Hydrol. Earth Syst. Sci. Discuss., 10, 1375–1422, 2013 www.hydrol-earth-syst-sci-discuss.net/10/1375/2013/ doi:10.5194/hessd-10-1375-2013 © Author(s) 2013. CC Attribution 3.0 License.



This discussion paper is/has been under review for the journal Hydrology and Earth System Sciences (HESS). Please refer to the corresponding final paper in HESS if available.

Characterization of physically based hydrologic model behaviour with temporal sensitivity analysis for flash floods in Mediterranean catchments

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Received: 4 December 2012 - Accepted: 3 January 2013 - Published: 28 January 2013

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Published by Copernicus Publications on behalf of the European Geosciences Union.

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Abstract

This paper presents a detailed analysis of 10 flash flood events in the Mediterranean region using the distributed hydrological model MARINE. Characterizing catchment's response during flash flood events may provide a new and valuable insight into the processes involved for extreme flood response and their dependency on catchment properties and flood severity. The main objective of this study is to analyze hydrologic model sensitivity in the case of flash floods with a new approach in hydrology, allowing model outputs variance decomposition for temporal patterns of parameter sensitivity analysis. Such approaches enable ranking of uncertainty sources for non-linear and non-monotonic mappings with a low computational cost. This study uses hydrologic model and sensitivity analysis as learning tools to derive temporal sensitivity analysis with a variance based method in the case of 10 flash floods that occurred in the French Pyrenees and Cévennes foothills. This constitutes a huge dataset given the scarcity of data about flash flood events. With Nash performances above 0.73 on average for

- this extended set of validation events, the five sensitive parameters of MARINE distributed physically based model are analyzed. This contribution shows that soil depth explains more than 80% of model output variance when most hydrographs are peaking. Moreover the lateral subsurface transfer is responsible for 80% of model variance for some catchment-flood events' hydrographs during slow declining limbs. The
- unexplained variance of model output representing interactions between parameters reveals to be very low during modeled flood peaks and informs that model parsimonious parameterization is appropriate to tackle the problem of flash floods. Interactions observed after model initialization or rainfall intensity peaks incite to improve water partition representation between flow components and initialization itself. This paper
- ²⁵ gives a practical framework for application of this method to other models, landscapes and climatic conditions, potentially helping to improve processes understanding and representation.



1 Introduction: problem framework

1.1 Flash flood modeling complexity

The Mediterranean climatic zone is prone to heavy rainfall events especially during the fall season. Either quasi-stationary mesoscale convective systems, which can last several hours, or frontal disturbances blocked by the mountains can produce high precipitation totals (Nuissier et al., 2008) that trigger severe flash floods. Precipitation is highly variable both in time and space and this variability increases with elevation (Moussa et al., 2007). This along with topography influence and spatial distribution of soil and land use properties makes hydrological processes largely variable both in time and space (Pilgrim et al., 1988). Flash floods are extreme catchment responses with high peak discharge often produced by severe localized thunderstorms. They are one of the most destructive hazard in the Mediterranean region and caused casualties and billions euros of damages in France over the last two decades (Gaume et al., 2009).

These events often reveal aspects of hydrological behavior that either were unex-¹⁵ pected on the basis of weaker responses or highlight anticipated but previously unobserved behavior (Delrieu et al., 2005). Characterizing the response of a catchment during flash flood events, thus, may provide a new and valuable insight into processes for extreme flood response and their dependency on catchment properties and flood severity (Borga et al., 2008).

In the literature, several approaches are proposed for flash flood events modeling and/or prediction, each with its specificities depending on perception and parameterization of the dominant hydrological processes (Roux et al., 2011; Saulnier and le Lay, 2009; Moussa et al., 2007; Braud et al., 2010, among the others for the Mediterranean region). These models often take advantage of available data in order to assign spa-

tially distributed forcing as well as distributed catchment parameters. However, increasing model complexity can lead to overparameterization and equifinality problems because of high dimensionality and multi-modal response surface. As a result, parameter values might not be identifiable in the calibration process (Beven, 1989). Sieber and



Uhlenbrook (2005) have highlighted that sensitivity analysis (SA) cannot only identify the most important parameters but also contribute to understanding and improving the structure of hydrologic model.

1.2 Understanding uncertainty with sensitivity analysis

- Sensitivity analysis (SA) assesses the impact of model parameters on the output, and is therefore a convenient tool to investigate model behavior and particularly the importance of particular parameterizations within the model. SA has become a popular tool in catchment modeling to explore high dimensional parameter spaces, assess parameter identifiability, and understand sources of uncertainty (Hornberger and Spear, 1981; Freer et al., 1996; Wagener et al., 2001; Tang et al., 2007; Hall et al., 2005; van Griensven et al., 2006). Some studies highlight the usefulness of sensitivity analysis for the improvement of hydrological models (Andréassian et al., 2001; Oudin et al., 2006; Tang et al., 2007; Pushpalatha et al., 2011; Castaings et al., 2007; Ratto et al., 2007b). Other studies used SA to better understand model behavior with respect to inputs such
- as precipitation (Meselhe et al., 2009; Xu et al., 2006).

With the current shift toward model complexification and/or real time hydrometeorological forecasts, of prior importance is the understanding of uncertainty and its sources. In catchment modeling it can be achieved with various methods, of which formal Bayesian methods (Kuczera and Parent, 1998) and GLUE method (Beven and

- ²⁰ Binley, 1992) are the most popular, and also recursive application of RSA for dynamic identifiability analysis (Wagener et al., 2003) or Bayesian total error analysis (BATEA) method for comprehensive calibration and uncertainty estimation. According to Saltelli et al. (2004) sensitivity analysis is the study of how uncertainty in the output of a model can be apportioned to different sources of uncertainty in the model input. Sensitiv-
- ity analysis is recognized as a helpful parameter space screening tool to identify key parameters controlling the performances. It can help in reducing problem dimensionality with factor fixing (FF) for non-influential parameters, and factor prioritization (FP) for those controlling the most model output uncertainty (Saltelli et al., 2000). Besides



the selection of the appropriate method for analyzing parameter sensitivity depends strongly on the goal of the sensitivity analysis (Saltelli et al., 2006). Of particular interest is the analysis of the dependence of the model output variance to simultaneously modified parameters; this can be achieved with methods based on variance decompo-

sition (Sobol, 1993; Efron and Stein, 1981). The application of three sensitivity analysis methods including Sobol's method by Massmann and Holzmann (2012) shows that the two most important parameters of their conceptual continuous rainfall-runoff model are correctly identified as being sensitive by all methods.

1.3 Variance based methods and temporal sensitivity analysis

¹⁰ Variance based methods result in reliable estimates of sensitivities even for non-linear and non-monotonic models, as was often demonstrated using examples where analytical solution can be calculated (Saltelli and Bolado, 1998). The price to be paid in order to relax all assumptions on model behavior is that the required number of model runs is relatively high (> 1000) for most approaches. Some variants of this method, in ¹⁵ terms of partial variances calculation, are Sobol's method (Sobol, 1990, 2001) and the extended Fourier amplitude sensitivity test ((E)FAST) (Cukier et al., 1973; Fang et al., 2003; Reusser et al., 2011; Saltelli and Bolado, 1998).

Variance-based sensitivity analysis methods aim to quantify the amount of variance that each parameter contributes to the unconditional variance of the model output.

- ²⁰ These amounts are characterized by first order or interactions effects expressed as sensitivity indices (S_i 's). Despite its high computational demands that contributions (Saltelli, 2002) try to make more effective, the powerful Sobol' SA technique for example has recently become more popular in environmental modeling (Pappenberger et al., 2007, 2008; Van Werkhoven et al., 2008; Jing, 2011; Li et al., 2012).
- Tang et al. (2006) compared state of the art in sensitivity analysis including Sobol's method and found it to be the most effective in estimating first-order parametric sensitivities and overall influence including interaction effects. Tang et al. (2007) make a stepwise analysis of a conceptual grid base distributed rainfall runoff model (HL-RDHM).



Their sensitivity analysis reveals the impact of rainfall distribution on spatial sensitivities and input variables mostly controlling HL-RDHM's behavior. The use of Sobol' indices for sensitivity analysis purposes is investigated by (Nossent et al., 2011) in the case of a SWAT model. They conclude that in general the Sobol' sensitivity analysis can be successfully applied for factor fixing and factor prioritization with respect to the input parameters of a SWAT model, even with a limited number of model evaluations. The analysis also supports the identification of model processes, parameter values and parameter interaction effects. Some of the recent studies applying SA to rainfall runoff, flood inundation, and water quality models are listed by Reusser et al. (2011), 8 out of

- the 18 studies use variance based methods. In 7 studies, on the order of 10 000 model runs were computed to calculate sensitivities, which is impossible for computationally expensive models. As highlighted by Reusser et al. (2011), analyzing temporal dynamics of parameter sensitivity (TEPADS) of model output variables, such as discharge, we can quantify which model components dominate the simulation response. Their anal-
- ¹⁵ ysis reveals that temporal dynamics of model parameter sensitivity can be a powerful tool for hydrological model analysis, especially to identify parameter interaction as well as the dominant hydrological response modes. Reusser and Zehe (2011) with TEPADS and TIGER (time series of grouped error) methods investigate parameter uncertainty for periods of poor model efficiencies. With modeling and temporal sensitivity analysis
 ²⁰ used as learning tools, WaSIM-ETH complex grid based model hypotheses are shown
- to be insufficient to describe Weisseritz headwater catchment behaviour and future developments seem required.

1.4 Scope of the paper

The core idea of this paper is to approach hydrologic model sensitivity with temporal sensitivity analysis, here in the case of quick and strong catchment flash flood responses. The originality lies in TEPADS analysis calculated from variance based decomposition that may reveal sensitivity peaks and so flow dynamics at key instants. This kind of analysis is new for hydrologic models especially for event physically based



distributed models. Using TEPADS as a diagnostic tool joins the idea of dynamic identifiability introduced by Wagener et al. (2003). But these two methods serve a different purpose since it is a necessary but not sufficient condition that parameters must be sensitive to be identifiable whereas sensitive parameters may not be identifiable.

- In this contribution, a temporal sensitivity analysis of the physically based spatially distributed MARINE model dedicated to flash floods is carried out. Based on the understanding of Mediterranean catchments hydrological processes the hydrological rainfallrunoff model MARINE (Modélisation et Anticipation du ruissellement et des inondations pour des évèNements Extrêmes) aims at (i) exploiting the potential of distributed
- ¹⁰ models (ii) using physically meaningful parameters while (iii) maintaining a simple and parsimonious parameterization Roux et al. (2011). Given a validated model structure for flash floods in the French Mediterranean region, the question of sensitivity is approached in a probabilistic framework. One parameter set for each of the 6 catchments is tested on validations events for which the analysis of TEPADS is performed. The
- procedure is implemented for contrasted hydrometeorological events in the Cévennes and the Pyrenean region (France) with the view to bring understanding in model behavior for contrasted catchments and flash flood events on steep terrains and complex geo-pedological formations.

The paper is organized as follows. Section 2 describes variance decomposition ²⁰ method and sensitivity indices calculation. MARINE model and the 6 Mediterranean catchments of interest are presented in Sect. 3. Catchment parameter sets and their efficiencies on 10 validation events are calculated and temporal sensitivity analysis hypotheses are tested in Sect. 4. Then TEPADS on these validation events are examined in Sect. 5. After a conclusion on the results, processes, variables and parameters that

require further description or observations are emphasized and the possibility of applying this method to improve the understanding of the major processes involved in flood events is discussed.



2 Background on model analysis with variance decomposition methods

Thoughtful descriptions of sensitivity analysis methods can be found in Saltelli et al. (2000). Variance-based methods are based on a decomposition of the model output variance.

Let $\Omega_k \in \mathfrak{R}^k$ denote the set of all possible values that the factors can assume. Let $X \in \Omega_k$ be a possible value of the *k* model inputs. We denote by $Y = g(\underline{X}) = g(X_1, \dots, \overline{X}_k)$ the relationship that links the model inputs to the model output. The input factors \underline{X} have a domain of validity linked to the uncertainty about their precise value.

Assuming that g is a square integrable function over $\Omega_k = \{X \mid 0 \le X_i \le 1; i = 1, ..., k\}$, it can be decomposed using an expansion with summand $g_{ip}(X_1, ..., X_p)$ of increasing dimensionality p < k.

$$Y = g(\underline{X}) = g_0 + \sum_{i=1}^{\kappa} g_i(X_i) + \sum_{i=1}^{\kappa} \sum_{i>j} g_{ij}(X_i, X_j) + \dots + g_{1,2,\dots,k}(X_1, X_2, \dots, X_k).$$
(1)

Sobol (1993) proved that this HDMR decomposition (High Dimension Model Representation) was unique if each term in the expansion has zero mean, then all the terms of the decomposition are orthogonal in pairs:

$$\int_{\Omega_k} g_{i_1,\ldots,i_p} g_{i_1,\ldots,i_s} \mathrm{d} \underline{X} = 0.$$

The total unconditional variance of model output can be defined as:

$$V(Y) = \int_{\Omega_k} g^2(\underline{X}) \mathrm{d}\underline{X} - f_0^2,$$

5

15



(2)

(3)

with $f_0^2 = \int_{\Omega_k} g_0^2 d\underline{X}$. The partial variances which are the components of the total variance decomposition are computed from each term in Eq. (1) as:

$$V_{i,\dots,p} = \int_{0}^{1} \int_{0}^{1} g_{i_1,\dots,i_p}^2 (X_{i_1},\dots,X_{i_p}) dX_{i_1},\dots,dX_{i_p}.$$
 (4)

The relation Eq. (3) expressed with Eqs. (1) and (4) leads to the so-called functional ⁵ ANOVA decomposition:

$$V(Y) = \sum_{i} V_{i} + \sum_{i} \sum_{j>i} V_{ij} + \ldots + V_{1,2,\ldots,k},$$
(5)

where V(Y) is the total variance, V_i is the variance caused by parameter X_i (first order variance), V_{ij} is the covariance caused by X_i and X_j (second order variance), and higher order terms show the variance contribution from multiple parameters. Two factors X_i and X_j are said to interact when their effect on Y cannot be expressed as the sum of their single effects V_i and V_j . Interactions may imply, for instance, that extreme values of the model output are uniquely associated with particular combinations of model inputs, in a way that is not described by first order effects S_i .

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From this relation (Eq. 5), sensitivity indices can be defined in order to assess the sensitivity of Y to \underline{X} when \underline{X} is uncertain. The first order effect representing the average output variance reduction that can be achieved when X_i is fixed is defined by:

$$S_{i} = \frac{V_{i}}{V} = \frac{V(Y) - E_{X_{i}} \left[V(Y|X_{i}) \right]}{V(Y)} = \frac{V_{X_{i}} \left[E(Y|X_{i}) \right]}{V(Y)}.$$
(6)

The partial variance V_i in Eq. (6) is given by the variance of the conditional expectation $V_i = V_{X_i} [E(Y|X_i)]$ and is also called the main effect of X_i on Y. The impact on the model output variance of the interactions between parameters X_i and X_i is given by



 $S_{ij} = V_{ij}/V$ and it can be generalized in interactions effects up to order k by replacing the index *i* by the corresponding set of input factors.

The estimation of partial variances could be very expensive with brute-force methods but a shortcut was proposed by Sobol' to reduce the calculation of the double loop integrals of Eq. (4). Efficient methods such as extended FAST from Saltelli (1999) or improved Sobol' from Saltelli (2002) where proposed in order to estimate both S_i and S_{Ti} for all inputs factors for a computational cost of N(k + 2). However, alternatives techniques were introduced recently allowing the estimation of $S'_i s$ and low interactions effects (up to order 3) for a computational cost independent from k (i.e. equal to N the sample size). They are based on Random Balance Designs and FAST (RBD-FAST from Tarantola et al., 2006; Mara, 2009) or on the advent of smooting/metamodelling techniques suitable for computationally demanding models building from the work of Sacks et al. (1989). This includes techniques based on non-parametric regression (Storlie and Helton, 2007), Gaussian process emulators (Oakley and O'Hagan, 2004),

¹⁵ Polynomial Chaos Expansion (Sudret, 2008; Crestaux et al., 2009).

The method used in this paper is the State Dependent Parameter method (Ratto et al., 2007a) which is based on recursive filtering and smoothing estimation. It is a very efficient method that do not require any specific rule for sampling inputs, and provides fastly accurate and unbiaised results for both sensitive and unsensi-²⁰ tive inputs according to Gatelli et al. (2009). Ratto et al. (2007a) show that even for a large number of parameters the SDP method allow a good estimation of variance based sensitivity indices with a mild computational cost for models with up to 20 input factors. In the following we use the routine (SS-ANOVA) available at (http://sensitivity-analysis.jrc.ec.europa.eu/software/index.htm). The recursive filtering

²⁵ and smoothing procedure provides standard errors of the estimated state dependant parameters and hence the relative significance of estimated HDMR terms and sensitivity indices.



3 Model and site description

3.1 Marine flash flood model

The modeling approach is the distributed model MARINE for flash flood forecasting (Roux et al., 2011) with subsurface transfer module. The predominant factor determining the formation of runoff is represented by the topography: slope and downhill 5 directions. MARINE runs on a fixed time step and is structured into three main modules (Fig. 1). The first module allows separating the precipitation into surface runoff and infiltration using the Green and Ampt model; the second module represents subsurface downhill flow with an approximation of the Darcy's law and the standard TOPMODEL transmissivity profile (Beven and Kirkby, 1979) and the third one the overland flow (over 10 hillslopes and in the drainage network): the transfer function component allows routing the rainfall excess to the catchment outlet using the kinematic wave approximation. Both infiltration excess and saturation excess are represented within MARINE. The spatial discretization of the catchment is performed using the Digital Elevation Model grid resolution, a regular grid of squared cells. Evapotranspiration is not represented 15 since the model purpose was to simulate individual flood events during which such process is negligible. Cell's soil moisture deficit is initialized from a continuous distributed water balance model output briefly described later. For a complete description of the

20 3.2 Study zone

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The proximity of the Mediterranean Sea and the steep surrounding orography can promote low level flow lifting in an unstable atmosphere, as for the Alps and Pyrenees (Davolio et al., 2009; Tarolli et al., 2012). The highest flooding risk is in autumn with wet soils and maximum rainfall rates. Summers are hot and dry; however summer storms also represent a non negligible flooding risk. The density of both hydrometeorological radar and hourly raingauge coverage offers interesting possibilities for flood

MARINE model the reader can refer to Roux et al. (2011).



triggering rainfall monitoring and quantitative precipitation estimation (Fig. 2), stream gauges and rating curves quality is good also for the catchments of interest. Thus the French Mediterranean region rather frequently affected by intense rainfalls represents an interesting area for flash flood study in a regional manner (Garambois et al.,

5 2012b). Indeed, the authors show how contrasted Pyrenean catchments' properties, rainfall distributions and hydrological responses characteristics are Garambois et al. (2012a).

From the Pyrenean foothills and the Corbières Mountains in the south to the Cévennes foothills and the Ardèche region, 6 flood prone catchments with areas rang-¹⁰ ing from 144 to 619 km² and contrasted physiographic properties are selected (Table 1). They present a highly marked topography with narrow valleys and steep hillslopes (Fig. 2). A DEM data file of the study site with a grid scale of 25 m was available from the National Geographic Institute (BD TOPO^{® ©}IGN – Paris – 2008. [©](SCHAPI)). The mean elevation ratios of the whole river basins are above 0.025 m m⁻¹.

¹⁵ Salz and Verdouble catchments area mainly develop on sedimentary formations contrarily to the other catchments which substrates develop on metamorphic and plutonic terrains. Soil thicknesses and textures were available from the BDSol-LR (Robbez-Masson et al., 2002) (IGCS-BDSol-LR-version no. 2006, INRA-Montpellier SupAgro) (Table 1). Soil saturated hydraulic conductivities, saturated water contents and soil suctions are determined with (Rawls and Brakensiek, 1985) pedotransfer functions as proposed by Manus et al. (2009). Braud et al. (2010) and Roux et al. (2011) highlight the importance of soil thickness and texture on hydrological process and catchment flood response. It has recently been shown with a comparative hydrologic study, that in Austria flood response is significantly controlled by geology (Gaál et al., 2012).

For the Gardon, Beaume and Ardèche catchments vegetation is dense and mainly composed of chestnut trees, pasture, holm oaks, conifers, waste land and garrigue. Chestnut trees are located in the upstream area and on the South-facing slopes (sunny sides or adret) while forested garrigues and holm oaks are located in the downstream



area and on the North-facing slopes (shady sides or ubac). Tech catchment's vegetation is rather dense also with broad leaved and coniferous forests. Mainly Mediterranean forest, garrigue, holm oaks and vineyards are encountered on the Salz and Verdouble catchments. A vegetation and land-use map (Corine Land Cover provided

⁵ by the Service de l'Observation et des Statistiques (SOeS) of the French Ministry of Environment, www.ifen.fr) is used to derive distributed surface roughness.

4 Preliminary analysis

Initialization is an important step in the case of flash flood event-based models running on a few days time window. Soils saturation at the beginning of each event is esti mated with SAFRAN-ISBA-MODCOU (SIM), a continuous hydrometeorological model (Habets et al., 2008). This continuous water balance model is run over the whole country on 8 km by 8 km cells and outputs such as soil moisture with at least daily time step are available. This systematically available spatial-temporal model outputs for catchment initial soil moisture accountancy is chosen. We keep in mind that soil moisture is related to soil parameters in defining catchments soil's infiltrability and storage capacity. But an accurate estimation of soil moisture at the catchment scale is still difficult even if combining spatialized superficial remotely sensed data and numerous in situ point measurements lead to promising results (Brocca et al., 2012; Albergel et al., 2012). An

estimation of uncertainty for soil moisture model outputs would be welcome but seems to require a very good knowledge of soil properties and structure which is another problematic task for extended catchment sets.

4.1 Calibrated parameter sets

In order to avoid a model over-parameterization, spatial patterns of several parameters are derived from soil surveys and a unique correction coefficient is then applied to each parameter map. This approach has been chosen for three parameters, namely



the distributed saturated hydraulic conductivity K, the lateral transmissivity T_0 and soil thicknesses Z. The calibration procedure consists in estimating: three coefficients of correction for spatialized data one for the saturated hydraulic conductivities, named C_K , another one C_{KSS} for the lateral subsurface flow transmissivity (T_0), and the last one

⁵ for the soil thicknesses, named C_Z , the Strickler roughness of the main channel K_{D1} and the Strickler roughness of the overbank of the drainage network K_{D2} (Roux et al., 2011; Garambois et al., 2012a). Concerning the transmissivity T_0 the spatial variability is taken from the hydraulic conductivity map. Catchment parameter sets for the whole catchment set in calibration on several flood events used in this paper are given in 10 Table 2.

4.2 Selected validation flash flood events for sensitivity analysis

Monitoring flash floods remains a hard exercise (Borga et al., 2008) since conventional measurement networks of rain and river discharges are not able to sample effectively because of scales problems. Hydrometric data are provided by the SCHAPI (Service

- ¹⁵ Central Hydrométéorologique d'Appui à la Prévision des Inondations French central flood forecast center) and the SPC Grand Delta located in Nîmes and the SPC Mediterranée Ouest located in Carcassone (Service de Prévision des Crues – Regional flood forecast center). Radar rainfall measurements (Météo France – Nîmes radar) combined with rain gauge data are available at five minutes time steps and 1 km by 1 km spatial
- ²⁰ resolution since 2002 for the whole French Mediterranean region and since 1994 on the Cévennes region. Few floods of several years return period have been experienced in the 6 catchments of interest catchment since 2002 (resp. 1994). In this paper an interesting set of 10 validation events is used. This constitutes a huge validation dataset given the scarcity of data about flash flood events in general.
- Validation events hydrographs with distinct shapes represent contrasted hydrological responses to different resonances between rainfall spatio-temporal distribution and catchment physiographic properties (cf. Table 3 and Figs. 5–8):



- single peak medium events (15 March 2011 at Pas-du-Loup, 20 December 2000 at Cassaignes, 28 September 2000 and 18 October 2006 with a slow rising limb at Anduze),
- single peak medium events with slow rising and/or declining limb (16 November 2006 at Rosières, 20 October 2008 and 31 October 2008 at Vogüé),
- multi peaks events (15 March 2011 at Tautavel, 3 November 2011 at Vogüé),
- and a 50 yr return period extreme event (8 September 2002 at Anduze).

In addition to classical normalized least squares criterion, L_{NP} (Table 3) considers features characterizing the flood peak (discharge value and time to peak) (Roux et al.,

¹⁰ 2011). Performance decrease is slight for the whole catchment set from calibration to validation with mean Nash values of 0.86 and 0.81 respectively. Validation efficiencies are high with $L_{\rm NP}$ efficiencies above 0.73 and of 0.83 on average. Mean peak flow discharge and timing relative errors are inferior to 0.17 and 0.13.

Most validation events present observed peak flow discharges ranging from 1.13 to 2.69 m³ s⁻¹ km⁻² (Fig. 3). The extreme event of September 2002 at Anduze has an estimated peak discharge of $6.08 \text{ m}^3 \text{ s}^{-1} \text{ km}^{-2}$. Differences between simulated and observed discharges are satisfactory with a R^2 of 0.87 with respect to the first bisector, so the model presents no significant bias for these catchments-floods.

4.3 Evaluation of temporal sensitivity analysis method

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²⁰ Using the variance-based sensitivity analysis method described in Sect. 2, a region of the parameter space around calibration point is explored and sensitivity indices are estimated to analyze the relative importance of MARINE model inputs for validation events. For each catchment we use the calibrated parameter sets of Sect. 4.1 for validation events and temporal parameter sensitivity ($S'_i s$) calculation. The tested input factors are the 5 calibrated ones: 3 coefficients of correction C_K , C_{KSS} , C_Z , the Strickler



roughness of the overbank of the drainage network $K_{\rm D2}$, the main channel roughness coefficient $K_{\rm D1}$.

We use a 1024 parameter sets quasi random Monte Carlo sample. The $S'_i s$ are calculated for MARINE discharge at each time step in a $\pm \alpha$ % interval around the nominal parameter value with the method described in section above. Ideally for each parameter the sampling range around nominal parameter value could be defined with information on parameter posterior distribution function with the strength of methods such as Markov chain Monte Carlo (Smith and Marshall, 2008; Vrugt et al., 2009; Kuczera et al., 2010). These methods are yet too computationally demanding for our extended study and the choice of α is tested here.

From 5 to 15 % around the nominal parameter values, the choice of α reveals to have a rather limited influence on temporal dynamics of parameter sensitivity (TEPADS) and their values (Table 5). S_{i1} the first order effect measures the relative importance of an individual input variable X_i , in driving the uncertainty. Parameter ranking remains the same with a total standard error lower than 0.03. Low error and high first order metamodel R^2 attest the good convergence of the SDR algorithm on the 1024 sample size. S_i values quantifying model output sensitivity to each parameter are quite similar with relative variation lower than 5% for the three alpha values.

Figure 4 shows a limited influence of sampling range on temporal sensitivity index to C_Z . S_{j_1} - C_Z estimation differences are lower than 15% after model initialization and during hydrograph late recession. It is yet the most sensitive parameter on average for this event especially when most hydrographs are peaking (t = 20 h to 27 h), and estimation differences are lower than 5%.

Small to large parameter sampling ranges show a limited influence on sensitivities calculations with similar event first order effects for each parameter. Observations made after testing S_i estimation lead us to select a ±10% sampling interval around the nominal parameter values for TEPADS calculation with small errors in the following.

Temporal sensitivity presents the same pattern for the different sampling ranges with low standard errors. Rapid oscillations (Fig. 4, bottom panel), can be apportioned to



strong temporal gradients in non zero simulated discharges and different trends between the 1024 hydrographs. Let us recall that MARINE model runs on a 200 m mesh and few seconds time step verifying CFL (Courant Friedrichs Lewy) conditions, it is then not surprising to obtain such temporal variations in sensitivities with 5 min time resolution radar rainfalls inputs and observed discharges. Important oscillations can also be remarked on TEPADS calculated for TOPMODEL and WaSIM-ETH models (Reusser et al., 2011).

5 Temporal analysis of flash flood model behavior

5.1 Event averaged first order effects

5

- ¹⁰ This measure indicates the relative importance of an individual input variable X_i , in driving the uncertainty. It is good to notice that sensitivities are not calculated with a cost function but with simulated discharge at the outlet. Different phases in catchment saturation and runoff generation are aggregated into discharge and their temporal variation is reported in terms of the partial variance explained by an input factor at this time step.
- ¹⁵ For example, a value of around 0.8 at 23 h when most hydrographs (1024 parameter sets sample) are peaking indicates that 80 % of the observed variation between model runs can be explained by this parameter (Fig. 4).

First order sensitivity indices and the related standard error and first order metamodel R^2 constitute basic outputs of SDR method; in a first time they are averaged on each validation event for the catchments of interest (Table 6). R^2 of first order metamodel are above 0.93 and indicate a good convergence of the method. Event averaged standard error on $S'_i s$ is 0.03 and the following parameter ranking $C_Z > C_{KSS} > K_{D1} > K_{D2} > C_K$ is obtained for the whole catchment flood dataset. According to results displayed in (Table 6), soil profile storage capacity controlled by parameter C_Z is the most important input factor for 8 of the 10 events considered. Soil storage capacity has a large impact on soil saturation dynamics and so runoff generation mechanisms. Relation between



soil profile storage capacity and flood event magnitude seems non-monotonous according to parameters sensitivities (Tables 3 and 6).

For the other parameters, relation is monotonous. The relative importance of catchment infiltrability (C_K) and friction in the drainage network (i.e. K_{D1} and K_{D2} , channel and overbank correction coefficients) is increasing with the magnitude of the event. On the contrary, given the reduction of the proportion subsurface flow represents, the influence of subsurface flow velocity (i.e. C_{KSS}) is decreasing with the magnitude of the event (Table 3). C_{KSS} is particularly sensitive for Ardèche and Salz catchments. Let us remark that the sum of first order effects $\sum_i S_i$ is lower than one with low standard errors (Table 6) and the equality would mean that the model is additive (Saltelli et al., 2000).

5.2 Temporal evolution of first order effects

In order to analyze the temporal dynamics of model input factors influence on the simulated discharge for the 10 flood events on 6 catchments, the explored variability of model response (top) and the temporally variable sensitivity indices (bottom) are represented on Fig. 5 to Fig. 8. Whatever the rainfall patterns and volume, for some aspects of the model response, catchments behaviors characterized by the first order sensitivity indices are similar. First, before rainfall starts, C_Z , C_{KSS} and K_{D1} , i.e. soil depths, lateral subsurface flow, main channel roughness explain most of the variability

²⁰ because the initial soil water content (above 48 % Table 3) is activating subsurface flow and exfiltration in the drainage network. Only the main channel represented by K_{D1} is concerned by these small amounts of water at the outlet (a few m³ s⁻¹).

Then we can distinguish 16 November 2006 event at Rosières (Fig. 8, right panel, the smallest one in terms of specific discharge, from the 9 others obviously activating

²⁵ all model's flow components. This event is underestimated by MARINE and is characterized by an important sensitivity to C_Z , especially at peak time and early recession (11 to 22 h) of the hydrograph. C_K and K_{D1} plays a small role during the rising limb. Moreover, while the influence of the parameter driving infiltrability (i.e. C_K) is low,



subsurface flow represented by parameter C_{KSS} plays an important role (10% of total variance). Only "minor flow components" are activated for that catchment and event, i.e. moderate solicitation of flow components without floodplain invasion.

- At the beginning of rainfalls, and during heavy rainfalls a similar general sensitivity ⁵ pattern can be found for the 9 other events (Figs. 5 to 8), most flow components are activated: infiltration, lateral subsurface flow, hillslope runoff, main channel and floodplain flow. The temporal evolution of parameter's influence involves in the following order the different processes: infiltrability, transfer and limitation by maximum soil storage capacity. In fact, C_K determining infiltration capacity is sensitive for significant rainfall intensity variations – Fig. 5, at 15 b, Fig. 6 at 47 b and 57 b, Fig. 7 (right papel) at 8b, Fig. 8 (loft
- ¹⁰ variations Fig. 5, at 15 h, Fig. 6 at 47 h and 57 h, Fig. 7 (right panel) at 8h, Fig. 8 (left panel) at 25 h. Before hydrographs rising limb K_{D1} , the main channel friction coefficient, is driving the uncertainty and then soil depth coefficient C_Z is sensitive, which defines cells total storage capacity. This highlights sensitivity to the soil volume which influences saturation dynamics and so on to water volumes partitioning among the catchment. Let us remark that C_Z explains more than 80 % of model outputs variance
 - when most hydrographs are peaking.

Besides the presence of some peaks of C_{KSS} influence during simulations – Fig. 5 around 10, 60 and 160 h, 210 and 240 h; Fig. 6 around 15, 55 and 87 h; Fig. 7 (left panel) around 50 h, (right panel) around 52 h; Fig. 8 (left panel) around 15 h – can be

- ²⁰ explained by a significant contribution of subsurface flow. Indeed $C_{\rm KSS}$ is the adjustment parameter of soil lateral conductivity for subsurface flow. It can have an impact on simulated discharge by modifying the repartition of soil water content and so infiltration dynamics. During recession $C_{\rm KSS}$ sensitivity generally increases which can show the role of subsurface in recession dynamics according to the model.
- Let us remark high C_{KSS} sensitivities explaining above 80 % of model output variance for slow recessions in the case of 31 October 2008 and 3 November 2011 floods on the Ardèche at Vogüé for instance (Fig. 5), for slow hydrograph rising limb in the case of 15 March 2011 flood on the Verdouble at Tautavel (Fig. 7, right panel, at 51 h) or



20 December 2000 flood on the Salz at Cassaignes (Fig. 8, left panel, between 10 and 20 h).

 K_{D2} the overbank roughness coefficient is sensitive during late rising and early falling limbs, when saturation is high and huge amount of water is transferred to the outlet by overbank flow – Fig. 5 around 20, 90 and between 170 and 225 h; Fig. 6 around 35, 67 and 115 h; Fig. 7 (left panel) around 90 h, (right panel) around 40, 65 and 100 h; Fig. 8 (left panel) around 45 h.

Finally it can be remarked that in the case of 8 September 2002 extreme event at Anduze a complex catchment behavior reflected by quickly variable and marked sen-¹⁰ sitivities is caused by an extreme storm in the very lower part of the catchment and so short response delays (more than 400 mm cumulated rainfall on half of the catchment with maxima greater than 700 mm located close to the outlet). On the contrary, 18 October 2006 and 28 September 2000 generating storms hit with less violence the medium or upper part of Gardon d'Anduze catchment. For these longer rain events the 15 temporal sensitivities vary more slowly. Moreover for sensitivity peaks of C_Z and then K_{D1} , K_{D2} , (Fig. 6 between 20 and 30 h, and between 95 and 122 h) corresponding to rainfall peaks responses, C_Z sensitivity stays above the other during the flood. This can be attributable to catchment spatio-temporal dampening effect: when storm hits catchment headwaters, a larger soil storage volume is involved in flood generation.

20 5.3 Analysis of temporal interaction effects

Using variance-based sensitivity analysis methods, an essential aspect is that the estimated $S'_i s$ have interesting normalization properties. Indeed from Eq. (5) normalized by V(Y) and with Eq. (6) the sum of sensitivity measures nicely scaled between 0 and 1 can be written as:

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$$1 = \sum_{i} S_{i} + \sum_{i} \sum_{j} S_{ij} + \ldots + S_{1,2,\ldots,k}.$$



(7)

Given that the sum of all sensitivity indices (up to order k, with k the number of input factors) sum up to one, it is possible to apprehend the importance of interactions using first order sensitivity indices. By definition the difference $1 - \sum S_i$ gives the sum of all

higher order sensitivities and is therefore an indicator of the presence of interactions according to Ratto et al. (2007a) and Saltelli et al. (2000). Indeed let us remark that by definitions given in Eqs. (4) and (5) each term of the variance decomposition is positive. The difference $1 - \sum_{i} S_{i}$ is presented for the validation events from Figs. 9 to 12. During hydrograph peaks less than 10% of model outputs variance is explained by correlations between the five sensitive MARINE model parameters. Moreover the highest correlation is at 0.6 after initialization – Fig. 9 at 1, 53, 141 h; Fig. 10 at 5, 42, 53, 77 h; Fig. 11 (left panel) at 8 h and Fig. 11 (right panel) at 1 h – or after first rainfall intensity peak – Fig. 10 at 12, 53, 84 h; Fig. 11 (right panel) at 12 h. These interactions reaching 60% of model variance, when the model starts running or when it is first raining might be due to water partition among the model components which can lead to parameter interactions. This indicates that water partitioning representation can be

improved and asks questions about initialization data.

This low interacting behavior of MARINE model in general, with some differences in function of the catchment-flood type and magnitude, can be interpreted as an indicator of a correct process parameterization especially during rising limbs. In summary, interactions play an important role at the initial stage of the rising limb (with peaks corresponding to the beginning of precipitation) and the model inputs influence tend to be

more orthogonal (few interactions) when the flow is significant.

6 Conclusions

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The aim of this contribution was to analyze hydrologic model sensitivity in the case of flash floods with a new approach in hydrological modeling, namely model outputs variance decomposition for temporal patterns of parameter sensitivity. Given a simple



and parsimonious parameterization of MARINE model structure, TEPADS are calculated on a significant number of contrasted validation flood events in the Cévennes and the Pyrenean region (France). Our results show the huge impact of soil depth on model sensitivity which is consistent with recent cévenols case studies (Braud et al.,

- ⁵ 2010; Roux et al., 2011). First order sensitivity to soil depth maps multiplicative constant C_Z explains more than 80% of model outputs variance when most hydrographs are peaking. Moreover, first order sensitivity to subsurface lateral transmisivity constant C_{KSS} is responsible for 80% of model output variance for slow recessions or multiple peak hydrograph rises. Using models as learning tools with TEPADS give information
- on the different processes giving rise to the flood hydrograph in the following order: infiltrability, transfer and limitation by maximum soil storage capacity. Concerning the transfer function, successive sensitivity to drainage network's main channel and flood plain roughness likely depends on event dynamics and amplitude.

The small part of variance explained by correlations between MARINE parameters probably stems from model parsimony. First hours of simulations or rainfall intensity peaks are yet determined as the instant when correlation occur pointing soil water content initialization or water repartition problems. Reduction of model uncertainty can be achieved by improving water partition between lateral flow components and other mechanisms such as exfiltration in the drainage network and its representation itself.

Measurements at different scales are yet necessary to better constrain these flow dynamics. Moreover in situ soil moisture measurements and smaller scale water balance modeling (Vincendon et al., 2010) would strengthen soil saturation dynamics representation and increase simulations realism for catchments flash flood responses.

A general pattern of model response is found for Mediterranean flash flood events ²⁵ but some peaks of sensitivity to infiltrability, subsurface during recession or friction coefficients for example can indicate particular process dynamics attributable to singular rainfall distributions. Combining variance based sensitivity analysis in a regionalization framework along with spatial and temporal sensitivities derived from variational methods, following the intent of Castaings et al. (2009), could bring understanding in spatial



temporal aggregation of flash flood generating processes and data/modeling uncertainties. The resonance between rainfall spatial and temporal distribution and catchment properties, in other words catchment temporal dampening effect could be accessible that way.

- Eventually it can be concluded that the variance based temporal sensitivity analysis method presented here can be successfully applied to distributed hydrological models allowing to:
 - analyze the model processes temporal dynamics for each flood event,
 - derive patterns of model responses according to the different characteristics of each event,
 - emphasize model structure or parameterization problems when an important part of the model variance is explained by correlations.

This method can be implemented with a very reasonable computational cost and studies for other litho-pedological conditions, landscapes and climatic regions could bring insight for example to help the number of possible hydrological process representations to converge.

Acknowledgements. Financial support and data supply by the SCHAPI are gratefully acknowledged by the authors.

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Table 1. Catchments physiographic properties, elevation ratio is height difference divided by longest flow path length, K mean is the mean soil saturated hydraulic conductivity.

Catchments	Area (km ²)	Height diff. (m)	Max. flow length (km)	Elevation ratio (mm ⁻¹)	Hsol min	Hsol max	Hsol mean	Hsol std	Soil volume (m ³)	<i>K</i> mean (mmh ⁻¹)
Tech (Pas-du-Loup)	250	2730	34.5	0.079	0.00	0.69	0.16	0.13	5.3E+07	2.5
Verdouble (Tautavel)	299	915	37.0	0.025	0.08	0.63	0.33	0.16	1.0E+08	2.4
Salz (Cassaignes)	144	995	17.2	0.058	0.00	0.74	0.31	0.19	4.2E+07	3.9
Gardon (Anduze)	543	1065	45.1	0.024	0.08	0.64	0.28	0.12	1.5E+08	5.0
Beaume (Rosières)	212	1360	29.0	0.047	0.05	0.49	0.25	0.07	5.2E+04	8.7
Ardèche (Vogüé)	619	1380	52.5	0.026	0.05	0.50	0.28	0.08	1.7E+08	8.7

Catchments	Area (km ²)	Cz	C_K	C_{KSS}	K _{D1}	K _{D2}	Global Nash
Tech (Pas-du-Loup)	250	4.34	11.00	1515	4.83	3.24	0.90
Verdouble (Tautavel)	299	1.30	15.00	4486	5.00	3.99	0.88
Salz (Cassaignes)	144	0.95	20.00	5595	5.00	2.54	0.89
Gardon (Anduze)	543	4.60	10.30	4540	11.70	9.70	0.88
Beaume (Rosière)	212	5.30	7.40	3712	21.40	14.70	0.80
Ardèche (Vogüé)	619	3.40	2.10	4891	10.00	19.10	0.80

 Table 2. Catchment parameter sets and Nash efficiencies for multiple events calibration.

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Table 3. Validation events characteristics. Mean initial soil moisture is the spatially averaged daily SIM output over a catchment.

Catchments	Area (km ²)	Validation events	Mean initial soil moisture (%)	Specific peak flow (m ³ s ⁻¹ km ⁻²)	Cumulated rainfall (mm)
Tech (Pas-du-Loup)	250	15 Mar 2011	62	2.2	270
Verdouble (Tautavel)	299	15 Mar 2011	52	1.2	217
Salz (Cassaignes)	144	20 Dec 2000	48	1.5	141
Gardon (Anduze)	543	28 Sep 2000	56	1.4	203
		8 Sep 2002	58	6.7	284
		18 Oct 2006	62	2.6	237
Beaume (Rosières)	212	16 Nov 2006	56	1.1	111
Ardèche (Vogüé)	619	20 Oct 2008	48	1.6	195
		31 Oct 2008	59	1.6	211
		3 Nov 2011	50	1.5	370
Average	344		55	2.1	224

Table 4. Validation events and efficiencies in terms of ΔQ comparing simulated and observed peak discharge $Q_{\rm P}^{\rm s}$ and $Q_{\rm P}^{\rm o}$, and ΔT comparing simulated and observed peak time normalized by concentration time $T_{\rm C}^{\rm o}$ (determined by averaging Bransby formula).

Catchments	Area	Validation	$\Delta Q =$	ΔΤ	Nash	$L_{\rm NP} =$
	(km²)	events	$\frac{\left Q_{\rm p}^{\rm s}-Q_{\rm p}^{\rm o}\right }{Q_{\rm p}^{\rm o}}$	$= \frac{\left T_{p}^{s} - T_{p}^{o}\right }{T_{c}^{o}}$		$\frac{1}{3}$ (Nash + (1 – ΔQ)
				-		$+(1 - \Delta T))$
Tech (Pas-du-Loup)	250	15 Mar 2011	0.15	0.32	0.70	0.73
Verdouble (Tautavel)	299	15 Mar 2011	0.13	0.32	0.82	0.79
Salz (Cassaignes)	144	20 Dec 2000	0.18	0.32	0.76	0.75
Gardon	543	28 Sep 2000	0.03	0.02	0.95	0.97
(Anduze)		8 Sep 2002	0.12	0.00	0.97	0.95
		18 Oct 2006	0.03	0.15	0.60	0.80
Beaume (Rosières)	212	16 Nov 2006	0.32	0.10	0.64	0.75
Ardèche	619	20 Oct 2008	0.02	0.02	0.93	0.96
(Vogüé)		31 Oct 2008	0.13	0.04	0.87	0.89
		3 Nov 2011	0.23	0.40	0.85	0.73
Average			0.13	0.17	0.81	0.83

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Table 5. Gardon d'Anduze 08 Sep 2002 flash flood event, first order effects and standard error averaged in time, and first order metamodel R^2 for different sampling ranges around nominal parameter values.

α	<i>S</i> _{<i>i</i>1} _ <i>C</i> _{<i>Z</i>}	S_{i1} _ C_K	S_{i1} - C_{KSS}	S_{i1} _ K_{D1}	<i>S_{i1}_K</i> _{D2}	Sum (<i>S_{i1}</i>)	Sum (S _{i1_std_err})	<i>R</i> ² Sum (<i>S_{i1}</i>)
±5%	0.392	0.183	0.119	0.198	0.091	0.983	0.020	0.972
±10%	0.413	0.170	0.109	0.195	0.079	0.967	0.028	0.975
±15%	0.376	0.169	0.117	0.196	0.081	0.940	0.030	0.971

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tandar et. 1024	d error 4 param	and first neter sets	order me s are ana	etamode alysed fo	el <i>R</i> ² a or each	veraged event.	in time	aper Discus	Temporal for hydrolo characte P. A. Garar	sensitivity ogic model erization nbois et al.
<i>S</i> _{<i>i</i>1} _ <i>C</i> _{<i>Z</i>}	S_{i1} - C_K	S_{i1} - C_{KSS}	S_{i1} _ K_{D1}	<i>S</i> _{<i>i</i>1} _ <i>K</i> _{D2}	Sum (<i>S_{i1}</i>)	Sum (<i>S_{i1_stdev}</i>)	R ²	sion F		
0.49	0.02	0.26	0.16	0.02	0.948	0.046	0.94	ap	Title	Page
0.36	0.06	0.22	0.13	0.22	0.992	0.021	0.97	er	Abstract	Introduction
0.29	0.03	0.42	0.11	0.09	0.941	0.038	0.99	_	Conclusions	References
0.49 0.41 0.57 0.73	0.01 0.17 0.00 0.00	0.17 0.11 0.15 0.23	0.20 0.20 0.14 0.00	0.07 0.08 0.08 0.00	0.943 0.966 0.947 0.979	0.056 0.028 0.035 0.038	0.94 0.98 0.93 0.99	Discussion	Tables	Figures
0.47 0.33 0.51 0.43	0.15 0.04 0.04 0.05	0.23 0.49 0.27 0.24	0.07 0.04 0.05 0.11	0.07 0.07 0.13 0.08	0.993 0.967 0.997 0.920	0.016 0.011 0.019 0.034	0.99 0.99 0.98 0.92	n Paper	I◄ ◀ Back	Close
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Table 6. First order effects (-), standa for each event of the validation set. 10

Catchments

(Pas du Loup)

Verdouble

(Tautavel) Salz

(Anduze)

Beaume

(Vogüé)

Average

(Rosières) Ardèche

(Cassaignes) Gardon

Tech

Area

(km²)

250

299

144

543

212

619

Validation

15 Mar 2011

15 Mar 2011

20 Dec 2000

28 Sep 2000

8 Sep 2002

18 Oct 2006

16 Nov 2006

20 Oct 2008

31 Oct 2008

3 Nov 2011

0.49

0.36

0.29

0.73

events

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Table A1. Notations

Notation	Unit	Meaning
a _d	[m ²]	Drainage area
C_{K}	[-]	Correction coefficient of the saturated hydraulic conductivity
C_Z	[—]	Correction coefficient of the soil thickness
1	[m]	Cumulative infiltration
T ₀	$[m^2 s^{-1}]$	Local transmissivity of fully saturated soil
C_{KSS}	[-]	Correction coefficient of local transmissivity of fully saturated soil
m	[-]	Transmissivity decay parameter
Н	[m]	Water depth
i	[ms ⁻ ']	Infiltration rate
r	[m s ⁻ ']	Rainfall rate
Κ	[ms ⁻¹]	Saturated hydraulic conductivity
L _{NP}	[]	Performance criterion
Ko	$[m^{-1/3}s^{-1}]$	Strickler roughness coefficient of the overland
K _{D1}	[m ^{-1/3} s]	Strickler coefficient of the main channel (drainage network)
K_{D2}	[m ^{-1/3} s]	Strickler coefficient of the overbanks (drainage network)
Q	[m ³ s ⁻¹]	Discharge
$S_{ m f}$	[mm ⁻¹]	Friction slope
S_0	[mm ⁻¹]	Bed slope
t _p	[s]	Time to ponding
Ż	[m]	Soil thickness
$ heta_{i}$	[m ³ m ⁻³]	Initial water content of the soil
$\theta_{\rm s}$	[m ³ m ⁻³]	Saturated water content of the soil
Ψ	[m]	Soil suction at wetting front

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Fig. 1. MARINE model structure, parameters and variables. Green and Ampt infiltration equation: infiltration rate *i* (ms⁻¹), cumulative infiltration *I* (mm), saturated hydraulic conductivity *K* (ms⁻¹), soil suction at wetting front ψ (m), saturated and initial water contents are respectively θ_s and θ_i (m³m⁻³). Subsurface flow: local transmissivity of fully saturated soil T_0 (m²s⁻¹), saturated and local water contents are θ_s and θ (m³m⁻³), transmissivity decay parameter is m (–), local slope angle β (rad). Kinematic wave: water depth *h* (m), time *t* (s), space variable *x* (m), rainfall rate *r* (ms⁻¹), infiltration rate *i* (ms⁻¹), bed slope *S* (mm⁻¹), Manning roughness coefficient *n* (m^{-1/3}s).





Fig. 2. Left panel: France topography (source: SRTM image, NASA/JPL). Right panel (white contour): topography of the six catchments of interest (France), BD TOPO[®] IGN, (concentric circles) Hydrometeorological radars, (white dots) operational raingauges.





Fig. 3. Simulated peak discharge versus observed peak discharge for validation events with first bisector (blue line).





Fig. 4. Top panel: Gardon d'Anduze 8 September 2002 flash flood event and quintiles Q_{10} and Q_{90} of simulated discharges for $\alpha = 5$, 10 and 15 %. Bottom panel: first order effects for C_Z and three sampling ranges.





Fig. 5. Top panel: Ardèche at Vogüé (619 km²), 20 October 2008, 31 October 2008 and 3 November 2011 flash flood events and quintiles Q_{10} and Q_{90} of simulated discharge. Bottom panel: first order effects representing first order contribution, partial variances out of model output variance (–).





Fig. 6. Top panel: Gardon at Anduze (543 km²), 28 September 2000, 8 September 2002 and 18 October 2006 flash flood events and quintiles Q_{10} and Q_{90} of simulated discharge. Bottom panel: first order effects.

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Fig. 7. Top panels: 15 March 2011 flash flood event and Q_{10} and Q_{90} quintiles of simulated discharge on (left panel) the Tech at Pas-du-Loup (250 km²) and (right panel) the Verdouble at Tautavel (299 km²). Bottom panels: first order effects.





Fig. 8. Top panels: Q_{10} and Q_{90} quintiles of simulated discharge on (left panel) 20 December 2000 for the Salz at Cassaignes (144 km²), (right panel) 16 November 2006 for the Beaume at Rosières (212 km²). Bottom panels: first order effects.





Fig. 9. Top panel: Ardèche at Vogüé, 20 October 2008, 31 October 2008 and 3 November 2011 flash flood events and quintiles Q_{10} and Q_{90} of simulated discharge. Bottom panel: $1 - \sum_{i} S_{i}$.





Fig. 10. Top panel: Gardon at Anduze, 28 September 2000, 8 September 2002 and 18 October 2006 flash flood events and quintiles Q_{10} and Q_{90} of simulated discharge. Bottom panel: $1 - \sum_{i} S_{i}$.





Fig. 11. Top panels: 15 March 2011 flash flood event and Q_{10} and Q_{90} quintiles of simulated discharge on (left panel) the Tech at Pas-du-Loup and (right panel) the Verdouble at Tautavel. Bottom panels: $1 - \sum_{i} S_{i}$.





Fig. 12. Top panels: Q_{10} and Q_{90} quintiles of simulated discharge on (left panel) 20 December 2000 for the Salz at Cassaignes, (right panel) 16 November 2006 for the Beaume at Rosières. Bottom panels: $1 - \sum_{i} S_{i}$.

