



Annual flood sensitivities to ENSO at the global scale

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Annual flood sensitivities to El Niño Southern Oscillation at the global scale

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Abstract

Floods are amongst the most dangerous natural hazards in terms of economic damage. Whilst a growing number of studies have examined how river floods are influenced by climate change, the role of natural modes of interannual climate variability remains poorly understood. Here, we present the first global assessment of the influence of El Niño Southern Oscillation (ENSO) on river floods. The analysis was carried out by simulating daily gridded discharges using the WaterGAP model, and examining statistical relationships between these discharges and ENSO indices. We found that, over the period 1958–1999, ENSO exerted a significant influence on annual floods in river basins covering over a third of the world's land surface, and that its influence on floods has been much greater than its influence on average flows. We show that there are more areas in which annual floods intensify with La Niña and decline with El Niño than vice versa. However, we also found that in many regions the strength of the relationships between ENSO and annual floods have been non-stationary, with either strengthening or weakening trends during the study period. We discuss the implications of these findings for science and management. Given the strong relationships between ENSO and annual floods, we suggest that more research is needed to assess relationships between ENSO and flood impacts (e.g. loss of lives or economic damage). Moreover, we suggest that in those regions where useful relationships exist, this information could be combined with ongoing advances in ENSO prediction research, in order to provide year-to-year probabilistic flood risk forecasts.

1 Introduction

Floods are one of the most dangerous natural hazards in terms of economic damage, causing billions of dollars of damage each year (Munich Re, 2012), and global flood damages have risen steeply over the past half century (UNISDR, 2011). At the same time, floods are essential for many wetland ecosystems and agricultural practices

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(Costanza et al., 1997). Hence, improvements in our understanding of global scale flood processes are generally beneficial.

In recent decades, a large number of studies have examined instrumental discharge records to identify possible changes in flood frequency and/or magnitude due to climate change (e.g. Allamano et al., 2009; Barros et al., 2004; Bhutiyani et al., 2008; Camilloni and Barros, 2003; Conway et al., 2009; Cunderlik and Ouarda, 2009; Delgado et al., 2009; Di Baldassarre et al., 2010; Douglas et al., 2000; Hannaford and Marsh, 2008; Hirsch and Ryberg, 2012; Jiang et al., 2008; Lindström and Bergström, 2004; Lins and Slack, 1999; Marengo et al., 2012; McCabe and Wolock, 2002; Mudelsee et al., 2003; Petrow and Merz, 2009; Renard et al., 2008; Robson et al., 1998; Shiklomanov et al., 2007; Tu et al., 2005; Villarini et al., 2009; Villarini and Smith, 2010; Yiou et al., 2006). There is also a growing literature on possible changes in flood frequency and/or magnitude based on future hydrological projections at different scales, including the local to basin-scale (Asokan and Dutta, 2008; Bell et al., 2007; Cameron, 2006; Charlton et al., 2006; Dairaku et al., 2008; Das et al., 2012; Fujihara et al., 2008; Graham et al., 2007; Kay et al., 2009; Kitoh et al., 2011; Lauri et al., 2012; Leander et al., 2008; Nakaegawa and Vergara, 2010; Poussin et al., 2012; Prudhomme and Davies, 2009; Raff et al., 2009; Shabalova et al., 2003; Taye et al., 2011; Te Linde et al., 2010; Thodsen, 2007; Van Pelt et al., 2009; Ward et al., 2011, 2013); continental scale (Dankers and Feyen, 2008, 2009; Feyen et al., 2012; Lehner et al., 2006); and global scale (Hirabayashi et al., 2008, 2013; Milly et al., 2002).

Despite this broad research attention to the possible influences of climate change on floods, there has been little attention to the role of present-day interannual climate variability at the global scale. As a result, the influence of this aspect on flooding is poorly understood, despite its importance for development and adaptation planning (IPCC, 2012). In this paper, we provide the first global assessment of the influences of ENSO-driven climate variations on historical flood discharges. We choose ENSO because it is the most dominant interannual climate signal on Earth (McPhaden et al., 2006) apart from the annual cycle.

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Many past studies have assessed ENSO's impacts on average river flows at the local to basin scale (see, for example, Dettinger et al., 2000, and references therein), and a few have examined global scale relationships between ENSO and average river flows (Chiew and McMahon, 2002; Dettinger and Diaz, 2000; Dettinger et al., 2000; Labat, 2010). In contrast, only a few studies have examined relationships between ENSO and peak flows. Most of these studies have focused on the United States (e.g. Bell and Janowiak, 1995; Cayan and Webb, 1992; Cayan et al., 1999), although studies have also been carried out in northern Peru (Waylen and Caviedes, 1986), South Asia (Mirza, 2011), and the Mekong Basin (Räsänen and Kumm, 2013). To a large extent, the lack of observed daily discharge data in many regions has hampered the kinds of consistent global scale assessments that are needed. Ward et al. (2010) examined the relationship between ENSO and observed annual peak discharge for 622 gauging stations, but the geographical coverage of those stations was highly biased towards a few regions (particularly North America and Central Europe), and for many regions data were limited or lacking.

In this paper, we address this problem by simulating daily discharges, allowing for the first fully global assessment of ENSO-driven climate variability's influence on annual floods. Here, we define the annual flood as the peak daily discharge in a given year. We then discuss key implications of the results for water management and practice.

2 Methods

In brief, we modelled global daily discharges using a global hydrological model forced by daily meteorological re-analysis data. We then identified statistical relationships between annual floods and indices of ENSO. In the following paragraphs, our methods are described in detail.

2.1 Simulating daily discharge

We simulated global gridded daily discharge at a spatial resolution of $0.5^\circ \times 0.5^\circ$ using the WaterGAP model (Alcamo et al., 2003; Döll et al., 2003), forced by daily meteorological fields (precipitation, temperature, and global radiation) for 1958–2000 from the EU-WATCH project (Weedon et al., 2011).

WaterGAP consists of two main components: (1) a water-balance model to simulate characteristic macro-scale behaviours of the terrestrial water cycle in order to estimate water availability; and (2) a water-use model to estimate water withdrawals and consumptive water uses. In principle, WaterGAP can account for human influences on the terrestrial water cycle by its inclusion of flow regulation by large dams and reservoirs as well as water withdrawals. For model validation, we used simulations that included these human influences on river discharge. However, as the main focus of this study is climate induced variability of river floods, we based the present assessment of ENSO influences on naturalised-flow simulations, i.e. human interferences were excluded.

The climate data used to force WATERGAP in this study were obtained from the EU-WATCH project (Weedon et al., 2011). WATCH developed a global dataset of sub-daily meteorological forcing data for the period 1958–2000 at a horizontal resolution of $0.5^\circ \times 0.5^\circ$ (WATCH forcing data; WFD). The time-series were derived from the ERA-40 reanalysis product (Uppala et al., 2005) via sequential interpolation to a horizontal resolution of $0.5^\circ \times 0.5^\circ$, with elevation corrections and monthly-scale adjustments of daily values to reflect CRU (corrected-temperature, diurnal temperature range, cloud-cover) and GPCC (precipitation) monthly observations combined with new corrections for varying atmospheric aerosol-loading and separate precipitation gauge corrections for rainfall and snowfall derived from the ERA-40 reanalysis product. Full details of the forcing data can be found in Weedon et al. (2011). WATCH also developed time-series of the WFD for the period 1901–1957, but these were developed by reordering of the ERA-40 data for the later 1958–2000 period. Hence, the extremes in the pre-1958

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dataset do not correspond to the extremes in actual years, which is essential for our research. Therefore, in this paper, we used data only from 1958 onwards.

While discharges were simulated on the grid scale, analyses were carried out at basin scales, i.e. correlations with ENSO (and other calculations) were based on values at the cells farthest downstream in each basin. Accounting for the fact that river basin areas vary significantly, the world's largest 34 river basins (in terms of surface area) were also divided into sub-basins.

2.2 Calculating mean and maximum annual discharges

For each grid cell and hydrological year, we calculated the maximum annual discharge, or annual flood discharge (Q_{\max}), and the mean annual discharge (Q_{ann}) from the simulated daily discharge time series for 1958–1999. In most cases, we used the standard hydrological year (October–September), as also used in several other global assessments (Dettinger et al., 2000; Dettinger and Diaz, 2000; Ward et al., 2010). However, this is problematic for the allocation of the maximum annual discharge to a given hydrological year for those areas in which that maximum occurs around the boreal autumn (September–November). Therefore, for the most downstream cell of each drainage basin, we calculated the month in which the maximum annual discharge occurred (over the period 1958–2000 in the WaterGAP results). For those basins in which the maximum annual discharge in the most downstream cell occurred in September, October, or November, we defined the hydrological year as July to June. A map showing the hydrological year used for each basin can be found in Fig. A1.

2.3 Relationships between discharge and ENSO

The time-series of annual-flood discharges and mean-annual discharges were validated against observed discharge time-series (Sect. 3). We then examined the correlation between the natural logarithm of Q_{\max} ($\ln Q_{\max}$) (and $\ln Q_{\text{ann}}$) and the Southern Oscillation Index (SOI; <http://www.cru.uea.ac.uk/cru/data/soi/>), as well as

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their sensitivities (β_1) to variations in SOI, fitting (Bouwer et al., 2008):

$$\ln(Q_i) = \beta_0 + \beta_1 a_i + \varepsilon_i, \quad (1)$$

whereby Q_i is the simulated Q_{\max} (or Q_{ann}) in hydrological year i , a_i is the 3-monthly mean value of the SOI (OND, NDJ, DJF, JFM), β_0 and β_1 are regression-derived coefficients, and ε_i is an error term. From this, it follows that a unit change in SOI (a_i) is associated with an average increase of $100 \cdot (\exp(\beta_1) - 1)$ in Q (Q_{\max} or Q_{ann}); we refer to this as the “sensitivity”. In the analyses, the correlations and sensitivity were estimated between $\ln Q_{\max}$ (and $\ln Q_{\text{ann}}$) time-series and each of the three-month mean SOI values (OND, NDJ, DJF, JFM) separately. In some of the tables and figures, we show results for the three-month period with the highest correlation coefficient. To assess the robustness of the results when using other indices of ENSO, we also repeated the analyses with the negative of the Multivariate ENSO Index (<http://www.esrl.noaa.gov/psd/enso/mei/>), negative NINO3.4 index (<http://www.cpc.ncep.noaa.gov/data/indices/>), and negative Global Sea–Surface Temperature (SST) ENSO index (http://www.jisao.washington.edu/data_sets/globalstenso/), where negatives were used to accommodate the difference in sign between SOI and SST-based ENSO indices.

The correlations between Q_{\max} and SOI, and between Q_{ann} and SOI, were carried out using the natural logarithms (\ln) of Q_{\max} and Q_{ann} , because the log discharge data are normally distributed for basins covering around 90% of land surfaces (normality was assessed using Lilliefors test; $\alpha = 0.05$). For the vast majority of the other basins, the data were not highly skewed. Given that the dataset is not small ($n \text{ yr} = 42$), we primarily assessed the correlations using Pearson’s r , given its greater power over non-parametric equivalents. For verification, we also examined correlations between the original Q_{\max} and Q_{ann} data (not the natural logarithms) using the non-parametric Spearman’s rank test, and found the results to be similar. The normality of SOI data was assessed using the Lilliefors test ($p = 0.24$).

We also examined the percentage anomalies in median Q_{\max} between El Niño (and, separately, La Niña) years compared to the median Q_{\max} of all years. We used

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the classification of ENSO years from the Center for Ocean–Atmospheric Prediction Studies (<http://coaps.fsu.edu/jma.shtml>), as shown in Table 1. The non-logarithmic Q_{\max} time-series used in this composite analysis are only normally distributed in basins covering 50 % of land surfaces. Therefore, when assessing differences in Q_{\max} between El Niño (La Niña) and all years, we used the non-parametric Mann–Whitney U test to assess the statistical difference in median values.

3 Validation

The general ability of global hydrological and land surface models, including WaterGAP, to reproduce various spatial and temporal characteristics of 20th century river discharge, using the WFD forcing data, has been evaluated extensively by the EU-WATCH project. Those analyses evaluated model performance for long-term mean runoff (Haddeland et al., 2011), as well as high and low flow indices (Gudmundsson et al., 2011; Prudhomme et al., 2011). WaterGAP was found to acceptably reproduce most regional characteristics of large-scale hydrological extremes.

However, these validations did not specifically assess the model’s performance in simulating differences in peak discharges between different phases of ENSO. Thus, in this study we validated model findings against observed discharge time-series from the GRDC database, using only those stations with upstream areas greater than 10 000 km² for which daily data are available for every day of the hydrological year in at least 15 hydrological years between 1958 and 1999. This yielded a set of 721 observed discharge time-series. For model validation, we used WaterGAP simulations including human influence.

From Fig. 1, it is clear that there are large biases between modelled and observed Q_{\max} for many stations. At 33 % of the stations, the percentage difference between modelled and observed median Q_{\max} is less than 25 % (positive or negative), but large positive biases (> 50 %) were found for 15 % of stations, and large negative biases for 25 %.

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For this study, though, we are most interested in the relative change in discharge magnitudes from year to year, and between different phases of ENSO, rather than absolute discharge values. In Fig. 2, we show that the correlation between modelled and observed $\ln Q_{\max}$ is generally very good. For 80% of stations, the correlation coefficient (r) is greater than 0.6, and greater than 0.4 for 92% of stations. We also carried out correlation analyses using the non-parametric Spearman's rank correlation coefficient (not shown here), using the original Q_{\max} data (instead of the natural logarithms), and found very similar values. Moreover, the majority of the stations with low correlation are located in upstream areas, whereas the analyses presented in this paper are based on values at the most downstream cell in each basin. This gives confidence that the model simulates interannual fluctuations in Q_{\max} similar to those in the observed records.

Finally, we examined the agreement between the modelled and observed data in terms of the relative change in Q_{\max} between El Niño and non-El Niño years (Fig. 3a) and between La Niña and non-La Niña years (Fig. 3b). For 90% (92%) of the stations, both modelled and observed median Q_{\max} show either no significant difference between El Niño (La Niña) and non-El Niño (non-La Niña) years, or significant differences of the same sign (indicated by green symbols; "1"). For the other stations there was a statistically significant difference in modelled median Q_{\max} between El Niño (La Niña) and non-El Niño (non-La Niña) years, but none for observed data (or vice versa). Finally, there are no stations at which modelled and observed median Q_{\max} show significant changes between El Niño/non-El Niño years or La Niña/non-La Niña years with different signs.

4 Results and discussion

In this section, we first show and discuss the relationships between annual floods and ENSO at global and regional scales, and then examine how these relationships have changed over time. We also show and discuss the anomalies of annual flood discharge

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Niña and decreases with El Niño conditions, than vice versa. This finding is important, since past studies examining relationships between the impacts of flood disasters and ENSO at the global scale have tended to only focus on El Niño episodes, and not La Niña. For example, a study of the relationship between ENSO and the frequency of global floods of sufficient magnitude to trigger international disasters (Dilley and Heyman, 1995) found no differences between El Niño and non-El Niño years, but La Niña years were not evaluated. Similarly, another study of the number of people affected by natural hazards also only examined differences between El Niño and non-El Niño years (finding strong relationships), but did not examine La Niña years (Bouma et al., 1997). Whilst the results here only show that annual floods (which come in all sizes) are correlated with ENSO, the relationships suggest that there may also be significant links between ENSO and floods large enough to lead to flood disasters. Indeed, fitting extreme value distributions (e.g. Gumbel, GEV, etc.) to the annual flood series leads to significantly different estimates of extreme floods when either El Niño or La Niña years are dropped from the time-series (not shown here).

We also found that Q_{\max} is more sensitive to changes in SOI than is Q_{ann} in basins covering the majority (76 %) of the Earth's land surface (Fig. 5), with sensitivity results for Q_{ann} shown in Fig. A3. If we only consider basins in which the correlation between Q_{\max} and SOI is statistically significant, the sensitivity of Q_{\max} is greater than that of Q_{ann} for basins covering 32 % of land areas, whilst Q_{ann} is more sensitive for basins covering 4 % of land areas. If we only consider basins in which the correlation between Q_{ann} and SOI is statistically significant, the sensitivity of Q_{\max} is greater than that of Q_{ann} for basins covering 31 % of land areas, whilst Q_{ann} is more sensitive for basins covering 16 % of land areas. In our earlier work (Ward et al., 2010) based on observed discharges at 622 gauging stations, we also found that, on average, ENSO has a greater impact on annual flood discharges than on mean discharges. Similarly, for observed discharges in the western USA, Cayan et al. (1999) found ENSO to have a greater impact on the number of days exceeding the 90 percentile values of streamflow as compared to the number of days exceeding the 50 percentile (i.e.

median) values. Our new results show that this pattern (annual floods generally more sensitive to ENSO than average flows) extends to the global scale. Similar results have been found for the sensitivity of discharge to the variability of atmospheric circulation over Europe (Bouwer et al., 2008).

4.2 Regional sensitivities of flood discharges to ENSO

There are several regions in which it is common knowledge that climate is affected by ENSO through teleconnections (Kiladis and Diaz, 1989), for example eastern Australia, Southeast Asia, parts of western South America, and western North America. However, little is known on the influence of ENSO teleconnections on annual floods at these large regional scales. In Table 2, we show the area-weighted percentage differences (unsigned) in Q_{\max} per unit change in SOI, per by geographical region (Kummu et al., 2010) and Köppen climate zone (Kottek et al., 2006), for those basins where the correlations in Fig. 4 are significant, and the percentage of land in each region/zone combination for which the correlations are significant. Globally, in those basins with significant correlation (i.e. basins covering 37% of global land surface), Q_{\max} varies by 27% for each unit change in SOI: this includes regions far removed from the classic ENSO-regions named above. In equatorial regions, Q_{\max} is significantly correlated with SOI in basins covering half of the land areas.

The highest sensitivities are found in arid regions, followed by equatorial regions. The sensitivity of discharge to ENSO in tropical regions has been widely reported (e.g. Dettinger and Diaz, 2000; Dettinger et al., 2001; Ward et al., 2010), since ENSO affects climate in tropical regions through perturbations in the Walker circulation (Kiladis and Diaz, 1989). However, less research has assessed the influence of ENSO on the hydroclimatology in arid regions. Whilst the paucity of observed discharge data in many of these regions limits the validation of our model results there, the strength of the signal provides motivations for enhancing research activities in those regions. This is especially the case since many arid regions of the developing world are expected to show some of the world's largest increases in population and asset values

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in coming years (Jongman et al., 2012). One arid region in which there are good records of discharge, and an abundance of studies on ENSO and hydroclimatology, is the southwest USA. Here, several studies have indeed found strong relationships between ENSO and annual or seasonal discharge (e.g. Hidalgo and Dracup, 2003; Piechota et al., 1997) or between ENSO and drought conditions (Piechota and Dracup, 1996). Generally, these studies found wetter conditions in El Niño years and drier conditions in La Niña years. Cayan and Webb (1992) and Cayan et al. (1999) also found relationships between ENSO and extreme discharges at a large number of locations in the arid regions of the southwest USA, with high flows conditions being more likely in El Niño years (see Sect. 4.4 for details).

In terms of geographic regions, the highest sensitivities of annual floods to ENSO are found in Central America (54 %) and the lowest in Western Europe (13 %), with large differences between climatic zone in the geographical regions. For example, floods in the equatorial zone of Australia and Oceania are far more sensitive than in the equatorial zone of North Africa. Moreover, the sensitivity is particularly high in several less developed regions (e.g. Africa, Indian Subcontinent, Central America) compared to highly developed regions (Western Europe, North America), although this does not hold for all cases (for example, sensitivity is also high in Australia and Oceania). Brown and Lall (2006) found significant correlation between the coefficient of variation of rainfall variability and per capita GDP at the country scale, and it might be useful to evaluate similar relationships between ENSO-driven hydrological variability and GDP or other development indicators.

4.3 Changes in ENSO-flood relationships through time

Whilst we have shown significant correlations between SOI and annual floods for the overall 1958–1999 period, it is known that the strength of ENSO has changed over time on timescales from millennia to decades (e.g. Cane, 2005; Li et al., 2013; Mann et al., 2005; McPhaden et al., 2006; Tudhope et al., 2001; Wunsch, 1999) and that its teleconnected influences to at least some distant regions (e.g. western North America

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and South America) have likewise varied (Gershunov and Barnett, 1998; Gershunov et al., 1999; McCabe and Dettinger, 1999; Dettinger et al., 2000). Hence, we examined whether we could find any indication of changes in the strength of the correlation between annual floods and ENSO through time. To do this, we assessed changes in the strength of the correlation between SOI_{DJF} (i.e. the mean SOI value for the months of December, January, and February) and $\ln Q_{max}$ using a 21 yr moving window, ranging from 1958–1978 to 1978–1998. DJF was chosen since the correlations are strongest for these months for the largest number of basins. In Fig. 6, we map the numbers of windows (out of 21) for which the flood-SOI correlations within the 21 yr windows are statistically significant (Pearson’s r , $\alpha = 0.10$). This figure gives an indication of the temporal stationarity (and thus long-term reliability) of the correlation between SOI and annual floods by river basin. In those basins shown in the darkest shade of red, 21 yr correlations are statistically significant throughout the 1958–1999 era. Basins with the most persistent or reliable correlations are found in southern Africa, several parts of South America, eastern Australia, the southwest USA, the Nile basin, northern India, and several basins in central and northern Asia.

In Fig. 7, we show how the strength of this correlation has changed over the period 1968–1988 (again based on the 21 yr moving windows described above, whereby 1968–1988 are the central years of the moving windows) for selected basins. The blue lines indicate the time-varying correlations in 21 yr windows, and the red dashed lines indicate the critical value ($r = 0.369$, $\alpha = 0.10$). We also show whether there are significant linear trends in the strength of the correlations (“no trend”; correlations growing “stronger” over time; or growing “weaker” over time), based on the Mann–Kendall test ($\alpha = 0.10$). The analyses were carried out for the 50 largest basins for which correlation over the entire period 1958–1999 proved significant. In order to make the figure more clear, we then removed several upstream sub-basins (e.g. several Amazon tributaries), where the overall signal was similar to that at the most downstream sub-basin.

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Several interesting regional patterns can be seen in Fig. 7. In South America, it appears that the strength of correlations between ENSO and flood discharges have strengthened over the period of analysis here in basins from Brazil southwards, except for the Tocantins. On the other hand, in northern South America, the strength of the correlation has weakened in the Essequibo, with no significant trend in the Magdalena. In North America, correlations for the Ohio and Lower Missouri (both tributaries of the Mississippi) have strengthened, albeit as an increasingly positive correlation in the former and an increasingly negative correlation in the latter. The results for the Colorado show no significant trend over the study period. The only basin shown here in North America for which the strength of the correlation weakened significantly is the Fraser River.

For basins in western and north central Eurasia, we found either weakening correlations, or no significant trends. For both the Yenisei and the Rhine, we found fairly strong negative correlations until the 21 yr period centred on ca. 1980, and much weaker correlation thereafter (reaching zero for 21 yr periods centred after 1996 in the case of the Rhine). In south Asia, the two basins shown both exhibit weakening correlations. On the other hand, the basins shown in eastern Asia (Chao Phraya, Yellow, Kolyma) all show trends of strengthening positive correlations over time. For the basins shown in Australia and Africa, a highly mixed picture in terms of trends emerges; however, it should be noted that the strength of the correlations remains rather strong in the majority of these basins throughout the study period.

On the whole, of the 35 basins highlighted in Fig. 7, correlations strengthened in 14 basins, weakened in 13, and exhibited no trend in 8. Thus, globally, there has been essentially no overall bias among the changing teleconnections in one direction or the other. This even global mix of strengthening vs. weakening teleconnections may suggest that the changes shown in Figs. 6 and 7 reflect changes in teleconnection strengths, rather than changes in the strength of the driving ENSO variations. The latter might more likely have yielded more universally consistent changes in flood correlations.

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In some regions, these long-term variations in ENSO teleconnections have been recognised in annual precipitation and streamflow records. For example, in the Mekong Räsänen and Kummu (2013) found epochal behaviour in ENSO-discharge correlations, with strongly negative correlations in the pre-1940s and after the mid-1970s, but a rather weak relationship between these periods (Räsänen and Kummu, 2013). Similar results were found by Zubair and Chandimala (2006), who investigated ENSO-seasonal stream flow relationships in Sri Lanka, and found that the correlations changed from positive (pre-1950) to strongly negative (post-1970). These findings are supported by other studies, which found similar epochal behaviour in relationship between ENSO and Asian-Australian Monsoon (Wang et al., 2008), and ENSO and Indian Monsoon (Torrence and Webster, 1999). Outside Asia, Beebee and Manga (2004) found that ENSO correlations with snowmelt runoff in Oregon, USA, were weaker between 1920–1950 than in periods before and after those decades. Such long-term variations in ENSO teleconnections have been associated with interferences and enhancements from multi-decadal climate modes, for example ENSO interactions or reflections of the Pacific Decadal Oscillation (PDO) as reported by Gershunov and Barnett (1998), Gershunov et al. (1999), and McCabe and Dettinger (1999). There are many such “low frequency” modes in the climate system, including also modes in the Atlantic (e.g. Apipattanavis et al., 2009; McCabe et al., 2004) and Indian (e.g. Hoerling et al., 2009, 2010) Ocean basins, so that the particular interferences at work in any given river basin may be complex and likely requires more research to identify. Indeed, it is also possible that human-caused multi-decadal climate trends themselves may be modifying some of these teleconnection strengths, or may do so in the future.

4.4 Flood discharge differences between ENSO phases

We also examined the differences in anomalies of annual flood discharge between ENSO phases. Figure 8 shows anomalies of median Q_{\max} in: (a) El Niño years compared to all years, and (b) La Niña years compared to all years. In a general sense, the patterns are similar to those shown in Fig. 4. However, this analysis allows

Knowledge of these El Niño-La Niña asymmetries can be useful for more precisely targeting (on one ENSO phase or the other, or both) plans and accommodations for the ENSO influences on flood magnitudes and, ultimately, flood risks in individual basins around the world.

5 4.5 Implications and recommendations

Given our finding that ENSO correlates significantly with annual flood discharge in basins covering over a third of global land surfaces, there is a clear need for more research on the influence of interannual and longer term climate variability on flood hydrology. This would complement (and lend greater practical urgency to) ongoing efforts to better understand the roles of ocean-atmosphere interactions on climate more generally, such as that carried out under CLIVAR (Climate Variability and Predictability Programme of the World Climate Research Programme). Specifically, if significant correlations exist between ENSO and even more extreme flood discharges (i.e. larger than most annual-flood levels), then the socioeconomic impacts of flooding (e.g. loss of life, displacement of people, economic damage) in some regions may also be related to, and predictable from, ENSO. To examine this, future research may assess the impacts of ENSO directly on flood risk, where risk is a product of the probability of flooding and the consequences of flooding.

With regards to such an assessment of ENSO's impacts on flood risks, we have already mentioned that two past studies have explored possible relationships between ENSO and the frequency of floods of sufficient magnitude to trigger international disasters (Dilley and Heyman, 1995) and the number of people affected by natural hazards (Bouma et al., 1997). However, these studies only examined differences between El Niño years and non-El Niño years, and did not consider differences in the La Niña phase of ENSO. The latter analysis is essential, since our results show that there are more basins where annual floods increase with La Niña and decrease with El Niño than vice versa. Moreover, the analyses of Bouma et al. (1997) and Dilley and Heyman (1995) were based on statistics at national scales. Our results

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show that analyses at such a scale will mask important relationships, since ENSO can have an opposite correlation with annual flood discharge within different regions of the same country (as in Brazil, China, and the USA). Finally, the studies named above were carried out almost twenty years ago, whilst reporting on flood impacts (especially deaths and economic losses) has since improved (Kron et al., 2012). An analysis of the relationships between ENSO and flood impacts could be considered using two approaches: (a) making use of reported impacts such as those documented in the EM-DAT database (The OFDA/CRED International Disaster Database – www.emdat.net – Université catholique de Louvain – Brussels – Belgium) or the NatCatService database (Munich Re, 2013); or (b) using recently developed flood risk models such as those of Hirabayashi et al. (2013), UNISDR (2011), or Winsemius et al. (2013).

Another promising research avenue would be to use the potential predictability of ENSO (Cheng et al., 2011) to provide probabilistic estimates of flood risk with lead times up to several seasons. The coupling of ENSO predictability with hydrometeorological variables such as precipitation and mean discharge has been on the research agenda for over a decade. However, also coupling such analyses with flood statistics and global risk models could provide probabilistic flood risk forecasts; enabling humanitarian and development agencies to prioritise short-term risk reduction efforts in the most at-risk regions; enabling (re-)insurance companies to accommodate anomalies in their risk portfolios in the coming seasons to years; and potentially enabling improved flood early warning and flood regulation by dam operators. However, in this study we have shown that the strength of the correlation between ENSO and annual floods is itself non-stationary through time. Hence, such analyses may be more suited to those regions where the temporal persistence of the ENSO-flood correlations is highest.

Technical flood-defences are designed to protect against floods with given return-periods, estimated from observed discharge records. However, should ENSO magnitude and frequency change over time, as has occurred in the recent and geological past (Mann et al., 1995), this would result in effective over- or underdesign of

flood protection infrastructure, such as dikes, for decades at a time. The present study identifies some areas where this may be most likely, i.e. those locations where floods are particularly sensitive to changes in ENSO.

Along with these extensions of the current research to potentially facilitate flood risk analyses, a number of analytical steps could be improved in future research. For example, future research could examine the relationships between ENSO and floods using additional climate indices or several different global hydrological models. Also, given the possible interaction between ENSO and other large-scale climate oscillations (such as PDO) that may serve to modulate ENSO relationships with flood discharge, analyses should be carried out using a wide range of interannual ocean-atmosphere interactions, in addition to ENSO. Finally, we recommend that future studies carry out detailed analyses of relationships between Q_{\max} and its climatological forcing to reveal regions in which climate dominates Q_{\max} variability, vs. those where this effect may be decreased or amplified by other factors (e.g. terrain, soil, cropping, or human flow management).

5 Conclusions

In this paper, we provide the first fully global assessment of ENSO-driven climate variability's influence on annual floods. This was achieved by simulating daily discharges over the period 1958–1999 using the WaterGAP model forced by global climate reanalysis data from the WATCH project. We first validated the simulated annual flood discharges by comparisons to observed discharges. We found that whilst there are large biases between modelled and simulated annual floods, they simulate similar relative changes in annual floods from year to year, and that their agreement is good in terms of the signal of change between different phases of ENSO. Whilst studies on the linkages between ENSO and flood discharge based on observations are limited, for those studies that are available, their findings are generally in line with our model results, adding confidence to our use of modelled data in these analyses.

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term risk reduction efforts in the most at risk regions; (re-)insurance companies to assess anomalies in their risk portfolios in the coming seasons to years; and potentially enable improved flood early warning and flood regulation of dams.

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Table 1. Hydrological years categorised as El Niño and La Niña, based on the ENSO classification of the Center for Ocean-Atmospheric Prediction Studies (COAPS) of Florida State University (<http://coaps.fsu.edu/jma.shtml>). Hydrological years are identified according to the calendar year at their beginning.

ENSO mode	Hydrological year
El Niño	1963, 1965, 1969, 1972, 1976, 1982, 1986, 1987, 1991, 1997
La Niña	1964, 1967, 1970, 1971, 1973, 1974, 1975, 1988, 1998, 1999

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Table 2. Sensitivity of annual flood discharge to SOI^a.

	Percentage change in Q_{\max} per unit change SOI (land area, %)				
	Equatorial	Arid	Warm temperate	Cold ^b	All Köppen zones
Australia and Oceania	36 % (43 %)	63 % (45 %)	28 % (29 %)	16 % (23 %)	45 % (39 %)
Central America	29 % (41 %)	98 % (62 %)	26 % (21 %)	N/A	54 % (46 %)
Eastern Asia	N/A	31 % (18 %)	11 % (51 %)	21 % (10 %)	16 % (29 %)
Eastern Europe & Central Asia	N/A	24 % (33 %)	15 % (12 %)	11 % (22 %)	14 % (23 %)
Indian Subcontinent	15 % (45 %)	63 % (30 %)	N/A	N/A	35 % (26 %)
Latin America	14 % (53 %)	49 % (27 %)	16 % (82 %)	205 % (25 %)	17 % (54 %)
Middle East	N/A	42 % (40 %)	11 % (12 %)	13 % (72 %)	40 % (38 %)
Middle and South Africa	21 % (55 %)	52 % (49 %)	64 % (17 %)	N/A	35 % (51 %)
North Africa	5 % (73 %)	39 % (38 %)	25 % (34 %)	N/A	32 % (41 %)
North America	N/A	50 % (55 %)	17 % (32 %)	11 % (19 %)	19 % (25 %)
Southeastern Asia	20 % (37 %)	N/A	8 % (37 %)	N/A	18 % (37 %)
Western Europe	N/A	N/A	13 % (28 %)	11 % (27 %)	13 % (28 %)
All regions	19 % (50 %)	49 % (40 %)	16 % (37 %)	12 % (21 %)	27 % (37 %)

^a Area-weighted percentage change in Q_{\max} per unit change SOI for basins with significant correlations between $\ln Q_{\max}$ and SOI, and percentage of land area (shown in brackets) with significant correlations. Results shown per geographical region and Köppen climate region. ^b Köppen regions “polar” and “snow” have been combined here into one “cold” region. Greenland and Antarctica are excluded from the analysis.

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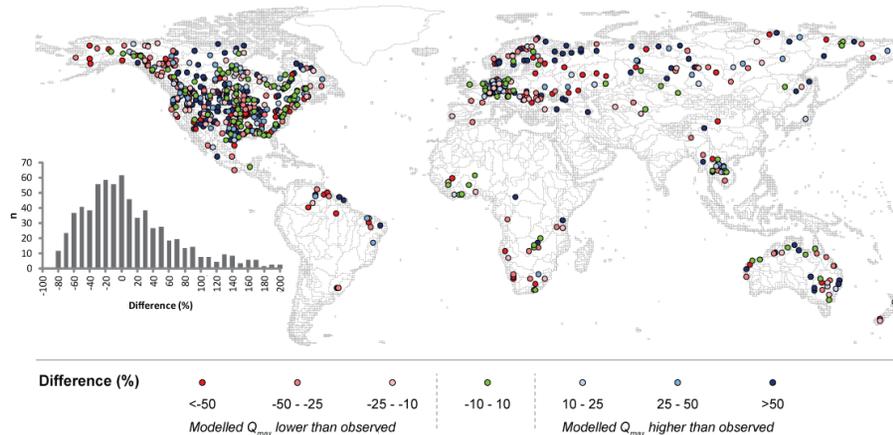


Fig. 1. Percentage difference between modelled and observed median Q_{max} over the period 1958–1999. Inset shows histogram of values for individual locations (difference > 200% for 58 locations).

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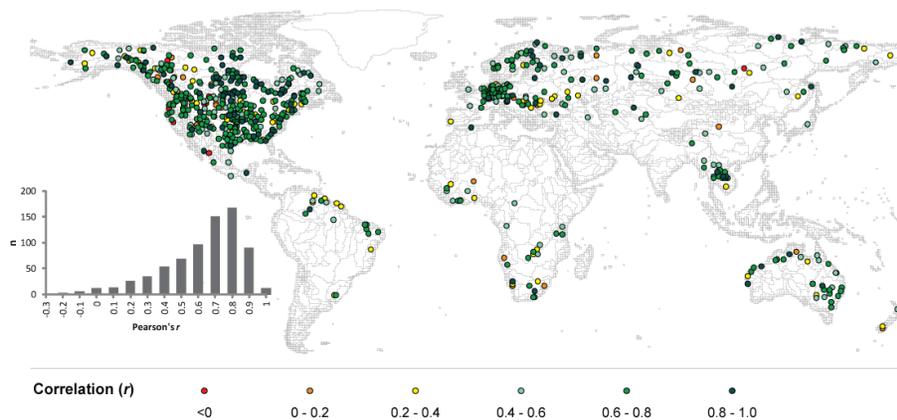


Fig. 2. Correlation (r) between modelled and observed $\ln Q_{\max}$ time-series over the period 1958–1999. Inset shows histogram of values for individual locations.

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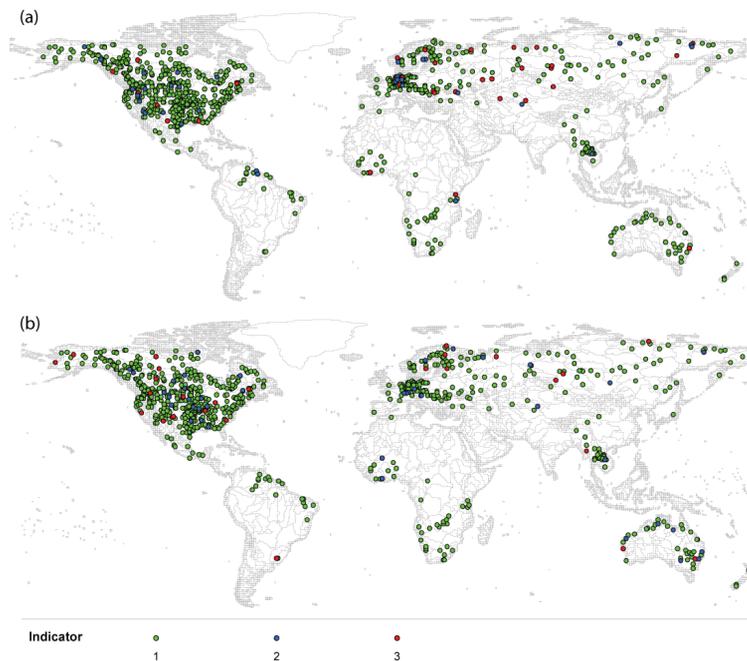


Fig. 3. Indicator of the (dis-)agreement between modelled and observed median Q_{\max} between **(a)** El Niño years and non-El Niño years and **(b)** La Niña years and non-La Niña years, whereby: (1, green) indicates that both modelled and observed median Q_{\max} show either no significant difference between **(a)** El Niño and non-El Niño years and **(b)** La Niña and non-La Niña years, whilst modelled and observed median Q_{\max} show significant difference between **(a)** El Niño and non-El Niño years with same sign and **(b)** La Niña and non-La Niña years with same sign; (2, blue) indicates that modelled median Q_{\max} shows no significant difference between **(a)** El Niño and non-El Niño years and **(b)** La Niña years and non-La Niña years, whilst observed shows a significant difference; and (3, red) indicates that observed median Q_{\max} shows no significant difference between **(a)** El Niño and non-El Niño years and **(b)** La Niña years and non-La Niña years, whilst modelled shows a significant difference. There are no stations for which modelled and observed median Q_{\max} show significant differences between El Niño and non-El Niño (or La Niña and non-La Niña years) years with the opposite sign. Statistical significance was assessed using a 2-tailed Mann–Whitney U (MWU) test, $\alpha = 0.05$.

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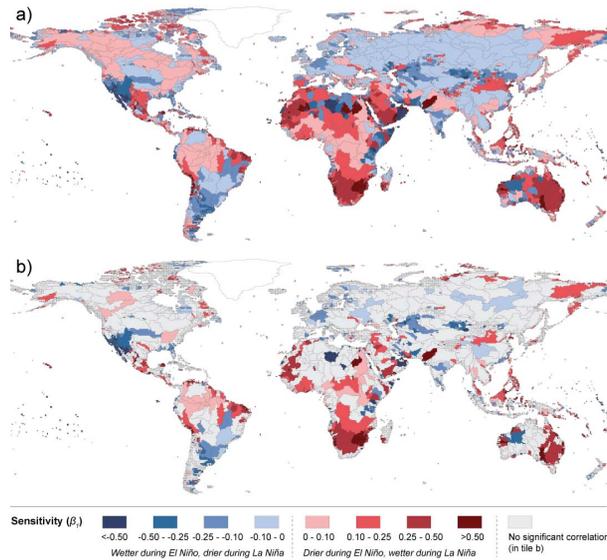


Fig. 4. Sensitivity (β_1) of $\ln Q_{\max}$ to variations in SOI. In the upper plate (a), the sensitivity is shown for all basins. In the lower plate (b), the sensitivity is only shown for basins with significant correlation (Pearson's r , t statistic, $\alpha = 0.10$) (basins where the correlation is not significant are shown in grey). Blue indicates negative correlation (higher annual floods in El Niño years/lower annual floods in La Niña years); and red indicates positive correlation (lower annual floods in Niño years/higher annual floods in La Niña years).

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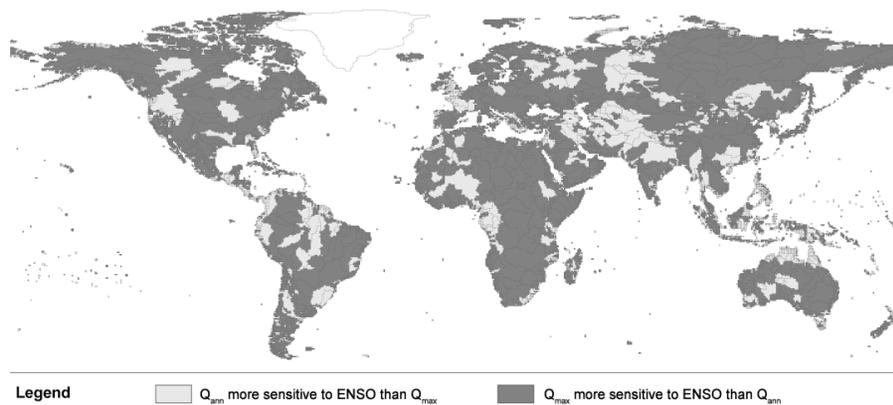


Fig. 5. Basins in which Q_{max} (Q_{ann}) is more sensitive than Q_{ann} (Q_{max}) to changes in SOI are shown in dark (light) grey.

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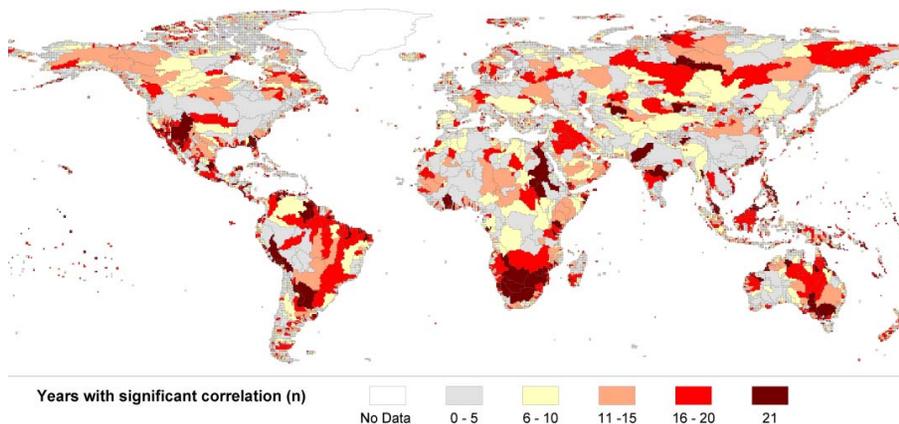


Fig. 6. Number of years (n) for which the correlation (Pearson's r) between $\ln Q_{\max}$ and SOI_{DJF} is statistically significant ($\alpha = 0.10$), based on 21 yr moving-windows centred on 1968 to 1988.

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Fig. 7. Correlation (Pearson's r) between $\ln Q_{\max}$ and SOI_{DJF} for 21 yr moving windows centred on the years 1968–1988. On the line graphs, the axes are unlabelled due to space constraints: the x-axes show years (1968–1998) and the y-axes show the strength of Pearson's r (+1 to -1). The blue line indicates the strength of the correlation, and the red dashed lines indicate the critical values of the significance test (critical $r = 0.369/-0.369$, $\alpha = 0.10$). The names of the basins are shown in text, as well as the trends in the strength of the correlation over time (Sect. 4.3).

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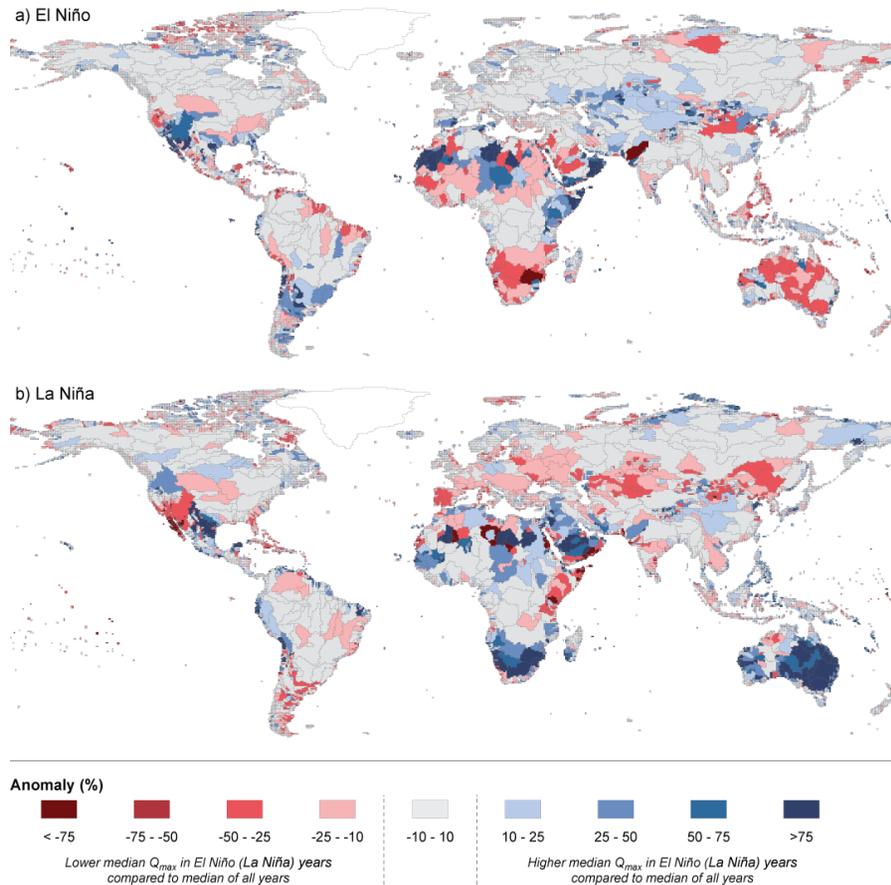


Fig. 8. Anomaly (percentage) in median Q_{max} between: **(a)** El Niño years and all years; and **(b)** La Niña years and all years. Blue indicates higher median Q_{max} (i.e. higher annual flood discharge) in El Niño (La Niña) years compared to the median of all years; and red the opposite.

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Annual flood sensitivities to ENSO at the global scale

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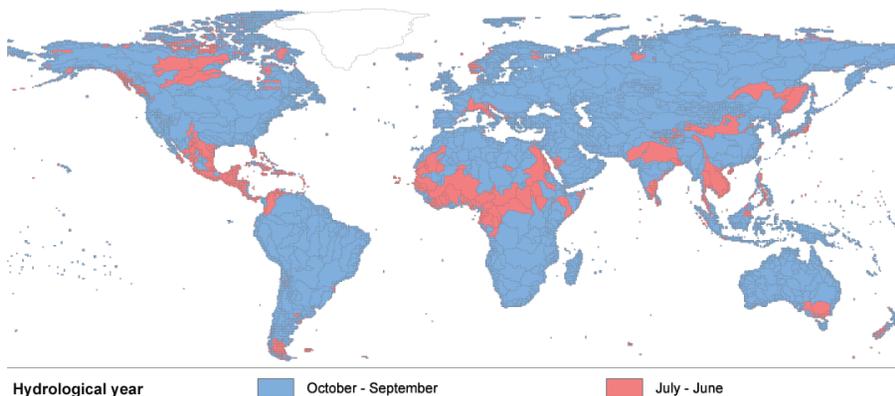


Fig. A1. Hydrological years used in this study for each basin. The standard hydrological year (October-September) was used as default, except for in those basins in which the mean Q_{\max} of the most downstream cell occurs in the months of September, October, or November. In the latter case, the hydrological year was set to July–June.

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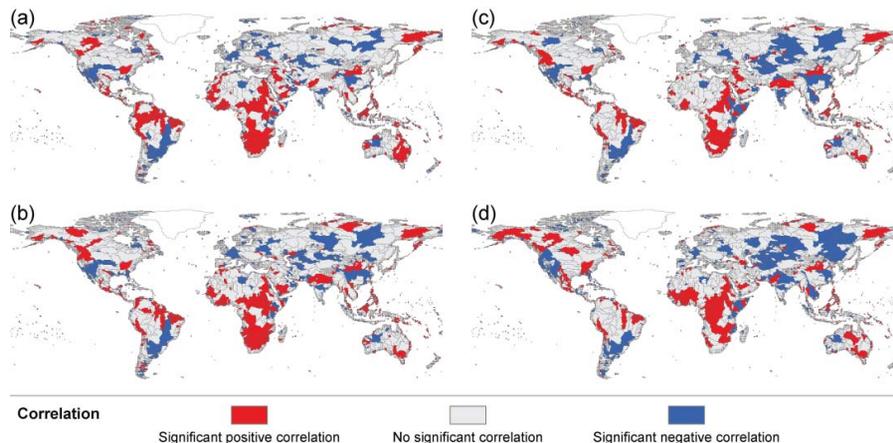


Fig. A2. Basins for which there is significant correlation (positive in red; negative in blue) between $\ln Q_{\max}$ and various indices of ENSO: **(a)** SOI; **(b)** inverse Multivariate ENSO Index (MEI); **(c)** inverse NINO3.4 index; and **(d)** inverse Global SST ENSO index. Basins with no significant correlation are shown in grey. Statistical significance was tested using the t -statistic ($\alpha = 0.10$).

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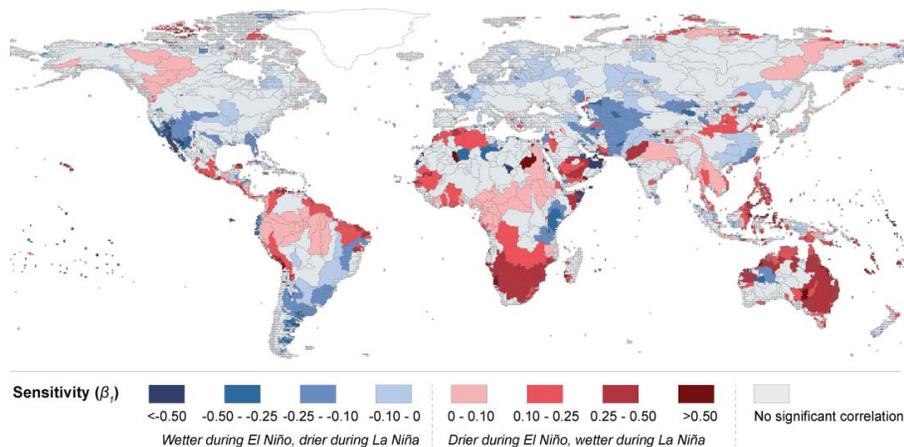


Fig. A3. Sensitivity (β_1) of $\ln Q_{\text{ann}}$ to variations in SOI. Basins with no significant correlation (Pearson's r , $\alpha = 0.10$) are shown in grey. Blue indicates negative correlation (wetter El Niño/drier La Niña); and red indicates positive correlation (drier El Niño/wetter La Niña).

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