus such legislation favours 115 simpler model structures. A further reason to favour simpler model structures lies in their 116 ease of application. Simpler models typically require less data than complex models, and 117 therefore should be easier to apply to the wide array of locations that are of interest to 118 insurance markets. It is also important to minimise the computational requirements of the 119 cascade due to the extremely large number of events that may need to be modelled in order to 120 estimate losses at very high return periods. The model structure used for this study was 121

- developed with such operational concerns in mind, and as such simple methods capable ofdelivering adequate performance against historical observations were favoured.
- 124 The following section of the literature review briefly explains the choice of model 125 components employed in this study. The methodology that follows explains in more detail 126 how each component functions within a CAT model framework.

## 127 **1.1 Stochastic Module**

Stochastic rainfall models are data-based approaches that use statistical information extracted 128 from observations to parameterise a mechanism used to generate synthetic rainfall records. 129 Such approaches are attractive in this context due to their relative simplicity and low 130 Stochastic rainfall models can generally be split into two computational costs. 131 methodological groups, namely profile-based and pulse-based, although there have been 132 attempts to test alternative approaches including chaotic (Rodriguez-Iturbe et al., 1989; 133 Sivakumar et al., 2001), artificial neural networks (Burian and Durran, 2002), simulated 134 annealing (Bárdossy, 1998) and multiplicative cascade disaggregation (Gaume et al., 2007). 135 Profile-based models typically use statistical distributions to characterise storms in terms of 136 intensity, duration and inter-arrival time, whereas pulse-based models use statistical 137 distributions to define raincells occurring within larger storm units characterised by duration 138 and inter-arrival time distributions. The raincells take the form of pulses with individual 139 durations and intensities, and the total storm intensity at a given time can therefore be 140 calculated through summation of all active cell intensities at that time. 141

For the purposes of building a flood catastrophe model, it is necessary to select a model 142 formulation that is able to reproduce the extreme events that drive flood risk. Several 143 comparison studies have noted that while pulse-based models are able to simulate storm 144 inter-arrival times and precipitation averages well, their ability to capture extreme statistics is 145 variable and often particularly poor over short timescales (Cameron et al., 2000; Khaliq and 146 Cunnane, 1996; Onof and Wheater, 1993; Verhoest et al., 1997). By comparison, the profile-147 based models have shown skill at simulating extreme events (Acreman, 1990; Blazkov and 148 149 Beven, 1997; Cameron et al., 2000), although their ability to perform well for such events is dependent on the length and quality of the historical record used for their calibration. Due to 150 its demonstrated ability to represent a range of different extreme precipitation events, this 151 152 study employs a model developed from the profile-based Cumulative Distribution Function Generalised Pareto Distribution Model (CDFGPDM) of Cameron et al. (1999). 153

# 154 **1.2 Hazard Module**

In order to convert the rainfall input from the stochastic module into an estimate of water depths across the spatial domain containing the insured assets, two components are required: a hydrological rainfall-runoff model to produce an estimate of river discharge and a hydraulic model to transform the estimate of river discharge into a map of water depths. Hydrological models vary in complexity from process-rich, spatially distributed models such as the Systeme Hydrologique Europeen (Abbott et al., 1986a, 1986b) and the US Department of Agriculture's Soil and Water Assessment Tool (Muleta and Nicklow, 2005), to simple,

spatially lumped conceptual models such as TOPMODEL (Beven and Kirkby, 1979) or 162 HBV (Bergstrom and Forsman, 1973). Increasing model complexity inevitably entails 163 increased dimensionality and data requirements, a situation that is often at odds with the 164 requirements of a CAT model. Furthermore, the fundamental argument as to how much 165 complexity is valuable in a model has not yet been conclusively answered in the literature 166 (Bai et al., 2009; Beven, 1989; Blöschl and Sivapalan, 1995), and a number of studies have 167 found that model performance does not necessarily improve with increased model complexity 168 (e.g. Butts et al., 2004; Reed et al., 2004). As a result, a simple variant of the HBV model 169 (Bergstrom and Forsman, 1973; Bergström and Singh, 1995; Seibert and Vis, 2012) was 170 chosen here thanks to its ease of application, low data and computation cost and 171 demonstrated performance across a large number of studies (Cloke et al., 2012; Deckers et 172 al., 2010; e.g. Seibert, 1999). 173

In order to translate estimates of river discharge into maps of water depth across a domain, an 174 additional hydraulic modelling component is required. The flow of water in urban areas is 175 176 inherently multi-dimensional and requires a model of commensurate dimensionality able to run at the fine spatial resolutions needed to represent urban environments where vulnerability 177 to losses will be most critical. The computational expense of such simulations has resulted in 178 a research drive to develop efficient methods of modelling high resolution two-dimensional 179 shallow water flows. Hunter et al. (2008) benchmarked a suite of commercial and research 180 2D codes on a small urban test scenario and found all to give plausible results, with predicted 181 water depths typically differing by less than the vertical error in the topographical error 182 despite the model governing equations varying from full 2D shallow-water equations to x-y 183 decoupled analytical approximations to the 2D diffusion wave. These results are supported 184 by further recent studies that have found highly efficient simplifications of the 2D shallow 185 water equations to be appropriate for a number of urban inundation modelling (Neal et al., 186 187 2011; Néelz and Pender, 2010). As a result, this study employs the latest inertial formulation of the highly efficient 2D storage cell inundation model LISFLOOD-FP (Bates et al., 2010). 188 This approach offers a more sophisticated representation of flow dynamics than the methods 189 adopted by most vendor CAT models; vendor models typically represent the channel and 190 floodplain using a 1D model, with a limited number of models also offering 2D modelling of 191 'off-floodplain' processes (AIR Worldwide, 2013; RMS, 2006). 192

## 193 **1.3 Vulnerability Module**

Flood damage models typically use water depths to predict damage based on a depth-damage 194 function derived from empirical data (Black et al., 2006; Merz and Thieken, 2009; Merz et 195 al., 2004), synthetic data (Penning-Rowsell et al., 2005), or a combination of both (ICPR, 196 2001). Studies have demonstrated significant variation in the curves produced by each 197 methodology (Merz and Thieken, 2009; Merz et al., 2010), with the greater accuracy of 198 empirical data compared to synthetic data (Gissing and Blong, 2004) being countered by the 199 limited transferability of empirical data between sites (Smith, 1994). 200 Depth damage 201 functions are inherently uncertain due to the large number of factors that may influence the level of damage that results from a water depth. These include, but are not limited to, 202 203 building type, building construction method, building age, building condition and

204 precautionary measures). Although there is ongoing research into the possibility of accounting for these factors explicitly within multivariate depth-damage functions (Kreibich 205 et al., 2010; Merz et al., 2013), such methods have not been widely adopted within the 206 insurance market as a lack of observed damage data in most regions prevents calibration of 207 such complex functions. Many commercial models instead attempt to represent much of the 208 209 total CAT model uncertainty within the vulnerability module by sampling around the depthdamage curve. This is typically done using beta distributions to represent the probabilities of 210 experiencing a range of damage ratios of between 0 and 1 for a given water depth. As the 211 focus of this study is on the uncertainty due to driving precipitation data, we employ fixed 212 depth-damage curves for most of our experiments. However, as recent studies (Jongman et 213 al., 2012; Moel and Aerts, 2010) have suggested that the vulnerability module may be the 214 dominant source of uncertainty, we also undertake a limited analysis using uncertain 215 vulnerability curves in section 3.4 in order to provide an indication of relative contributions 216 217 to modelled uncertainty. The curves and distribution parameters were supplied by Willis Global Analytics and were derived from a combination of synthetic and empirical data, 218 claims data, and industry expertise. 219

## 220 **1.4 Financial Module**

Due to their proprietary nature, public domain literature describing the financial component 221 of CAT models is very limited. Generally the role of financial modules is to transform 222 damage estimates from the vulnerability module into estimates of insured ground up loss (i.e. 223 loss before application of deductibles and/or reinsurance) before aggregating the location-224 specific losses to produce portfolio-wide loss estimates for a given event. These can then be 225 transformed into estimates of gross insured loss by applying policy conditions such as 226 227 deductibles, coverage limits, triggers, reinsurance terms, etc. (Grossi et al., 2005). Where the hazard module is computationally expensive, the financial module is often used to fit curves 228 to the loss distributions generated by calculation chain, allowing much larger synthetic 229 databases of event losses to be generated by subsequent resampling of the distributions. The 230 primary output of a financial model takes the form of a curve that describes the probability of 231 exceeding a certain level of loss within a fixed time period (typically annual). The two most 232 common exceedence probability (EP) curves are the annual occurrence exceedence 233 probability (OEP), representing the probability of a single event loss exceeding a certain level 234 in a given year, and the aggregate exceedence probability (AEP), representing the probability 235 of aggregate losses exceeding a certain level in a given year. Details of the financial module 236 237 employed in this study are shown in section 2.2.4.

#### 238 2.0 Study Site, Data and Methodology

Dublin, Ireland, was selected as the test site for this study due to its flood prone nature and the availability of suitable data sources. Historically, Dublin has been prone to fluvial, pluvial and tidal flooding, with fluvial risk being largely concentrated along two rivers, namely the River Dodder and the River Tolka. The River Dodder has its source in the Wicklow Mountains to the South of the city and drains an area of approximately 113 km<sup>2</sup>. High rainfall intensities over the peaks of the Wicklow Mountains (annual totals can reach

2000 mm) coupled with steep gradients results in the River Dodder exhibiting flashy 245 responses to storm events, with a typical time to peak of less than 24 hours. The River Tolka 246 has its source in gently sloping farmland to the North West of the city and drains an area of 247 approximately 150 km<sup>2</sup>; it exhibits a slightly less flashy response than the Dodder with a time 248 to peak of approximately 24 hours. As a result of the short catchment response times, sub-249 250 daily (ideally hourly) rainfall data are required to drive hydrological models of the rivers. Both catchments contain a mixture of urban and rural land use. Figure 1 is a map showing 251 the location of these rivers and their respective catchment boundaries upstream of their 252 gauging stations, as well as the boundary of the hydraulic model, the location of river 253 gauging stations and the location of rain gauges. The calculation chain uses hydrological 254 models of the Dodder and Tolka catchments to drive a hydraulic model of the rivers as they 255 flow through the city and out into Dublin Bay. A third major river, the River Liffey, is also 256 shown. The Liffey is not modelled in this study as its flow is controlled by three reservoirs 257 that supply a hydroelectric generator upstream; serious flooding downstream of these features 258 has not been observed since their construction was completed in 1949. River flow records 259 are available from 1986 to present on the River Dodder and 1999 to present on the River 260 261 Tolka.

## 262 [FIGURE 1 AROUND HERE]

In section 2.1, the four types of precipitation data (ground rain gauge, radar, meteorological 263 reanalysis and satellite) used to drive the model are introduced along with the methods used 264 265 to derive a catchment average precipitation series from each type of data. This step was required as using the stochastic module to generate extremely long (>500,000 years) spatial 266 rainfall fields on an hourly time step would not have been computationally feasible, nor was 267 268 it necessary given the input requirements of the simple hydrological model used here. The four types of precipitation data were chosen to represent the range of rainfall products 269 available, from the high resolution localised gauge and radar data to the coarser (but globally 270 available) reanalysis and satellite products. The record lengths of the different data sources 271 were variable, but all four were available for the period January 2002 - May 2009; for 272 experiments comparing the different data sources this was the period used. 273

In section 2.2, the components and data used to build and calibrate the stochastic, hazard,vulnerability and financial modules are presented.

#### 276 2.1.1 Rain Gauge Record

The catchments surrounding Dublin are relatively well served by a network of rain gauges operated by Dublin City Council and the Irish weather service, Met Éireann. The gauges are primarily daily, with hourly weather stations sited at Dublin airport and Casement aerodrome. However, the network is subject to the usual limitations of gauge data which include missing data and inconsistent recording periods across the network. While some of the daily rain gauges have been operating for over 100 years, others were recently installed or retired. The gauges shown in figure 1 are the ones selected for use in this study following a significant pre-processing effort to check the availability of uninterrupted records from each gauge forperiods coinciding with the available river flow records.

The daily catchment average time series were constructed by generating a gridded 286 precipitation record at 50 m resolution for each of the catchments; the relatively fine grid was 287 chosen due to the negligible computational cost of this process. The contribution of each 288 daily gauge within a catchment to a given grid cell was calculated using an inverse distance 289 290 weighting function. The difference in altitude between a given gauge and grid cell was also accounted for by correction using a precipitation-altitude gradient derived from the gauge 291 record. Once the precipitation in all cells within a catchment was calculated, the catchment 292 average precipitation was obtaining by averaging the value across all cells. The daily record 293 294 was then distributed according to the nearest hourly station (Casement Aerodrome in the Dodder; Dublin Airport in the Tolka) to produce an hourly catchment average record. 295

#### 296 2.1.2 Radar Record

297 The radar rainfall data were provided by the Met Éireann from a C-band radar located at Dublin Airport. A number of different products are produced for this radar, and the 1 km pre-298 gridded 15 minute Precipitation Accumulation (PAC) product is used in this study. The PAC 299 product estimates the rainfall intensity at 1 km above the topographical surface, and the data 300 were supplied for the period 2002 - 2009. Pre-processing was required to remove an echo 301 signal present over mountainous parts of the Dodder catchment that was expressed in the data 302 as anomalous near-continuous low intensity rainfall. An hourly timestep catchment average 303 series was generated by averaging the cells that fell within the boundaries of a catchment. 304 305 Whilst radar data are able to provide an estimate of the spatial distribution of precipitation, correction using ground-based observations is required in order for reasonable estimates of 306 rainfall intensities (Borga, 2002; Germann et al., 2006; O'Loughlin et al., 2013; Steiner et al., 307 1999). Adjustment factors were therefore used to match the radar-derived catchment rainfall 308 309 volume to the gauge-derived catchment rainfall volume on a three-monthly basis. The 310 adjustment factor values were assumed to be time invariant for the duration of each three month period (Gjertsen et al., 2004). 311

#### 312 **2.1.3 ECMWF ERA-Interim Reanalysis**

ERA-Interim is a global atmospheric reanalysis produced by the European Centre for 313 Medium-Range Weather Forecasts (ECMWF) (Dee et al., 2011). The reanalysis covers the 314 period 1979-present and produces gridded surface parameters. The ERAI configuration has a 315 spectral T255 horizontal resolution, which corresponds to approximately 79 km spacing on a 316 reduced Gaussian grid. The vertical resolution is using 60 model levels with the top of the 317 atmosphere located at 0.1 hPa. ERA Interim data have been used in a wide range of 318 applications such as mapping of drought, fire, flood and health risk (Pappenberger et al., 319 Precipitation data are available in the form of 3-hour rainfall accumulation 320 2013). totals. Three-hourly timestep catchment average precipitation time series were produced 321 using a weighted average of the ERA-Interim cells that covered the catchment, where weights 322 323 were assigned based on the fraction of the catchment covered by each cell.

## 324 2.1.4 CMORPH Satellite Precipitation

The Climate Prediction Center morphing method (CMORPH) precipitation record is produced by using motion vectors derived from half-hourly interval geostationary satellite infrared imagery to propagate passive microwave precipitation estimates (Joyce et al., 2004). Data are available from 1998 – present day at a 3 hourly timestep on a 0.25 degree spatial grid. Three-hourly timestep catchment average precipitation time series were produced in the same way as with the ERA-Interim reanalysis data.

## 331 2.2.0 Catastrophe Model Framework

The CAT model framework employed in this study replicates the logic used by proprietary 332 commercial models but uses detailed and transparent components that allow us to experiment 333 in a controlled and repeatable fashion. The stochastic event generator creates a long time 334 series of rainfall events that are used to drive the hazard module. When a flood event occurs, 335 the predicted water depths are input into the vulnerability module to produce an estimate of 336 337 loss. The event ID and loss ratio (event loss expressed as a percentage of the total sum insured across the portfolio) are recorded in an event loss table. The number of events 338 occurring in each year is also recorded. Finally, the financial module resamples the event 339 loss table in order to produce an aggregate annual loss exceedence probability (AEP) curve. 340 Table 1 summarises the implications of a number of key uncertainties and assumptions 341 present in the four modules. 342

#### 343 [TABLE 1 AROUND HERE]

344 As we demonstrate in section 3.0, the sampling uncertainty associated with extreme events can be large. This is because different realisations of events with a common return period 345 produce different losses, and multiple stochastic model runs of a given length may generate 346 very different sets of extreme events. Whilst it is possible to handle this uncertainty by 347 producing an extremely large stochastic event set, using the hazard module to simulate every 348 small scale event that occurs in such a large event set is not computationally feasible. This 349 computational restraint requires that a simple event similarity criterion based on hydrograph 350 peak and hydrograph volume is used to test for similar previously simulated events. Events 351 352 are only simulated with the hydraulic model if the hydrograph peak or hydrograph volume on either river differs from a previously simulated event by more than a preset threshold of 10%. 353 If this requirement is not met then it is assumed that a similar event has already been 354 simulated, and the calculated loss from this earlier simulation is selected and added again to 355 the event loss table. 356

## 357 2.2.1 Stochastic Rainfall Module

The Cumulative Distribution Function Generalised Pareto Distribution Model CDFGPDM employed here uses statistical distributions to define storms in terms of mean durations, intensities and inter-arrival times. The CDFGPDM is a profile-based stochastic rainfall model that generates a series of independent rainstorms and 'inter-arrival' periods (dryspells) via a Monte Carlo sampling procedure. The model retains the Eagleson (1972)

approach of characterising a storm in terms of inter-arrival time, duration and mean intensity 363 whilst incorporating a profiling component to distribute the total precipitation throughout the 364 duration of the storm. Storms in the observational record are classed by duration and their 365 intensities are recorded using empirical cumulative distribution functions (CDFs). In order to 366 enable the simulation of storms of greater duration or intensity than in the observational 367 record, the tails of the CDFs are modelled using maximum likelihood Generalised Pareto 368 Distributions (GPD). The threshold above which the GPD was fitted depended on the 369 number of observations in each class and ranged from the 75<sup>th</sup> to 95<sup>th</sup> quantile. The empirical 370 CDFs are then combined with their modelled GPD tails to generate hybrid distributions from 371 which storm characteristics can be sampled. Previous studies have argued that rainfall runoff 372 models can be realistically driven by such a model structure as the shape parameter within the 373 GPD allows a wide range of upper tail shapes to be adequately captured (Cameron et al., 374 2000, 1999). Following Cameron et al. (1999) we here define a rainstorm as any event with 375 an intensity of  $\geq 0.1$  mm/hour, a duration of  $\geq 1$  hour and an inter-arrival time of  $\geq 1$  hour, 376 where no zero-rainfall periods are permitted within a storm. It should be noted that for the 377 ERA-Interim and CMORPH driven models, the minimum duration and inter-arrival times 378 379 were 3 hours due to the 3 hour timestep of these products. This definition encapsulates all recorded precipitation in the 1 hour interval historical records available for Dublin, making it 380 appropriate for characterisation and subsequent generation of continuous rainfall records. 381 The rainstorm generation procedure is identical to the method detailed in Cameron et al. 382 (1999). In order to evaluate the model's ability to recreate the extremes seen in the observed 383 series, a total of 50 synthetic series of 40 years length were simulated using the rain gauge 384 derived series for the Dodder catchment. The annual maximum rainfall totals (ANNMAX) 385 for each duration class were extracted from the synthetic series and plotted against their 386 counterparts from the observed catchment average series (figure 2). The reduced variate 387 plots show that the observed ANNMAX values are well bracketed by those from the 50 388 synthetic series, indicating the ability of the model to recreate a reasonable distribution of 389 extreme events suited to a study of flood risk. 390

#### 391 [FIGURE 2 AROUND HERE]

Due to the need to limit model complexity and computational expense, it was necessary to 392 assume a spatially uniform rainfall across the modelled catchments. Such an assumption may 393 be justified for Dublin as the modelled catchments are relatively small (<130 km<sup>2</sup>) and floods 394 in this region are driven by large weather systems such as frontal depressions and decaying 395 hurricanes rather than by small scale convective cells. The gauge-based catchment average 396 397 records produced for the Dodder and Tolka catchments were tested for correlation, yielding a Pearson's linear correlation coefficient of 0.89 and a Kendall tau of 0.69. These values 398 indicate that rainfall in the two catchments is indeed strongly correlated; however the lack of 399 perfect correlations implies that the approach will result in a slight overestimation of domain-400 total rainfall for a given event. The assumption allows a spatially uniform, time varying 401 rainfall series to be generated for all catchments by training the CDFGPDM on a single, 402 centrally located, observation site. However, due to significant variation in altitude across the 403 domain, it was necessary to correct the rainfall intensities of the generated series for each 404

405 catchment as the observed precipitation intensity distributions varied between the catchment mean records and the central training site. To achieve this, a quantile-quantile bias correction 406 method (Boé et al., 2007) was used on each observed record type in turn, where adjustment 407 factors for each quantile bin were obtained by comparing the observed time series at the 408 training site to the observed catchment average rainfall series. Therefore, for each of the 409 modelled catchments, a different set of adjustment factor values were generated for the 410 ground gauge, radar, ERA-Interim and CMORPH data, allowing precipitation time series to 411 412 be generated in which the correct precipitation intensity distributions of each individual catchment are persevered despite all catchments sharing a common temporal rainfall pattern. 413

#### 414 2.2.2 Hazard Module

The hazard module consists of a hydrological model and a hydraulic model. The 415 hydrological model employed here is the widely used conceptual rainfall runoff model HBV 416 417 (Bergstrom and Forsman, 1973; Bergström and Singh, 1995). While there are many variants of the HBV model, the one used for this study is most closely related to HBV Light (Seibert 418 and Vis, 2012). The model uses precipitation, temperature and potential evaporation as 419 inputs, the latter of which is calculated from extraterrestrial radiation and temperature using 420 the McGuiness model (McGuinness and Bordne, 1972), to produce an estimate of river 421 discharge at the gauge station locations shown in figure one with an hourly timestep. Model 422 calibration was undertaken to generate behavioural parameter sets for each precipitation data 423 source in each catchment. Initially, the 15-parameter space was explored using Monte Carlo 424 425 simulation and parameter ranges were set by visually identifying upper and lower limits from the resultant simulations. Where the model did not exhibit detectable parameter range limits, 426 ranges from previous studies were employed (Abebe et al., 2010; Cloke et al., 2012; Shrestha 427 428 et al., 2009). Once defined, the parameter ranges were sampled using Latin hypercube Monte Carlo sampling to produce 100,000 parameter sets, a number of samples which proved 429 computationally feasible whilst providing adequate exploration of the parameter space. The 430 parameter sets were then used to simulate discharge during a period for which observations 431 were available, and those that failed to produce behavioural simulations, defined by a Nash-432 Sutcliffe (NS) score exceeding a threshold of 0.7 (Nash and Sutcliffe, 1970), were discarded. 433 The choice of performance measure and threshold used to define what constitutes a 434 behavioural simulation is necessarily subjective (Beven and Freer, 2001); NS was chosen as 435 it is particularly influenced by high flow performance, and the threshold of 0.7 was selected 436 following visual inspection of hydrographs generated from a preliminary sample of parameter 437 sets. In order to assign weights, the behavioural parameter sets were then ranked and 438 weighted by their ability to minimise error in the top 0.1% of the flow duration curve. Due to 439 computational constraints imposed by the subsequent hydraulic model, the number of 440 behavioural parameter sets was limited to the 100 highest ranked sets. Weighting was 441 performed by calculating the inverse sum of absolute errors between the simulated and 442 observed series in the top 0.1% of the flow duration curve for each of the behavioural 443 parameter sets. These values were then normalised to give the best performing parameter set 444 a weight of 1 and the worst a weight of 0. This approach favours behavioural parameter sets 445

that best simulate high-flow periods and is therefore appropriate for a study concerned withflood risk.

448 Initially, attempts were made to calibrate HBV using each precipitation data type. However, only those simulations driven using the gauge-derived precipitation data were able to satisfy 449 the behavioural NS threshold in all catchments. Models driven using ECMWF and 450 CMORPH data were especially poor; this may be explained by their reduced spatial and 451 452 temporal resolution compared to the gauge and radar data. As the model was only able to adequately represent observed catchment flow characteristics using the behavioural 453 parameter sets identified using gauge data, it was therefore decided to employ these 454 parameter sets for all simulations. The very large number of event simulations required to 455 produce an EP curve precluded HBV parametric uncertainty from being incorporated directly 456 into the CAT model; such an approach would have further increased the required 457 computational resource to an unfeasible level. Due to this limitation, the highest ranked 458 parameter set produced using gauge data was used to generate the EP curves. The impact of 459 460 parametric uncertainty is addressed separately on an event basis in section 3.3, where the weighted behavioural parameter sets are used to produce uncertain loss estimates with 5-95% 461 confidence intervals for four synthetic flood events. 462

The hydraulic model LISFLOOD-FP (Bates and De Roo, 2000) is used to generate flood 463 inundation maps from the event hydrographs produced by HBV. The configuration 464 employed here uses a subgrid representation of the channel (Neal et al., 2012b) coupled to a 465 2D flood plain model that uses a simplified 'inertial formulation' of the shallow water 466 equations (Bates et al., 2010) solved using the numerical method of de Almeida et al. (2012). 467 The channel models include weirs and were constructed using surveyed river cross sections 468 supplied by Dublin City Council, and the digital elevation model (DEM) for the 144 km<sup>2</sup> 2D 469 hydraulic model was constructed from 2 m resolution bare-earth LiDAR data that was 470 coarsened to 10 m and 50 m resolution (1,440,000 and 57600 cells respectively) using 471 bilinear resampling (Fewtrell et al., 2008). Where >50% of the surface area of a cell was 472 occupied by building(s), identified through Ordinance Survey Ireland data, the cell elevation 473 was increased by 10 m to become a 'building cell'. Model calibration of channel floodplain 474 friction was undertaken by driving the hydraulic model with observed discharges and 475 comparing the observed and simulated flood inundation extents for the August 1986 476 Hurricane Charlie and the November 2002 flood events. These are the largest events for 477 478 which observed discharge and inundation data are available, with the 2002 event generating 479 \$47.2 million in unindexed losses (AXCO, 2013), and have been attributed with ~700 and ~100 year return periods respectively (RPS Consulting Engineers, 2008; RPS MCOS, 2003). 480 The extent of the larger 1986 event was digitised from hand drawn post-event flood outline 481 maps, which included indications of dominant flow directions, although the completeness of 482 these maps is uncertain. The November 2002 flood outlines were supplied by Dublin City 483 Council. Both of these datasets will be subject to considerable uncertainty as they were 484 constructed from eye witness accounts and post-event ground based observations; they 485 should therefore be considered as approximations of the true maximum extents. Observed 486 and simulated flood outlines for the calibration events are shown in figure 3. The quantitative 487

488 F-squared performance measure (Werner et al., 2005) was calculated for each calibration run, with the optimised model yielding values of 0.62 and 0.44 for the 10 m and 50 m resolution 489 models respectively. Some of the variation between the observed and simulated extents may 490 be explained by errors in the observed data; some may also be explained by land 491 492 development and engineering works that occurred between the events and the date on which the modern DEM terrain data were collected; this latter factor may have an especially strong 493 influence for the 1986 event results. Nevertheless, the F-squared values still compare 494 495 favourably with a previous study of urban inundation modelling (Fewtrell et al., 2008), in which it is noted that performance of models in urban areas is strongly affected by the ability 496 of the DEM to represent urban structures; subsequent studies have also highlighted the 497 influence of detailed terrain features on urban inundation processes (Fewtrell et al., 2011; 498 Sampson et al., 2012). These findings are further evidenced here, as the reduced 499 representation of buildings on the 50 m DEM removes flow restrictions and results in an 500 overestimation of flood extents with a corresponding reduction in water depths near the 501 channel. Despite this, qualitative assessment of the modelled dynamics with reference to the 502 observations suggests that, at both resolutions, the model is capturing the dominant process 503 well, with water entering the floodplain in the correct areas. Unfortunately, the 504 505 computational expense of the 10 m resolution model was several orders of magnitude greater than the 50 m model, resulting in simulation times of several hours compared to ~ 20 seconds 506 for a 48 hour event. Due to this cost, the 50 m model was adopted for use within the CAT 507 model. Whilst this will result in some lost predictive skill relative to the 10 m model, the 508 representation of 2D flow both on and off the floodplain ensures the model remains more 509 sophisticated than the 1D or quasi-2D approaches typically employed by vendor CAT 510 models. The implication to loss estimates of this decision is briefly discussed in section 3.3. 511

512 [FIGURE 3 AROUND HERE]

513

#### 514 2.2.3 Vulnerability Module

A synthetic portfolio of insured properties, modelled on real data, was provided by Willis 515 Global Analytics for use in this study. This was necessary to preserve the anonymity of real 516 policy holders, and the portfolio was built by resampling a distribution of asset values for the 517 region. As is common for insurance portfolios, the data were aggregated to postcode level. 518 The portfolio took the form of an insured sum for three lines of business (residential, 519 commercial and industrial) for each postcode area. It is common practice in industry to 520 disaggregate such datasets using proxy data (Scott, 2009), and the approach adopted here to 521 use the National Oceanic and Atmospheric Administration (NOAA) Impervious Surface Area 522 (ISA) dataset as a proxy for built area (Elvidge et al., 2007). This method assumes a linear 523 relationship between the percentage of a grid cell that is impervious and its insured value, and 524 allows the sum insured within each postcode to be distributed around the postcode area based 525 on ISA pixel values. From these data we built a simple industry exposure database (IED) that 526 contained the values of insured assets for each line of business within each grid cell. 527

When a cell is flooded, the damage sustained within the cell is calculated using depth-damage 528 functions supplied by Willis Global Analytics that were derived from historical data of floods 529 in European cities. In this paper we employ both a simplified deterministic depth damage 530 curve approach and a more sophisticated uncertain vulnerability function. The simplified 531 approach involves separate curves for the residential, commercial and industrial lines of 532 business that relate the water depth within a cell to the percentage of the cell's insured value 533 that is lost. These simple curves therefore represent a mean damage ratio and were used for 534 all experiments other than the vulnerability uncertainty analysis in order to reduce 535 computational cost and better isolate the subject of each experiment. The more sophisticated 536 functions used in the vulnerability uncertainty analysis sample around the fixed curves using 537 modified beta distributions. Here, the depth in a cell determines the mean damage ratio as 538 well as the probabilities of zero damage (P0) and total loss (P1). A stratified antithetic 539 sample of values between 0 and 1 is performed, with all values below P0 being assigned a 540 damage ratio of 0 and all values above P1 being assigned a damage ratio of 1. The values 541 between P0 and P1 are rescaled to between 0 and 1 and used to sample from a beta 542 distribution whose parameters are calculated based on the mean damage ratio, P0, P1 and an 543 544 assumed variance. The result is a sample of damage ratios, with a mass of values at zero, a mass of values at one, and an intermediary range drawn from a beta distribution. As the 545 water depth in a cell increases, the mass of zero damages becomes smaller, the mass of total 546 losses becomes larger, and the mean of the intermediary sampled beta distribution moves 547 towards one (total loss). This method is currently used by Willis on an operational basis and 548 therefore represents industry practice at the date of publication. 549

#### 550 2.2.4 Financial Module

The financial module employed here is used to aggregate simulated losses from the hazard 551 module across a specified aerial unit (here the entire domain) before generating and 552 resampling occurrence and loss distributions from the results. The occurrence distribution 553 represents the distribution of event counts for a given time period (here defined as one year) 554 using an empirical CDF. The main body of the loss distribution is modelled using an 555 empirical CDF, with a GPD fitted to the tail to produce a smooth curve where data are sparse. 556 A synthetic series can then be rapidly generated by adopting a Monte Carlo resampling 557 method. This procedure samples first from the occurrence distribution to find the number (n)558 of events occurring in a given year. The loss distribution is then sampled *n* times to assign a 559 loss to each event. Finally, the annual aggregate loss is found by summing the losses for that 560 year. By repeating this process a large number of times, multiple synthetic series can be 561 562 generated. From these series, an annual AEP curve can be generated that includes confidence intervals derived from the spread of values at any given return period. The annual AEP curve 563 is a standard insurance tool that is used to express the expected probability of exceeding a 564 given level of loss over a one year period, i.e. the expected '1 in 100 year loss' is equivalent 565 to a loss with an annual exceedence probability (AEP) of 0.01. 566

#### 567 **3.0 Results - Event Sampling Uncertainty**

A known source of uncertainty within a CAT model originates in the event generation 568 procedure used to build an event set. This is referred to as 'primary uncertainty' by the 569 insurance industry (Guin, 2010). A key difficulty in calculating the expected loss at a given 570 AEP is that the predicted insured loss will vary from one model run to another due to the 571 572 random component of the stochastic module. One method of reducing this 'sampling uncertainty' is to simulate a series that is considerably longer than the desired recurrence 573 interval (Neal et al., 2012a). Alternatively a large number of realisations can be simulated, 574 575 and the expected loss can then be defined by the mean loss across the realisations. The second method also allows the sampling uncertainty to be investigated by looking at the 576 spread of values across the realisations. The number of realisations that it is feasible to 577 simulate is determined by the required series length and the available computational resource. 578 Here the stochastic module is trained using the rain gauge record and used to generate 500 579 realisations of a 1000 year rainfall series in order to investigate the effect of sampling 580 uncertainty on the 1-in-1000 year loss. 581

582 The object of this experiment is to determine the number of realisations required to adequately capture the range of possible losses at a given event scale. One way to examine 583 such 'sampling uncertainty is to assemble batches of realisations and observe how key 584 descriptors (such as the mean loss or standard deviation of losses) vary between batches. By 585 586 altering the number of realisations in each batch, it is possible to observe how the variation of descriptors between batches changes as the batch size changes. It is then possible to predict 587 the expected average variation, in terms of the descriptors, between the simulated batch of n588 realisations and any other batch of *n* realisations. 589

To do this, the maximum losses recorded in each of the 500 realisations were randomly sampled to produce batches containing 5, 10, 25, 50, 100 or 250 loss ratios ('batch A'). The process was repeated to produce a second batch ('batch B') of identical size to batch A. The mean and standard deviation of loss ratios in batch A ( $\bar{L}_A$  and  $s_A$ ) were then calculated and compared to their equivalent values in Batch B ( $\bar{L}_B$  and  $s_B$ ), yielding two simple measures:

595	$M =  \overline{L}_A - \overline{L}_B $	(1)
596	$S =  s_A - s_B $	(2)

#### 597 [FIGURE 4 AROUND HERE]

By repeating this process a large number of times (10,000 for each batch size), the expected 598 uncertainty due to sampling variability can be assessed. The results of this experiment are 599 shown in figure 4a, where M is expressed as a percentage of the mean 1-in-1000 year loss 600 across all 500 realisations and S is equivalently expressed as a percentage of the standard 601 deviation across all 500 realisations. The plots show that differences between batches A and 602 B decrease as the number of samples within a batch increases, with the median value of M603 604 decreasing from 23.0% to 3.8% as the batch size increases from 5 to 250. This finding can be explained by the underlying distribution of loss ratios being increasingly well represented as 605 the sample size is increased; this is observed in the diminishing value of S as sample size 606 increases. By transforming the median values of M with reciprocal  $1/M^2$  and fitting a linear 607

608 regression model, the expected value of M for the 500 realisations was calculated as 2.7%. This indicates that the mean loss ratio of any 500 simulated realisations will typically differ 609 from any other batch of 500 realisations by ~ 3% of the mean loss ratio itself; the same 610 process yields a value of 2.3% for the standard deviations (figure 4b). Primary uncertainty is 611 an accepted facet of catastrophe modelling and, relative to inherent aleatory uncertainty, 612 uncertainty of this order due to sampling variability is reasonable (Guin, 2010). Whilst the 613 uncertainty caused by sampling variability could be reduced by significantly increasing the 614 number of realisations simulated, the additional computational cost of such an increase would 615 be large and the benefit questionable in the presence of other uncertainties within the 616 calculation chain. For the purpose of this study we identify 50 realisations as the minimum 617 required; at this level the mean and median values of M and S are <10% of the mean and 618 standard deviation of all 500 realisations respectively. The practical implication of this 619 analysis is that it is necessary for the hazard module to simulate >50 time series of length 620 621 equal to the return period of interest.

#### 622 3.1 Variability across data sources

The availability and quality of observed precipitation records varies greatly between sites. In 623 order to investigate how the use of different types of precipitation data might affect predicted 624 losses, each of the data types described in section 2.1 was used to train the stochastic module. 625 The training record length was defined by the longest period for which a continuous record 626 was available from all data sources; this ran from the 1<sup>st</sup> January 2002 to the 1<sup>st</sup> May 2009. 627 This period is clearly shorter than ideal and it is likely that the true variability within each 628 data source is underrepresented as a result; however it was necessary to ensure that the 629 records were of equal length over the same period in order to fairly compare between data 630 631 types. All parameters in the hazard, vulnerability and financial modules were identical across the simulations. Taking a maximum return period of interest to be the 1-in-10,000 year 632 event, 500,000 years' worth of simulations was performed for each data type (giving the 633 required 50 realisations of the 1-in-10,000 year event). The annual aggregate EP curves 634 resulting from these model runs are shown in figure 5, with uncertainty bounds that represent 635 the 5 - 95% confidence intervals generated by the financial module. Also plotted are the 636 modelled losses of two observed historical floods (August 1986 and November 2002), 637 produced by driving the hydraulic and vulnerability components with observed river 638 discharges. 639

#### 640 [FIGURE 5 AROUND HERE]

It is immediately apparent from figure 5 that the different precipitation data sets produce very 641 different EP curves despite the fact that each record covered the same spatial area over a 642 common period of time. At certain points the difference can be as great as an order of 643 magnitude – for example, the ERA-Interim driven model predicts a 1-in-100 year (AEP =  $10^{-10}$ 644  $^{2}$ ) loss ratio of 0.02% whereas the CMORPH driven model predicts a loss ratio of 0.17%. 645 The pronounced differences between the curves can be explained in terms of the ability of 646 647 each of the data sources to represent the local rainfall patterns. The gauge and radar driven models produced EP curves of similar shape, with losses from the radar driven model being 648

slightly lower than from the gauge record. Their relative similarity compared to the ERA-Interim and CMORPH driven models was expected as both are detailed local data sources rather than global products. Furthermore the adjustment factors for radar rainfall intensity were derived from the gauge record so that the two records had equal 3-monthly rainfall volumes. As a result, storms were usually captured in both records and attributed with similar rainfall totals, yielding similar stochastic model calibrations and therefore similar loss projections.

The curves produced by the ERA-Interim and CMORPH driven models differ greatly from 656 those produced by the local gauge and radar datasets. The ERA-Interim curve shows only 657 gradual growth in losses as the return period increases to the maximum modelled value of the 658 1-in-10,000 year event, and at all return periods the ERA-Interim model under predicts 659 compared to the other data sources. By contrast, the losses predicted by the CMORPH driven 660 model are consistently higher than the others, especially at lower return periods. Figure 6a 661 shows cumulative daily precipitation for all four data types. As previously found by Kidd et 662 663 al. (2012) in a study of rainfall products over Northwest Europe, CMORPH is found to consistently underestimate rainfall totals compared to the local data whereas ERA-Interim 664 consistently overestimates rainfall totals. Given the pattern of cumulative rainfall totals, the 665 opposite pattern found in the loss projections is initially surprising. However, once hourly 666 rainfall intensities are considered (figure 6b) the findings can be explained. CMORPH is 667 found to underestimate rainfall totals in this region because of the limited sensitivity of 668 satellite products to very low intensity rainfall ('drizzle') (Kidd et al., 2012). However, it 669 exhibits higher rainfall intensities in the upper (>95<sup>th</sup>) quantiles of rainfall intensity than the 670 other records. Severe storms in the CMORPH record typically had slightly higher rainfall 671 volumes than the same storms in other records, the result of which is an increased expected 672 loss at all return periods. ERA-Interim has the opposite problem whereby the frequency of 673 674 low intensity precipitation is over predicted and high intensity precipitation is severely underestimated. 675

#### 676 [FIGURE 6 AROUND HERE]

#### 677 **3.2 Uncertainty due to record length**

A similar approach to the above comparison between data sources was adopted to examine 678 679 the sensitivity of projected losses to the length of record used to train the stochastic module. For this test the gauge precipitation data were cropped to produce training records of 5, 10, 20 680 and 40 years in length. The training records share a common end date (September 2011) and 681 therefore the longer records extend further into the past. As with the data sources test, all 682 other parameters were held constant across the other components, and the resulting EP curves 683 are plotted in figure 7. The EP curves demonstrate that altering the training record length has 684 a significant impact on the projected losses for a given return period. At  $AEP = 10^{-2}$ , the 685 median expected loss ratio ranges from 0.05 to 0.28; at AEP= $10^{-3}$ , representing the 1-in-1000 686 year event, the expected loss ratios vary from 0.12 to 0.60. The relative overestimation of 687 loss ratios by the 5 year training data set demonstrates how the presence of a large event in a 688 689 short training set is able to skew the results. There are two storms that generate exceptionally

high precipitation volumes in the 40 year observed record, and the second of these falls within the final five years that form the 5 year training record. When trained with this short record, the stochastic module inevitably over predicts the rate of occurrence of such storms, leading to an overestimation of expected flood losses. Modelled uncertainty increases as the return period increases; in the case of the 10 year training period, the range of modelled losses at the  $10^{-4}$  AEP level is greater than the median estimate of 0.36%.

#### 696 [FIGURE 7 AROUND HERE]

#### 697 **3.3 Hazard module uncertainty**

In order to provide some context for the uncertainty associated with the choice of driving 698 data, the uncertainty resulting from the choice of parameter set used with HBV was also 699 investigated. Due to computational limitations it was not feasible to produce EP curves for a 700 large number of parameter sets, so instead we focussed on individual events. The largest 701 event was extracted from each of four 500 year runs of the stochastic module. Each event 702 703 was then simulated using the 100 best performing HBV parameter sets, all of which had previously been selected and assigned weights as described in section 2.2.2. The resulting 704 hydrographs were then used to drive the hydraulic model, and the event loss from each 705 simulation was calculated and weighted according to their respective parameter set weights. 706 Figure 8 shows each event hyetograph, the range of hydrographs produced by the different 707 parameter sets on both the Dodder and Tolka rivers, and the resulting weighted CDF of loss 708 ratios. The weighted 95% confidence interval values for peak discharge, hydrograph volume 709 710 and loss ratio are shown in table 2.

#### 711 [FIGURE 8 AROUND HERE]

## 712 [TABLE 2 AROUND HERE]

The results of this exercise demonstrate the impact of parametric uncertainty within the hydrological model on expected losses. For the smallest of the events (event 3), the ratio of the 95<sup>th</sup> to 5<sup>th</sup> quantile peak discharges for the Dodder and Tolka was ~1.1. Despite these relatively modest increases, the ratio of 95<sup>th</sup> to 5<sup>th</sup> quantile losses across the whole domain was ~1.7. For a larger event (event 4), the equivalent 95<sup>th</sup> to 5<sup>th</sup> quantile peak discharge ratio increased to ~1.2 and yielded a ratio of losses of ~3.25.

The high sensitivity of expected losses to relatively smaller percentage changes in 719 hydrograph peak or volume is due to the fact that losses are only affected by the part of the 720 hydrograph that drives flood inundation – namely the portion of flow that is out-of-bank. 721 This region of the hydrograph is clearly sensitive to parametric uncertainty, leading to the 722 high degree of uncertainty in modelled losses exhibited here. It should also be noted that 723 these results are sensitive to the subjective choice of behavioural threshold and performance 724 725 measures employed. Had a higher threshold been chosen, the available parameter space from 726 which behavioural sets could be selected would be smaller, leading to a reduction in the modelled loss ratio uncertainty. However, despite parametric uncertainty clearly being 727

important, in the context of this study the choice of driving precipitation data source remainsthe greater source of uncertainty in modelled losses.

As noted in the hazard module description (section 2.2.2), the high computational cost of 730 731 hydraulic simulations on a 10m grid prevented the finer resolution model from being adopted. The earlier qualitative assessment of the hydraulic model at 50 m relative to 10 m indicated 732 that both exhibited similar first order dynamics, with the coarser model producing a greater 733 734 simulation extent with reduced water depths as a result of the reduced building blockages and terrain smoothing. In order to provide a general indication as to how this might affect loss 735 estimates, the losses from the 10 m and 50 m calibration simulations were calculated. These 736 calculations yielded loss ratios of 0.101 and 0.146 respectively, indicating that areas of deep 737 localised flooding present in the 10 m simulations were generating high losses not adequately 738 captured by the 50 m model. However, although a more detailed study is required before 739 firm conclusions can be drawn regarding the importance of hydraulic model resolution in this 740 context, this result does suggest that the contribution of the hydraulic model to the total 741 742 hazard model uncertainty may be small relative to the hydrological model.

## 743 **3.4 Vulnerability module uncertainty**

Contemporary CAT models typically account for uncertainty within the vulnerability module by using historical claims data to develop a distribution of damage ratios for any given water depth as described in sections 1.3 and 2.2.3. In order to investigate the uncertainty imparted onto the EP curves by the vulnerability module, the 500,000 years' worth of hazard module simulations performed for section 3.1 were coupled to the uncertain vulnerability module. This process generated EP curves for each data source in which the 5-95% confidence intervals are defined by uncertainty within the vulnerability module (figure 9).

## 751 [FIGURE 9 AROUND HERE]

752 Figure 9 demonstrates that the uncertainty imparted by the vulnerability module is large relative to uncertainty generated by the financial model (figure 5) for small to moderate event 753 scales (1 in 10 to 1 in ~250 year). However, for the more extreme events the two contribute 754 uncertainty of a broadly similar magnitude. This is due to the nature of uncertainty within the 755 756 vulnerability module. At small event scales the vulnerability module is able to generate a 757 wide range of loss ratios even when water depths are relatively low. This produces significant uncertainty within the EP curve relative to a model that uses fixed depth-damage 758 759 curves, as loss ratios from the fixed curves will typically be low when water depths are shallow. However, during more extreme events where high loss ratios dominate the curve 760 761 due to increased water depths, the relative uncertainty of the vulnerability model is seen to decrease as both the uncertain and fixed vulnerability methods cannot generate losses 762 exceeding 1 (total loss). This exhibition of asymptotic behaviour highlights the fact that 763 764 uncertainties vary both in absolute terms and relatively to each other as event scale changes.

## 765 **4.0 Discussion**

The results presented above examine how the loss estimates produced by a flood catastrophe 766 model are affected by the choice of data used to drive the model's stochastic component. 767 Parametric uncertainty from the hydrological model has also been examined on an event 768 basis to contextualise the scale of uncertainty induced by the stochastic component and 769 770 uncertainty from the vulnerability module has also been modelled. The findings highlight the 771 difficulty in producing robust EP curves using a cascade methodology, as the uncertainty associated with each component is large and increases as event scale increases. Furthermore, 772 not all sources of uncertainty have been considered – for example flood defence failure rates. 773 Despite this, the model presented here is very detailed compared to standard industry 774 practice, and contains detailed local information (such as river channel geometry and 775 features) that would often be unavailable under the time and financial constraints of most 776 commercial catastrophe modelling activities. The required computational resource would 777 also exceed what is practicably available if models of this detail were extended to cover 778 779 entire national territories. As a result, the uncertainty estimates made in this study are likely to be conservative. The CMORPH and ERA-Interim precipitation records have global 780 coverage and are typical of the kind of product that could be used to drive a commercial CAT 781 782 model. However, the hydrological model was unable to generate behavioural results when driven by these data sources, indicating their inability to produce realistic storm precipitation 783 and thus runoff in the modelled catchments. It is therefore unsurprising that they generated 784 EP curves that were both very different to each other and to the curves produced using more 785 detailed local records. Examination of the observed precipitation records reveals that the 786 precipitation intensity distributions vary significantly between the data sources. 787 The observed records are relatively short; a common record across all four data sources was only 788 available for a little over seven years due to the short length of radar records and gaps in the 789 ground gauge data. The divergence in estimates of precipitation totals for heavy storms 790 between the observational records is reflected in the synthetic series produced by the 791 792 stochastic module, and this divergence inevitably continues as the simulated event scale 793 increases. This results in the pronounced differences in higher return period loss estimates produced by the model when trained with each of the data sources in turn. Whilst access to 794 longer overlapping records might have reduced the severity of this divergence, the 795 796 consistently different storm rainfall intensities recorded by the four data types means that the stochastic module would still be expected to generate very different estimates of high return 797 period rainfall events depending on which data it was driven with. It is also worth noting at 798 this point that we did not consider the parametric uncertainty associated with fitting GPDs to 799 the precipitation intensity and duration tails; this source of epistemic uncertainty is likely to 800 be large given the relatively short rainfall records to which the GPDs are fitted and therefore 801 the true uncertainty is most likely greater than reported here. Unfortunately, investigating the 802 impact of this on modelled losses would have required a number of runs of the entire model 803 804 cascade that was computationally prohibitive.

The EP curves were also found to be sensitive to the length of record used to train the stochastic module. Unfortunately, satellite and model reanalysis precipitation records are typically short (CMORPH runs from the mid-1990's; ERA-Interim from 1979) and the results presented here demonstrated significant differences between the EP curves produced 809 by records of 5, 10, 20 and 40 years in length. Lack of available data prevented longer records from being tested, but our results do indicate that extra care is required when using 810 short (<10 years) records due to the ability of a single extreme observation to skew results. 811 Furthermore, the fact that there is an appreciable difference between the 20 and 40 year 812 813 curves suggests that records of at least 40 years in length should be used where possible. 814 Future reanalysis products hoping to extend records further back in time may help to alleviate this issue; the European Reanalysis of Global Climate Observations (ERA-CLIM) project 815 aims to provide a 100 year record dating back to the early 20<sup>th</sup> century. The impact of 816 parametric uncertainty within HBV should also be of concern to practitioners. The model in 817 this study was calibrated with detailed precipitation and discharge records and might 818 therefore be considered tightly constrained compared to commercial models that will have to 819 operate at national scales. Despite this, the variation in predicted loss ratios over a range of 820 behavioural parameter sets for individual events was very large. Due to computational 821 822 constraints we were unable to also consider uncertainty in the hydraulic model component of the hazard module, although it is believed that the hydraulic model is a relatively minor 823 source of uncertainty in this context (Apel et al., 2008a). Former studies have indicated that 824 825 topography is the dominant driver of uncertainty within hydraulic models if we consider the inflow boundary condition uncertainty to be associated with the hydrological model (Fewtrell 826 et al., 2011; Gallegos et al., 2009; Schubert et al., 2008; Yu and Lane, 2006), and given the 827 differences seen between the calibration runs at 10 m and 50 m resolution (figure 3) it is very 828 829 likely that the uncertainty reported in this study is an underestimate of the total uncertainty present within the hazard module. 830

831 The final uncertainty source considered was the vulnerability module. This module was 832 found to contribute significantly to the uncertainty at smaller event scales but, due to the inherently asymptotic nature of a damage function, its relative contribution was shown to 833 834 decrease as event scale increased. Of particular interest is the fact that, in contrast to some previous studies (e.g. Moel and Aerts, 2010), the vulnerability module uncertainty is smaller 835 than the uncertainty resulting from choice of data used to drive the hazard module. This is 836 likely due to such studies using relatively constrained event scenarios in which under which 837 hazard uncertainty is more limited than in a stochastic model. Studies which considered a 838 wider range of events (Apel et al., 2008b; Merz and Thieken, 2009) have found uncertainty in 839 840 the features controlling the occurrence and magnitude of events (e.g. stage discharge relationships, flood frequency analysis) to be similar to or greater than the vulnerability 841 uncertainty, especially at larger event scales. 842

843 Spatial scales are an important consideration in the context of this study. The catchments modelled in this study are relatively small, and it is reasonable to suggest that the relatively 844 coarse reanalysis and satellite products might perform better for major rivers where fluvial 845 floods are driven by rainfall accumulations over longer time periods and large spatial areas. 846 Some of their inherent traits, such as tendency for the reanalysis product to persistently 847 'drizzle' while underestimating storm rainfall accumulations, will negatively impact their 848 applicability to flood modelling across most catchment scales although the severity of the 849 effect may reduce as catchment sizes increase. However, it is wrong to assume that the 850

dominant driver of flood risk is always large events on major rivers. A significant proportion of insurance losses resulting from the 2007 UK floods and 2013 Central European Floods can be classified 'off-floodplain' – that is to say they occurred either as a result of surface water (pluvial) flooding or as a result of fluvial flooding in small catchments (Willis, personal communication). This suggests that even when considering large events, the ability to produce realistic hazard footprints in small catchments remains critical and thus for practitioners concerned about such events, the findings of this paper remain relevant.

When considered together, the above findings make it difficult to commend a stochastic flood 858 model driven by precipitation data as a robust tool for producing EP curves for use in 859 portfolio analysis. The sensitivity of the stochastic component to the driving data is of 860 fundamental concern due to the high degree of uncertainty in observed precipitation 861 extremes, suggesting that alternative driving mechanisms such as flood frequency analysis 862 should be evaluated in this context. Furthermore, the results demonstrate sensitivity to model 863 parametric uncertainty that will be difficult to overcome. However, these shortcomings do 864 865 not mean that such a model has no value. Although it may be difficult to use such a system to project accurately how often events of a certain magnitude will occur, and thus estimate 866 probable losses over a given time window, the model could still be used to assess the relative 867 risk of assets within a portfolio. We argue that understanding and quantifying the 868 uncertainties generated by the stochastic and hazard modules for a given portfolio may be 869 870 important to managing assets effectively. Although the computational demand of the hazard module in particular will likely render this unfeasible on an operational basis, studies such as 871 this may be used to inform judgments regarding the total uncertainty within such model 872 structures. A valuable exercise for users of commercial models may be to compare such 873 874 findings to the uncertainty generated by their own models, many of which may attempt to account for hazard uncertainty via sampling widened distributions within the vulnerability 875 876 module.

## 877 **5.0 Conclusions**

In this study, stochastic, hazard, vulnerability and loss modules have been assembled into a 878 cascade framework that follows the same principles as an insurance catastrophe model. The 879 model operates by generating a large synthetic series of events in the stochastic component 880 which is then simulated by the hazard component. The vulnerability component assesses the 881 damage and loss caused by each event, building up a database of occurrence intervals and 882 883 event losses. Finally, the loss component resamples from the modelled occurrence and loss distributions, producing exceedence probability curves that estimate the expected annual 884 aggregate loss for a range of return periods. The model simulates fluvial flood risk in Dublin, 885 Ireland, and the components were calibrated using local historical observations where 886 appropriate data were available. 887

A number of different precipitation datasets were tested with the model, including high resolution local gauge and radar records, model reanalysis records (ERA-Interim) and satellite records (CMORPH). The exceedence probability curves produced by the model were found to be very sensitive to the choice of driving precipitation data, with different 892 driving datasets producing loss estimates that varied by more than an order of magnitude in some instances. Examination of the observational records reveals that the precipitation 893 intensity distributions over a common period vary markedly between the different data types. 894 These differences are inevitably reflected in the output produced by the stochastic module 895 and result in large differences in the modelled magnitude of high return period events. The 896 897 calculation chain was also found to be sensitive to the length of observational record available, with the presence of a large event in a short training set resulting in severe 898 899 overestimation of losses relative to models driven by a longer record. The sensitivity of the model to parameterisation of the hydrological model was tested on an event basis. Modelled 900 loss ratios were found to be highly sensitive to the choice of parameter set. Despite all being 901 classified as behavioural, the loss ratios for one event varied by up to six times dependent on 902 the parameter set selected. Finally uncertainty in the vulnerability module was considered. 903 Due to the asymptotic nature of damage functions it was found to be a larger relative 904 905 contributor at small event scales than large, although even at large scales its contribution remained high. However, the impact of both hydrological parameter uncertainty and 906 vulnerability uncertainty were both smaller than the impact of uncertainty within the driving 907 908 precipitation data.

909 Considered together, the results of this study illustrate the difficulty in producing robust 910 estimates of extreme events. The uncertainty in the observed record, along with the short 911 length of records relative to return periods of interest, is of particular concern as observed differences diverge when the event scale is extrapolated far beyond what has historically been 912 observed. A lack of suitable observational data for model calibration makes it challenging to 913 envisage how similar methods to those employed in this study could be used to produce the 914 national scale models required by industry without uncertainty bounds becoming 915 unmanageably high. Further issues that will compound these problems are the scarcity of 916 data relating to the condition and location of flood defences, another important source of 917 uncertainty (Gouldby et al., 2008), and the requirement to build models in data-poor 918 919 developing regions where insurance market growth is greatest. The results of this study have 920 emphasised the dramatic impact of data uncertainties on loss estimates, and it is important that the users and developers of catastrophe models bare such results in mind when assessing 921 the validity of the uncertainty mechanisms within their models. At present, the combination 922 of short record lengths and highly uncertain precipitation intensities during storm events 923 924 make it difficult to recommend the use of rainfall-driven model cascades to estimate fluvial flood risk, especially where estimates of return period are necessary. Looking forward, 925 increased resolution regional reanalysis products with improved rainfall process 926 representation may help to reduce these uncertainties as may the assimilation of local data 927 into global observational datasets to produce improved regional calibrations for rainfall 928 products (Dinku et al., 2013). Further effort should also be concentrated on developing 929 alternative means of characterising the loss driving properties of river basins. One such 930 alternative may be to revisit methods based on geomorphology and flood frequency analysis 931 (Leopold and Maddock, 1953; Meigh et al., 1997) in conjunction with modern observational 932 databases (such as the Global Runoff Data Centre) and remotely sensed data. 933 As 934 supercomputing power continues to grow exponentially, large ensemble stochastic

935 frameworks that combine such approaches will likely become tenable projects over the 936 coming decade.

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## 1212 Tables

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	Key Data	Кеу	Кеу	Additional	Кеу
	Rey Data	Uncertainties	Implications	Assumptions	Implications
Stochastic Module	- Rainfall - DEM	Short observational record lengths Precipitation intensities vary	Limited data to constrain GPD fits to tails of rainfall intensity/duration distributions Modelled losses highly sensitive to	Uniform rainfall pattern	Likely over- estimation of modelled losses
		between data sources chosen data source	chosen data source		
Hazard Module: Hydrological Model	- Rainfall - Temperature - PET - Discharge	Parametric uncertainty Small number of flood events in discharge records Observed flood discharges uncertain	Modelled losses sensitive to parameterisation and calibration	River Dodder artificial reservoirs not modelled	Overestimation of losses, especially when antecedent conditions are dry and reservoir level would be low
		Small number of flood events in discharge records choice of behavioural performance measure	Uncertainty range dependent on performance measure		
Hazard Module: Hydraulic Model	- Discharge - Flood extents - River channel geometry - DEM	Errors in observed extents Observed flood discharges uncertain	Unknown sensitivity of modelled losses to hydraulic model structure	No significant flood defence additions since observed events	Model may simulate losses in newly defended areas
		Roughness coefficients Unrepresented channel features		Choice of which events to simulate based on hydrograph peak and volume	Relationship between hydrograph properties and loss may be oversimplified
		DEM resolution	Flood extents and depths influenced by DEM; losses not grid independent	Depth in building cell assumed to be mean of surround cell depths	Likely over- estimation of modelled losses
Vulnerability Module	<ul> <li>Water depths</li> <li>Postcode areas</li> <li>Depth damage curves</li> <li>ISA data</li> </ul>	ISA data is low resolution	Likely errors in loss calculation as true location of assets is unknown	Fixed damage fixed for a given flood depth	Loss for a given depth would really vary depending on building type
		Depth damage curves highly uncertain	Loss estimates vary depending on choice of depth damage curve	Period of inundation not considered	Possible Over/under estimation of losses for short/long duration events respectively
Financial Module	- Event loss table	-	-	Policy terms such as deductibles and limits not included	Overestimation of losses compared to 'real' portfolios

**Table 1**. Table showing the required data sources for each module, along with keyuncertainties, assumptions and their respective implications for modelled losses.

	Measure	Event 1	Event 2	Event 3	Event 4
Dodder	Peak Discharge (m <sup>3</sup> )	212 - 256	185 - 226	185 - 203	250 - 291
	Volume $(x \ 10^7 \ m^3)$	1.69 – 1.89	1.66 – 1.84	1.76 – 1.97	1.74 – 1.92
Tolka	Peak Discharge (m <sup>3</sup> )	125 - 150	130 - 147	113 - 124	118 – 139
	Volume $(x \ 10^7 \ m^3)$	1.50 - 1.64	1.54 – 1.64	1.49 – 1.60	1.35 – 1.47
Entire Domain	Loss Ratio (%)	0.03 - 0.14	0.04 - 0.07	0.03 - 0.05	0.04 - 0.13

1219	Table 2.	Weighted $5^{th} - 95^{t}$	<sup>h</sup> quantile values	for event based	HBV uncerta	ainty simulations.
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Figure 1. Map of Dublin region. Modelled rivers are shown by thick blue lines.
Hydrological model boundaries are shown in red. Hydraulic model boundary is shown in
yellow. Rain gauge locations shown by black crosses. River flow gauges are shown by
yellow triangles.



Figure 2. Modelled annual maximum rainfall totals for each duration class compared to
observational record for the Dodder catchment. The annual maxima for each class in the 40
year catchment average observed record are ranked and plotted using Gringorten plotting
positions (black circles). The process was repeated for 50 x 40 year simulated series (grey
crosses).



Figure 3. Hydraulic model calibration results. Red shaded area shows observed flood extent.
Blue outline shows flood outline from 10 m resolution model. Yellow outline shows flood
outline from 50 m resolution model. Underlying DEM is 10 m resolution.



Figure 4. Box plots that show the variation between two batches of simulations reducing as the number of simulations in each batch increases. The top plot (4a) shows the difference between the means of the two batches, expressed as a percentage of the mean loss across all 500 simulations. The bottom plot (4b) shows the difference between the standard deviations of the two batches, expressed as a percentage of the standard deviations simulations.





**Figure 5.** Exceedence probability plots produced by the catastrophe model when trained using the four different precipitation datasets. The grey shaded area denotes the 5-95% confidence intervals generated by the financial model. The losses simulated when the hydraulic and vulnerability modules are driven with observed flows for two historical events are shown for reference.



Figure 6. Top plot (6a) showing cumulative precipitation for each source. Bottom plot (6b)
shows anomalies in >90<sup>th</sup> quantile precipitation intensities between gauge and other sources.





1276 Figure 7. Exceedence probability plots produced by the catastrophe model when trained 1277 using the gauge record cropped to four different lengths. The grey shaded area denotes the 5-1278 95% confidence intervals generated by the financial model. The losses simulated when the 1279 hydraulic and vulnerability models are driven with observed flows for two historical events 1280 are shown for reference.



**Figure 8.** Plots showing event hyetographs and hydrographs for the River Dodder (rows 1 and 2) and River Tolka (rows 3 and 4), and cumulative distribution function plots of modelled losses across the entire domain (row 5). The number of parameter sets simulating discharge at or above a given level at time *t* is represented by the hydrograph colour, ranging from all 100 (dark blue) to 1 (dark red). The weighted  $5^{th} - 95^{th}$  quantile values from these plots are shown in table 2.



Figure 9. Exceedence probability plots produced by the model when trained using the four
different precipitation datasets. The grey shaded area denotes the 5-95% confidence intervals
generated by uncertainty within the vulnerability model.